## RANDOM TIME TRANSFORMATIONS OF SEMI-MARKOV PROCESSES

## By RICHARD F. SERFOZO

Syracuse University

1. Introduction. Let  $\{X_t:t\geq 0\}$  be a right continuous stochastic process taking values in some space  $(E,\mathscr{E})$  and adapted to a right continuous family of  $\sigma$ -algebras  $\{\mathscr{F}_t:t\geq 0\}$ , and let  $\{\tau_t:t\geq 0\}$  be a nonnegative real-valued right continuous stochastic process such that each  $\tau_t$  is an  $\mathscr{F}_t$ -stopping time. Then the process  $X_t'=X(\tau_t)$   $t\geq 0$ , which is adapted to  $\{\mathscr{F}_{\tau_t}\}$ , is called the random time transformation (RTT) of  $\{X_t\}$  determined by the change of time  $\{\tau_t\}$ .

Two classical RTT's that transform Markov processes into Markov processes are found in Dynkin [3] (or in Blumenthal and Getoor [2]) and in Feller [4]. For the RTT in Dynkin,  $\{\tau_t\}$  is the inverse mapping of a strictly increasing continuous additive functional of  $\{X_t\}$ . This RTT is used in constructing generalized Brownian motion from a Wiener process. Feller discusses RTT's under the heading of subordination of processes. He shows that if  $\{X_t\}$  is a Markov process with continuous transition probabilities and  $\{\tau_t\}$  has stationary, independent, nonnegative increments and is independent of  $\{X_t\}$ , then  $\{X_t'\}$  is a Markov process. This type of RTT was used by Bochner to construct symmetric, stable processes from a Wiener process. In addition to their theoretical significance, RTT's are useful in applications where  $\{X_u: u \ge 0\}$  describes some phenomena as a function of some parameter u, which increases in time according to a process  $\{\tau_t\}$ , and one is interested in  $\{X_t'\}$  which depicts the phenomena as a function of time.

In this paper we present several RTT's of semi-Markov step processes (SMP's) (Definition 2). In Theorem 1, our major result, we identify some general conditions on a time process  $\{\tau_t\}$  under which  $\{X_t'\}$  is an SMP whenever  $\{X_t\}$  is an SMP. We use this result in Section 4 to identify four types of RTT's. Two of these RTT's are analogous to those discussed by Dynkin and Feller for Markov processes, and a special case of another is similar to the RTT presented by Yackel [15]. Special cases of the RTT's presented transform Markov chains or regular step Markov processes or SMP's into any one of these three classes of processes. We conclude in Section 5 by presenting two other RTT's for special SMP's  $\{X_t\}$  when  $\{\tau_t\}$  is a step process independent of  $\{X_t\}$ .

**2. Preliminaries.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space: all stochastic processes introduced herein will be defined on this space. Let E be a locally compact Hausdorff space with a countable base, and set  $\mathscr{E} = \mathscr{B}(E)$ , the smallest  $\sigma$ -algebra containing the Borel sets of E. Let  $R_+ = [0, \infty)$  and set  $\mathscr{B}_+ = \mathscr{B}(R_+)$ .

Let Q(x, G) be a real-valued function defined for each  $x \in E$  and  $G \in \mathscr{E} \times \mathscr{B}_+$  with the properties:

Received September 15, 1969; revised July 24, 1970.

176

- (i)  $G \to Q(x, G)$  is a probability measure on  $\mathscr{E} \times \mathscr{B}_+$  for each  $x \in E$ , and
- (ii)  $x \to Q(x, G)$  is  $\mathscr{E}$ -measurable for each  $G \in \mathscr{E} \times \mathscr{B}_+$ .

For each  $n \ge 0$ ,  $x \in E$  and  $G \in \mathscr{E} \times \mathscr{B}_+$  we define  $Q^n(x, G)$  by setting  $Q^0(x, G) = 1$  or 0 accordingly as  $\{x\} \times \emptyset$  is or is not in G, and for  $n \ge 0$  we set

$$Q^{n+1}(x,G) = \int_0^\infty Q(x,dy)Q^n(y,G).$$

Note that Q is uniquely determined by the values it takes on the cylinder sets of  $\mathscr{E} \times \mathscr{B}_+$ . That is, if (i) and (ii) hold for cylinder sets of  $\mathscr{E} \times \mathscr{B}_+$ , then for each  $x \in E$  we let  $G \to \hat{Q}(x, G)$  be the unique extension, of the measure  $G \to Q(x, G)$ , to the  $\sigma$ -algebra  $\mathscr{E} \times \mathscr{B}_+$ . The mapping  $x \to \hat{Q}(x, G)$  turns out to be  $\mathscr{E}$ -measurable for each  $G \in \mathscr{E} \times \mathscr{B}_+$ . This follows since  $(E \times R_+, \mathscr{E} \times \mathscr{B}_+)$  being a locally compact Hausdorff space with a countable base implies that  $G \to \hat{Q}(x, G)$  is a regular probability measure (see [8]), and so for each  $G \in \mathscr{E} \times \mathscr{B}_+$  we can write

$$\widehat{Q}(x,G) = \lim_{n \to \infty} \sum_{k=1}^{\infty} Q(x, A_{nk} \times B_{nk})$$

for an appropriate sequence of cylinder sets  $\{A_{nk} \times B_{nk}\}$  of  $\mathscr{E} \times \mathscr{B}_+$ .

DEFINITION 1. Let  $\{\mathscr{G}_n : n \geq 0\}$  be a non-decreasing family of sub- $\sigma$ -algebras of  $\mathscr{F}$  (which are complete with respect to P) and let Q be as above. Let  $\{Y_n, W_n : n \geq 0\}$  be a Markov chain (on  $(\Omega, \mathscr{F}, P)$ ) with respect to  $\{\mathscr{G}_n\}$  taking values in  $(E \times R_+, \mathscr{E} \times \mathscr{B}_+)$  and with kernel Q. That is:

- (i) Each  $(Y_n, W_n)$  is  $\mathcal{G}_n$ -measurable.
- (ii) For each n and  $G \in \mathscr{E} \times \mathscr{B}_+$

(2.1) 
$$P[(Y_{n+1}, W_{n+1}) \in G \mid \mathcal{G}_n] = Q(Y_n, G)$$
 a.s.

We call  $\{Y_n, W_n\}$  a Markov renewal process (MRP) with respect to  $\{\mathcal{G}_n\}$  with kernel O.

