## ON THE EXISTENCE OF THE OPTIMAL STOPPING RULE IN THE $S_n/n$ PROBLEM WHEN THE SECOND MOMENT IS INFINITE<sup>1</sup>

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Let  $X_1, X_2, \cdots$  be i.i.d. random variables with mean 0, and let  $S_n = \sum_{i=1}^n X_i$ . The  $S_n | n$  optimal stopping problem is to maximize  $E(S_\tau / \tau)$  among finite-valued stopping times  $\tau$  relative to the process  $(S_n, n \ge 1)$ . In this paper we prove partially Dvoretzky's (1967) conjecture that an optimal stopping time should exist when  $E|X_1|^{\beta} < \infty$  for some  $\beta > 1$ , by showing that the result holds if  $\limsup_{n\to\infty} P(S_n \ge c ||S_n||) > 0$  for some c > 0, where  $||S_n|| = (E|S_n|^{\beta})^{1/\beta}$ . This condition is shown to hold in some special cases, including the case where the  $X_i$  are in the domain of attraction of a stable distribution with exponent greater than one.

**1. Introduction.** Let  $X_1, X_2, \cdots$  be independent and identically distributed random variables with mean 0, and let  $S_n = \sum_{i=1}^n X_i$ . If  $\mathcal{M}$  is the collection of finite valued stopping times  $\tau$  relative to the process  $(S_n, n \ge 1)$  for which  $E(S_{\tau}/\tau)$  is defined (possibly infinite), then the  $S_n/n$  optimal stopping problem is to find if possible  $\sigma \in \mathcal{M}$  such that

(1.1) 
$$E(S_{\sigma}/\sigma) = \sup \left[ E(S_{\tau}/\tau) : \tau \in \mathcal{M} \right].$$

Burgess Davis (1971) has shown that if  $E(X_1 \log^+ X_1) = \infty$  (where  $\log^+ a = \log a$  if  $a \ge 1$ , and 0 if a < 1), then there is a  $\sigma \in \mathcal{M}$  for which  $E(S_{\sigma}/\sigma) = \infty$ . This  $\sigma$  is clearly optimal in the sense of (1.1). On the other hand, it is also well known (see Davis (1971)) that if  $E(X_1 \log^+ X_1) < \infty$ , then  $E(\sup_{n\ge 1} (S_n^+/n)) < \infty$ . Therefore, if  $E(X_1 \log^+ X_1) < \infty$ ,  $\mathcal{M}$  is the class of all finite stopping times relative to  $(S_n, n \ge 1)$ ; and general optimal stopping theory tells us a good deal more.

The process  $X = ((S_n, n), n \ge 1)$  may be regarded as a Markov process with state space  $R \times [0, \infty)$ , stationary transition probabilities, and initial distribution that of  $X_1$  on the line  $\{(x, 1): x \in R\}$ . (Here R denotes the real numbers.) Suppose that  $E(\sup_n (S_n^+/n)) < \infty$ , and for each  $(x, s) \in R \times (0, \infty)$  let

(1.2) 
$$h(x,s) = (x/s)^+ \vee \sup \left[ E((x+S_\tau)/(s+\tau))^+ : \tau \in \mathcal{M} \right].$$

Let  $D = \{(x, s) \in R \times (0, \infty) : h(x, s) = x^+/s\}$ . It can be proved as in Theorem 8 of Chow and Robbins (1967) that  $(h(x+S_n, s+n), n \ge 1)$  is the minimal supermartin-

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gale above  $((x+S_n)^+/(s+n), n \ge 1)$  for each  $(x, s) \in R \times [0, \infty)$ . If  $\tau \in \mathcal{M}$  and  $\bar{\tau} = \inf [n \ge \tau : S_n > -x]$ , then  $\bar{\tau} \in \mathcal{M}$  (Feller (1966), page 380) and

(1.3) 
$$E\left(\frac{x+S_{\bar{\tau}}}{s+\bar{\tau}}\right) \leq E\left(\frac{x+S_{\bar{\tau}}}{s+\bar{\tau}}\right) = E\left(\left(\frac{x+S_{\bar{\tau}}}{s+\bar{\tau}}\right)^{+}\right).$$

Hence  $D \subset (0, \infty) \times (0, \infty)$ , and it follows by the Corollary to Theorem 6 of Chow and Robbins (1967) that if for  $(x, s) \in R \times [0, \infty)$  the stopping time  $\tau(x, s) = \inf [n \ge 1 : (x + S_n, s + n) \in D]$  is finite, it maximizes  $E(x + S_n, s + n)$  among  $\tau \in \mathcal{M}$ .

