

# A stochastic approximation method approximating the roots of time varying regression functions\*

By

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## 1. Introduction

Let  $M(x)$  be a regression function on the real line  $R$  and  $\theta$  be the unique root of the equation  $M(x)=0$ . In practical problems there exist some cases where the use of a process which converges to  $\theta$  from below (in some sense) is advantageous to us. For instance,  $\theta$  may be an optimal level in operating a system where the costs caused by operating at a level above  $\theta$  are considerably greater than those caused by operating at a level below  $\theta$ . EICHHORN and ZACKS [3] considered this situation in the problem of finding an optimal dosage of drugs which have secondary harmful effects in addition to their therapeutic effects. ANBAR [1] has proposed the modified Robbins-Monro (R-M) procedure  $X_n$  and has proved that  $X_n$  converges almost surely (a. s.) to  $\theta$  as  $n$  tends to infinity and that, with probability one,  $X_n$  exceeds  $\theta$  only finitely many times. ISOGAI [6] has obtained the same results as above for the case where there exist errors in setting the  $x$ -levels.

However some situation may occur where the regression function  $M(x)$  varies with the time. This situation has been treated by several authors (for example, by DUPAČ [2] and WATANABE [9]). In this paper we also consider this situation. Let  $M_n(x)$  be a regression function on  $R$  at time  $n$  and  $\theta_n$  be the unique root of the equation  $M_n(x)=0$ . The modified R-M procedure  $X_n$  defined by ANBAR [1] will be used. The aim of this paper is to derive the rate of convergence of  $X_n-\theta_n$  to zero by using the method due to HEYDE [4], and to show that, with probability one,  $X_n-\theta_n$  exceeds zero only finitely many times.

In Section 2 we shall give notations and prove several auxiliary results which are needed for Section 3. In Section 3 the main results will be proved.

## 2. Preliminaries and auxiliary results

In this section we shall give notations which are used throughout this paper and

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prove several auxiliary results which are needed for the next section.

The modified R-M procedure by ANBAR [1] is defined as follows:

$$(2.1) \quad \begin{aligned} &X_1 \text{ is any random variable with } E[X_1^2] < \infty, \\ &X_{n+1} = X_n - a_n(y_n + b_n) \quad \text{for } n=1, 2, \dots, \end{aligned}$$

where  $\{a_n\}$  is a sequence of positive numbers,  $\{b_n\}$  is a sequence of real numbers and  $\{y_n\}$  is a sequence of random variables. In this paper we assume the following assumptions be valid.

*Assumption.* (i)  $\{Y_n(x)\}$  is a sequence of random variables which depend on parameter  $x \in R$ , and for each  $n$  and  $x$  the expectation of  $Y_n(x)$  exists, and the regression function  $M_n(x)$  is Borel measurable on  $R$  and has the unique root  $\theta_n$  of the equation  $M_n(x)=0$ , where

$$(2.2) \quad M_n(x) = E[Y_n(x)].$$

(ii)  $y_n$  in (2.1) is a random variable which has conditional distribution given  $X_1, \dots, X_n$  equal to that of  $Y_n(X_n)$  given  $X_n$ .

By Assumption (ii) we get

$$(2.3) \quad E[y_n | X_1, \dots, X_n] = M_n(X_n) \quad a.s.,$$

where  $E[\cdot | \cdot]$  denotes the conditional expectation operator. We can rewrite (2.1) as

$$(2.4) \quad X_{n+1} = X_n - a_n M_n(X_n) - a_n b_n + a_n v_n \quad \text{for } n=1, 2, \dots,$$

where

$$v_n = M_n(X_n) - y_n \quad \text{for } n=1, 2, \dots$$

Clearly

$$(2.5) \quad E[v_n | X_1, \dots, X_n] = 0 \quad a.s.,$$

so that  $\{v_n\}$  is a martingale difference. In this paper  $C_1, C_2, \dots$  denote appropriate positive constants.

LEMMA 2.1. Let  $\{\alpha_n\}$  be a sequence of positive numbers with  $\lim_{n \rightarrow \infty} \alpha_n = \alpha > \frac{1}{2}$ . Then

$$\sum_{m=n_0}^n (m\gamma_m)^{-2} \sim (2\alpha-1)^{-1} (n\gamma_n^2)^{-1} \quad \text{as } n \rightarrow \infty,$$

where  $n_0$  is a positive integer such that

$$\begin{aligned} 1 - \alpha_n n^{-1} &> 0 \quad \text{for all } n \geq n_0, \\ \gamma_n &= \prod_{j=n_0}^n (1 - \alpha_j j^{-1}) \quad \text{for all } n \geq n_0 \end{aligned}$$

and “ $a_n \sim b_n$  as  $n \rightarrow \infty$ ” means  $\lim_{n \rightarrow \infty} a_n/b_n = 1$ .

PROOF. It suffices to show that

$$(2.1.1) \quad (2\alpha-1)n\gamma_n^2 \sum_{m=n_0}^n (m\gamma_m)^{-2} \rightarrow 1 \quad \text{as } n \rightarrow \infty.$$

For any  $\varepsilon > 0$  with  $2(\alpha-\varepsilon) > 1$  there exists a positive integer  $n_1 \geq n_0$  such that for all  $j \geq n_1$

$$\alpha + \varepsilon > \alpha_j > \alpha - \varepsilon \quad \text{and} \quad 1 - (\alpha + \varepsilon)j^{-1} > 0,$$

which implies

$$(2.1.2) \quad \prod_{j=m+1}^n (1 - (\alpha + \varepsilon)j^{-1}) < \prod_{j=m+1}^n (1 - \alpha_j j^{-1}) < \prod_{j=m+1}^n (1 - (\alpha - \varepsilon)j^{-1})$$

for all  $n > m \geq n_1$ . Let any  $\eta (0 < \eta < 1)$  be fixed. By using (2.3) of SACKS [8] there exists a positive integer  $n_2 \geq n_1$  such that for all  $n > m \geq n_2$

$$(2.1.3) \quad \prod_{j=m+1}^n (1 - (\alpha + \varepsilon)j^{-1}) > (1 - \eta)n^{-(\alpha + \varepsilon)}m^{\alpha + \varepsilon}$$

$$\prod_{j=m+1}^n (1 - (\alpha - \varepsilon)j^{-1}) < (1 + \eta)n^{-(\alpha - \varepsilon)}m^{\alpha - \varepsilon},$$

