

## ON THE DISTRIBUTION OF THE MAXIMUM AND MINIMUM OF RATIOS OF ORDER STATISTICS<sup>1</sup>

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**1. Introduction and summary.** Let  $X_i$  ( $i = 0, 1, \dots, p$ ) be  $(p + 1)$  independent and identically distributed nonnegative random variables each representing the  $j$ th order statistic in a random sample of size  $n$  from a continuous distribution  $G(x)$  of a nonnegative random variable. Let  $G_{j,n}(x)$  be the cumulative distribution function of  $X_i$  ( $i = 0, 1, \dots, p$ ). Consider the ratios  $Y_i = X_i/X_0$  ( $i = 1, 2, \dots, p$ ). The random variables  $Y_i$  ( $i = 1, 2, \dots, p$ ) are correlated and the distribution of the maximum and the minimum is of interest in problems of selection and ranking for restricted families of distribution. The distribution-free subset selection rules using the percentage points of these order statistics are investigated in a companion paper by Barlow and Gupta (1969). In the present paper, we discuss the distribution of these statistics, in general, for any  $G(x)$  and then derive specific results for  $G(x) = 1 - e^{-x/\theta}$ ,  $x > 0$ ,  $\theta > 0$ . Section 2 deals with the distribution of the maximum while Section 3 discusses the distribution of the minimum. Some asymptotic results are given in Section 4, while Section 5 describes the tables of the percentage points of the two statistics.

**2. Distribution of  $Y_{\max}$ .** First we derive the joint distribution of  $Y_i$  ( $i = 1, 2, \dots, p$ ). The joint density function for  $X_0, X_1, \dots, X_p$  is given by

$$(2.1) \quad f(x_0, x_1, \dots, x_p) = [j \binom{n}{j}]^{p+1} \prod_{i=0}^p G^{j-1}(x_i) [1 - G(x_i)]^{n-j} g(x_i)$$

where  $g(x) = dG(x)/dx$ .

Making appropriate transformations the joint density of  $Y_1, Y_2, \dots, Y_p$  can be written as

$$(2.2) \quad f_1(y_1, y_2, \dots, y_p) = [j \binom{n}{j}]^{p+1} \int_0^\infty y_0^p [G(y_0) \prod_1^p G(y_i y_0)]^{j-1} \cdot [(1 - G(y_0)) \prod_1^p (1 - G(y_i y_0))]^{n-j} g(y_0) \prod_1^p g(y_i y_0) dy_0.$$

For  $G(X) = 1 - e^{-x/\theta}$ , (2.2) reduces to

$$(2.3) \quad f_1(y_1, y_2, \dots, y_p) = [j \binom{n}{j}]^{p+1} \int_0^1 (-\log u)^p [(1 - u)(1 - u^{y_1}) \cdots (1 - u^{y_p})]^{j-1} u^{(n-j+1)(1+y_1+\cdots+y_p)-1} du.$$

Received 13 May 1968.

<sup>1</sup> This research was partly supported by the Office of Naval Research contract NONR-1100(26) and the Aerospace Research Laboratories contract AF 33(615)67C1244 at Purdue University and the Office of Naval Research contract NONR-3656(18) at the University of California at Berkeley. Reproduction in whole or in part is permitted for any purposes of the United States Government.

For  $j = 1$ , we obtain

$$(2.4) \quad f_1(y_1, y_2, \dots, y_p) = \Gamma(p + 1)(1 + y_1 + \dots + y_p)^{-(p+1)}, \quad 0 \leq y_1, \dots, y_p < \infty.$$

It should be noted that the distribution (2.4) is independent of  $n$ , as was also pointed out in the companion paper by Barlow and Gupta (1969). Again, (2.4) gives the joint density functions of several correlated  $F$  random variables each with (2, 2) degrees of freedom. In this case,  $Y_{\max}$  and  $Y_{\min}$  are the largest and the smallest of several correlated  $F$  statistics with degrees of freedom (2, 2). The distribution of the largest and the smallest of several correlated  $F$  statistics with different degrees of freedom have been discussed by Gupta (1963a) and Gupta and Sobel (1962) respectively.

The cumulative distribution function of  $Y_{\max}$  can be obtained directly (without using (2.2)) as follows.

$$(2.5) \quad P\{Y_{\max} \leq y\} \equiv H_1(y) = \int_0^{\infty} G_{j,n}^p(yx)g_{j,n}(x) dx$$

where

$$(2.6) \quad g_{j,n}(x) = j \binom{n}{j} G^{j-1}(x)[1 - G(x)]^{n-j}g(x),$$

and

$$(2.7) \quad G_{j,n}(x) \equiv \int_0^x g_{j,n}(t) dt = I_{G(x)}(j, n - j + 1) = \sum_{t=j}^n \binom{n}{t} G^t(x)(1 - G(x))^{n-t}$$

The density of  $Y_{\max}$  is

$$(2.8) \quad h_1(y) = p \int_0^{\infty} x G_{j,n}^{p-1}(yx)g_{j,n}(yx)g_{j,n}(x) dx.$$

By expanding  $G_{j,n}^p(yx)$  in powers of  $1 - G(yx)$ , we can express (2.5) as

$$(2.9) \quad H_1(y) = j \binom{n}{j} \sum_{r=0}^{n-j} \int_0^{\infty} b(r, p; n, j)[1 - G(xy)]^r G^{j-1}(x) \cdot [1 - G(x)]^{n-j}g(x) dx$$

where  $b(r, p; n, j)$  is the coefficient of  $y^r$  in  $[\sum_{t=j}^n \binom{n}{t}(1 - y)^t y^{n-t}]^p$  and is given by the following recursion relations.