The name MRP is motivated by properties (a) and (b) below. Several authors also use this name for the processes  $\{N_t\}$  and  $\{X_t\}$  in Definition 2. Note that without loss of generality we can assume that O is of the form

(2.2) 
$$Q(x, A \times [0, t]) = \int_A K(x, dy) F_{xy}(t),$$

where K is a Markov kernel on  $(E, \mathscr{E})$  and  $\{F_{xy}(\cdot): x, y \in E\}$  is a family of distribution functions such that  $(x, y) \to F_{xy}(t)$  is  $\mathscr{E} \times \mathscr{E}$ -measurable for each t. This follows since

$$Q(x, A \times [0, t]) = P[Y_1 \in A, W_1 \le t \mid Y_0 = x]$$
  
=  $\int_A P[Y_1 \in dy \mid Y_0 = x] P[W_1 \le t \mid Y_0 = x, Y_1 = y],$ 

and we can take K(x, A) and  $F_{xy}(t)$  to be the regular conditional distribution functions  $P[Y_1 \in A \mid Y_0 = x]$  and  $P[W_1 \le t \mid Y_0 = x, Y_1 = y]$  (see [8]) respectively.

An MRP  $\{Y_n, W_n\}$  with respect to  $\{\mathscr{G}_n\}$  with kernel Q has the following properties.

- (a)  $\{Y_n\}$  is a Markov chain with respect to  $\{\mathscr{G}_n\}$  with kernel K.
- (b) For any  $n_1, \dots, n_k$  and  $t_1, \dots, t_k$

$$P[W_{n_1} \le t_1, \dots, W_{n_k} \le t_k \mid Y_n : n \ge 0] = \prod_{i=1}^k F_{Y_{n_i-1}Y_{n_i}}(t_i).$$

That is, the  $\{W_n\}$  are conditionally independent given the  $\sigma$ -algebra  $\sigma(Y_n: n \ge 0)$ .

(c) Set  $T_0 = 0$  a.s. and  $T_n = \sum_{k=1}^n W_k$ . Then  $\{Y_n, T_n\}$  is a Markov chain with respect to  $\{\mathscr{G}_n\}$  and for any  $n \ge 1$ , and  $G \in \mathscr{E} \times \mathscr{B}_+$ ,

$$P[(Y_n, T_n) \in G \mid \mathscr{G}_0] = Q^n(Y_0, G)$$
 a.s.

(d) Set  $\zeta = \sup_{n} T_{n}$ . Then a necessary and sufficient condition for

$$P[\zeta = +\infty | X_0 = x] = 1$$
 for each  $x \in E$ ,

is that for each  $\lambda > 0$  and  $x \in E$ 

(2.3) 
$$\int_0^\infty e^{-\lambda t} Q^n(x, E \times dt) \to 0 \quad \text{as } n \to \infty.$$

Properties (a)–(c) follow directly from (2.1) and (2.2). Property (d) follows since, by (c), the integral in (2.3) represents  $E[\exp(-\lambda T_n) \mid X_0 = x]$ . Another useful property of MRP's is as follows.

PROPOSITION 1. Let  $\{Y_n, W_n\}$  be an MRP with respect to  $\{\mathcal{G}_n\}$  with kernel Q which satisfies (2.3). Set  $v_0 = 0$  and for each  $n \ge 0$  let  $v_{n+1} = \inf\{m > v_n \colon W_m > 0\}$ . Set  $\widetilde{Y}_n = Y_{v_n}$ ,  $\widetilde{W}_n = W_{v_n}$  and  $\widetilde{\mathcal{G}}_n = \mathcal{G}_{v_n}$ . Then  $\{\widetilde{Y}_n, \widetilde{W}_n\}$  is an MRP with respect to  $\{\widetilde{\mathcal{G}}_n\}$  with kernel  $\widetilde{Q}$  satisfying

$$\widetilde{Q}(x, A \times [0, t]) = \sum_{k=0}^{\infty} \int_0^{\infty} Q^k(x, dy \times \{0\}) Q(y, A \times (0, t]),$$

for each  $x \in E$ ,  $A \in \mathcal{E}$  and  $t \ge 0$ .

PROOF. The variables  $\{v_n\}$  are well defined and finite valued due to condition (2.3). We see by induction that each  $v_n$  is a  $\mathcal{G}_n$ -stopping time. Obviously  $v_0 = 0$  is a  $\mathcal{G}_n$ -stopping time and assuming that  $v_0, \dots, v_n$  also are, we have

$$\{v_{n+1} > m\} = \bigcup_{k=1}^{m} \{v_n = k, T_m = T_k\} \cup \{v_n > m\} \in \mathcal{G}_m,$$

and so  $v_{n+1}$  is a  $\mathcal{G}_n$ -stopping time.

By the strong Markov property for  $\{Y_n, W_n\}$ , for each n, t and  $A \in \mathscr{E}$ 

$$\begin{split} P\big[\widetilde{Y}_{n+1} \in A, \, \widehat{W}_{n+1} & \leq t \, \big| \, \widetilde{\mathscr{G}}_n \big] \\ &= \sum_{k=0}^{\infty} \int_0^{\infty} P\big[Y_{\nu_n+k+1} \in A, \, W_{\nu_n+k+1} \in (0,t], \, Y_{\nu_n+k} \in dy, \, T_{\nu_n+k} = T_{\nu_n} \, \big| \, Y_{\nu_n} \big] \\ &= \sum_{k=0}^{\infty} \int_0^{\infty} Q^k(\widetilde{Y}_n, dy \times \{0\}) Q(y, A \times (0,t]) \\ &= \widetilde{Q}(\widetilde{Y}_n, A \times [0,t]) \quad \text{a.s.} \end{split}$$

This proves the assertion.  $\square$ 

Definition 2. Let  $\{Y_n, W_n : n \ge 0\}$  be a stochastic process (on  $(\Omega, \mathcal{F}, P)$ ), taking values in  $(E \times R_+, \mathcal{E} \times \mathcal{B}_+)$  such that

(2.4) 
$$W_0 = 0 \text{ and } W_n > 0$$
 for  $n \ge 1$ .

For each  $n \ge 0$  set  $T_n = \sum_{k=0}^n W_k$ , and let  $\zeta = \sup_n T_n$ . For each  $0 \le t < \zeta$  let  $N_t = n$  if  $T_n \le t < T_{n+1}$  and  $X_t = Y_n$  if  $T_n \le t < T_{n+1}$ . Let  $\{\mathscr{F}_t : t \in R_+\}$  be a non-decreasing right continuous family of sub- $\sigma$ -algebras of  $\mathscr{F}$  (which are complete with respect to P). Assume that  $\{X_t, N_t\}$  is adapted to  $\{\mathscr{F}_t\}$ , i.e., for each n, t and  $A \in \mathscr{E}$ 

$$\{X_t \in A, N_t = n\} \in \mathscr{F}_t.$$

If  $\{Y_n, W_n\}$  is an MRP with respect to  $\mathscr{G}_n = \mathscr{F}_{T_n}$ ,  $n \ge 0$ , then we call  $\{X_t, \zeta\}$  a semi-Markov step process (SMP) with respect to  $\{\mathscr{F}_t\}$ . We say that  $\{X_t, \zeta\}$  is induced by the MRP  $\{Y_n, W_n\}$ . The process  $\{N_t, \zeta\}$  is called the counting process associated with  $\{Y_n, W_n\}$ , or with  $\{X_t, \zeta\}$ .