Dvoretzky's proof ((1967), Section 3) of the following theorem, describing the set D, is valid, as it does not use the general assumption of his paper that  $EX_1^2 < \infty$ .

THEOREM 1.1. Let  $EX_1 \log^+ X_1$  be finite. Then there is a strictly increasing positive function f(s) on  $[0, \infty)$  such that for  $(x, s) \in R \times [0, \infty)$ 

$$\frac{x}{s} < \sup \left[ E\left(\frac{x+S_{\tau}}{s+\tau}\right) : \tau \in \mathcal{M} \right] \quad \text{if} \quad x < f(s),$$

$$\frac{x}{s} = \sup \left[ E\left(\frac{x+S_{\tau}}{s+\tau}\right) : \tau \in \mathcal{M} \right] \quad \text{if} \quad x = f(s),$$

$$\frac{x}{s} > \sup \left[ E\left(\frac{x+S_{\tau}}{s+\tau}\right) : \tau \in \mathcal{M} \right] \quad \text{if} \quad x > f(s).$$

Dvoretzky (1967) and Teicher and Wolfowitz (1966) showed that if  $E{X_1}^2 < \infty$ , then  $\tau(0,0)$  is finite, and therefore optimal. Dvoretzky conjectured that the same result ought to hold when  $E|X_1|^{\beta} < \infty$  for some  $\beta > 1$ . In this paper we prove a partial result in this direction; we establish the existence of an optimal stopping rule when the truncated variance

(1.5) 
$$U(x) = \int_{(-x)^{-}}^{x} y^{2} dP(X_{1} \le y)$$

is of "dominated variation" as  $x \to \infty$ , in a sense defined precisely in Section 3 below. A particular case is the situation when  $X_1$  belongs to the domain of attraction of some stable random variable with exponent greater than one.

**2.** An upper bound on f. Let  $\overline{\mathcal{M}}$  denote the class of all (possibly infinite) stopping times relative to  $(S_n, n \ge 1)$ . If  $\tau \in \overline{\mathcal{M}}$ , we shall set

(2.1) 
$$\frac{x+S_{\tau}}{s+\tau} = \lim_{n\to\infty} \frac{x+S_n}{s+n} = 0$$

on the set  $\{\tau = \infty\}$  for any  $(x, s) \in R \times [0, \infty)$ . The theorem we shall prove in this section is the following:

THEOREM 2.1. If there is a  $\beta > 1$  such that  $E|X_1|^{\beta} < \infty$ , then there is a finite constant  $K_0$  such that for every  $s \ge 1$  and  $x \ge K_0 ||S_{[s]}||$ , where  $||S_{[s]}|| = (E|S_{[s]}|^{\beta})^{1/\beta}$ , we have

(2.2) 
$$\frac{x}{s} \ge \sup \left[ E\left(\frac{x+S_{\tau}}{s+\tau}\right) : \tau \in \overline{\mathcal{M}} \right].$$

The proof will be broken down into a sequence of lemmas, in which it will always be assumed that  $E|X_1|^{\beta} < \infty$  for some  $\beta > 1$ . The first one is an analogue of Lemma 2 of Dvoretzky (1967).

Lemma 2.1. If  $s \ge 1$  then

(2.3) 
$$\sup \left[ E\left(\frac{S_{\tau}}{\tau + s}\right) : \tau \in \overline{\mathcal{M}} \right] < \frac{2\beta' \|S_{[s]}\|}{\lceil s \rceil}$$

where  $\beta' = (1 - 1/\beta)^{-1}$ .

Proof. We have

$$(2.4) E\left[\sup_{n\geq 1}\left|\frac{S_n}{n+s}\right|^{\beta}\right] \leq E\left[\sup_{[s]\geq n\geq 1}\left|\frac{S_n}{n+s}\right|^{\beta}\right] + E\left[\sup_{n>[s]}\left|\frac{S_n}{n+s}\right|^{\beta}\right] \\ \leq E\left[\sup_{[s]\geq n\geq 1}\left|\frac{S_n}{n+s}\right|^{\beta}\right] + E\left[\sup_{n>[s]}\left|\frac{S_n}{n}\right|^{\beta}\right].$$