$$(2.10) \quad \begin{aligned} b(r, 1; n, j) &= 1, & r &= 0, \\ &= 0, & 1 &\leq r \leq n - j, \\ &= \binom{n}{r} \sum_{t=0}^{n-j} (-1)^{r-t} \binom{r}{t}, & n - j + 1 &\leq r \leq n, \\ &= 0, & n < r < \infty, \\ b(r, p; n, j) &= 1, & r &= 0, \\ &= 0, & 1 &\leq r \leq n - j, \\ &= b(r, p - 1; n, j) \end{aligned}$$

$$\begin{aligned}
 & + \sum_{t=n-j+1}^r b(t, 1; n, j)b(r - t, p - 1; n, j), \\
 & \hspace{15em} n - j + 1 \leq r \leq n, \\
 (2.11) \quad & = b(r, p - 1; n, j) \\
 & + \sum_{t=n-j+1}^n b(t, 1; n, j)b(r - t, p - 1; n, j), \\
 & \hspace{15em} n + 1 \leq r \leq np - n, \\
 & = \sum_{t=\max(n-j+1, r-np+n)}^n b(t, 1; n, j)b(r - t, p - 1; n, j), \\
 & \hspace{15em} np - n + 1 \leq r \leq np, \\
 & = 0, \hspace{15em} np < r < \infty.
 \end{aligned}$$

The density  $h(y)$  can be written in a similar way. For the special case  $G(x) = 1 - e^{-x/\theta}$ , we obtain

$$\begin{aligned}
 (2.12) \quad H_1(y) & = 1 + \sum_{r=n-j+1}^{np} b(r, p; n, j) \\
 & \cdot [(1 + ry/n)(1 + ry/(n - 1)) \cdots (1 + ry/(n - j + 1))]^{-1}.
 \end{aligned}$$

For  $j = 1$ , the coefficients  $b(r, p; n, 1) = (-1)^l \binom{p}{l}$  if  $r = nl, l = 0, 1, 2, \dots, p$  and zero otherwise. It follows that

$$(2.13) \quad H_1(y) = \sum_{l=0}^p (-1)^l \binom{p}{l} (1 + ly)^{-1}$$

which is independent of  $n$  as it should be, and

$$(2.14) \quad h_1(y) = \sum_{l=0}^p (-1)^{l+1} \binom{p}{l} l (1 + ly)^{-2}.$$

Incidentally, one can obtain inequalities on the right hand sides of (2.12) and (2.13) by using the fact that  $H_1(y)$  and  $h_1(y)$  are the cdf and the density function.

**3. Distribution of  $Y_{\min}$ .** The cdf  $H_2(y)$  of  $Y_{\min}$  is given by

$$(3.1) \quad H_2(y) \equiv P\{Y_{\min} \leq y\} = 1 - \int_0^\infty [1 - G_{j,n}(yx)]^p g_{j,n}(x) dx,$$

where  $g_{j,n}(x)$  and  $G_{j,n}(x)$  are given by (2.6) and (2.7).

The density of  $Y_{\min}$  is

$$(3.2) \quad h_2(y) = p \int_0^\infty x [1 - G_{j,n}(yx)]^{p-1} g_{j,n}(yx) g_{j,n}(x) dx.$$

Let  $1 - H_2(y) = F(y)$ . By expanding  $[1 - G_{j,n}(yx)]^p$  in powers of  $1 - G(yx)$ , we can write

$$\begin{aligned}
 (3.3) \quad F(y) & = j \binom{n}{j} \sum_{r=0}^{np} \int_0^\infty b'(r, p; n, j) [1 - G(xy)]^r G^{j-1}(x) \\
 & \quad \cdot [1 - G(x)]^{n-j} g(x) dx
 \end{aligned}$$

where  $b'(r, p; n, j)$  is the coefficient of  $y^r$  in  $[\sum_{t=0}^{j-1} \binom{n}{t} (1 - y)^t y^{n-t}]^p$  and is given by the following recursion relations:

$$\begin{aligned}
 (3.4) \quad b'(r, 1; n, j) &= 0, & 0 \leq r \leq n - j, \\
 &= \sum_{k=0}^{j-1-n+r} (-1)^k \binom{n}{n-r+k} \binom{n-r+k}{k}, \\
 &= 0, & n - j + 1 \leq r \leq n, \\
 & & n < r < \infty;
 \end{aligned}$$

$$\begin{aligned}
 (3.5) \quad b'(r, p; n, j) &= 0, & 0 \leq r \leq (n - j + 1)p - 1, \\
 &= \sum_{m=\max(n-j+1, r-n(p-1))}^{\min(n, r-(p-1)(n-j+1))} b'(m, 1; n, j) b'(r - m, p - 1; n, j), \\
 & & (n - j + 1)p \leq r \leq np \\
 &= 0, & np < r < \infty.
 \end{aligned}$$

The density  $h_2(y)$  can be written similarly. For the special case  $G(x) = 1 - e^{-x/\theta}$ , we obtain

$$\begin{aligned}
 (3.6) \quad F(y) &= \sum_{r=(n-j+1)p}^{np} b'(r, p; n, j) \\
 &\quad \cdot [(1 + ry/n)(1 + ry/(n - 1)) \cdots (1 + ry/(n - j + 1))]^{-1}.
 \end{aligned}$$

For  $j = 1$ , the coefficients  $b(r, p; n, 1) = 1$  if  $r = np$  and zero otherwise. So (3.6) reduces to

$$(3.7) \quad F(y) = (1 + yp)^{-1},$$

which is independent of  $n$ . In this case

$$(3.8) \quad h_2(y) = (1 + yp)^{-2}.$$

**4. Asymtotic results.** Let  $\xi_\alpha$  denote the quantile of order  $\alpha$  of the distribution  $G(x)$ , i.e. the root (assumed unique) of the equation  $G(\xi) = \alpha$ , where  $0 < \alpha < 1$ . We assume that, in some neighbourhood of  $x = \xi_\alpha$ , the density function  $g(x)$  is continuous and has a continuous derivative  $g'(x)$ . In a sample of size  $n$  from the distribution  $G(x)$ , we take the  $j$ th smallest observation such that  $j \leq (n + 1)\alpha < j + 1$ . Then  $X_{j,n}$  is asymptotically normal with mean  $\xi$  and standard deviation  $(g(\xi_\alpha))^{-1}(\alpha\bar{\alpha}/n)^{1/2}$ , where  $\bar{\alpha} = 1 - \alpha$ .