Calling  $\{N_t, \zeta\}$  and  $\{X_t, \zeta\}$  stochastic processes (on  $(\Omega, \mathcal{F}, P)$ ) is justified for

$$\{N_t = n, t < \zeta\} = \{T_n \le t < T_{n+1}\} \in \mathscr{F}$$

and

$$\{X_t \in A, t < \zeta\} = \bigcup_n \{T_n \le t < T_{n+1}, Y_n \in A\} \in \mathscr{F}.$$

(Herein set equality is taken to mean equality up to a set of P-measure zero.) The assumption that  $\{N_t\}$  is adapted to  $\{\mathscr{F}_t\}$  guarantees that each  $T_n$  is an  $\mathscr{F}_t$ -stopping time, since  $\{T_n \leq t\} = \{N_t > n, \, t < \zeta\} \in \mathscr{F}_t$  for each  $t \geq 0$ . This in turn justifies the definition of the  $\mathscr{G}_n$ . Without this assumption, the  $T_n$  may not be  $\mathscr{F}_t$ -stopping times.

The SMP  $\{X_t,\zeta\}$  is a continuous time step process, which proceeds through its state space  $(E,\mathscr{E})$  according to the Markov chain  $\{Y_n\}$ , at a speed which is given by the  $\{T_n\}$ . It makes  $N_t$  jumps in the interval (0,t], the nth jump occurs at time  $T_n$  and takes the process to state  $Y_n$  where it remains for a time  $W_{n+1} > 0$  a.s. Note that some of these jumps may not be actual jumps in that we may have  $P[Y_n = Y_{n+1}] > 0$ . (For this reason an SMP may be induced by more than one MRP.) The lifetime of the SMP is given by the random variable  $\zeta$ : note that  $\zeta = +\infty$  a.s. if (2.3) holds. Our results can be formulated in an obvious manner (see [3]) for SMP's with random lifetimes, but for clarity of exposition we will assume  $\zeta = +\infty$  a.s. The basic properties of SMP's on countable state spaces appear in Lévy [6], Smith [14], Pyke [9], [10], Pyke and Schaufele [11], [12] and Feller [5].

REMARK 1. An SMP induced by an MRP with kernel  $Q(x, A \times [0, t]) = K(x, A)H_1(t)$ , where  $H_1$  is the degenerate distribution with unit mass at 1, is a Markov chain with kernel K. An SMP induced by an MRP with kernel

$$Q(x, A \times [0, t]) = K(x, A)[1 - e^{-\lambda(x)t}],$$

where K satisfies  $K(x, \{x\}) = 0$  for each  $x \in E$ , and  $\lambda$  is a positive real valued  $\mathscr{E}$ -measurable function, is a regular step Markov process [1]).

REMARK 2. For each  $t \ge 0$  let  $X_t^+ = Y_{N_t+1}$ ,  $U_t = T_{N_t}$  and  $U_t^+ = T_{N_t+1}$ . Then it can be shown that the processes  $\{X_t, U_t\}$  and  $\{X_t, X_t^+, U_t, U_t^+\}$  are regular step

Markov processes. They in general will not be Markov processes with respect to  $\{\mathcal{F}_t\}$ .

For our major result we require the following.

LEMMA 1. Let f be a function from  $R_+$  into  $R_+$  which is non-decreasing, right continuous and f(o) = 0 and  $f(t) \uparrow \infty$  as  $t \to \infty$ . For each  $t \in R_+$  let  $g(t) = \inf \{s: f(s) > t\}$ . Then the following are satisfied:

- (a) The function g is non-decreasing, right continuous and for each  $s \in R_+$   $f(s) = \inf\{t: g(t) > s\}$ .
  - (b) If f is continuous at s, then f(s) > t is equivalent to g(t) < s for any  $t \in R_+$ .

PROOF. Part (a) is proved on page 108 of Meyer [7]. For part (b) first assume f(s) > t. Under the hypothesis, there exists a  $\delta > 0$  such that  $f(s) > f(s - \delta) > t$ , and so

$$g(t) = \inf\{u : f(u) > t\} \le s - \delta < s.$$

Conversely, if g(t) < s, then since g is right continuous (by part (a)) there exists a  $\delta > 0$  such that  $g(t) \le g(t+\delta) < s$ , and so from part (a),

$$f(s) = \inf \{u : g(u) > s\} \ge t + \delta > t.$$

This completes the proof of (b).

3. The main result. For the remainder of this paper we take  $\{X_t\}$  to be an SMP with respect to  $\{\mathscr{F}_t\}$  (as in Definition 2), which is induced by the MRP  $\{Y_n, W_n\}$  having kernel Q satisfying (2.3).

Let  $\{\varphi_t: t \in R_+\}$  be a stochastic process on  $(\Omega, \mathcal{F}, P)$ , taking values in  $(R_+, \mathcal{B}_+)$ , adapted to  $\{\mathcal{F}_t\}$  and with right continuous, non-decreasing paths, which for convenience, satisfy  $\varphi_0 = 0$  a.s. and  $\varphi_t \uparrow \infty$  a.s. as  $t \uparrow \infty$ . Assume the following:

- (A1) For a.a.  $\omega \in \Omega$ , the mapping  $s \to \varphi_s(\omega)$  is continuous at the points  $T_n(\omega)$ ,  $n \ge 0$  (this is obviously satisfied if  $\{\varphi_t\}$  has continuous paths).
- (A2) For any n let  $\tilde{\varphi}_t = \varphi(T_n + t) \varphi(T_n)$ . Then  $\sigma(\tilde{\varphi}_t : t \ge 0)$  is conditionally independent of  $\mathscr{G}_n = \mathscr{F}_{T_n}$  given  $\sigma(Y_n, Y_{n+1}, W_{n+1})$ ; and for any t, u, w and  $x, y \in E$

$$P[\tilde{\varphi}_t \le u \mid Y_n = x, Y_{n+1} = y, W_{n+1} = w] = P[\varphi_t \le u \mid Y_0 = x, Y_1 = y, W_1 = w].$$

For each  $t \ge 0$  let  $\tau_t = \inf\{s : \varphi_s > t\}$ . We call  $\{\tau_t\}$  the *inverse* of  $\{\varphi_t\}$ . For each  $t \ge 0$  let  $X_t' = X(\tau_t)$  and  $\mathscr{F}_t' = \mathscr{F}_{\tau_t}$ . Our main result is as follows.

