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ON STOCHASTIC APPROXIMATION METHODS¹

BY J. WOLFOWITZ

Cornell University

In [1] A. Dvoretzky proved the theorem quoted below, which implies all previous results on the convergence to a limit of stochastic approximation methods. (For a description of these results see [1].) In the present note we give a simple and, we think, perspicuous proof of this theorem which may be of help in further work. The present note is entirely self-contained and may be read without reference to [1].

THEOREM. (Dvoretzky) *Let α_n , β_n and γ_n ($n = 1, 2, \dots$) be non-negative real numbers satisfying*

$$(1) \quad \lim_{n \rightarrow \infty} \alpha_n = 0,$$

$$(2) \quad \sum_{n=1}^{\infty} \beta_n < \infty,$$

and

$$(3) \quad \sum_{n=1}^{\infty} \gamma_n = \infty.$$

Let θ be a real number and T_n ($n = 1, 2, \dots$) be measurable transformations satisfying

$$(4) \quad |T_n(r_1, \dots, r_n) - \theta| \leq \max[\alpha_n, (1 + \beta_n)|r_n - \theta| - \gamma_n]$$

for all real r_1, \dots, r_n . Let X_1 and Y_n ($n = 1, 2, \dots$) be random variables and define²

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² In the proof of the theorem we will, for the sake of brevity, write $T_n(X_n)$ for

$$T_n(X_1, \dots, X_n),$$

just as is done in [1]. No ambiguity will be caused by this.

$$(5) \quad X_{n+1}(\omega) = T_n(X_1(\omega), \dots, X_n(\omega)) + Y_n$$

for $n \geq 1$.

Then the conditions $E\{X_1^2\} < \infty$,

$$(6) \quad \sum_{n=1}^{\infty} E\{Y_n^2\} < \infty$$

and

$$(7) \quad E\{Y_n | X_1, \dots, X_n\} = 0$$

with probability 1 for all n , imply

$$(8) \quad \lim_{n \rightarrow \infty} E\{(X_n - \theta)^2\} = 0$$

and

$$(9) \quad P\{\lim_{n \rightarrow \infty} X_n = \theta\} = 1.$$

EXTENSION. The theorem remains valid if α_n and β_n in (4) are replaced by non-negative functions $\alpha_n(r_1, \dots, r_n)$ and $\beta_n(r_1, \dots, r_n)$ respectively, provided: The functions $\alpha_n(r_1, \dots, r_n)$ are uniformly bounded and

$$(10) \quad \lim_{n \rightarrow \infty} \alpha_n(r_1, \dots, r_n) = 0$$

uniformly for all sequences r_1, \dots, r_n, \dots ; the functions $\beta_n(r_1, \dots, r_n)$ are measurable and

$$(11) \quad \sum_{n=1}^{\infty} \beta_n(r_1, \dots, r_n)$$

is uniformly bounded and uniformly convergent for all sequences r_1, \dots, r_n, \dots ; and for any $L > 0$ there exist non-negative functions $\gamma_n(r_1, \dots, r_n)$ satisfying (4), and

$$(12) \quad \sum_{n=1}^{\infty} \gamma_n(r_1, \dots, r_n) = \infty$$

holds uniformly for all sequences r_1, \dots, r_n, \dots for which

$$(13) \quad \sup_{n=1,2,\dots} |r_n| < L.$$

PROOF: Without loss of generality we may take $\theta = 0$.

I. From (4) and (6) it follows readily that $EX_n^2 < \infty$ for any n .

II. Define $s(n)$ to be the sign of $[T_n(X_n)][X_n]$ if neither factor is zero, and $s(n) = 1$ if either factor is zero. Define $\pi(m, n) = \prod_{j=m}^n s(j)$, $Y'_n = \pi(1, n)Y_n$. The series $\sum_1^{\infty} Y'_n$ converges w.p.1, by Loève ([2], p. 387, D) and (6) and (7). Let

$$Z(m, n) = \sum_{j=m}^n Y'_j$$

For any δ and ϵ both >0 , there exists $M'(\delta, \epsilon)$ such that

$$(14) \quad P \left\{ \sup_{M' \leq m \leq n} |Z(m, n)| > \frac{\delta}{48} \right\} < \frac{\epsilon}{2}.$$

III. Let $d(m, m - 1) = 1$ and, for $n \geq m$,

$$d(m, n) = \prod_{j=m}^n (1 + \beta_j).$$

Consider the sum

$$S(m, n) = \sum_{j=m}^{n+1} d(j, n) Y'_{j-1},$$

which is equal to

$$(15) \quad \sum_{j=m}^{n-1} Z((m-2), (j-1)) [d(j, n) - d(j+1, n)] \\ - Y'_{m-2} d(m, n) + Z((m-2), (n-1)) d(n, n) + Y'_n.$$