which, together with (2.1.2), yields

$$(2.1.4) \quad \gamma_n^2 \sum_{m=n_0}^n (m\gamma_m)^{-2}$$

$$\geq \sum_{m=n_2}^n m^{-2}(\gamma_n \gamma_m^{-1})^2 > (1 - \eta)^2 n^{-2(\alpha + \varepsilon)} \sum_{m=n_2}^n m^{2(\alpha + \varepsilon) - 2}$$

for all  $n \geq n_2$ . According to Lemma 4 of SACKS [8] there exists a positive integer  $n_3 \geq n_2$  such that for all  $n \geq n_3$

$$(2.1.5) \quad \sum_{m=n_2}^n m^{2(\alpha + \varepsilon) - 2} > (1 - \eta)(2(\alpha + \varepsilon) - 1)^{-1} n^{2(\alpha + \varepsilon) - 1}$$

$$\sum_{m=n_2}^n m^{2(\alpha - \varepsilon) - 2} < (1 + \eta)(2(\alpha - \varepsilon) - 1)^{-1} n^{2(\alpha - \varepsilon) - 1},$$

which, together with (2.1.4), implies

$$(2.1.6) \quad \liminf_{n \rightarrow \infty} (2\alpha - 1)n\gamma_n^2 \sum_{m=n_0}^n (m\gamma_m)^{-2} \geq 1.$$

Since by (2.1.2) and (2.1.3)  $n\gamma_n^2 \leq C_1 n^{1-2(\alpha-\varepsilon)}$  for all  $n \geq n_3$  and some constant  $C_1 > 0$ , using  $2(\alpha-\varepsilon) > 1$  we get

$$(2.1.7) \quad n\gamma_n^2 \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

It follows from (2.1.2), (2.1.3) and (2.1.5) that

$$\gamma_n^2 \sum_{m=n_0}^n (m\gamma_m)^{-2}$$

$$\leq \gamma_n^2 \sum_{m=n_0}^{n_3-1} (m\gamma_m)^{-2} + (1 + \eta)^3 (2(\alpha - \varepsilon) - 1)^{-1} n^{-1} \quad \text{for all } n \geq n_3,$$

which, together with (2.1.7), yields

$$(2.1.8) \quad \limsup_{n \rightarrow \infty} (2\alpha - 1)n\gamma_n^2 \sum_{m=n_0}^n (m\gamma_m)^{-2} \leq 1.$$

Thus by (2.1.6) and (2.1.8) we obtain (2.1.1). This completes the proof.

The following proposition gives the a.s. and mean square convergences of  $X_n - \theta_n$  to zero.

PROPOSITION 2.2. *Suppose that the following conditions are fulfilled:*

*There exist three sequences of positive numbers  $\{\varepsilon_n\}$ ,  $\{\rho_n\}$  and  $\{A_n\}$ , and a sequence of real numbers  $\{b_n\}$  such that*

$$(2.2.1) \quad (x - \theta_n)M_n(x) > 0 \quad \text{if } |x - \theta_n| \geq \varepsilon_n \text{ for all } n \geq 1,$$

$$(2.2.2) \quad \inf \{|M_n(x)| : \varepsilon_n < |x - \theta_n| < \varepsilon_n^{-1}\} \geq \rho_n \quad \text{for all } n \geq 1,$$

$$(2.2.3) \quad |M_n(x)| \leq A_n(1 + |x - \theta_n|) \quad \text{for all } x \in R \text{ and all } n \geq 1,$$

$$(2.2.4) \quad \sum_{n=1}^{\infty} a_n^2 E[(M_n(X_n) - Y_n(X_n))^2] < \infty,$$

$$(2.2.5) \quad \lim_{n \rightarrow \infty} \varepsilon_n = 0, \quad 0 < \varepsilon_n \leq 1 \quad \text{for all } n \geq 1, \quad \lim_{n \rightarrow \infty} a_n A_n = 0,$$

$$\sum_{n=1}^{\infty} a_n \rho_n = \infty, \quad \sum_{n=1}^{\infty} a_n |b_n| < \infty$$

and

$$(2.2.6) \quad \sum_{n=1}^{\infty} |\theta_n - \theta_{n+1}| < \infty.$$

Then we have

$$\lim_{n \rightarrow \infty} |X_n - \theta_n| = 0 \quad \text{a.s.}$$

and

$$\lim_{n \rightarrow \infty} E[(X_n - \theta_n)^2] = 0.$$

PROOF. Let  $Z_n = X_n - \theta_n$  for each  $n \geq 1$ . By (2.4), (2.2.3), (2.2.4) and  $E[X_1^2] < \infty$  we get

$$(2.2.7) \quad E[X_n^2] < \infty \quad \text{for all } n \geq 1,$$

which implies

$$(2.2.8) \quad E[Z_n^2] < \infty \quad \text{for all } n \geq 1.$$

It follows from (2.4) that

$$(2.2.9) \quad Z_{n+1} = T_n(Z_1, \dots, Z_n) + U_n \quad \text{for each } n \geq 1,$$

where

$$(2.2.10) \quad T_n(r_1, \dots, r_n) = T_n(r_n) = r_n + (\theta_n - \theta_{n+1}) - a_n M_n(r_n + \theta_n) - a_n b_n$$

for  $(r_1, \dots, r_n) \in R^n$  and  $n \geq 1$ , and  $U_n = a_n v_n$  for each  $n \geq 1$ .

By (2.5) and (2.2.4) we have

$$(2.2.11) \quad E[U_n | Z_1, \dots, Z_n] = 0 \quad \text{a.s.} \quad \text{for all } n \geq 1$$

and

$$(2.2.12) \quad \sum_{n=1}^{\infty} E[U_n^2] < \infty.$$

According to (2.2.5) there exists a positive integer  $N$  such that

$$(2.2.13) \quad 1 - a_n A_n \geq 0 \quad \text{for all } n \geq N.$$

We shall estimate  $T_n(r_1, \dots, r_n)$  given by (2.2.10). Let  $n (\geq N)$  be fixed.

Case 1 where  $|r_n| \leq \varepsilon_n$ .

By (2.2.3) we get

$$|T_n(r_n)| \leq |r_n| + |\theta_n - \theta_{n+1}| + a_n A_n (1 + |r_n|) + a_n |b_n|,$$

which, together with  $|r_n| \leq \varepsilon_n \leq 1$ , implies

$$(2.2.14) \quad |T_n(r_n)| \leq \varepsilon_n + |\theta_n - \theta_{n+1}| + 2a_n A_n + a_n |b_n|.$$

Case 2 where  $\varepsilon_n < |r_n| < \varepsilon_n^{-1}$ .