THEOREM 1. The process  $\{X_t'\}$  is an SMP with respect to  $\{\mathscr{F}_t'\}$ .

PROOF. Set  $v_0 = 0$  and for  $n \ge 0$  let

$$v_{n+1} = \inf\{m > v_n : \varphi(T_m) > \varphi(T_{v_n})\}.$$

We see by induction that each  $v_n$  is a  $\mathscr{G}_n$ -stopping time. Obviously  $v_0 = 0$  is a  $\mathscr{G}_n$ -stopping time. Assume that  $v_0, \dots, v_n$  are  $\mathscr{G}_n$ -stopping times. Since  $\{\varphi_t\}$  is

adapted to  $\{\mathscr{F}_t\}$  and  $T_n$  is a  $\mathscr{F}_t$ -stopping time  $\{\varphi(T_n) \leq t\} \in \mathscr{G}_n$  for each t and n, and so for each m

$$\{v_{n+1} > m\} = \bigcup_{k=1}^{m} \{v_n = k, \varphi(T_m) = \varphi(T_k)\} \cup \{v_n > m\} \in \mathcal{G}_m$$

that is,  $v_{n+1}$  is a  $\mathscr{G}_n$ -stopping time. This completes the induction argument. For each  $n \ge 0$  let

$$(3.1) Y_n' = Y_{\nu_n}, T_n' = T_{\nu_n} \text{and} \mathscr{G}_n' = \mathscr{G}_{\nu_n},$$

and set

(3.2) 
$$W_0' = 0$$
, and  $W_n' = T_n' - T_{n-1}'$  for  $n \ge 1$ .

We will show that  $\{Y_n', W_n'\}$  is an MRP which induces  $\{Y_t'\}$ .

Since  $\{\mathcal{F}_t\}$  is right continuous, for any s, t

$$\{\tau_t < s\} = \bigcup_n \{\varphi_{s-n^{-1}} > t\} \in \mathscr{F}_s,$$

and so each  $\tau_t$  is a  $\mathscr{F}_t$ -stopping time. Moreover by Lemma 1(a),  $\{\tau_t\}$  is right continuous. It follows that  $\{X_t'\}$  is right continuous; and it is adapted to  $\{\mathscr{F}_t'\}$ , since for any t, u and  $A \in \mathscr{E}$ ,

$$\{X_t' \in A\} \cap \{\tau_t \leq u\} = \bigcup_{r \leq u: r \text{ rational}} \{\tau_t = r, X_r \in A\} \in \mathscr{F}_u$$

For each t we can write

$$(3.3) X_t' = Y_n if T_n \le \tau_t < T_{n+1}.$$

Under assumption (A1) it follows from Lemma 1(b) that for any n and t

$$\{T_n \le \tau_t < T_{n+1}\} = \{\varphi(T_n) \le t < \varphi(T_{n+1})\},$$

and so (3.3) is equivalent to

$$(3.5) X_t' = Y_n if \varphi(T_n) \le t < \varphi(T_{n+1}).$$

Then by the definition of the  $\{v_n\}$ , (3.5) is equivalent to  $X_t' = Y_n'$  if  $T_n' \le t < T_{n+1}'$ . The counting process,  $N_t' = n$  if  $T_n' \le t < T_{n+1}'$ , associated with  $\{X_t'\}$  is obviously adapted to  $\{\mathscr{F}_t'\}$  if each  $T_n'$  is a  $\mathscr{F}_t'$ -stopping time. We prove the latter by induction. First note that by an argument similar to the above it can be shown that  $\{N(\tau_t)\}$  is a right continuous process adapted to  $\{\mathscr{F}_t'\}$ , and  $N(\tau_t) = v_n$  if  $T_n' \le t < T_{n+1}'$ . (Note that each  $T_n'$  is a jump point of  $N(\tau_t)$  since  $v_n > v_{n-1}$ .) Now obviously  $T_0' = 0$  is a  $\mathscr{F}_t'$ -stopping time, and assuming that  $T_0', \dots, T_n'$  are  $\mathscr{F}_t'$ -stopping times, for each  $t \ge 0$ 

$$\left\{T_{n+1}'>t\right\} = \bigcup_{r \leq t; \ r \text{ rational}} \left\{T_{n}'=r, N(\tau_{t})=N(\tau_{r})\right\} \cup \left\{T_{n}'>t\right\} \in \mathscr{F}_{t}'$$

which completes the induction argument.

It remains to show that  $\{Y_n', W_n'\}$  is an MRP with respect to  $\{\mathscr{F}'_{T_n'}\}$ . One can show by their definitions that

(3.6) 
$$\mathscr{F}'_{T_n'} = \mathscr{F}_{\tau(T_n')}$$
 and  $\mathscr{G}_{\nu_n} = \mathscr{F}_{T_{\nu_n}}$ 

By assumption (A1), for any n

$$\tau(T_n') = \inf \{u : \varphi_u > T_n'\}$$

$$= \max \{u : \varphi_u = \varphi(T_{\nu_n})\}$$

$$= T_{\nu_n},$$

and so from (3.6) we have

$$\mathscr{F}'_{T_{n'}} = \mathscr{G}_{\nu_n} = \mathscr{G}'_{n'}.$$

To show that  $\{Y_n, W_n'\}$  is an MRP with respect to  $\{\mathscr{G}_n'\}$  it suffices by Proposition 1 to show that  $\{Y_n, \varphi(T_n) - \varphi(T_{n-1})\}$  is an MRP with respect to  $\mathscr{G}_n$ . The latter follows, since by the strong Markov property of  $\{Y_n, W_n\}$ , and (A2) we have for any  $x \in E$ ,  $A \in \mathscr{E}$  and any n, t

(3.7) 
$$P[Y_{n+1} \in A, \varphi(T_{n+1}) - \varphi(T_n) \leq t \mid \mathcal{G}_n, Y_n = x]$$

$$= \int_{A \times R_+} Q(x, dy \times dw) P[\varphi(T_n + W_{n+1}) - \varphi(T_n) \leq t \mid Y_n = x, Y_{n+1} = y,$$

$$W_{n+1} = w]$$

$$= \int_{A \times R_+} Q(x, dy \times dw) P[\varphi(W_1) \leq t \mid Y_0 = x, Y_1 = y, W_1 = w]$$

$$= P[Y_1 \in A, \varphi(T_1) \leq t \mid Y_0 = x].$$

Thus the proof is complete.

