

**ON THE CONSTITUENT ITEMS OF THE REDUCTION AND THE
REMAINDER IN THE METHOD OF LEAST SQUARES**

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1. Consider a set of variates y_i , ($i = 1, 2, \dots, n$), which are normally and independently distributed with variance 1. Let also a matrix (x_{ik}) with $i = 1, 2, \dots, n; k = 1, 2, \dots, s$ and rank s be given. Find b_1, \dots, b_s in terms of y_i so that

$$\psi^2 = \sum_i (y_i - \sum_k x_{ik} b_k)^2$$

is a minimum. This minimum value shall be denoted by ψ_{\min}^2 .

It is known (see e.g. R. A. Fisher, "Applications of Student's distribution", *Metron* Vol. 5, Part 3 (1925)) that ψ_{\min}^2 varies as does χ^2 with $n - s$ degrees of freedom and that it is possible to express ψ_{\min}^2 as the sum of $n - s$ squares of linear functions of the y_i . In the following lines $\sum_i y_i^2$ will be expressed as the sum of n squares of such functions which are independent and of variance 1. The sum of the first s squares will equal $\sum_i y_i^2 - \psi_{\min}^2$ and therefore the remaining $n - s$ squares equal ψ_{\min}^2 .

Thus a simple way will be found of writing down explicitly the linear functions, whose existence only was proved by Professor Fisher in *Metron*.

2. We first calculate ψ_{\min}^2 .

$\frac{\partial \psi^2}{\partial b_l} = 0$, for $l = 1, 2, \dots, s$, gives the normal equations

$$(1) \quad \sum_{i=1}^n x_{il} y_i = \sum_{i=1}^n \sum_{k=1}^s x_{il} x_{ik} b_k,$$

which can be written

$$(2) \quad \sum_{i=1}^n x_{il} y_i = \sum_{k=1}^s X_{lk} b_k$$

with

$$X_{lk} = \sum_{i=1}^n x_{il} x_{ik}.$$

It follows from (1) that

$$(A) \quad \psi_{\min}^2 = \sum_{i=1}^n y_i^2 - \sum_{i=1}^n \sum_{l=1}^s \sum_{k=1}^s x_{il} x_{ik} b_l b_k = \sum_{i=1}^n y_i^2 - \sum_{l=1}^s \sum_{k=1}^s X_{lk} b_l b_k,$$

where the b are solutions of (1).

3. A second expression for ψ_{\min}^2 can be found as follows:
Introducing

$$c_i = \sum_{k=1}^s x_{ik} b_k$$

we obtain from (1)

$$(3) \quad \sum_{i=1}^n x_{il} c_i = \sum_{i=1}^n x_{il} y_i, \quad (l = 1, 2, \dots, s).$$

Now if z_{iu} , ($u = s + 1, \dots, n$), are any $n - s$ independent solutions of

$$\sum_{i=1}^n z_{iu} x_{il} = 0, \quad (l = 1, 2, \dots, s),$$

then the c_i satisfy also

$$(4) \quad \sum_{i=1}^n z_{iu} c_i = 0, \quad (u = s + 1, \dots, n).$$

Let such a set of z_{iu} be chosen. Then (3) will be solved by

$$(5) \quad c_i = y_i - \sum_{v=s+1}^n \lambda_v z_{iv}$$

with λ_v as indefinite factors and these c_i satisfy (4), if

$$(6) \quad \begin{aligned} \sum_{i=1}^n z_{iu} y_i &= \sum_{v=s+1}^n \sum_{i=1}^n z_{iu} z_{iv} \lambda_v, \quad (u = s + 1, \dots, n), \quad \text{or} \quad \sum_{i=1}^n z_{iu} y_i \\ &= \sum_{v=s+1}^n Z_{uv} \lambda_v \end{aligned}$$

with

$$Z_{uv} = \sum_{i=1}^n z_{iu} z_{iv}.$$

Because of (2) the equation (A) can be transformed into

$$\psi_{\min}^2 = \sum_{i=1}^n y_i^2 - \sum_{l=1}^s \sum_{i=1}^n x_{il} y_i b_l = \sum_{i=1}^n y_i^2 - \sum_{i=1}^n y_i c_i = \sum_{i=1}^n \sum_{v=s+1}^n \lambda_v z_{iv} y_i$$

which is, because of (6)

$$(B) \quad \psi_{\min}^2 = \sum_{u=s+1}^n \sum_{v=s+1}^n Z_{uv} \lambda_u \lambda_v,$$

where the λ are solutions of (6).

The comparison of (A) and (B) gives

$$\sum_{i=1}^n y_i^2 = \sum_{l=1}^s \sum_{k=1}^s X_{lk} b_l b_k + \sum_{u=s+1}^n \sum_{v=s+1}^n Z_{uv} \lambda_u \lambda_v$$

where the first form on the r.h.s. shows the reduction of $\sum_{i=1}^n y_i^2$ by the method of least squares and the second form constitutes the remainder.