Assume that  $r_n > 0$ . Since by (2.2.1)  $M_n(r_n + \theta_n) = |M_n(r_n + \theta_n)|$ , it follows from (2.2.2) that

$$\begin{aligned} T_n(r_n) &\leq |r_n| + |\theta_n - \theta_{n+1}| - a_n |M_n(r_n + \theta_n)| + a_n |b_n| \\ &\leq |r_n| + |\theta_n - \theta_{n+1}| - a_n \rho_n + a_n |b_n|. \end{aligned}$$

On the other hand, by (2.2.3) and (2.2.13) we get

$$\begin{aligned} T_n(r_n) &\geq |r_n| - |\theta_n - \theta_{n+1}| - a_n |M_n(r_n + \theta_n)| - a_n |b_n| \\ &\geq |r_n| - |\theta_n - \theta_{n+1}| - a_n A_n (1 + |r_n|) - a_n |b_n| \\ &\geq -\{|\theta_n - \theta_{n+1}| + a_n A_n + a_n |b_n|\}. \end{aligned}$$

Thus the above relations imply

$$(2.2.15) \quad |T_n(r_n)| \leq \max\{|\theta_n - \theta_{n+1}| + a_n A_n + a_n |b_n|, |r_n| - a_n \rho_n + |\theta_n - \theta_{n+1}| + a_n |b_n|\}.$$

When  $r_n < 0$ , in the same way as above we get (2.2.15).

Case 3 where  $|r_n| \geq \varepsilon_n^{-1}$ .

In the same way as case 2 we have

$$(2.2.16) \quad |T_n(r_n)| \leq \max\{|\theta_n - \theta_{n+1}| + a_n A_n + a_n |b_n|, |r_n| + |\theta_n - \theta_{n+1}| + a_n |b_n|\}.$$

Define  $\gamma_n(r_1, \dots, r_n)$  as follows: for  $(r_1, \dots, r_n) \in R^n$

$$\gamma_n(r_1, \dots, r_n) = \gamma_n(r_n) = \begin{cases} a_n \rho_n & \text{if } |r_n| < \varepsilon_n^{-1} \\ 0 & \text{otherwise.} \end{cases}$$

By (2.2.14), (2.2.15) and (2.2.16) we obtain

$$(2.2.17) \quad |T_n(r_1, \dots, r_n)| \leq \max\{\varepsilon_n + |\theta_n - \theta_{n+1}| + 2a_n A_n + a_n |b_n|, |r_n| - \gamma_n(r_1, \dots, r_n) + |\theta_n - \theta_{n+1}| + a_n |b_n|\}$$

for all  $(r_1, \dots, r_n) \in R^n$  and all  $n \geq N$ . It follows from (2. 2. 5) and (2. 2. 6) that

$$(2. 2. 18) \quad \lim_{n \rightarrow \infty} (\varepsilon_n + |\theta_n - \theta_{n+1}| + 2a_n A_n + a_n |b_n|) = 0$$

and

$$(2. 2. 19) \quad \sum_{n=1}^{\infty} \{|\theta_n - \theta_{n+1}| + a_n |b_n|\} < \infty.$$

Since  $\lim_{n \rightarrow \infty} \varepsilon_n^{-1} = \infty$ , for any sequence of real numbers  $\{r_m\}$  with  $\sup_{m \geq 1} |r_m| < \infty$  there exists a positive integer  $N_1 (\geq N)$  such that  $\sup_{m \geq 1} |r_m| < \varepsilon_n^{-1}$  for all  $n \geq N_1$ . Hence by the definition of  $\gamma_n$  we get  $\gamma_n(r_1, \dots, r_n) = a_n \rho_n$  for all  $n \geq N_1$ , which, together with (2. 2. 5), yields

$$(2. 2. 20) \quad \sum_{n=1}^{\infty} \gamma_n(r_1, \dots, r_n) = \infty.$$

We note that Lemma 1 of WATANABE [9] remains valid by replacing any constant  $X_1$  in [9] by any random variable  $X_1$  with  $E[X_1^2] < \infty$ . Thus by (2. 2. 8), (2. 2. 9), (2. 2. 11), (2. 2. 12), (2. 2. 17) to (2. 2. 20) and Lemma 1 of WATANABE [9] we obtain the result. This completes the proof.

The following lemma gives the rate of convergence of  $E[(X_n - \theta_n)^2]$  to zero, which will be useful for proving theorems in the next section.

LEMMA 2. 3. *Assume the following conditions:*

$$(2. 3. 1) \quad (x - \theta_n)M_n(x) \geq 0 \quad \text{for all } x \in R \text{ and all } n \geq 1,$$

*there exist two sequences of positive constants  $\{K_{n1}\}$  and  $\{K_{n2}\}$  such that*

$$(2. 3. 2) \quad K_{n1}|x - \theta_n| \leq |M_n(x)| \leq K_{n2}|x - \theta_n| \quad \text{for all } x \in R \text{ and all } n \geq 1,$$

$$(2. 3. 3) \quad \lim_{n \rightarrow \infty} K_{n2}^2/n = 0$$

and

$$(2. 3. 4) \quad K_1 = \inf_{n \geq 1} K_{n1} > 0,$$

*there exists a sequence of nonnegative valued Borel measurable function  $L_n(x)$  on  $R$  such that*

$$(2. 3. 5) \quad E[(M_n(x) - Y_n(x))^2] \leq L_n(x) \quad \text{for all } x \in R \text{ and all } n \geq 1$$

and

$$(2. 3. 6) \quad l_n = o(\log_2 n),$$

where  $l_n = E[L_n(X_n)] < \infty$  and  $\log_2 n = \log(\log n)$ ,

$$(2. 3. 7) \quad |\theta_n - \theta_{n+1}| = o(n^{-3/2}(\log_2 n)^{1/2})$$

and  $a_n = An^{-1}$  where  $A$  is an arbitrary positive number with

$$(2. 3. 8) \quad 2AK_1 > 1.$$

Let  $\{b_n\}$  be any sequence of real numbers satisfying

$$(2.3.9) \quad b_n^2 \leq C_1(\log_2 n)/n \quad \text{for all } n \geq 3$$

$$b_1 = b_2 = 0,$$

where  $C_1$  is some positive constant. Then, Proposition 2.2 holds and there exists a positive constant  $C$  such that

$$(2.3.10) \quad E[(X_n - \theta_n)^2] \leq C(\log_2 n)/n \quad \text{for all } n \geq 3.$$