Thus we have, as  $n \rightarrow \infty$  and  $j/n \rightarrow \alpha$

$$(4.1) \quad H_1(y) \approx \int_{-\infty}^{\infty} \Phi^p(xy + (y - 1)\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}}(\alpha\bar{\alpha})^{-\frac{1}{2}}) d\Phi(x)$$

and

$$(4.2) \quad H_2(y) \approx 1 - \int_{-\infty}^{\infty} [1 - \Phi(xy + (y - 1)\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}}(\alpha\bar{\alpha})^{-\frac{1}{2}})]^p d\Phi(x)$$

where

$$\Phi(x) = \int_{-\infty}^x (2\pi)^{-\frac{1}{2}} e^{-t^2/2} dt.$$

(Note:  $a_n \approx b_n$  means  $\lim_{n \rightarrow \infty} a_n/b_n = 1$ .) For  $p = 1$ , we get

$$(4.3) \quad H_1(y) = H_2(y) \approx \int_{-\infty}^{\infty} \Phi(xy + (y - 1)\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}}(\alpha\bar{\alpha})^{-\frac{1}{2}}) d\Phi(x).$$

Using the result

$$\int_{-\infty}^{\infty} \Phi(\alpha x + \beta) d\Phi(x) = \Phi(\beta)(1 + \alpha^2)^{-\frac{1}{2}},$$

this reduces to

$$(4.4) \quad H_1(y) = H_2(y) \approx \Phi((y - 1)\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}}(1 + y^2)^{-\frac{1}{2}}(\alpha\bar{\alpha})^{-\frac{1}{2}}).$$

So if  $H_1(y) = P^*$

$$(4.5) \quad (y - 1)\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}}(1 + y^2)^{-\frac{1}{2}}(\alpha\bar{\alpha})^{-\frac{1}{2}} = \Phi^{-1}(P^*),$$

which can be written as

$$(4.6) \quad y^2(1 - B^2) - 2y + (1 - B^2) = 0,$$

where

$$B = \Phi^{-1}(P^*)(\alpha\bar{\alpha})^{\frac{1}{2}}(\xi_\alpha g(\xi_\alpha)n^{\frac{1}{2}})^{-1}.$$

Obviously, this quadratic equation in  $y$ , has two positive roots which are reciprocals of each other. The appropriate roots can be determined using the fact that  $H_1(y)$  is increasing in  $y$  and  $H_1(1) = \frac{1}{2}$ . So for  $P^* > \frac{1}{2}$  (which will be the case for the selection procedures discussed in the companion paper),  $y > 1$ .

For the special case  $G(x) = 1 - e^{-x/\theta}$ ,  $e^{-\xi_\alpha/\theta} = 1 - \alpha$  and  $g(\xi_\alpha) = \theta^{-1}e^{-\xi_\alpha/\theta} = (1 - \alpha)\theta^{-1}$ . So (4.1) and (4.2) reduce to

$$(4.7) \quad H_1(y) \approx \int_{-\infty}^{\infty} \Phi^p(xy - (n\bar{\alpha}/\alpha)^{\frac{1}{2}}(y - 1) \log \bar{\alpha}) d\Phi(x)$$

and

$$(4.8) \quad H_2(y) \approx 1 - \int_{-\infty}^{\infty} [1 - \Phi(xy - (n\bar{\alpha}/\alpha)^{\frac{1}{2}}(y - 1) \log \bar{\alpha})]^p d\Phi(x).$$

For a general  $p$ , to solve for  $y$  from  $H_1(y) = P^*$  or  $H_2(y) = P^*$  using (4.1) and (4.2), the Table II of Gupta (1963b) can be used with interpolations if necessary.

**5. Description of the tables.** Table 1 provides for the case  $j = 1$  the reciprocals of the percentage points of the distribution of  $Y_{\max}$  corresponding to the probability levels  $\alpha = P^* = .75, .90$  and  $.95$  and the percentage points of the distribution of  $Y_{\min}$  corresponding to the probability levels  $\alpha = 1 - P^* = .05, .10$  and  $.25$  for  $p = 1(1)10$ . We note that when  $j = 1$ , the statistics  $Y_{\max}$  and  $Y_{\min}$  are the maximum and the minimum of several correlated  $F$  statistics with degrees of freedom (2, 2) and hence the entries in Table 1 are the same as those for  $\nu = 2$  in Tables 1 A, B, C of Gupta (1963a) in the case of  $Y_{\max}$  and same as those for  $\nu = 2$  in Tables 3 A, B, C of Gupta and Sobel (1962) in the case of  $Y_{\min}$ , but are given for more places of decimals.

Tables 2 A through 2 E give the reciprocals of the percentage points of the distribution of  $Y_{\max}$  corresponding to the probability levels  $\alpha = P^* = .75, .90, .95$  for  $p = 1$  through 5 respectively. The ranges of  $n$  are: 5(1)15 in Tables 2 A, B, C, 5(1)12 in Table 2D and 5(1)10 in Table 2E.

Tables 3A through 3E present the percentage points of the distribution of  $Y_{\min}$  corresponding to the probability levels  $\alpha = 1 - P^* = .25, .10, .05$  for  $p = 1$  through 5 respectively. The ranges of  $n$ : 5(1)15 in Tables 3A-D & 5(1)13 in Table 3E.

TABLE I

A. Reciprocals of  $100\alpha$  percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_0$  for  $j = 1$  and all  $n$

B.  $100\alpha$  percentage points of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_0$  for  $j = 1$  and all  $n$

$p$	$\alpha = P^*$			$p$	$1 - \alpha = P^*$		
	0.75	0.90	0.95		0.75	0.90	0.95
1	.33333	.11111	.05263	1	.33333	.11111	.05263
2	.20783	.07233	.03469	2	.16667	.05556	.02632
3	.16631	.05871	.02827	3	.11111	.03704	.01754
4	.14472	.05145	.02483	4	.08333	.02778	.01316
5	.13115	.04683	.02263	5	.06667	.02222	.01053
6	.12166	.04357	.02107	6	.05556	.01852	.00877
7	.11456	.04112	.01990	7	.04762	.01587	.00752
8	.10901	.03919	.01897	8	.04167	.01389	.00658
9	.10451	.03762	.01822	9	.03704	.01235	.00585
10	.10078	.03631	.01759	10	.03333	.01111	.00526