4. These two forms must now be expressed in terms of the y_i .

We introduce the notations

$$X^{(1)} = X_{11}, \quad X^{(2)} = \begin{vmatrix} X_{11} & X_{12} \\ X_{21} & X_{22} \end{vmatrix}, \quad \dots \quad X^{(s)} = \begin{vmatrix} X_{11} & \dots & X_{1s} \\ \dots & \dots & \dots \\ X_{s1} & \dots & X_{ss} \end{vmatrix}$$

and

$$Z^{(s+1)} = Z_{s+1, s+1}, \quad Z^{(s+2)} = \begin{vmatrix} Z_{s+1, s+1} & Z_{s+1, s+2} \\ Z_{s+2, s+1} & Z_{s+2, s+2} \end{vmatrix} \quad \text{etc.}$$

It is well known (and can easily be verified) that

$$\begin{aligned} \sum_{l=1}^s \sum_{k=1}^s X_{lk} b_l b_k &= \frac{1}{X^{(1)}} (X_{11} b_1 + \dots + X_{1s} b_s)^2 \\ &+ \frac{1}{X^{(1)} X^{(2)}} \left(\begin{vmatrix} X_{11} & X_{12} \\ X_{21} & X_{22} \end{vmatrix} b_2 + \dots + \begin{vmatrix} X_{11} & X_{1s} \\ X_{21} & X_{2s} \end{vmatrix} b_s \right)^2 \\ &+ \dots + \frac{1}{X^{(s-1)} X^{(s)}} X^{(s)^2} b_s^2 \end{aligned}$$

which may be written

$$\begin{aligned} \frac{1}{X^{(1)}} \left(\sum_{k=1}^s X_{1k} b_k \right)^2 &+ \frac{1}{X^{(1)} X^{(2)}} \begin{vmatrix} X_{11} & \sum_{k=1}^s X_{1k} b_k \\ X_{21} & \sum_{k=1}^s X_{2k} b_k \end{vmatrix}^2 \\ &+ \dots + \frac{1}{X^{(s-1)} X^{(s)}} \begin{vmatrix} X_{11} X_{12} \dots \sum_{k=1}^s X_{1k} b_k \\ \dots & \dots & \dots \\ X_{s1} X_{s2} \dots \sum_{k=1}^s X_{sk} b_k \end{vmatrix}^2. \end{aligned}$$

Using (2), this can be expressed in terms of the y_i instead of b_k as follows:

$$\begin{aligned} \frac{1}{X^{(1)}} \left(\sum_{i=1}^n x_{i1} y_i \right)^2 &+ \frac{1}{X^{(1)} X^{(2)}} \begin{vmatrix} X_{11} & \sum_{i=1}^n x_{i1} y_i \\ X_{21} & \sum_{i=1}^n x_{i2} y_i \end{vmatrix}^2 \\ (7) \quad &+ \dots + \frac{1}{X^{(s+1)} X^{(s)}} \begin{vmatrix} X_{11} X_{12} \dots \sum_{i=1}^n x_{i1} y_i \\ \dots & \dots & \dots \\ X_{s1} X_{s2} \dots \sum_{i=1}^n x_{is} y_i \end{vmatrix}^2. \end{aligned}$$

If, however, $s = 2, n = 3$, then easy calculations lead to

$$\begin{aligned} \psi_{\min}^2 &= \sum_{i=1}^n y_i^2 - \frac{(x_{11}y_1 + x_{21}y_2 + x_{31}y_3)^2}{x_{11}^2 + x_{21}^2 + x_{31}^2} \\ &\quad - \frac{\begin{vmatrix} x_{11}^2 + x_{21}^2 + x_{31}^2 & x_{11}y_1 + x_{21}y_2 + x_{31}y_3 \\ x_{11}x_{12} + x_{21}x_{22} + x_{31}x_{32} & x_{12}y_1 + x_{22}y_2 + x_{32}y_3 \end{vmatrix}}{\begin{vmatrix} x_{11}^2 + x_{21}^2 + x_{31}^2 & x_{11}x_{12} + x_{21}x_{22} + x_{31}x_{32} \\ x_{12}x_{11} + x_{22}x_{21} + x_{32}x_{31} & x_{12}^2 + x_{22}^2 + x_{32}^2 \end{vmatrix}} \\ &= \left(\begin{vmatrix} x_{21}x_{31} \\ x_{22}x_{32} \end{vmatrix} y_1 + \begin{vmatrix} x_{31}x_{11} \\ x_{32}x_{12} \end{vmatrix} y_2 + \begin{vmatrix} x_{11}x_{21} \\ x_{12}x_{22} \end{vmatrix} y_3 \right)^2 \\ &\quad \div \left(\begin{vmatrix} x_{21}x_{31} \\ x_{22}x_{32} \end{vmatrix}^2 + \begin{vmatrix} x_{31}x_{11} \\ x_{32}x_{12} \end{vmatrix}^2 + \begin{vmatrix} x_{11}x_{21} \\ x_{12}x_{22} \end{vmatrix}^2 \right). \end{aligned}$$

As a specialized case consider $s = 1$, and $x_{11} = x_{21} = \dots = x_{n1} = 1$. Then the Z are

$$\begin{matrix} 2 & 1 & 1 & 1 & \dots & 1 & 1 \\ 1 & 2 & 1 & 1 & \dots & 1 & 1 \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ 1 & 1 & 1 & 1 & \dots & 2 & 1 \\ 1 & 1 & 1 & 1 & \dots & 1 & 2 \end{matrix}$$

and

$$\psi_{\min}^2 = \sum_{i=1}^m y_i^2 = \frac{1}{n} \left(\sum_{i=1}^n y_i \right)^2 = \sum_{i=1}^n \left(y_i - \frac{\sum_{i=1}^n y_i}{n} \right)^2.$$

The sum of squares into which ψ_{\min}^2 can be transformed is then found to be

$$\begin{aligned} \frac{1}{2} (-y_1 + y_2)^2 + \frac{1}{2 \cdot 3} (-y_1 - y_2 + 2y_3)^2 \\ + \frac{1}{3 \cdot 4} (-y_1 - y_2 - y_3 + 3y_4)^2 + \dots.^1 \end{aligned}$$

¹ This is the result contained in a paper by J. O. Irwin, "Independence of the constituent items in the analysis of variance" *Suppl. Roy. Stat. Soc. Jour.* Vol. 1 (1934).