PROOF. Since  $\sum_{n=1}^{\infty} a_n = \infty$  and  $\sum_{n=1}^{\infty} K_{n1}/n = \infty$  by (2.3.4), there exists a sequence of positive numbers  $\{\pi_n\}$  such that

$$\pi_n \leq 1 \quad \text{for all } n \geq 1, \quad \lim_{n \rightarrow \infty} \pi_n = 0, \quad \sum_{n=1}^{\infty} (K_{n1}\pi_n)/n = \infty$$

and  $\sum_{n=1}^{\infty} a_n \pi_n = \infty$ . In Proposition 2.2 let  $\varepsilon_n = \pi_n$ ,  $\rho_n = K_{n1}\pi_n$ ,  $A_n = K_{n2}$ . Then by the conditions of this lemma all the conditions of Proposition 2.2 are fulfilled. Thus Proposition 2.2 holds. Next, we shall show (2.3.10). Let  $Z_n = X_n - \theta_n$  for each  $n \geq 1$ . By (2.4) and (2.5) we get

$$(2.3.11) \quad \begin{aligned} E[Z_{n+1}^2] = & E[Z_n^2] + (\theta_n - \theta_{n+1})^2 + a_n^2 E[M_n^2(X_n)] + a_n^2 E[v_n^2] + a_n^2 b_n^2 \\ & + 2(\theta_n - \theta_{n+1})E[Z_n] + 2a_n^2 b_n E[M_n(X_n)] - 2a_n E[Z_n M_n(X_n)] \\ & - 2a_n b_n E[Z_n] - 2a_n(\theta_n - \theta_{n+1}) E[M_n(X_n)] - 2a_n b_n(\theta_n - \theta_{n+1}) \end{aligned}$$

for all  $n \geq 1$ . Define  $Q_n(x)$  as follows: for each  $n \geq 1$

$$Q_n(x) = \begin{cases} M_n(x)/(x - \theta_n) & \text{if } x \neq \theta_n \\ \alpha_n & \text{if } x = \theta_n, \end{cases}$$

where  $\alpha_n$  is an arbitrary constant with  $K_{n1} \leq \alpha_n \leq K_{n2}$ . Hence

$$(2.3.12) \quad M_n(x) = Q_n(x)(x - \theta_n) \quad \text{for all } x \in R \text{ and all } n \geq 1.$$

Here, by (2.3.1) and (2.3.2) we get

$$(2.3.13) \quad K_{n1} \leq Q_n(x) \leq K_{n2} \quad \text{for all } x \in R \text{ and all } n \geq 1.$$

It follows from (2.3.12) and (2.3.13) that

$$(2.3.14) \quad E[M_n^2(X_n)] \leq K_{n2}^2 E[Z_n^2] \quad \text{for all } n \geq 1$$

and

$$(2.3.15) \quad |E[M_n(X_n)]| \leq K_{n2} E[|Z_n|] \quad \text{for all } n \geq 1.$$

By (2.3.15) and the inequality

$$(2.3.16) \quad 2|ab| \leq ka^2 + k^{-1}b^2 \quad \text{for any } a, b \in R \text{ and any } k > 0,$$

we have

$$(2.3.17) \quad 2a_n^2 |b_n E[M_n(X_n)]| \leq K_{n2} a_n^2 E[Z_n^2] + K_{n2} a_n^2 b_n^2 \quad \text{for all } n \geq 1.$$

By (2.3.5) we get

$$(2.3.18) \quad E[v_n^2] \leq l_n \quad \text{for all } n \geq 1.$$

Since by (2.3.12) and (2.3.13)  $E[Z_n M_n(X_n)] \geq K_{n1} E[Z_n^2]$ , we have

$$(2.3.19) \quad -2a_n E[Z_n M_n(X_n)] \leq -2K_{n1} a_n E[Z_n^2] \quad \text{for all } n \geq 1.$$

From (2.3.8) there exists a positive number  $\varepsilon$  such that  $2AK_1(1-\varepsilon) > 1$ . It follows from (2.3.12), (2.3.13) and (2.3.16) that for all  $n \geq 1$

$$(2.3.20) \quad 2a_n |b_n E[Z_n]| \leq 2^{-1} K_1 \varepsilon a_n E[Z_n^2] + 2(K_1 \varepsilon)^{-1} a_n b_n^2,$$

$$(2.3.21) \quad 2a_n |\theta_n - \theta_{n+1}| |E[M_n(X_n)]| \\ \leq 2^{-1} K_1 \varepsilon a_n E[Z_n^2] + 2K_{n2}^2 (K_1 \varepsilon)^{-1} a_n (\theta_n - \theta_{n+1})^2,$$

$$(2.3.22) \quad 2a_n |b_n| |\theta_n - \theta_{n+1}| \leq a_n b_n^2 + a_n (\theta_n - \theta_{n+1})^2$$

and

$$(2.3.23) \quad 2|\theta_n - \theta_{n+1}| |E[Z_n]| \leq 2^{-1} K_1 \varepsilon a_n E[Z_n^2] + 2(K_1 \varepsilon)^{-1} a_n^{-1} (\theta_n - \theta_{n+1})^2.$$

Hence by (2.3.11), (2.3.14), (2.3.17) to (2.3.23) and the assumption that  $K_{n1} \geq K_1$  for all  $n \geq 1$ , we have

$$(2.3.24) \quad E[Z_{n+1}^2] \leq (1 - 2AK_1 n^{-1} \delta_n) E[Z_n^2] \\ + \{1 + 2K_{n2}^2 (K_1 \varepsilon)^{-1} a_n + a_n + 2(K_1 \varepsilon)^{-1} a_n^{-1}\} (\theta_n - \theta_{n+1})^2 \\ + a_n^2 l_n + (1 + K_{n2}) a_n^2 b_n^2 + (1 + 2(K_1 \varepsilon)^{-1}) a_n b_n^2,$$

where  $\delta_n = 1 - (3\varepsilon/4) - K_{n2}^2 (2K_1)^{-1} a_n - K_{n2} (2K_1)^{-1} a_n$ . Since by (2.3.3) and (2.3.4)  $\lim_{n \rightarrow \infty} K_{n2}/n = 0$ , from (2.3.3) there exist a positive integer  $n_0 \geq 3$  and some constant  $C_2 > 0$  such that

$$(2.3.25) \quad (K_{n2}^2 + K_{n2}) (2K_1)^{-1} a_n < \varepsilon/4 \quad \text{and} \quad a_n \leq 1 \quad \text{for all } n \geq n_0$$

and

$$(2.3.26) \quad 1 + 2K_{n2}^2 (K_1 \varepsilon)^{-1} a_n + a_n + 2(K_1 \varepsilon)^{-1} a_n^{-1} \leq C_2 a_n^{-1} \quad \text{for all } n \geq 1.$$