For given  $p$  and  $P^*$ , the entries in Tables A and B are respectively the values of  $c$  and  $d$  ( $c$  and  $d$  are independent of  $n$ ) for which

$$\int_0^\infty G_{1,n}^p(x/c) dG_{1,n}(x) = P^* \quad \text{and} \quad \int_0^\infty [1 - G_{1,n}(xd)]^p dG_{1,n}(x) = P^*$$

where  $G_{1,n}(\cdot)$  is the cdf of the smallest order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 2A  
 Reciprocals of the percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_0$  for  $p = 1$

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.48307	.55596	.59583	.60462										
6	.24225	.32197	.37004	.38251										
7	.15560	.22871	.27582	.28958										
8	.48353	.55788	.60192	.62654	.62516									
9	.24266	.32397	.37699	.40851	.40827									
10	.15591	.23045	.28228	.31441	.31548									
11	.48379	.55889	.60473	.63401	.64944	.64125								
12	.24289	.32503	.38021	.41755	.43822	.42890								
13	.15609	.23138	.28527	.32314	.34485	.33649								
14	.48396	.55949	.60626	.63757	.65802	.66737	.65431							
15	.24303	.32566	.38198	.42189	.44902	.46206	.44593							
16	.15620	.23193	.28692	.32735	.35558	.36965	.35399							
17	.48407	.55988	.60720	.63958	.66222	.67685	.68189	.66520						
18	.24313	.32607	.38306	.42434	.45436	.47436	.48175	.46032						
19	.15627	.23229	.28793	.32973	.36089	.38211	.39036	.36890						
20	.48414	.56014	.60782	.64082	.66463	.68159	.69211	.69396	.67447					
21	.24320	.32635	.38377	.42587	.45743	.48057	.49534	.49837	.47271					
22	.15633	.23253	.28860	.33121	.36396	.38843	.40435	.40801	.38182					
23	.48420	.56033	.60824	.64166	.66616	.68436	.69731	.70481	.70421	.68250				
24	.24325	.32654	.38427	.42689	.45939	.48422	.50232	.51307	.51266	.48353				
25	.15637	.23270	.28906	.33220	.36592	.39215	.41157	.42334	.42330	.39316				
26	.48424	.56047	.60855	.64224	.66720	.68614	.70040	.71041	.71558	.71305	.68954			
27	.24329	.32669	.38462	.42761	.46071	.48657	.50648	.52074	.52833	.52512	.49311			
28	.15639	.23283	.28939	.33290	.36724	.39455	.41588	.43136	.43979	.43671	.40323			
29	.48427	.56057	.60878	.64267	.66793	.68736	.70240	.71377	.72153	.72488	.72078	.69578		
30	.24331	.32680	.38489	.42813	.46165	.48817	.50919	.52535	.53659	.54163	.53612	.50166		
31	.15642	.23293	.28964	.33341	.36818	.39619	.41869	.43622	.44855	.45425	.44861	.41227		
32	.48430	.56066	.60896	.64299	.66847	.68823	.70379	.71597	.72513	.73112	.73300	.72762	.70138	
33	.24334	.32689	.38509	.42852	.46234	.48933	.51106	.52839	.54163	.55044	.55338	.54593	.50936	
34	.15643	.23300	.28982	.33380	.36888	.39737	.42064	.43941	.45389	.46365	.46707	.45926	.42044	
35	.48432	.56072	.60909	.64324	.66888	.68888	.70479	.71751	.72751	.73494	.73951	.74018	.73373	.70643
36	.24336	.32696	.38525	.42883	.46287	.49019	.51242	.53052	.54497	.55585	.56266	.56384	.55474	.51636
37	.15645	.23306	.28997	.33409	.36940	.39825	.42205	.44165	.45744	.46944	.47707	.47856	.46887	.42787

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $c$  for which  $\int_0^\infty G_{j,n}^p(x/c) dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 2B  
 Reciprocals of the percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_0$  for  $p = 2$

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.34892	.42508	.46809	.47527										
	.18181	.25464	.29996	.31044										
	.11853	.18353	.22682	.23872										
6	.34945	.42738	.47564	.50261	.49813									
	.18220	.25665	.30712	.33734	.33528									
	.11882	.18521	.23315	.26322	.26304									
7	.34975	.42860	.47913	.51207	.52891	.51631								
	.18242	.25772	.31045	.34682	.36670	.35541								
	.11898	.18610	.23611	.27194	.29238	.28297								
8	.34994	.42934	.48104	.51659	.53993	.54982	.53125							
	.18256	.25835	.31228	.35139	.37819	.39056	.37217							
	.11908	.18663	.23774	.27616	.30024	.31642	.29971							
9	.35006	.42980	.48221	.51915	.54535	.56214	.56697	.54382						
	.18265	.25876	.31340	.35398	.38388	.40378	.41046	.38644						
	.11915	.18697	.23873	.27855	.30865	.32917	.33668	.31405						
10	.35015	.43012	.48298	.52074	.54847	.56833	.58038	.58137	.55461					
	.18272	.25904	.31414	.35559	.38718	.41049	.42519	.42740	.39880					
	.11920	.18720	.23939	.28005	.31178	.33567	.35111	.35408	.32655					
11	.35022	.43035	.48352	.52180	.55045	.57195	.58724	.59572	.59371	.56402				
	.18276	.25924	.31465	.35667	.38927	.41443	.43279	.44345	.44206	.40966				
	.11923	.18737	.23984	.28105	.31377	.33949	.35858	.36998	.36923	.33757				
12	.35026	.43051	.48390	.52254	.55179	.57428	.59131	.60316	.60885	.60442	.57232			
	.18280	.25939	.31502	.35743	.39069	.41697	.43732	.45184	.45926	.45492	.41930			
	.11926	.18749	.24017	.28175	.31512	.34196	.36305	.37834	.38644	.38259	.34740			
13	.35030	.43064	.48419	.52309	.55274	.57588	.59396	.60762	.61679	.62026	.61386	.57973		
	.18283	.25950	.31529	.35798	.39169	.41871	.44028	.45692	.46837	.47314	.46633	.42796		
	.11928	.18758	.24042	.28227	.31608	.34365	.36597	.38341	.39562	.40098	.39450	.35625		
14	.35033	.43074	.48441	.52350	.55344	.57702	.59579	.61056	.62162	.62865	.63030	.62224	.58639	
	.18285	.25958	.31550	.35840	.39244	.41996	.44233	.46026	.47393	.48289	.48544	.47654	.43578	
	.11930	.18766	.24060	.28265	.31679	.34487	.36800	.38675	.40123	.41088	.41394	.40520	.36427	
15	.35035	.43082	.48458	.52382	.55397	.57787	.59711	.61260	.62481	.63379	.63909	.63922	.62977	.59243
	.18287	.25965	.31567	.35872	.39300	.42089	.44381	.46259	.47763	.48889	.49578	.49647	.48575	.44291
	.11931	.18771	.24075	.28295	.31734	.34578	.36947	.38909	.40496	.41700	.42451	.42559	.41489	.37160

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $c$  for which  $\int_0^\infty G_{j,n}^p(x/c) dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.