Since $d(j, n) \geq d(j+1, n)$ we have that the absolute value of (15) is not greater than

$$2 \left[\sup_{m-1 \leq j \leq n} |Z((m-2), (j-1))| \right] (d(m, n)) + |Y'_n|.$$

Hence, from (11) and (14) it follows that, for δ and ϵ both >0 , there exists an $M''(\delta, \epsilon) \geq M'(\delta, \epsilon)$ such that $d(m, \infty) < \frac{3}{2}$ for $m \geq M''$ and

$$(16) \quad P \left\{ \sup_{M'' \leq m \leq n} |Z(m, n)| < \frac{\delta}{48}, \sup_{M'' \leq m \leq n} |S(m, n)| < \frac{\delta}{8} \right\} > 1 - \frac{\epsilon}{2}.$$

Proof of (9) under the conditions of the extension. Let ϵ and δ be positive and arbitrary. It is sufficient to prove that

$$(17) \quad P\{|X_n| < \delta \text{ for all } n \text{ sufficiently large}\} > 1 - \epsilon.$$

Let $M \geq M''(\delta, \epsilon)$ be so large that, for $n \geq M$, $\alpha_n < \delta/8$. Let L be so large that $L > \delta$ and

$$(18) \quad \max_{1 \leq j \leq M} EX_j^2 < \frac{\epsilon L^2}{32M}.$$

We take this to be the L for which (12) holds. It also follows that

$$(19) \quad P \left\{ \max_{1 \leq j \leq M} |X_j| \leq \frac{L}{4} \right\} > 1 - \frac{\epsilon}{2}.$$

Suppose that the following four conditions are fulfilled:

$$(20) \quad \text{The relations in curly brackets in (16);}$$

$$(21) \quad |X_m| \leq \frac{\delta}{4} \text{ for some } m \geq M;$$

$$(22) \quad |X_{m+j}| > \frac{\delta}{4}, \quad 1 \leq j \leq k;$$

$$(23) \quad |X_{m+k+1}| \leq \frac{\delta}{4}.$$

Here $1 \leq k \leq \infty$. In case $k = \infty$, (22) is to hold for all $j \geq 1$ and (23) is vacuous. (It will be clear by the time the proof is finished that k cannot = ∞ .) Because $\alpha_n < \delta/8$ for $n \geq M$ and because of (20), (21), and (22) it follows that

$$(24) \quad |T_{m+j}(X_{m+j})| > \alpha_{m+j}, \quad 0 \leq j \leq k - 1,$$

$$(25) \quad \text{sign } X_{m+j+1} = \text{sign } T_{m+j}(X_{m+j}), \quad 0 \leq j \leq k - 1.$$

Applying (4) (with the γ 's zero) we obtain that X_{m+1} lies between zero and

$$(26) \quad s(m)(1 + \beta_m)X_m + Y_m.$$

Repeating this argument, we obtain that, for $1 \leq j \leq k$, X_{m+j} lies between 0 and

$$(27) \quad \begin{aligned} & s(m+j-1)s(m+j-2) \cdots s(m) d(m, m+j-1)X_m \\ & + s(m+j-1) \cdots s(m+1) d(m+1, m+j-1)Y_m + \cdots \\ & + s(m+j-1) d(m+j-1, m+j-1)Y_{m+j-2} + Y_{m+j-1}. \end{aligned}$$

The absolute value of (27) is not greater than

$$(28) \quad |X_m| d(m, m+j-1) + |S(m+1, m+j-1)|.$$

Hence

$$(29) \quad |X_{m+j}| < \delta, \quad 1 \leq j \leq k.$$

To prove (17) it remains only to show that the following conditions cannot both hold:

$$(30) \quad \text{the relations in curly brackets in (16) and (19);}$$

$$(31) \quad |X_n| > \frac{\delta}{4} \text{ for all } n \geq M.$$

Applying the argument of the previous paragraph with δ replaced by L we obtain that

$$(32) \quad |X_n| < L \text{ for all } n \geq 1.$$

Hence (12) holds. In view of (30) and (31) it follows that

$$(33) \quad |T(X_n)| > \alpha_n \text{ for all } n \geq M - 1,$$

$$(34) \quad \text{sign } T_n(X_n) = \text{sign } X_{n+1} \text{ for all } n \geq M - 1.$$

We may now, and do, apply the argument which led to (28), but with the γ 's which satisfy (12). We conclude that, for all $n > M$, the absolute value of $|X_n|$ is not greater than

$$(35) \quad |X_M| d(M, n - 1) + |S(M + 1, n - 1)| - \sum_{j=M}^{n-1} \gamma_j$$

For n sufficiently large this becomes negative, contradicting (33) and hence (31). This completes the proof of (9).