By (2.3.9), (2.3.24), (2.3.25) and (2.3.26) we get

$$(2.3.27) \quad E[Z_{n+1}^2] \leq (1 - \xi n^{-1}) E[Z_n^2] + C_3 n (\theta_n - \theta_{n+1})^2 + C_3 n^{-2} l_n \\ + C_3 n^{-2} \log_2 n \quad \text{for all } n \geq n_0,$$

where  $\xi = 2AK_1(1-\varepsilon) (> 1)$ . Let  $d_n = E[Z_n^2]$  and

$$\tilde{\beta}_{mn} = \begin{cases} \prod_{j=m+1}^n (1 - \xi j^{-1}) & \text{if } n > m \\ 1 & \text{if } n = m. \end{cases}$$

It is clear that there exist a positive integer  $n_1 (> n_0)$  and a positive constant  $C_4$  such that

$$(2.3.28) \quad 0 < \tilde{\beta}_{mn} \leq C_4 n^{-\xi m^\varepsilon} \quad \text{for all } n \geq m \geq n_1 - 1.$$



Since by (2. 2. 8) and (2. 3. 27)

$$d_{n+1} \leq \tilde{\beta}_{n_1-1, n} d_{n_1} + C_3 \sum_{m=n_1}^n \tilde{\beta}_{mn} m (\theta_m - \theta_{m+1})^2 \\ + C_3 \sum_{m=n_1}^n \tilde{\beta}_{mn} m^{-2} l_m + C_3 \sum_{m=n_1}^n \tilde{\beta}_{mn} m^{-2} \log_2 m \quad \text{for all } n \geq n_1$$

and  $d_{n_1} < \infty$ , using (2. 3. 28) we have

$$(2. 3. 29) \quad d_{n+1} \leq C_5 \{n^{-\xi} + n^{-\xi} \sum_{m=n_1}^n m^{\xi+1} (\theta_m - \theta_{m+1})^2 \\ + n^{-\xi} \sum_{m=n_1}^n m^{\xi-2} l_m + n^{-\xi} \sum_{m=n_1}^n m^{\xi-2} \log_2 m\} \quad \text{for all } n \geq n_1.$$

According to Lemma 4 of SACKS [8], (2. 3. 6), (2. 3. 7) and the fact that  $\xi > 1$  we get

$$n^{1-\xi} (\log_2 n)^{-1} \sum_{m=n_1}^n m^{\xi-2} \log_2 m \leq C_6, \\ n^{1-\xi} (\log_2 n)^{-1} \sum_{m=n_1}^n m^{\xi-2} l_m \leq C_6, \\ n^{1-\xi} (\log_2 n)^{-1} \sum_{m=n_1}^n m^{\xi+1} (\theta_m - \theta_{m+1})^2 \leq C_6$$

and

$$n^{1-\xi} (\log_2 n)^{-1} \leq C_6$$

for all  $n \geq n_1$ , which, together with (2. 3. 29), imply

$$d_{n+1} / (n^{-1} \log_2 n) \leq C_7 \quad \text{for all } n \geq n_1.$$

Thus there exists a positive constant  $C$  such that

$$d_n \leq C n^{-1} \log_2 n \quad \text{for all } n \geq 3.$$

This completes the proof.

### 3. Results

In this section the results of the previous section are used to prove the main results. Assume the following conditions:

$$(3. 1) \quad |\theta_n - \theta_{n+1}| = o(n^{-3/2} (\log_2 n)^{1/2}),$$

$$(3. 2) \quad M_n(x) = \alpha_1(n)(x - \theta_n) + \alpha_2(n)(x - \theta_n)^2 + o((x - \theta_n)^2)$$

for each  $n \geq 1$ , where  $\alpha_1(n) > 0$  and  $\alpha_2(n)$  are constants depending only on  $n$  with

$$(3. 3) \quad \lim_{n \rightarrow \infty} \alpha_1(n) = \alpha_1 > 0$$

and

$$(3. 4) \quad \sum_{n=3}^{\infty} (n^{-3} \log_2 n)^{1/2} |\alpha_2(n)| < \infty,$$

and  $o(x^2)$  means " $o(x^2)/x^2 \rightarrow 0$  as  $|x| \rightarrow 0$ ",

there exists a sequence of positive numbers  $\{\eta_n\}$  such that

$$(3.5) \quad \bar{R} = \sup_{n \geq 1} \sup_{x \in R} E[|V_n(x)|^{2+\eta_n}] < \infty$$

and

$$(3.6) \quad 0 < \underline{\eta} \leq \bar{\eta} < \infty,$$

where  $V_n(x) = M_n(x) - Y_n(x)$ ,  $\underline{\eta} = \inf_{n \geq 1} \eta_n$  and  $\bar{\eta} = \sup_{n \geq 1} \eta_n$ ,

there exists a finite positive constant  $\sigma^2$  such that for any sequence of real numbers  $\{x_n\}$  with  $x_n - \theta_n \rightarrow 0$  as  $n \rightarrow \infty$

$$(3.7) \quad E[V_n^2(x_n)] \rightarrow \sigma^2 \quad \text{as } n \rightarrow \infty.$$

The following result concerning the rate of  $x_n - \theta_n$  to zero is analogous to that of HEYDE [4].

**THEOREM 3.1.** *Let  $\{b_n\}$  be any sequence of real numbers satisfying*

$$(3.1.1) \quad b_n = o((n^{-1} \log_2 n)^{1/2}) \\ b_1 = b_2 = 0.$$

*Assume all the conditions of Lemma 2.3 except for (2.3.7) and (2.3.9), and the conditions (3.1) to (3.7). Then we have*

$$X_n = \theta_n + A\sigma(2A\alpha_1(n) - 1)^{-1/2} \zeta_n (2n^{-1} \log_2 n)^{1/2} \quad \text{a.s. for all } n \geq 3,$$

*where the sequence of random variables  $\{\zeta_n, n \geq 3\}$  has its set of a.s. limit points confined to  $[-1, 1]$  with*

$$\limsup_{n \rightarrow \infty} \zeta_n = 1 \quad \text{a.s., and} \quad \liminf_{n \rightarrow \infty} \zeta_n = -1 \quad \text{a.s.}$$