But the processes  $(|S_1|^{\beta}, |S_2|^{\beta}, \cdots)$  and  $(\cdots, |S_3/3|^{\beta}, |S_2/2|^{\beta}, |S_1|^{\beta})$  are both submartingales. Therefore (see Doob (1953), page 317),

(2.5) 
$$E \left[ \sup_{n \ge 1} \left| \frac{S_n}{n+s} \right|^{\beta} \right] \le 2(\beta')^{\beta} E \left[ \left| \frac{S_{[s]}}{\lceil s \rceil} \right|^{\beta} \right].$$

From Jensen's inequality

(2.6) 
$$E \left[ \sup_{n \ge 1} \left| \frac{S_n}{n+s} \right| \right] \le 2\beta' \frac{\left\| S_{[s]} \right\|}{[s]},$$

and the lemma follows.

LEMMA 2.2. Let  $s' \ge s > 0$ ,  $x' \le x$  and  $\tau \in \overline{\mathcal{M}}$  satisfy

(2.7) 
$$E\left(\frac{x+S_{\tau}}{s+\tau}\right) \ge \frac{x}{s}.$$

Then for  $m=0, 1, 2, \cdots$  there is a stopping time  $\tau(m) \in \overline{\mathcal{M}}$  satisfying

$$(2.8) S_{\tau(m)} > x - x' \text{ a.s. on } \{\tau(m) \leq m\},$$

and

(2.9) 
$$E\left(\frac{x'+S_{\tau(m)}}{s'+\tau(m)}\right) \ge \frac{x'}{s'}.$$

PROOF. This is Lemma 6 of Dvoretzky (1967).

LEMMA 2.3. Let (2.7) hold for some x > 0, s > 0,  $\tau \in \overline{\mathcal{M}}$ . Then there is a  $\tau^* \in \overline{\mathcal{M}}$  satisfying

(2.10) 
$$E\left(\frac{x/2 + S_{\tau^*}}{s + \tau^*}\right) \ge \frac{x}{2s}$$

and

(2.11) 
$$E\left(\frac{1}{s+\tau^*}\right) \leq \frac{1}{2s} + \frac{\|S_{[s]}\|}{sx}.$$

PROOF. Let  $\tau^*$  be  $\tau([s])$  from Lemma 2.2, with s'=s and x'=x/2. Then (2.10) holds, and max  $(S_1, S_2, \dots, S_{[s]}) > x/2$  a.s. on  $\{\tau^* \leq s\}$ . Thus  $P(\tau^* \leq s) \leq 2E|S_{[s]}|/x$ , and hence by Jensen's inequality,  $P(\tau^* \leq s) \leq 2|S_{[s]}|/x$ . But

(2.12) 
$$E\left(\frac{1}{s+\tau^*}\right) \le \frac{P(\tau^* \le s)}{s+1} + \frac{P(\tau^* > s)}{s+\lceil s \rceil + 1}$$
$$\le \frac{P(\tau^* \le s)}{s} + \frac{P(\tau^* > s)}{2s}$$
$$= \frac{1}{2s} + \frac{P(\tau^* \le s)}{2s},$$

and the result follows.

To prove Theorem 2.1 we observe that if (2.7) holds for some  $\tau \in \overline{\mathcal{M}}$ , then by Lemmas 2.3 and 2.1,

(2.13) 
$$\frac{x}{2s} \leq E\left(\frac{x/2 + S_{\tau^*}}{s + \tau^*}\right) = \frac{x}{2}E\left(\frac{1}{s + \tau^*}\right) + E\left(\frac{S_{\tau^*}}{s + \tau^*}\right)$$
$$\leq \frac{x}{4s} + \frac{\|S_{[s]}\|}{2s} + \frac{2\beta'\|S_{[s]}\|}{[s]},$$

or

(2.14) 
$$\frac{x}{4s} \le (2\beta' + \frac{1}{2}) \frac{\|S_{[s]}\|}{[s]} \le (4\beta' + 1) \frac{\|S_{[s]}\|}{s}.$$

Theorem 2.1 now follows with  $K_0 = 16\beta' + 4$ .

COROLLARY. If  $E|X_1|^{\beta} < \infty$  for some  $\beta > 1$ , then  $f(s) \leq K_0 ||S_{[s]}||$  for  $s \geq 1$ ; and if in addition

(2.15) 
$$\lim \sup_{n \to \infty} S_n / ||S_n|| > K_0 \text{ a.s.},$$

then an optimal stopping time exists for the  $S_n/n$  problem.