TABLE 2C  
*Reciprocals of the percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_0$  for  $p = 3$*

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.29806	.37258	.41523	.42096										
	.15749	.22607	.26937	.27858										
	.10324	.16388	.20487	.21565										
6	.29860	.37502	.42329	.45018	.44395									
	.15786	.22808	.27658	.30566	.30259									
	.10352	.16551	.21113	.23985	.23894									
7	.29891	.37631	.42703	.46038	.47708	.46236								
	.15808	.22914	.27994	.31528	.33440	.32214								
	.10367	.16639	.21406	.24854	.26810	.25813								
8	.29911	.37708	.42909	.46528	.48904	.49862	.47757							
	.15821	.22978	.28179	.31994	.34614	.35791	.33850							
	.10377	.16691	.21568	.25276	.27899	.29153	.27432							
9	.29924	.37758	.43035	.46804	.49495	.51207	.51640	.49043						
	.15830	.23018	.28292	.32257	.35198	.37149	.37762	.35248						
	.10384	.16724	.21667	.25515	.28442	.30437	.31137	.28825						
10	.29933	.37792	.43117	.46976	.49835	.51885	.53111	.53141	.50151					
	.15837	.23047	.28366	.32422	.35535	.37840	.39281	.39447	.36463					
	.10388	.16747	.21732	.25665	.28757	.31093	.32595	.32846	.30040					
11	.29939	.37816	.43174	.47091	.50051	.52283	.53865	.54720	.54431	.51121				
	.15841	.23066	.28418	.32532	.35750	.38246	.40067	.41107	.40911	.37753				
	.10392	.16763	.21777	.25765	.28958	.31479	.33353	.34459	.34341	.31115				
12	.29944	.37834	.43216	.47172	.50197	.52538	.54314	.55541	.56103	.55556	.51979			
	.15845	.23081	.28455	.32609	.35896	.38508	.40536	.41978	.42694	.42198	.38486			
	.10394	.16775	.21809	.25835	.29094	.31729	.33806	.35310	.36091	.35662	.32075			
13	.29948	.37847	.43246	.47231	.50301	.52713	.54605	.56034	.56983	.57310	.56550	.52746		
	.15848	.23092	.28483	.32666	.35999	.38688	.40842	.42504	.43642	.44091	.43342	.39342		
	.10396	.16784	.21834	.25887	.29191	.31901	.34103	.35825	.37026	.37536	.36842	.32942		
14	.29951	.37858	.43270	.47276	.50378	.52839	.54807	.56359	.57517	.58241	.58374	.57435	.53438	
	.15850	.23101	.28505	.32709	.36076	.38817	.41054	.42851	.44221	.45108	.45334	.44369	.40119	
	.10398	.16791	.21852	.25925	.29262	.32024	.34309	.36166	.37599	.38547	.38827	.37904	.33729	
15	.29954	.37866	.43289	.47310	.50436	.52932	.54953	.56585	.57871	.58812	.59352	.59323	.58232	.54068
	.15851	.23108	.28521	.32742	.36134	.38913	.41208	.43094	.44606	.45735	.46413	.46449	.45298	.40827
	.10399	.16797	.21867	.25956	.29317	.32116	.34458	.36404	.37980	.39173	.39908	.39990	.38869	.34449

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $c$  for which  $\int_0^\infty G_{j,n}^p(x/c) dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 2D  
 Reciprocals of the percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_n$  for  $p = 4$

n	j										
	2	3	4	5	6	7	8	9	10	11	12
5	.26969	.34243	.38435	.38901							
	.14355	.20924	.25105	.25937							
	.09440	.15215	.19156	.20155							
6	.27023	.34493	.39269	.41919	.41181						
	.14392	.21123	.25827	.28647	.28272						
	.09466	.15377	.19778	.22552	.22411						
7	.27055	.34626	.39657	.42980	.44617	.43014					
	.14414	.21229	.26164	.29617	.31468	.30181					
	.09481	.15463	.20068	.23418	.25309	.24275					
8	.27075	.34705	.39870	.43490	.45867	.46788	.44534				
	.14427	.21293	.26351	.30086	.32656	.33786	.31783				
	.09491	.15514	.20229	.23839	.26399	.27605	.25853				
9	.27088	.34757	.40004	.43779	.46485	.48198	.48587	.45823			
	.14436	.21334	.26464	.30353	.33247	.35164	.35735	.33155			
	.09497	.15547	.20328	.24079	.26943	.28893	.29555	.27213			
10	.27097	.34791	.40086	.43958	.46841	.48910	.50133	.50110	.46937		
	.14442	.21361	.26539	.30519	.33590	.35865	.37280	.37405	.34350		
	.09502	.15570	.20392	.24228	.27259	.29552	.31021	.31240	.28402		
11	.27104	.34816	.40146	.44079	.47068	.49328	.50927	.51774	.51423	.47913	
	.14447	.21381	.26591	.30631	.33808	.36279	.38080	.39097	.38860	.35404	
	.09505	.15586	.20437	.24328	.27461	.29942	.31785	.32864	.32715	.29456	
12	.27109	.34835	.40188	.44163	.47220	.49596	.51399	.52640	.53188	.52572	.48779
	.14450	.21396	.26629	.30709	.33956	.36545	.38559	.39987	.40681	.40142	.36344
	.09508	.15598	.20470	.24399	.27597	.30193	.32243	.33723	.34481	.34022	.30399