**PROOF.** We shall proceed the proof along that of Theorem 2.1 in HEYDE [4]. We note that under the conditions of this theorem all the conditions of Lemma 2.3 are fulfilled. It follows from (2.3.2), (2.3.8), (3.2) and (3.3) that

$$(3.1.2) \quad \alpha_1(n) \geq K_1 \quad \text{for all } n \geq 1$$

and

$$(3.1.3) \quad 2A\alpha_1 \geq 2AK_1 > 1.$$

Let  $\delta_n = \alpha_2(n)(X_n - \theta_n)^2 + o((X_n - \theta_n)^2)$  and  $Z_n = X_n - \theta_n$ . By (2.4) and (3.2) we get

$$(3.1.4) \quad Z_{n+1} = \beta_{2n} Z_3 + \sum_{m=3}^n \beta_{mn} (\theta_m - \theta_{m+1}) - A \sum_{m=3}^n \beta_{mn} m^{-1} \delta_m \\ + A \sum_{m=3}^n \beta_{mn} m^{-1} v_m - AD_n \quad \text{for all } n \geq 3,$$

where

$$\beta_{mn} = \begin{cases} \prod_{j=m+1}^n (1 - A\alpha_1(j)j^{-1}) & \text{if } n > m \\ 1 & \text{if } n = m \end{cases}$$

and

$$(3.1.5) \quad D_n = \sum_{m=3}^n \beta_{mn} m^{-1} b_m.$$

From (3.3) and (3.1.2) there exists a positive integer  $n_0$  such that

$$0 < 1 - A\alpha_1(j)j^{-1} \leq 1 - AK_1j^{-1} \quad \text{for all } j \geq n_0,$$

by which we get

$$(3.1.6) \quad |\beta_{mn}| \leq C_2 m^{AK_1} n^{-AK_1} \quad \text{for all } n \geq m \geq 1$$

and

$$(3.1.7) \quad \beta_{mn} = \gamma_n \gamma_m^{-1} \quad \text{for all } n \geq m \geq n_0,$$

where

$$\gamma_n = \prod_{j=n_0}^n (1 - A\alpha_1(j)j^{-1}) (> 0) \quad \text{for all } n \geq n_0.$$

Since by (2.2.8)  $|Z_3| < \infty$  a.s., using (3.1.3) and (3.1.6) we have

$$(3.1.8) \quad \beta_{2n} Z_3 = o((n^{-1} \log_2 n)^{1/2}) \quad \text{a.s.}$$

It follows from (3.1.6) that

$$(3.1.9) \quad \begin{aligned} & n^{1/2} (\log_2 n)^{-1/2} \sum_{m=3}^n |\beta_{mn}| |\theta_m - \theta_{m+1}| \\ & \leq C_2 n^{1/2 - AK_1} (\log_2 n)^{-1/2} \sum_{m=3}^n m^{AK_1} |\theta_m - \theta_{m+1}|. \end{aligned}$$

According to Lemma 4 of SACKS [8] and (3.1.3) we get

$$\sum_{m=3}^n m^{AK_1 - 3/2} (\log_2 m)^{1/2} \sim \left( AK_1 - \frac{1}{2} \right)^{-1} n^{AK_1 - 1/2} (\log_2 n)^{1/2} \quad \text{as } n \rightarrow \infty,$$

which, together with (3.1) and the Toeplitz lemma (LOÈVE [7], page 238), implies that the right hand side (R. H. S.) of (3.1.9) converges to zero as  $n$  tends to infinity.

Hence

$$(3.1.10) \quad \sum_{m=3}^n \beta_{mn} (\theta_m - \theta_{m+1}) = o((n^{-1} \log_2 n)^{1/2}).$$

In the same manner as (3.1.10), by using (3.1.1) we get

$$(3.1.11) \quad D_n = o((n^{-1} \log_2 n)^{1/2}).$$

By (3.1.6) we have

$$(3.1.12) \quad \begin{aligned} & n^{1/2} (\log_2 n)^{-1/2} \sum_{m=3}^n |\beta_{mn}| m^{-1} |\alpha_2(m)| Z_m^2 \\ & \leq C_3 n^{1/2 - AK_1} (\log_2 n)^{-1/2} \sum_{m=3}^n m^{AK_1 - 1/2} (\log_2 m)^{1/2} \times (m \log_2 m)^{-1/2} |\alpha_2(m)| Z_m^2. \end{aligned}$$

Since by (2.3.10) and (3.4)

$$\sum_{n=3}^{\infty} (n \log_2 n)^{-1/2} |\alpha_2(n)| E[Z_n^2] < \infty,$$

it holds that

$$\sum_{n=3}^{\infty} (n \log_2 n)^{-1/2} |\alpha_2(n)| Z_n^2 < \infty \quad \text{a.s.},$$

which, together with (3.1.3), (3.1.12) and the Kronecker lemma (LOÈVE [7], page 238), implies that the R. H. S. of (3.1.12) converges to zero as  $n$  tends to infinity. Hence

$$(3.1.13) \quad \sum_{m=3}^n \beta_{mn} m^{-1} \alpha_2(m) Z_m^2 = o((n^{-1} \log_2 n)^{1/2}) \quad \text{a.s.}$$

Since by Lemma 2.3  $Z_m \rightarrow 0$  a.s. as  $m \rightarrow \infty$ , there exists a positive constant  $C_4$ , possibly depending on  $\omega$ , such that

$$|o(Z_m^2)| \leq C_4 Z_m^2 \quad \text{a.s.} \quad \text{for all } m \geq 1,$$

which, together with (3.1.6), yields

$$\begin{aligned} & n^{1/2} (\log_2 n)^{-1/2} \sum_{m=3}^n |\beta_{mn} m^{-1} o(Z_m^2)| \\ & \leq C_5 n^{1/2 - AK_1} (\log_2 n)^{-1/2} \sum_{m=3}^n m^{AK_1 - 1/2} (\log_2 m)^{1/2} \times (m \log_2 m)^{-1/2} Z_m^2. \end{aligned}$$

It follows from (2.3.10) that

$$\sum_{n=3}^{\infty} (n \log_2 n)^{-1/2} E[Z_n^2] < \infty,$$

which implies

$$\sum_{n=3}^{\infty} (n \log_2 n)^{-1/2} Z_n^2 < \infty \quad \text{a.s.}$$

Thus, in the same manner as (3.1.13) we get

$$\sum_{m=3}^n \beta_{mn} m^{-1} o(Z_m^2) = o((n^{-1} \log_2 n)^{1/2}) \quad \text{a.s.},$$

which, together with (3.1.13), yields

$$(3.1.14) \quad \sum_{m=3}^n \beta_{mn} m^{-1} \delta_m = o((n^{-1} \log_2 n)^{1/2}) \quad \text{a.s.}$$