REMARK 3. The SMP  $\{X_t'\}$  with respect to  $\{\mathscr{F}_t'\}$  is induced by the MRP  $\{Y_n', W_n'\}$  (defined in (3.1) and (3.2)) which by Proposition 1 and the above proof, has a kernel Q' satisfying

$$Q'(x, A \times [0, t]) = \sum_{k=0}^{\infty} \int_{0}^{\infty} \widetilde{Q}^{k}(x, dy \times \{0\}) \widetilde{Q}(y, A \times (0, t])$$

where

$$\widetilde{Q}(x, A \times [0, t]) = \int_A K(x, dy) P[\varphi(W_1) \le t \mid Y_0 = x, Y_1 = y].$$

REMARK 4. The above theorem also holds if we replace (A2) by: (A2') There exists a sequence  $\{\alpha_n\}$  of finite  $\mathcal{G}_n$ -stopping times that satisfy

$$(3.8) \alpha_{n+1} = \alpha_n + \alpha_1 \circ \theta_{\alpha_n},$$

where  $\{\mathcal{O}_n\}$  is the family of translation operators associated with  $\{Y_n, W_n\}$  (see [2]), and in addition

(3.9) 
$$\varphi(T_{\alpha_{n+1}}) - \varphi(T_{\alpha_{n+1}}) = 0 \qquad \text{for any } n \ge 0.$$

Set

(3.10) 
$$\widetilde{Y}_n = Y_{\alpha_n}$$
,  $\widetilde{T}_n = T_{\alpha_n}$ ,  $\widetilde{\mathscr{G}}_n = \mathscr{G}_{\alpha_n}$ ,  $\widetilde{V}_0 = 0$  and  $\widetilde{V}_n = \widetilde{T}_n - \widetilde{T}_{n-1}$ ; and assume that (A2) holds with  $Y_n$ ,  $T_n$  and  $\mathscr{G}_n$  replaced by  $\widetilde{Y}_n$ ,  $\widetilde{T}_n$  and  $\widetilde{\mathscr{G}}_n$  respectively.

Note that (A2) was only invoked in proving that  $\{Y_n', W_n'\}$  is an MRP with respect to  $\{\mathscr{G}_n'\}$ . Under (A2') it can be shown that the  $\{v_n\}$  defined in the above proof satisfy

$$v_{n+1} = \inf\{m > v_n : \varphi(\tilde{T}_m) > \varphi(\tilde{T}_{v_n})\}$$

and so

$$Y_n' = \widetilde{Y}_{\nu_n}, \qquad W_n' = \varphi(\widetilde{T}_{\nu_n}) - \varphi(\widetilde{T}_{\nu_{n-1}}) \quad \text{and} \quad \mathscr{G}_n' = \widetilde{\mathscr{G}}_{\nu_n}.$$

Then to show that  $\{Y_n', W_n'\}$  is an MRP with respect to  $\{\mathscr{G}_n'\}$ , it suffices by Proposition 1 to show that  $\{\widetilde{Y}_n, \varphi(\widetilde{T}_n) - \varphi(\widetilde{T}_{n-1})\}$  is an MRP with respect to  $\{\widetilde{\mathscr{G}}_n\}$ . The latter follows since it can be shown (see [13]) that  $\{\widetilde{Y}_n, \widetilde{W}_n\}$  is an MRP with respect to  $\{\widetilde{\mathscr{G}}_n\}$ , and so under (A2'), (3.7) holds with  $Y_n$ ,  $T_n$  and  $\mathscr{G}_n$  replaced by  $\widetilde{Y}_n$ ,  $\widetilde{T}_n$  and  $\widetilde{\mathscr{G}}_n$  respectively.

**4. Examples.** We now list several classes of processes  $\{\varphi_t\}$ , which according to the above, determine RTT's that transform SMP's into SMP's.

Functionals of  $\{X_t\}$ . Using the notation of Remark 2, set  $Z_t = (X_t, X_t^+, U_t, U_t^+)$  for each  $t \ge 0$ , and assume  $\mathscr{F}_t = \sigma(Z_s; s \le t)$ . For each  $0 \le s \le t$  let  $\varphi_t^s$  be a (nonnegative) random variable such that  $\{\varphi_t^s \le u\} \in \mathscr{F}_t$ . The family  $\{\varphi_t^s : 0 \le s \le t\}$  is called a (nonnegative) functional of the Markov process  $\{Z_t\}$ .

We will take  $\{\varphi_t^s\}$  to be continuous, that is for a.a.  $\omega \in \Omega$ ,  $\iota \to \varphi_t^s(\omega)$  is continuous for each s; and additive, that is for a.a.  $\omega \in \Omega$ 

(4.1) 
$$\varphi_{t}^{s}(\omega) + \varphi_{u}^{t}(\omega) = \varphi_{u}^{s}(\omega).$$

We will also assume that for any n and  $0 \le s \le t$ 

$$\varphi_t^s \circ \mathscr{O}_{T_n} = \varphi_{T_n + t}^{T_n + s},$$

where  $\{\ell_t\}$  is the usual family of translation operators associated with the Markov process  $\{Z_t\}$ . (This resembles the homogeneity property of functionals, see page 173 of [3].) We will call  $\{\varphi_t^s\}$  a nonnegative, homogeneous, continuous, additive functional of the SMP  $\{X_t\}$ .

Define the process  $\{\varphi_t\}$  by setting  $\varphi_t = \varphi_t^0$  for each  $t \ge 0$ . Clearly  $\{\varphi_t\}$  satisfies (A1), and (A2) is satisfied, since by (4.1), (4.2) and the strong Markov property for  $\{Z_t\}$ , for any  $n \ge 0$ ,  $x, y \in E$  and  $t, u, w \ge 0$ 

$$P[\varphi(T_n+t)-\varphi(T_n) \leq u \mid \mathscr{G}_n, Y_{n+1}, W_{n+1}] = P[\varphi_{T_n+t}^{T_n} \leq u \mid \mathscr{G}_n]$$
$$= P[\varphi_t \circ \mathscr{O}_{T_n} \leq u \mid \mathscr{F}_{T_n}]$$

where the latter is a  $\sigma(Y_0, Y_1, W_1)$ -measurable function for each u.

The RTT's determined by functionals of this sort are analogous to the RTT's described in Dynkin [3].

An example of the above type of functional is given by

$$\varphi_t^s = \int_s^t f(X_v, X_v^+, v - U_v, U_v^+ - v) dv,$$

where f(x, y, u, v) is a nonnegative real-valued measurable function defined for each  $x, y \in E$  and  $u, v \ge 0$ , and satisfies

That this functional is a.s. finite valued follows since for each t,

$$\varphi_t \leq \sum_{k=1}^{N_t+1} \varphi_{T_k}^{T_{k-1}} < \infty$$
 a.s.,

the above being true since  $N_t < \infty$  a.s., and by (4.3) and the fact that  $W_k < \infty$  a.s.,

$$\varphi_{T_{k-1}}^{T_{k-1}} = \int_{T_{k-1}}^{T_{k-1}+W_k} f(Y_{k-1}, Y_k, u - T_{k-1}, T_k - u) du < \infty$$
 a.s.

