## MONOTONE CONVERGENCE OF MOMENTS IN AGE DEPENDENT BRANCHING PROCESSES

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**1.** Introduction. Let Z(t) be the number of cells at time t of a branching process starting at t=0 with one new cell. Let N(t) be the total number of such cells born by time t. Each cell has lifetime distribution function G(t) with G(0)=0. At the end of its life the cell disappears and is replaced by k cells with probability  $p_k$ ,  $k=0,1,2,\cdots$ , where  $p_k \geq 0$  and  $\sum_{k=0}^{\infty} p_k = 1$ . Each cell has lifetime distribution function G(t) and proceeds independently of the state of the system, and identically as the parent cell. Such a process, for this general G, is called an age-dependent branching process and is extensively treated in [1].

Define  $h(s) = \sum_{k=0}^{\infty} p_k s^k$ . For  $h'(1) \equiv m > 1$ , there is an increasing population with probability one. For this case, by use of Smith's key renewal theorem, it is shown ([1], Ch. 6) that for  $\int_0^{\infty} u \, dG(u) < \infty$ , as  $t \to \infty$ ,

(1) 
$$E[Z(t)] \sim K_1 \exp(\alpha t)$$

(2) 
$$E[N(t)] \sim K_2 \exp(\alpha t)$$

where  $K_1$ ,  $K_2$  and  $\alpha$  man be evaluated.

Necessary and sufficient conditions for the monotone convergence of  $E[Z(t) \exp(-\alpha t) \text{ and } E[N(t)] \exp(-\alpha t)$  were given in [2].

It is the purpose of this note to show that for m > 1 that

(3) 
$$M_n(t) \exp(-n\alpha t) \uparrow d_n < \infty$$
, for  $n \ge 2$ ,

where  $M_n(t)$  is the *n*th factorial moment of Z(t). This will be done by exhibiting the monotone nature of the solution of an integral equation in Section 2, a special case of which determines the corresponding monotone behavior in (3), as is given in Section 3.

**2. Integral equation.** A special case of the following theorem was given in [2]. Theorem. Let Q(t), defined on the positive axis, satisfy the following equation

(4) 
$$Q(t) = \int_0^t [(f(t-u) + b(u))Q(t-u) + k(u) + l(t-u)]g(u) du$$

where f, b, k, l are non-negative functions defined on the positive axis with f and l non-decreasing,  $f + b \leq 1$ , and g is a probability density strictly positive on the entire positive axis. Then Q(t) is non-decreasing.

PROOF. Let  $\{X_i\}_1^{\infty}$  be independent and identically distributed random variables on the positive axis, each with density g(u). Let  $S_n = \sum_{i=1}^n X_i$ ,  $n \ge 1$ . Let  $Y_i = k(X_i)$ ,  $Z_i = l(t - S_i)$  and let  $N(t) = \max\{n: S_n \le t\}$ .

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Let

(5) 
$$W(t) = \sum_{i=1}^{N(t)} (Y_i + Z_i) \prod_{j=1}^{i-1} (f(t - S_j) + b(X_j))$$
 and

(6) 
$$Q_0(t) \equiv E[W(t)],$$

where  $\sum_{i=1}^{0} = 0$  and  $\prod_{j=1}^{0} = 1$  by convention

Then one may write, following a suggestion of Prof. R. Pyke,

(7) 
$$W(t) = Y_1 + Z_1 + W_1(t - X_1)[f(t - X_1) + b(X_1)]$$

where W and  $W_1$  are independent and identically distributed. From (7) it is clear that (6) is a solution to (4).

To show uniqueness, let  $h(t) \equiv |Q_0(t) - Q_1(t)|$  be the absolute difference of two solutions to (4). Then clearly  $h(t) \leq \int_0^t h(t-u)g(u) du$ . Suppose  $0 < a_l = \sup_{0 \leq t \leq l} h(t)$ . Then it follows that  $a_l \leq a_l G(l) < a_l$ , a contradiction. Hence (6) is the unique solution to (4) and the result follows.

## 3. Monotonicity of normalized moments.

THEOREM. Let m>1 in a simple age-dependent branching process defined in the introduction. Then, defining  $\alpha>0$  by  $m\int_0^\infty \exp{(-\alpha u)}g(u)\,du=1$ , where g(u)=G'(u) exists a.e., it follows that  $M_n(t)\exp{(-n\alpha t)}$  is monotone increasing to a finite limit as  $t\to\infty$  for all  $n\geq 1$  if in addition,

$$g(u) > \alpha [1 - G(u)][m - 1]^{-1}$$
 for  $u \ge 0$ .

PROOF. The theorem holds for n=1 by the theorem in Section 3 of [2]. Let  $F(x,t)=E[s^{z(t)}]$ . Then ([1], Ch. 6)

$$F(s, t) = s(1 - G(t)) + \int_0^t h(F(s, t - u))g(u) du.$$

By successive differentiation, for  $n \ge 2$ ,  $M_n(t) \exp(-n\alpha t) = M_n^*(t)$  satisfies an equation of the form

$$M_n^*(t) = \int_0^t [a_n M_n^*(t-u) + k_n(t-u)] g_n^*(u) du$$

where  $g_n^*(u)$  is a density (different from g), and where  $0 < a_n = m \int_0^\infty \exp(-\alpha u n) g(u) du < 1$  and  $k_n$  is a positive, increasing function. The theorem of Section 2 establishes the monotonicity.

By induction and use of approximations by writing  $\int_0^t k_n(t-u)g_n^*(u) du = \int_0^{t\gamma} + \int_{t\gamma}^t k_n(t-u)g_n^*(u) du$  for some  $0 < \gamma < 1$ , it is established that  $\lim_{t\to\infty} \int_0^t k_n(t-u)g_n^*(u) du \uparrow c_n < \infty$ , since by [2] and the integral equation for  $M_2(t)$ , it is readily seen to hold for n=1, 2.

That  $M_n(t) \exp(-n\alpha t) \uparrow d_n < \infty$  for  $n \ge 2$  now follows by standard Abelian and Tauberian theorems on p. 182 and p. 192 of [3], respectively, or directly by Lemma 4, pp. 161, 163 of [1].

## 4. Remarks. By differentiation of

$$E[\exp(sZ(t))] \equiv R(s,t) = \exp(s)(1 - G(t)) + \int_0^t h(R(s,t-u))g(u) du$$

a sufficient condition that  $E[Z^n(t)] \exp(-n\alpha t)$  be monotone increasing for  $n \ge 2$  may be obtained similarly, although it will be of more complex form than the simple condition given in Section 3 for the factorial moments. It would be of interest to determine when the condition in Section 3 would suffice for the monotone increase of the normalized moments  $E[Z^n(t)] \exp(-n\alpha t)$ ,  $n \ge 2$ .

## REFERENCES

- [1] HARRIS, T. E. (1963). The Theory of Branching Processes. Springer, Berlin.
- [2] Weiner, H. J. An integral equation in age dependent branching processes. Ann. Math. Stat. (to appear).
- [3] WIDDER, D. V. (1941). The Laplace Transform. Princeton Univ. Press.