For given  $p, n, j$  and  $P^* = .75(\text{top}), .90(\text{middle}), .95(\text{bottom})$ , the entries in this table are the values of  $c$  for which  $\int_0^\infty G_{j,n}^p(x/c) dG_{j,n}(x) = P^*$ , where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 2E  
*Reciprocals of the percentage points of  $Y_{\max} = \max_{1 \leq i \leq p} X_i/X_0$  for  $p = 5$*

<i>n</i>	<i>j</i>									
	2	3	4	5	6	7	8	9	10	
5	.25101	.32219	.36340	.36723						
	.13424	.19776	.23842	.24607						
	.08844	.14410	.18232	.19171						
6	.25156	.32473	.37191	.39800	.38978					
	.13461	.19975	.24565	.27314	.26889					
	.08870	.14570	.18850	.21549	.21370					
7	.25188	.32609	.37587	.40887	.42490	.40796				
	.13482	.20081	.24903	.28288	.30091	.28760				
	.08885	.14655	.19139	.22413	.24253	.23193				
8	.25208	.32690	.37806	.41411	.43776	.44664	.42308			
	.13495	.20144	.25099	.28761	.31287	.32379	.30334			
	.08895	.14706	.19299	.22834	.25342	.26511	.24738			
9	.25221	.32742	.37940	.41707	.44412	.46117	.46469	.43592		
	.13504	.20185	.25204	.29029	.31884	.33770	.34308	.31684		
	.08901	.14739	.19397	.23073	.25888	.27802	.28433	.26072		
10	.25231	.32777	.38027	.41894	.44779	.46852	.48065	.48001	.44704	
	.13510	.20213	.25279	.29197	.32230	.34480	.35870	.35965	.32861	
	.08905	.14762	.19462	.23223	.26205	.28463	.29905	.30097	.27240	

For given  $p, n, j$  and  $P^* = .75$ (top),  $.90$ (middle),  $.95$ (bottom), the entries in this table are the values of  $c$  for which  $\int_0^\infty G_{j,n}^p(x/c) dG_{j,n}(x) = P^*$ , where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 3A  
 Percentage points of the distribution of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_0$  for  $p = 1$

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.48307	.55596	.59583	.60462										
	.24225	.32197	.37004	.38251										
	.15560	.22871	.27582	.28958										
6	.48353	.55788	.60192	.62654	.62516									
	.24266	.32397	.37699	.40851	.40827									
	.15591	.23045	.28228	.31441	.31548									
7	.48379	.55889	.60473	.63401	.64944	.64125								
	.24289	.32503	.38021	.41755	.43822	.42890								
	.15609	.23138	.28527	.32314	.34485	.33648								
8	.48396	.55949	.60626	.63757	.65802	.66737	.65431							
	.24303	.32566	.38198	.42189	.44902	.46206	.44593							
	.15620	.23193	.28692	.32735	.35557	.36964	.35399							
9	.48407	.55988	.60720	.63958	.66222	.67685	.68189	.66520						
	.24313	.32607	.38306	.42434	.45436	.47436	.48175	.46032						
	.15627	.23229	.28793	.32972	.36089	.38211	.39036	.36890						
10	.48414	.56014	.60782	.64082	.66463	.68159	.69211	.69396	.67447					
	.24320	.32635	.38377	.42587	.45743	.48057	.49534	.49837	.47271					
	.15633	.23253	.28859	.33121	.36396	.38843	.40435	.40801	.38182					
11	.48420	.56033	.60824	.64166	.66616	.68436	.69731	.70481	.70421	.68250				
	.24325	.32654	.38427	.42689	.45939	.48422	.50232	.51307	.51266	.48353				
	.15637	.23270	.28905	.33220	.36592	.39215	.41157	.42333	.42330	.39316				
12	.48424	.56047	.60855	.64224	.66720	.68614	.70040	.71041	.71558	.71305	.68954			
	.24329	.32669	.38462	.42761	.46071	.48657	.50648	.52074	.52833	.52512	.49310			
	.15639	.23283	.28939	.33290	.36724	.39455	.41588	.43136	.43980	.43671	.40323			
13	.48427	.56057	.60878	.64267	.66793	.68736	.70240	.71377	.72153	.72488	.72078	.69578		
	.24331	.32680	.38489	.42813	.46165	.48817	.50919	.52535	.53659	.54163	.53612	.50166		
	.15642	.23293	.28964	.33341	.36818	.39619	.41869	.43621	.44855	.45425	.44860	.41226		
14	.48430	.56066	.60896	.64299	.66847	.68823	.70379	.71597	.72513	.73112	.73300	.72762	.70138	
	.24334	.32689	.38509	.42852	.46234	.48933	.51106	.52839	.54163	.55044	.55338	.54593	.50936	
	.15643	.23300	.28983	.33379	.36888	.39737	.42064	.43941	.45389	.46366	.46707	.45926	.42044	
15	.48432	.56072	.60909	.64324	.66888	.68888	.70479	.71751	.72751	.73494	.73951	.74018	.73373	.70643
	.24336	.32696	.38525	.42883	.46287	.49019	.51242	.53052	.54497	.55585	.56266	.56384	.55474	.51636
	.15645	.23306	.28997	.33409	.36940	.39825	.42205	.44165	.45745	.46945	.47707	.47856	.46887	.42787

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $d$  for which  $\int_0^\infty [1 - G_{j,n}(x; d)]^p dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 3B  
 Percentage points of the distribution of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_0$  for  $p = 2$