In view of (3.1.4), (3.1.8), (3.1.10), (3.1.11) and (3.1.14), it suffices, in order to prove the theorem, to show the following:

$$(3.1.15) \quad \limsup_{n \rightarrow \infty} \zeta'_n = 1 \quad \text{a.s.},$$

$$(3.1.16) \quad \liminf_{n \rightarrow \infty} \zeta'_n = -1 \quad \text{a.s.}$$

and  $\{\zeta'_n, n \geq 3\}$  has its set of a.s. limit points confined to  $[-1, 1]$ , where

$$\zeta'_n = \sigma^{-1} (2A \alpha_1(n) - 1)^{1/2} (2n^{-1} \log_2 n)^{-1/2} \sum_{m=3}^n \beta_{mn} m^{-1} v_m \quad \text{for } n \geq 3.$$

Let  $U_n = n^{-1} \gamma_n^{-1} v_n$  for all  $n \geq n_0$ . Since  $\{v_n\}$  is a martingale difference,  $\{U_n\}$  is also a martingale difference.

Let

$$(3.1.17) \quad s_n^2 = E \left[ \left( \sum_{m=n_0}^n U_m \right)^2 \right] = \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} E[v_m^2] \quad \text{for all } n \geq n_0.$$

Since  $X_m - \theta_m \rightarrow 0$  a.s. as  $m \rightarrow \infty$ , it follows from (3.7) that

$$(3.1.18) \quad \lim_{m \rightarrow \infty} E[v_m^2 | \mathcal{F}_{m-1}] = \lim_{m \rightarrow \infty} E[V_m^2(X_m) | X_m] = \sigma^2 \quad \text{a.s.},$$

where  $\mathcal{F}_m$  is the  $\sigma$ -field generated by  $X_1, \dots, X_{m+1}$ . On the other hand, by (3.5) and the Hölder inequality we get

$$E[v_m^2 | \mathcal{F}_{m-1}] \leq \bar{R}^{2/(2+\eta_m)} \leq \max\{1, \bar{R}\} < \infty \quad \text{for all } m \geq 1,$$

which implies

$$\sup_{m \geq 1} E[v_m^2 | \mathcal{F}_{m-1}] \leq \max\{1, \bar{R}\}.$$

Hence by the bounded convergence theorem and (3.1.18) we have

$$(3.1.19) \quad \lim_{n \rightarrow \infty} E[v_n^2] = \sigma^2.$$

From (3.1.3) and Lemma 2.1 we get

$$(3.1.20) \quad \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} \sim (2A\alpha_1 - 1)^{-1} (n\gamma_n^2)^{-1} \quad \text{as } n \rightarrow \infty,$$

which, together with (3.1.3), (3.1.6) and (3.1.7), implies

$$(3.1.21) \quad \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} \rightarrow \infty \quad \text{as } n \rightarrow \infty.$$

It follows from (3.1.17), (3.1.19), (3.1.20), (3.1.21) and the Toeplitz lemma that

$$(3.1.22) \quad s_n^2 \sim \sigma^2 \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} \sim \sigma^2 (2A\alpha_1 - 1)^{-1} (n\gamma_n^2)^{-1} \quad \text{as } n \rightarrow \infty.$$

Set

$$\zeta_n'' = (2s_n^2 \log_2 s_n^2)^{-1/2} \sum_{m=n_0}^n U_m \quad \text{for } n \geq n_0.$$

By Theorem 1 of HEYDE and SCOTT [5] it will hold that

$$(3.1.23) \quad \limsup_{n \rightarrow \infty} \zeta_n'' = 1 \quad \text{a.s.},$$

$$(3.1.24) \quad \liminf_{n \rightarrow \infty} \zeta_n'' = -1 \quad \text{a.s.}$$

and  $\{\zeta_n'', n \geq n_0\}$  has its set of a.s. limit points confined to  $[-1, 1]$  if the following conditions are fulfilled:

$$(3.1.25) \quad \sum_{n=n_0}^{\infty} s_n^{-4} E[U_n^4 I(|U_n| < \delta s_n)] < \infty \quad \text{for some } \delta > 0,$$

$$(3.1.26) \quad \sum_{n=n_0}^{\infty} s_n^{-1} E[|U_n| I(|U_n| \geq \varepsilon s_n)] < \infty \quad \text{for all } \varepsilon > 0$$

and

$$(3.1.27) \quad s_n^{-2} \sum_{m=n_0}^n U_m^2 \rightarrow 1 \quad \text{a.s. as } n \rightarrow \infty,$$

where  $I(A)$  denotes the indicator function of the set  $A$ . First we shall show (3.1.25).

By (3. 1. 22) it suffices to show that

$$(3. 1. 28) \quad \sum_{n=n_0}^{\infty} n^{-2} E [v_n^4 I(|v_n| < \delta n^{1/2})] < \infty \quad \text{for some } \delta > 0.$$

We may assume that  $0 < \eta_n < 2$  for all  $n \geq 1$  in (3. 5). Let  $\delta (0 < \delta < 1)$  be fixed. From (3. 5) and (3. 6) it is clear that

$$\begin{aligned} & \sum_{n=n_0}^{\infty} n^{-2} E [v_n^4 I(|v_n| < \delta n^{1/2})] \\ & \leq \sum_{n=n_0}^{\infty} \delta^{2-\eta_n} n^{-(1+\eta_n/2)} E[|v_n|^{2+\eta_n}] \\ & \leq \bar{R} \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} < \infty, \end{aligned}$$

which concludes (3. 1. 28). Next we shall show (3. 1. 26). By (3. 1. 22) it suffices to show that

$$(3. 1. 29) \quad \sum_{n=n_0}^n n^{-1/2} E [ |v_n| I(|v_n| \geq \varepsilon n^{1/2}) ] < \infty \quad \text{for all } \varepsilon > 0.$$

From (3. 5) and (3. 6) it is clear that

$$\begin{aligned} & \sum_{n=n_0}^n n^{-1/2} E [ |v_n| I(|v_n| \geq \varepsilon n^{1/2}) ] \\ & \leq \max \{1, \varepsilon^{-(1+\bar{\eta})}\} \bar{R} \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} < \infty, \end{aligned}$$

which concludes (3. 1. 29). Finally we shall show (3. 1. 27). It is clear that

$$(3. 1. 30) \quad s_n^{-2} \sum_{m=n_0}^n U_m^2 = s_n^{-2} \sum_{m=n_0}^n (U_m^2 - E[U_m^2 | \mathcal{F}_{m-1}]) + s_n^{-2} \sum_{m=n_0}^n E[U_m^2 | \mathcal{F}_{m-1}].$$