The fact that $EX_1^2 < \infty$ is used in the above proof only in order that $EX_n^2 < \infty$ for all n , and this latter fact is needed only for (8), and not for (9). For in the proof above we used the fact that $EX_n^2 < \infty$ only to obtain explicitly an L for which (19) holds. Such an L obviously exists whether or not $EX_n^2 < \infty$.

Proof of (8) under the conditions of the extension. Let $K = \max_{1 \leq j < \infty} \alpha_j$. Let N be an integer to be chosen later. In view of (9) we have only to prove that $\lim_{n \rightarrow \infty} E\{|X_n| - K\}^2 = 0$. Let P denote probability measure and A be any set in the sample space which can be defined in terms of X_1, \dots, X_m . We use the inequality

$$(36) \quad \begin{aligned} H_{m+1}(A) &= \int_A ((|X_{m+1}| - K)^+)^2 dP = \int_A ((|T_m(X_m) + Y_m| - K)^+)^2 dP \\ &\leq \int_A [Y_m^2 + ((|T_m(X_m)| - K)^+)^2] dP \\ &\leq \int_A [Y_m^2 + K\beta_m(1 + K\beta_m) + (1 + \beta_m)^2(1 + K\beta_m)((|X_m| - K)^+)^2] dP \end{aligned}$$

which is in [1] and can be deduced from (4) and (7). Let $B(j)$ be the set $\{|X_{N+j}| \leq K, |X_{N+i}| > K \text{ for } 0 \leq i < j\}$, $D(j)$ the complement of

$$B(0) + B(1) + \dots + B(j).$$

Iterate the inequality (36) to obtain an upper bound on $H_n(A)$, $n > N$, beginning the iteration at $m = N, N + 1, \dots, n - 1$, respectively, and using as A the sets $B(0), B(1), \dots, B(n - N - 1)$, respectively. In each case the last term of the integrand of the right member of (36) vanishes. Adding, we obtain that $H_n(B(0) + \dots + B(n - N))$ can be made arbitrarily small by making N sufficiently large.

It remains only to consider $H_n(D(n - N))$. For any point in $D(n - N)$ we have, as in (27), that

$$(37) \quad |X_n| \leq |\pi(1, N - 1) d(N, n - 1)X_N + S(N + 1, n - 1)|$$

Hence, by Minkowski's inequality

$$(38) \quad \begin{aligned} &\left(\int_{D(n-N)} (X_n)^2 dP \right)^{\frac{1}{2}} \\ &\leq [d(1, \infty)] \left(\int_{D(n-N)} (X_N)^2 dP \right)^{\frac{1}{2}} + [d(1, \infty)] \left(\sum_{j=N}^{\infty} EY_j^2 \right)^{\frac{1}{2}} \end{aligned}$$

The second term on the right of (38) can be made arbitrarily small by making N sufficiently large. The first term can be made arbitrarily small by making n sufficiently large, since $P\{D(n - N)\} \rightarrow 0$ as $n \rightarrow \infty$. This completes the proof of (8).

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ON THE DERIVATIVES OF A CHARACTERISTIC FUNCTION AT THE ORIGIN

BY E. J. G. PITMAN

University of Tasmania

1. Introduction. Let $F(x)$, $-\infty < x < \infty$, be a distribution function, and

$$\phi(t) = \int_{-\infty}^{\infty} e^{itx} dF(x)$$

its characteristic function, defined and continuous for all real t . Let k be a positive integer. If the k th moment of $F(x)$,

$$\mu_k = \int_{-\infty}^{\infty} x^k dF(x),$$

exists and is finite (integral absolutely convergent), $\phi(t)$ has a finite k th derivative for all real t given by

$$\phi^{(k)}(t) = i^k \int_{-\infty}^{\infty} x^k e^{itx} dF(x).$$

In particular,

$$\phi^{(k)}(0) = i^k \mu_k.$$

The existence and finiteness of μ_k is a sufficient condition for the existence and finiteness of $\phi^{(k)}(0)$. It can be shown (see [1]) that when k is even, this condition is also necessary; but when k is odd this is not so. Zygmund [2] has given a necessary and sufficient condition for the existence of $\phi'(0)$ and also one for the existence of a symmetric derivative of higher odd order at $t = 0$; but he imposes a certain condition (smoothness) on the characteristic function. In the following theorem the conditions are on the distribution function only.