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.31553	.39895	.44798	.46243										
	.16203	.23711	.28562	.30102										
	.10551	.17100	.21617	.23160										
6	.31594	.40084	.45438	.48616	.48824									
	.16234	.23881	.29187	.32510	.32778									
	.10574	.17244	.22180	.25390	.25725									
7	.31617	.40184	.45733	.49430	.51519	.50871								
	.16251	.23971	.29477	.33353	.35619	.34946								
	.10587	.17320	.22441	.26178	.28424	.27829								
8	.31631	.40243	.45894	.49820	.52479	.53822	.52548							
	.16262	.24024	.29636	.33758	.36652	.38147	.36752							
	.10595	.17365	.22585	.26559	.29416	.30929	.29598							
9	.31641	.40282	.45993	.50040	.52952	.54905	.55708	.53956						
	.16230	.24059	.29734	.33988	.37163	.39346	.40257	.38289						
	.10601	.17395	.22673	.26775	.29910	.32104	.33042	.31115						
10	.31648	.40308	.46058	.50177	.53225	.55450	.56894	.57289	.55161					
	.16275	.24082	.29798	.34131	.37459	.39954	.41600	.42053	.39620					
	.10604	.17415	.22731	.26909	.30196	.32701	.34379	.34857	.32436					
11	.31653	.40326	.46103	.50269	.53398	.55770	.57502	.58563	.58639	.56209				
	.16279	.24099	.29842	.34227	.37647	.40311	.42293	.43522	.43606	.40787				
	.10607	.17429	.22772	.27000	.30378	.33054	.35071	.36337	.36438	.33601				
12	.31656	.40340	.46135	.50333	.53515	.55976	.57864	.59226	.59990	.59811	.57132			
	.16281	.24111	.29875	.34294	.37775	.40542	.42708	.44292	.45187	.44968	.41824			
	.10609	.17439	.22800	.27063	.30501	.33281	.35486	.37116	.38046	.37833	.34639			
13	.31659	.40350	.46160	.50380	.53598	.56117	.58099	.59625	.60701	.61227	.60840	.57953		
	.16283	.24121	.29899	.34343	.37865	.40700	.42977	.44758	.46026	.46647	.46175	.42753		
	.10611	.17447	.22822	.27109	.30588	.33437	.35756	.37588	.38903	.39555	.39075	.35574		
14	.31662	.40358	.46178	.50415	.53659	.56218	.58262	.59888	.61134	.61981	.62314	.61754	.58691	
	.16285	.24128	.29917	.34380	.37932	.40813	.43165	.45065	.46538	.47548	.47941	.47255	.43592	
	.10612	.17453	.22839	.27144	.30653	.33549	.35944	.37899	.39428	.40483	.40900	.40191	.36421	
15	.31663	.40365	.46193	.50442	.53706	.56293	.58381	.60071	.61421	.62444	.63105	.63278	.62572	.59359
	.16287	.24134	.29931	.34409	.37983	.40898	.43300	.45279	.46879	.48103	.48898	.49100	.48229	.44356
	.10613	.17459	.22852	.27171	.30702	.33633	.36080	.38117	.39778	.41057	.41893	.42110	.41202	.37194

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $d$  for which  $\int_0^\infty [1 - G_{j,n}(x d)]^p dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 3C  
 Percentage points of the distribution of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_n$  for  $p = 3$

n	i													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.24830	.33214	.38352	.40058										
	.12895	.19983	.24751	.26412										
	.08451	.14516	.18868	.20473										
6	.24865	.33390	.38968	.42388	.42810									
	.12920	.20134	.25329	.28677	.29094									
	.08470	.14643	.19383	.22551	.22998									
7	.24885	.33483	.39253	.43191	.45493	.45008								
	.12935	.20214	.25597	.29472	.31803	.31284								
	.08481	.14710	.19623	.23287	.25546	.25084								
8	.24898	.33538	.39409	.43576	.46455	.47978	.46818							
	.12944	.20262	.25744	.29855	.32792	.34367	.33118							
	.08488	.14750	.19755	.23644	.26486	.28038	.26847							
9	.24906	.33574	.39505	.43794	.46930	.49075	.50024	.48344						
	.12950	.20293	.25834	.30072	.33283	.35527	.36519	.34685						
	.08492	.14776	.19835	.23845	.26955	.29163	.30154	.28365						
10	.24912	.33598	.39568	.43929	.47204	.49628	.51236	.51747	.49655					
	.12955	.20314	.25894	.30208	.33567	.36117	.37830	.38359	.36046					
	.08495	.14794	.19888	.23971	.27226	.29736	.31443	.31979	.29690					
11	.24916	.33616	.39611	.44020	.47378	.49954	.51859	.53059	.53225	.50798				
	.12958	.20329	.25935	.30298	.33748	.36464	.38508	.39802	.39957	.37244				
	.08498	.14806	.19925	.24056	.27399	.30075	.32113	.33415	.33574	.50798				
12	.24920	.33629	.39642	.44084	.47496	.50163	.52231	.53743	.54622	.54511	.51807			
	.12960	.20340	.25965	.30362	.33871	.36688	.38913	.40561	.41517	.41362	.38309			
	.08499	.14816	.19952	.24115	.27516	.30293	.32514	.34173	.35143	.34985	.31910			
13	.24922	.33638	.39666	.44130	.47580	.50307	.52473	.54156	.55361	.55984	.55644	.52707		
	.12962	.20348	.25987	.30409	.33958	.36841	.39178	.41020	.42348	.43027	.42610	.39266		
	.08501	.14823	.19972	.24159	.27600	.30443	.32766	.34633	.35981	.36672	.36245	.32853		
14	.24924	.33646	.39684	.44165	.47642	.50410	.52641	.54428	.55812	.56771	.57182	.56652	.53516	
	.12963	.20355	.26004	.30444	.34022	.36952	.39362	.41323	.42856	.43922	.44368	.43730	.40132	
	.08502	.14828	.19987	.24191	.27661	.30551	.32958	.34936	.36495	.37584	.38040	.37380	.33711	
15	.24926	.33652	.39698	.44193	.47689	.50487	.52762	.54619	.56112	.57255	.58013	.58249	.57556	.54250
	.12964	.20360	.26017	.30471	.34071	.37034	.39495	.41535	.43195	.44476	.45323	.45572	.44741	.40921
	.08503	.14833	.19999	.24217	.27708	.30631	.33090	.35149	.36838	.38149	.39019	.39272	.38409	.34494