It is easy to show using the Hewitt-Savage Zero-One Law (see Feller (1966), page 122) that inequality (2.15) holds if  $\limsup_{n\to\infty} P(S_n > K_0 ||S_n||) > 0$ . Moreover, by using an inequality of Marcinkiewicz and Zygmund (1938) (see Theorem 5) we can strengthen this result as follows.

COROLLARY. If 
$$E|X_1|^{\beta} < \infty$$
 for some  $\beta > 1$ , and if (2.16) 
$$\limsup_{n \to \infty} P(S_n \ge c ||S_n||) > 0$$

for some c > 0, then an optimal stopping time exists for the  $S_n/n$  problem.

3. Dominated variation and the relative growth of  $S_n$  and  $||S_n||$ . Let F be the common distribution function of  $X_1, X_2, \cdots$  and define the function U on  $[0, \infty)$  by

(3.1) 
$$U(x) = \int_{(-x)^{-}}^{x} y^{2} dF(y).$$

DEFINITION. (Feller (1967), Section 8). The truncated variance U is of dominated variation if there exist constants v > 0, C and T > 0 such that

$$(3.2) U(tx)/U(t) < Cx^{2-\nu}$$

for all x > 1 and t > T.

The following lemma will relate the behavior of the other truncated moments of  $X_1$  with that of U, when U is of dominated variation.

LEMMA 3.1. (Feller (1967), Section 8, Theorem 2). Let U be of dominated variation and let

$$(3.3) V_a(x) = \int_x^\infty y^{-q} dU(y).$$

If q > 2 - v, then for t > T

(3.4) 
$$t^{q}V_{q}(t)/U(t) \leq L_{q} = -1 + \frac{Cq}{q-2+\nu}.$$

We observe that  $V_2(x) = 1 - F(x) + F[(-x) - ]$ , and

$$V_1(x) = \int_{(-\infty, -x) \cup (x, \infty)} |y| dF(y).$$

A trivial consequence of (3.2) is that  $\lim_{x\to\infty} x^{-2}U(x) = 0$ . Therefore, there is a sequence  $(a_n)$  of positive numbers tending to infinity and a finite constant  $K_1$  such that

$$(3.5) na_n^{-2}U(a_n) \le K_1$$

for all n. From (3.2) we then have

$$(3.6) na_n^{-2}U(a_nx) < K_1Cx^{2-\nu}$$

for x > 1 and n large. The relations

$$nV_2(a_{-}x) \le L_2K_1Cx^{-\nu}$$

and (for v > 1)

$$(3.8) nV_1(a_n x) \le L_1 K_1 C a_n x^{1-\nu},$$

again for x > 1 and n large, now follow from (3.4).

LEMMA 3.2. Let U be of dominated variation with v > 1, and let  $(a_n)$  be a sequence of numbers tending to infinity and satisfying (3.5). For n and x large, there is a constant  $C_0$  independent of n and x such that

$$(3.9) P(|S_n| \ge xa_n) \le C_0 x^{-\nu}.$$

**PROOF.** Fix x > 1, and define the truncated variables

(3.10) 
$$X_{kn} = X_k \quad \text{if} \quad -xa_n \le X_k \le xa_n,$$
$$= 0 \quad \text{otherwise.}$$

Let 
$$\mu_n = EX_{kn}$$
 and  $S_{mn} = \sum_{k=1}^m X_{kn}$ . Then

$$(3.11) \quad P(|S_n| \ge xa_n) \le n[P(X_1 < -xa_n) + P(X_1 > xa_n)] + P(|S_{nn}| \ge xa_n).$$

Since  $EX_{1} = 0$ ,

(3.12) 
$$|n\mu_n| \le nV_1(xa_n) \le L_1 K_1 Ca_n x^{1-\nu} \le \frac{1}{2} a_n x^{1-\nu}$$

for x and n large. Therefore

(3.13) 
$$P(|S_{nn}| > xa_n) \leq P(|S_{nn} - n\mu_n| > \frac{1}{2}xa_n)$$
$$\leq 4x^{-2}a_n^{-2}nU(a_nx)$$
$$\leq 4K_1Cx^{-\nu}$$

for x and n large. Moreover,

$$(3.14) n[P(X_1 < -a_n x) + P(X_1 > a_n x)] = nV_2(a_n x) \le L_2 K_1 C x^{-\nu}$$

for x and n large. The conclusion of the lemma is now immediate.