Random functions of  $\{X_t\}$ . Let  $\{V_n(x,y): n \ge 0, x, y \in E\}$  be a family of independent random variables which is independent of  $\{Y_n, W_n\}$  and satisfies:

- (i)  $(x, y, \omega) \rightarrow V_n(x, y, \omega)$  is  $\mathscr{E} \times \mathscr{E} \times \mathscr{F}$ -measurable for each n.
- (ii) For each  $x, y \in E$ ,  $V_0(x, y) = 0$  a.s. and

$$P[V_n(x, y) \le u] = P[V_1(x, y) \le u]$$
 for  $n \ge 1$ .

Set  $V_t = V_{n+1}(X_n, X_{n+1})$  if  $T_n \le t < T_{n+1}$ . It can be shown that  $Z_t = (X_t, X_t^+, U_t, U_t^+, V_t)$ ,  $t \ge 0$  is a regular step Markov process. Assume that for each  $t \ge 0$ 

$$\mathscr{F}_t = \sigma(Z_s: s \leq t).$$

Let  $\{\varphi_t^s : 0 \le s \le t\}$  be a nonnegative, continuous additive functional of  $\{Z_t\}$  satisfying (4.2) and set  $\varphi_t = \varphi_t^0$  for each  $t \ge 0$ . As above it can be shown that  $\{\varphi_t\}$  (called a random function of  $\{X_t\}$ ) satisfies (A1) and (A2).

The following special case of the above determines a RTT that transforms SMP's into regular step Markov processes. Suppose that the SMP  $\{X_t\}$  has a kernel Q which satisfies  $Q(x, \{x\} \times R_+) = 0$  for each  $x \in E$ . Assume that

$$P[V_1(x, y) \le t] = 1 - e^{-\lambda(x)t}$$
 for  $t \ge 0$ ,  
= 0 otherwise;

where  $\lambda$  is a positive  $\mathscr E$ -measurable function. For each  $t \geq 0$  set  $\rho_t = V_t/W_{N_t+1}$  and  $\varphi_t = \int_0^t \rho_u du$ . This random function determines a RTT of  $\{X_t\}$  such that, according to Remark 3,  $\{X_t'\}$  is an SMP induced by the MRP  $\{Y_n', W_n'\}$  (in this case  $\{Y_n, V_n(X_{n-1}, X_n)\}$ ) which has kernel

$$Q'(x, A \times [0, t]) = K(x, A)(1 - e^{-\lambda(x)t}).$$

That is,  $\{X_t'\}$  is a regular step Markov process. This type of RTT is presented in Yackel [15]. He uses it to show that an SMP with a countable state space (and with some instantaneous states) can be transformed into a Markov process having the same succession of states, and the same instantaneous states as the original SMP.

Random functions based on MRP's imbedded in  $\{Y_n, W_n\}$ . As in Remark 4 we let  $\{\alpha_n\}$  be a sequence of  $\mathcal{G}_n$ -stopping times satisfying (3.8) and let  $\{\tilde{Y}_n, \tilde{W}_n\}$  be the

MRP with respect to  $\widetilde{\mathscr{G}}_n$  defined by (3.9) and (3.10). Setting  $\widetilde{X}_t = \widetilde{Y}_n$  if  $\widetilde{T}_n \leq t < \widetilde{T}_{n+1}$ , it follows that  $\{\widetilde{X}_t\}$  is an SMP with respect to  $\{\mathscr{F}_t\}$  and induced by  $\{\widetilde{Y}_n, \widetilde{W}_n\}$ . Let  $\{\widetilde{X}_t, \widetilde{X}_t^+, \widetilde{U}_t, \widetilde{U}_t^+\}$  be the Markov process associated with  $\{\widetilde{X}_t\}$  as in Remark 2. Let  $\{V_n(x,y): n \geq 0, x, y \in E\}$  be a family of random variables as above and set

$$\widetilde{V}_t = V_{n+1}(\widetilde{Y}_n, \widetilde{Y}_{n+1})$$
 if  $\widetilde{T}_n \leq t < \widetilde{T}_{n+1}$ .

For each  $t \ge 0$  set

$$Z_t = (X_t, X_t^+, U_t, U_t^+, \widetilde{X}_t, \widetilde{X}_t^+, \widetilde{U}_t, \widetilde{U}_t^+, \widetilde{V}_t)$$

and  $\mathscr{F}_t = \sigma(Z_s : s \le t)$ . It can be shown that  $\{Z_t\}$  is a regular step Markov process. Let  $\{\varphi_t^s : 0 \le s \le t\}$  be a nonnegative continuous additive functional of  $\{Z_t\}$ , which in addition satisfies:

(i) For each  $n \ge 0$ 

(4.4) 
$$\varphi_{T_{n+1}}^{T_{\alpha_{n+1}}} = 0$$
 a.s.

(ii) For each  $n \ge 0$  and  $0 \le s \le t$ 

$$\varphi_t^s \circ \mathscr{O}_{\tilde{T}_n} = \varphi_{\tilde{T}_n + t}^{\tilde{T}_n + s}$$

where  $\{\mathcal{O}_t\}$  is the family of translation operators of  $\{Z_t\}$ .

Then define  $\{\varphi_t\}$  by setting  $\varphi_t = \varphi_t^0$  for each t. This process satisfies (A1) and it satisfies (A2') since (3.9) follows from (4.4), and by (4.5) and the strong Markov property of  $\{Z_t\}$ , for any  $n \ge 0$ ,  $x, y \in E$  and  $t, u, w \ge 0$ .

$$P\big[\varphi(\widetilde{T}_n+t)-\varphi(\widetilde{T}_n)\leqq u\ \big|\ \widetilde{\mathcal{G}}_n,\ \widetilde{Y}_{n+1},\ \widetilde{W}_{n+1}\big]=P\big[\varphi_t\circ\theta_{\widetilde{T}_n}\leqq u\ \big|\ \mathcal{F}_{T_n}\big]$$

where the latter is a  $\sigma(Y_0, Y_1, W_1)$ -measurable function for each u.