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $d$  for which  $\int_0^\infty [1 - G_{j,n}(x, d)]^p dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 3D  
 Percentage points of the distribution of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_0$  for  $p = 4$

n	j													
	2	3	4	5	6	7	8	9	10	11	12	13	14	15
5	.21024	.29285	.34497	.36357										
	.10994	.17752	.22430	.24157										
	.07234	.12953	.17176	.18812										
6	.21055	.29650	.35081	.38624	.39188									
	.11017	.17891	.22974	.26316	.26830									
	.07251	.13060	.17658	.20783	.21301									
7	.21073	.29537	.35364	.39407	.41825	.41462								
	.11029	.17965	.23226	.27076	.29436	.29023								
	.07260	.13131	.17883	.21483	.23741	.23368								
8	.21084	.29589	.35514	.39784	.42774	.44401	.43341							
	.11037	.18009	.23365	.27442	.30391	.32010	.30866							
	.07266	.13168	.18007	.21822	.24644	.26216	.25121							
9	.21092	.29622	.35606	.39996	.43243	.45491	.46532	.44930						
	.11043	.18037	.23450	.27650	.30865	.33138	.34179	.32446						
	.07270	.13192	.18082	.22014	.25094	.27303	.28325	.26635						
10	.21097	.29645	.35666	.40129	.43514	.46043	.47744	.48332	.46299					
	.11046	.18056	.23506	.27779	.31140	.33711	.35460	.36040	.33821					
	.07273	.13208	.18132	.22134	.25355	.27858	.29579	.30151	.27960					
11	.21101	.29661	.35708	.40218	.43686	.46367	.48368	.49650	.49881	.47494				
	.11049	.18070	.23545	.27866	.31315	.34050	.36124	.37457	.37660	.35034				
	.07275	.13219	.18167	.22214	.25521	.28186	.30231	.31554	.31751	.29134				
12	.21104	.29673	.35738	.40280	.43802	.46576	.48741	.50339	.51290	.51231	.48551			
	.11051	.18080	.23573	.27927	.31433	.34268	.36522	.38203	.39198	.39087	.36114			
	.07276	.13228	.18192	.22271	.25634	.28398	.30622	.32295	.33288	.33169	.30184			
13	.21106	.29682	.35761	.40326	.43885	.46719	.48984	.50756	.52037	.52721	.52422	.49495		
	.11053	.18088	.23594	.27972	.31518	.34418	.36782	.38655	.40018	.40733	.40358	.37085		
	.07277	.13234	.18210	.22312	.25714	.28543	.30877	.32745	.34111	.34827	.34438	.31132		
14	.21108	.29689	.35778	.40360	.43947	.46822	.49153	.51031	.52494	.53520	.53984	.53484	.50345	
	.11054	.18094	.23610	.28005	.31580	.34526	.36962	.38954	.40521	.41621	.42101	.41499	.37966	
	.07278	.13239	.18224	.22343	.25773	.28648	.31055	.33042	.34616	.35725	.36206	.35582	.31994	
15	.21109	.29695	.35792	.40387	.43993	.46899	.49275	.51223	.52798	.54012	.54829	.55109	.54438	.51116
	.11055	.18099	.23623	.28031	.31627	.34606	.37092	.39163	.40856	.42170	.43050	.43329	.42532	.38769
	.07279	.13243	.18236	.22368	.25818	.28726	.31184	.33250	.34953	.36281	.37172	.37449	.36621	.32783

For given  $p, n, j$  and  $P^* = .75$  (top),  $.90$  (middle),  $.95$  (bottom), the entries in this table are the values of  $d$  for which  $\int_0^\infty [1 - G_{j,n}(x d)]^p dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.

TABLE 3E  
 Percentage points of the distribution of  $Y_{\min} = \min_{1 \leq i \leq p} X_i/X_0$  for  $p = 5$

n	j										
	2	3	4	5	6	7	8	9	10	11	12
5	.18512	.26616	.31847	.33808							
	.09728	.16219	.20805	.22583							
	.06419	.11872	.15989	.17643							
6	.18541	.26772	.32418	.36014	.36682						
	.09748	.16350	.21333	.24659	.25241						
	.06433	.11981	.16447	.19533	.20101						
7	.18557	.26855	.32682	.36779	.39268	.39000					
	.09760	.16419	.21573	.25391	.27766	.27431					
	.06442	.12038	.16660	.20205	.22457	.22150					
8	.18567	.26904	.32827	.37146	.40201	.41899	.40921				
	.09767	.16460	.21706	.25744	.28693	.30340	.29278				
	.06447	.12073	.16778	.20531	.23331	.24915	.23894				
9	.18574	.26935	.32916	.37354	.40663	.42977	.44082	.42550			
	.09772	.16486	.21787	.25944	.29154	.31441	.32517	.30864			
	.06451	.12095	.16850	.20716	.23767	.25974	.27017	.25402			
10	.18579	.26957	.32974	.37484	.40930	.43524	.45286	.45931	.43955		
	.09775	.16504	.21840	.26070	.29421	.32002	.33773	.34389	.32246		
	.06453	.12110	.16897	.20831	.24020	.26514	.28242	.28840	.26725		
11	.18582	.26972	.33015	.37570	.41100	.43845	.45908	.47245	.47525	.45184	
	.09777	.16517	.21877	.26153	.29591	.32333	.34425	.35783	.36022	.33468	
	.06455	.12120	.16930	.20908	.24181	.26833	.28880	.30215	.30442	.27899	
12	.18585	.26984	.33044	.37631	.41215	.44053	.46279	.47934	.48935	.48917	.46272
	.09779	.16527	.21904	.26212	.29706	.32546	.34815	.36518	.37539	.37463	.34557
	.06456	.12128	.16954	.20963	.24290	.27040	.29262	.30943	.31952	.31863	.28951

For given  $p, n, j$  and  $P^* = .75(\text{top}), .90(\text{middle}), .95(\text{bottom})$ , the entries in this table are the values of  $d$  for which  $\int_0^\infty [1 - G_{j,n}(x d)]^p dG_{j,n}(x) = P^*$  where  $G_{j,n}(\cdot)$  is the cdf of the  $j$ th order statistic in a sample of size  $n$  from the exponential distribution.



In all these tables the probability levels are chosen such that  $P^*$ , the infimum of the probability of correct selection in the companion paper by Barlow and Gupta is .75, .90 and .95 and the entries are either the percentage points or the reciprocals of the percentage points so that they will be the values of the constants  $d$  or  $c$  ( $0 < c, d < 1$ ) to be used in the selection procedures discussed in the companion paper.

The authors wish to thank Mrs. Louise Lui of the Statistical Consulting Section, Purdue University, for programming and carrying out the computations of the tables.

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