COROLLARY. Let U be of dominated variation with v > 1, and choose  $(a_n)$  to satisfy (3.5). Let  $F_n$  be the distribution of  $S_n/a_n$ . If  $1 < \beta < v$ , and some subsequence  $(F_n)$  of  $(F_n)$  converges to G, then

(3.15) 
$$\lim_{n'\to\infty} \int_{-\infty}^{\infty} |y|^{\beta} dF_{n'}(y) = \int_{-\infty}^{\infty} |y|^{\beta} dG(y).$$

Feller (1967) proves that since U is of dominated variation, we can choose  $(a_n)$  so that the sequence  $na_n^{-2}U(a_n)$  is bounded below by some constant  $\rho > 0$ , in which case the sequence  $(F_n)$  is stochastically compact in the following sense.

DEFINITION. If  $(G_n)$  is a sequence of distribution functions and every subsequence of  $(G_n)$  has a further subsequence converging to a nondegenerate distribution function, the sequence  $(G_n)$  is called *stochastically compact*.

In view of the criterion (2.16), the following theorem is immediate from these remarks and (3.15).

THEOREM 3.1. If U is of dominated variation with v > 1, then an optimal stopping time exists for the  $S_n/n$  problem.

**4.** Structure of the optimal stopping rule. It is an easy consequence of Theorem 2.1 and (3.15) that when U is of dominated variation with v > 1 and  $(a_n)$  is a sequence for which  $(F_n)$  is stochastically compact, the function f satisfies

$$(4.1) f(s) \le K' a_{[s]}$$

for some finite positive K'. We may also prove the following in direct analogy with the second part of Dvoretzky's (1967) Theorem 2.

LEMMA 4.1. Under the conditions stated above

$$(4.2) f(s) \ge ka_{[s]}$$

for some positive k and all s.

PROOF. Let c > 0 be given. For any integer r > 0, let  $\tau(r) = \inf[n \ge r : S_n > -ca_r]$ . The stopping time  $\tau(r)$  is a.s. finite. Then

(4.3) 
$$E\left(\frac{ca_r + S_{\tau(r)}}{r + \tau(r)}\right) \ge \frac{1}{2r} \int_{-ca_r}^{\infty} (ca_r + u) dP(S_r \le u)$$
$$= \frac{a_r}{2r} \int_{-c}^{\infty} (c + w) dP(S_r / a_r \le w).$$

If  $(F_{n'})$  is a subsequence of  $(F_n)$  converging to G

(4.4) 
$$\lim \inf_{n' \to \infty} \frac{n'}{ca_{n'}} E\left(\frac{ca_{n'} + S_{\tau(n')}}{n' + \tau(n')}\right) \ge \frac{1}{2c} \int_{-c}^{\infty} (c + w) dG(w).$$

Now  $2c - \int_{-c}^{\infty} (c+w)dG(w)$  is negative for c=0, and continuous. Hence for some finite  $c_0 > 0$ ,  $2c_0 < \int_{-c_0}^{\infty} (c_0 + w)dG(w)$ , and

(4.5) 
$$\liminf_{n' \to \infty} \frac{n'}{c_0 a_{n'}} E\left(\frac{c_0 a_{n'} + S_{\tau(n')}}{n' + \tau(n')}\right) > 1.$$

Thus  $c_0 a_{n'} < f(n')$  for all but finitely many n'. The lemma now follows by the stochastic compactness of the family  $(F_n)$ .

It is well known (see Feller (1966), page 305) that if  $X_1$  belongs to the domain of attraction of a nondegenerate random variable Y, with norming constants  $(a_n)$ , then there exists a number  $\alpha > 0$  such that  $a_{rn}/a_n \to r^{1/\alpha}$  as  $n \to \infty$ , and Y is strictly stable with exponent  $\alpha$ . If  $\alpha > 1$ , then  $X_1$  satisfies the hypothesis of Theorem 3.1 with  $1 < v < \alpha$  (see Feller (1966), page 303). Because  $a_{rn}/a_n \to r^{1/\alpha}$  as  $n \to \infty$ ,  $a_n$  is of the form  $n^{1/\alpha}L(n)$  where L is a slowly varying function of its argument (Feller (1966), page 269), and equations (4.1) and (4.2) suggest that f may have the same asymptotic behavior. That this is true will be shown in a subsequent paper.

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