This type of random function can be used as follows. Given any Markov chain  $\{\tilde{Y}_n\}$  (as above) imbedded in  $\{Y_n\}$  and any family of distribution functions  $\{G_{xy}(t): x, y \in E\}$ , where  $G_{xy}(0) = 0$  and  $(x, y) \to G_{xy}(t)$  is  $\mathscr{E} \times \mathscr{E}$ -measurable for each t, one can construct an RTT of  $\{X_t\}$  such that  $\{X_t'\}$  is an SMP whose succession of states  $\{Y_n'\}$  is given by  $\{\tilde{Y}_n\}$ , and whose sojourn times  $\{W_n'\}$  satisfy

$$P[W'_{n+1} \le t \mid Y'_{n+1} = x, Y'_{n} = y] = G_{xy}(t).$$

A RTT of this type is determined, according to the above and Remark 3 and Remark 4, by the process  $\varphi_t = \int_0^t \rho_u du$  where

$$\rho_{t} = V_{n+1}(\widetilde{Y}_{n}, \widetilde{Y}_{n+1})/\widetilde{W}_{n+1}' \quad \text{if } T_{\alpha_{n}} \leq t < T_{\alpha_{n+1}}$$

$$= 0 \quad \text{otherwise};$$

and where  $P[V_1(x, y) \le t] = G_{xy}(t)$ .

Processes  $\{\varphi_t\}$  independent of  $\{X_t\}$ . Let  $\{\varphi_t\}$  be a non-decreasing right continuous process (with  $\varphi(0) = 0$ ) taking values in  $(R_+, \mathcal{B}_+)$ , adapted to  $\{\mathcal{F}_t\}$ , and which is continuous in probability and is independent of  $\{Y_n, W_n\}$ . Assume that  $\{\varphi_t\}$  has

stationary independent nonnegative increments with respect to  $\{\mathcal{F}_t\}$ , that is for any  $s, t \ge 0$ ,  $\sigma(\varphi(t+s) - \varphi(s))$  is independent of  $\mathcal{F}_s$  and

$$P[\varphi(t-s) - \varphi(s) \le u \mid \mathscr{F}_s] = P[\varphi(t) \le u].$$

Since  $\{\varphi_t\}$  is continuous in probability, for any  $t \ge 0$ 

$$P[\varphi(t) - \varphi(t^{-}) > 0] \le P[\varphi(t) - \varphi(t - n^{-1}) > 0] \to 0$$
 as  $n \to \infty$ ,

and since  $\{\varphi_t\}$  is independent of  $\{Y_n, W_n\}$ 

$$P[\varphi(T_n) - \varphi(T_n^-) > 0 \mid Y_0 = x] = \int_{-\infty}^{\infty} P[\varphi(t) - \varphi(t^-) > 0] Q^n(x, E \times dt) = 0.$$

Thus (A1) is satisfied.

Assumption (A2) is also satisfied for it can be shown (see [1]) that if  $\tau$  is any finite  $\mathcal{F}_{t}$ -stopping time, setting

$$\tilde{\varphi}_t = \varphi(t+\tau) - \varphi(\tau)$$
 and  $\tilde{\mathscr{G}}_t = \sigma(\tilde{\varphi}_s; s \leq t)$  for each  $t \geq 0$ ,

the  $\sigma$ -algebra  $\mathcal{G}_t$  is independent of  $\mathcal{F}_t$  and

$$(4.6) P[\tilde{\varphi}(t) \leq u \mid \mathcal{F}_{\tau}] = P[\varphi(t) \leq u] a.s.$$

Then (A2) follows from (4.6) and the independence of  $\{\varphi_t\}$  and  $\{Y_n, W_n\}$ .

RTT's determined by this class of processes are analogous to the RTT's of Markov Processes as presented in Feller [4] (which we noted in the introduction). For his RTT's,  $\{\tau_t\}$  is a process with stationary independent nonnegative increments, which is independent of  $\{X_t\}$ ; while in this case  $\{\tau_t\}$  is the inverse of such a process.

5. Some other random time transformations. In this section we take  $\{\tau_t\}$  to be a step function of the form

(5.1) 
$$\tau_t = S_n \quad \text{if } Z_n \le t < Z_{n+1},$$

where  $\{S_n\}$  and  $\{Z_n\}$  are sequences of strictly increasing random variables such that  $S_0 = Z_0 = 0$  a.s. The inverse  $\{\varphi_t\}$  of  $\{\tau_t\}$  obviously satisfies  $\varphi_t = Z_{n+1}$  if  $S_n \le t < S_{n+1}$ . Assume the following:

- (i)  $\{S_n\}$  is independent of  $\{Z_n\}$ .
- (ii)  $\{\tau_i\}$  is independent of  $\{Y_n, W_n\}$ .
- (iii)  $\{\varphi_t\}$  is adapted to  $\{\mathcal{F}_t\}$ .
- (iv) There exist distribution functions G and H such that for any n

$$P[S_{n+1} - S_n \le t \mid \mathscr{F}_{S_n}] = G(t)$$

and

$$P[Z_{n+2}-Z_{n+1} \le t \, \big| \, \mathscr{F}_{S_n}] = H(t).$$

(Note that each  $S_n$  is an  $\mathscr{F}_t$ -stopping time, since these are the jump points of  $\{\varphi_t\}$  which is adapted to  $\{\mathscr{F}_t\}$ .)

Let  $\{X_t'\}$  be the RTT of  $\{X_t\}$  by the change of time  $\{\tau_t\}$ . Then we have the following.

THEOREM 2. Suppose that  $\{X_t\}$  is a regular step Markov process with respect to  $\{\mathcal{F}_t\}$ . Then  $\{X_t'\}$  is an SMP with respect to  $\{\mathcal{F}_t'\}$ , which is induced by the MRP  $\{X(S_n), Z_{n+1} : n \geq 0\}$  with kernel  $\tilde{Q}(x, A \times [0, t]) = \tilde{K}(x, A)H(t)$  where

$$\widetilde{K}(x,A) = \int_0^\infty P[X_u \in A \mid X_0 = x]G(du).$$

Hence, if  $\{\tau_t\}$  is a compound Poisson process (i.e.,  $H(t) = 1 - e^{-\mu t}$  for  $\mu > 0$ ), then  $\{X_t'\}$  is a regular step Markov process with respect to  $\{\mathcal{F}_t'\}$ .

**PROOF.** Just as in the proof of Theorem 1 it follows that  $\{X_t'\}$  is a right continuous process adapted to  $\{\mathscr{F}_t'\}$ , and directly from (5.1) we have  $X_t' = X(S_n)$  if  $Z_n \leq t < Z_{n+1}$ . The counting process of  $\{X_t'\}$  is given by  $N_t' = n$  if  $Z_n \leq t < Z_{n+1}$ . Each  $Z_n$  is a  $\mathscr{F}_t'$ -stopping time since it is a jump point of the process  $\{\tau_t\}$ , which is adapted to  $\{\mathscr{F}_t'\}$ . Thus it follows that  $\{N_t'\}$  is adapted to  $\{\mathscr{F}_t'\}$ .

Under Assumptions (i)-(iv) and the strong Markov property for  $\{X_t\}$  it follows that for each  $n \ge 0$ ,  $x \in E$ ,  $A \in \mathscr{E}$  and  $t \ge 0$ 

$$\begin{split} P\big[X(S_{n+1}) \in A, Z_{n+2} - Z_{n+1} &\leq t \, \big| \, \mathscr{F}_{S_n}, X(S_n) = x \big] \\ &= P\big[X(S_{n+1}) \in A \, \big| \, \mathscr{F}_{S_n}, X(S_n) = x \big] P\big[Z_{n+2} - Z_{n+1} &\leq t \big] \\ &= \widetilde{K}(x, A) H(t). \end{split}$$

Thus  $\{X(S_n), Z_{n+1}\}$  is an MRP with respect to  $\mathscr{F}_{S_n}$  and this proves the first statement of the theorem. The second statement follows from Remark 1.  $\square$ 

Note that the second statement in the above theorem is a special case of the above mentioned result in Feller. It can be shown (see [13]) that the result in Feller for Markov processes does not hold for SMP's. However, we do have the following special case.

THEOREM 3. Suppose that the jump points of the SMP  $\{X_t\}$  occur a.s. at integer time points, i.e., each  $F_{xy}(t)$  is an arithmetic distribution with an integer valued span. Let  $(\tau_t)$  be a Poisson process where, using the above notation, for each  $t \ge 0$ ,  $\tau_t = n$  if  $Z_n \le t < Z_{n+1}$  and

$$P[Z_1 \le t] = 1 - e^{-\lambda t}$$
 for some  $\lambda > 0$ .

Set  $\tilde{Z}_n = Z_{T_n}$  for each n. Then  $\{X_t'\}$  is an SMP with respect to  $\{\mathcal{F}_t'\}$  and is induced by the MRP  $\{Y_n, \tilde{Z}_n\}$  which has the kernel

$$\widetilde{Q}(x, A \times [0, t]) = \sum_{k=1}^{\infty} Q(x, A \times \{k\}) H_{\lambda, k}(t)$$

where  $H_{\lambda,k}$  is a gamma distribution with parameter  $\lambda$  and order k.

**PROOF.** Just as in the proof of Theorem 1 it follows that  $\{X_t'\}$  is a right continuous process adapted to  $\{\mathcal{F}_t'\}$ , and that

(5.2) 
$$X_{t}' = Y_{n}$$
 if  $T_{n} \le \tau_{t} < T_{n+1}$ .

Under the hypotheses it follows that

$$\begin{split} & \{T_n \leq \tau_t < T_{n+1}\} \\ & = \{T_n = \tau_t\} \cup \{T_n + 1 = \tau_t\} \cup \dots \cup \{T_{n+1} - 1 = \tau_t\} \\ & = \{Z_{T_n} \leq t < Z_{T_{n+1}}\} \cup \{Z_{T_{n+1}} \leq t < Z_{T_{n+2}}\} \cup \dots \cup \{Z_{T_{n+1} - 1} \leq t < Z_{T_{n+1}}\} \\ & = \{\widetilde{Z}_n \leq t < Z_{n+1}\}. \end{split}$$

Thus (5.2) is equivalent to  $X_t' = Y_n$  if  $\tilde{Z}_n \leq t < \tilde{Z}_{n+1}$ .

Just as in the proof of Theorem 2, it follows that the counting process associated with  $\{X_t'\}$  is adapted to  $\{\mathscr{F}_t'\}$ . In addition, each  $\widetilde{Z}_n$  is a  $\mathscr{F}_t$ -stopping time. Thus  $\{Y_n, \widetilde{Z}_n\}$  is an MRP with respect to  $\{\mathscr{F}_{Z_n}\}$ , since for any  $x \in E$ ,  $A \in \mathscr{E}$  and  $t \ge 0$ 

$$\begin{split} P\big[Y_{n+1} \in A, \widetilde{Z}_{n+1} - \widetilde{Z}_n &\leq t \, \big| \, \mathscr{F}_{Z_n}, Y_n = x \big] \\ &= \sum_{k=1}^{\infty} P\big[Y_{n+1} \in A, T_{n+1} - T_n = k, Z_{T_n+k} - Z_{T_n} \leq t \, \big| \, \mathscr{F}_{Z_n}, Y_n = x \big] \\ &= \sum_{k=1}^{\infty} Q(x, A \times \{k\}) P\big[Z_k \leq t \big] \\ &= \sum_{k=1}^{\infty} Q(x, A \times \{k\}) H_{\lambda, t}(t). \end{split}$$

This completes the proof.

Acknowledgment. This paper represents a generalization of the heart of the author's dissertation written under Professor Erhan Çinlar at Northwestern University.

## REFERENCES

- [1] Breiman, L. (1968). Probability. Addison-Wesley, Reading.
- [2] BLUMENTHAL, R. M. and GETOOR, R. K. (1968). Markov Processes and Potential Theory. Academic Press, New York.
- [3] DYNKIN, E. (1965). Markov Processes-I. Springer-Verlag, Berlin.
- [4] FELLER, W. (1966). An Introduction to Probability Theory and its Applications 2. Wiley, New York.
- [5] Feller, W. (1964). On semi-Markov processes. Proc. Nat. Acad. Sci. USA 51 653-659.
- [6] LÉVY, P. (1954). Processus semi-Markoviens. Proc. Internat. Congress Math. (Amsterdam) 3 416-426.
- [7] MEYER, P. A. (1966). Probability and Potentials. Ginn, Boston.
- [8] PARTHASARATHY, K. R. (1967). Probability Measures on Metric Spaces. Academic Press, New York.
- [9] PYKE, R. (1961a). Markov renewal processes: definitions and preliminary properties. Ann. Math. Statist. 32 1231-1242.
- [10] PYKE, R. (1961b). Markov renewal processes with finitely many states. Ann. Math. Statist. 32 1243-1259.
- [11] PYKE, R. and SCHAUFELE, R. A. (1964). Limit theorems for Markov renewal processes. Ann. Math. Statist. 35 1746-1764.
- [12] PYKE, R. and SCHAUFELE, R. A. (1966). The existence and uniqueness of stationary measures for Markov renewal processes. Ann. Math. Statist. 37 1439–1462.
- [13] Serfozo, R. F. (1969). Time and space transformations of semi-Markov processes. Ph.D. dissertation, Northwestern Univ.
- [14] SMITH, W. L. (1955). Regenerative stochastic processes. Proc. Roy. Soc. London Ser. A 232 6-31.
- [15] YACKEL, J. (1968). A random time change relating semi-Markov and Markov processes. Ann. Math. Statist. 39 358-364.