By (3. 1. 18), (3. 1. 21), (3. 1. 22) and the Toeplitz lemma we have

$$(3. 1. 31) \quad s_n^{-2} \sum_{m=n_0}^n E[U_m^2 | \mathcal{F}_{m-1}] \rightarrow 1 \quad \text{a.s. as } n \rightarrow \infty.$$

Set  $v'_m = v_m I(|v_m| \leq m^{1/2})$  for each  $m \geq n_0$ . By making use of the Kolmogorov strong law of large numbers of martingale we shall obtain

$$(3. 1. 32) \quad s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} ((v'_m)^2 - E[(v'_m)^2 | \mathcal{F}_{m-1}]) \rightarrow 0 \quad \text{a.s. as } n \rightarrow \infty$$

if it holds that

$$(3. 1. 33) \quad \sum_{n=n_0}^{\infty} s_n^{-4} n^{-4} \gamma_n^{-4} \text{Var} [(v'_n)^2] < \infty,$$

where  $\text{Var} [X]$  denotes variance of  $X$ . In order to show (3. 1. 33) it suffices, from (3. 1. 22), to show

$$\sum_{n=n_0}^{\infty} n^{-2} E[(v'_n)^4] < \infty,$$

which is satisfied by the fact that

$$\sum_{n=n_0}^{\infty} n^{-2} E[(v'_n)^4] \leq \bar{R} \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} < \infty.$$

Hence the relation (3. 1. 32) holds. By the Markov inequality we have

$$\begin{aligned} \sum_{n=n_0}^{\infty} P(v_n \neq v'_n) &\leq \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} E[|v_n|^{2+\eta_n}] \\ &\leq \bar{R} \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} < \infty, \end{aligned}$$

which implies  $P(v_n \neq v'_n \text{ i.o.})=0$ . Hence by (3. 1. 32) we get

$$(3. 1. 34) \quad s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} (v_m^2 - E[(v'_m)^2 | \mathcal{F}_{m-1}]) \rightarrow 0 \quad \text{a.s.} \quad \text{as } n \rightarrow \infty.$$

As  $v_m^2 = (v'_m)^2 + v_m^2 I(|v_m| > m^{1/2})$ , we get

$$\begin{aligned} (3. 1. 35) \quad &s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} (v_m^2 - E[v_m^2 | \mathcal{F}_{m-1}]) \\ &= s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} (v_m^2 - E[(v'_m)^2 | \mathcal{F}_{m-1}]) \\ &\quad - s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} E[v_m^2 I(|V_m| > m^{1/2}) | \mathcal{F}_{m-1}]. \end{aligned}$$

Since

$$\sum_{n=n_0}^{\infty} n^{-1} E[v_n^2 I(|v_n| > n^{1/2})] \leq \bar{R} \sum_{n=n_0}^{\infty} n^{-(1+\eta/2)} < \infty,$$

it follows from (3. 1. 22) that

$$\sum_{n=n_0}^{\infty} s_n^{-2} n^{-2} \gamma_n^{-2} E[v_n^2 I(|v_n| > n^{1/2})] < \infty,$$

which implies

$$\sum_{n=n_0}^{\infty} s_n^{-2} n^{-2} \gamma_n^{-2} E[v_n^2 I(|v_n| > n^{1/2}) | \mathcal{F}_{n-1}] < \infty \quad \text{a.s.}$$

Thus, by the Kronecker lemma we obtain

$$s_n^{-2} \sum_{m=n_0}^n m^{-2} \gamma_m^{-2} E[v_m^2 I(|v_m| > m^{1/2}) | \mathcal{F}_{m-1}] \rightarrow 0 \quad \text{a.s.} \quad \text{as } n \rightarrow \infty,$$

which, together with (3. 1. 34) and (3. 1. 35), yields

$$(3. 1. 36) \quad s_n^{-2} \sum_{m=n_0}^n (U_m^2 - E[U_m^2 | \mathcal{F}_{m-1}]) \rightarrow 0 \quad \text{a.s.} \quad \text{as } n \rightarrow \infty.$$

Combining (3. 1. 30), (3. 1. 31) and (3. 1. 36) we have (3. 1. 27). It follows from (3. 3) and the Toeplitz lemma that

$$\sum_{j=n_0}^n \alpha_1(j) j^{-1} \sim \alpha_1 \log n \quad \text{as } n \rightarrow \infty.$$

After some calculations, using the above result, (3. 1. 22) and the Taylor theorem, we obtain

$$(3. 1. 37) \quad \log_2 s_n^2 \sim \log_2 n \quad \text{as } n \rightarrow \infty.$$

By (3. 5) we get  $\sup_{n \geq 1} E[v_n^2] < \infty$ , which implies that  $|v_n| < \infty$  a.s. for all  $n \geq 1$ . Hence

using (3. 3), (3. 1. 3) and (3. 1. 6) we have

$$(3. 1. 38) \quad \sigma^{-1}(2A\alpha_1(n)-1)^{1/2} (2n^{-1}\log_2 n)^{-1/2} \sum_{m=3}^{n_0-1} m^{-1} \beta_{mn} v_m = 0 \quad \text{a.s.}$$

Combining (3. 3), (3. 1. 7), (3. 1. 22), (3. 1. 23), (3. 1. 24), (3. 1. 37), (3. 1. 38) and the fact that  $\{\zeta_n'', n \geq n_0\}$  has its set of a.s. limit points confined to  $[-1, 1]$ , we obtain (3. 1. 15), (3. 1. 16) and the fact that  $\{\zeta_n', n \geq 3\}$  has its set of a.s. limit points confined to  $[-1, 1]$ . This completes the proof.

Let

$$(3. 8) \quad D_n \geq D(2n^{-1} \log_2 n)^{1/2} + o((n^{-1} \log_2 n)^{1/2}),$$

where  $b_n$ 's are nonnegative numbers satisfying (3. 1. 1),  $D_n$  is defined as (3. 1. 5) and  $D$  is any positive number such that

$$(3. 9) \quad D > \sigma(2A\alpha_1-1)^{-1/2} \quad \text{for } \sigma \text{ and } \alpha_1 \text{ in (3. 3) and (3. 7).}$$

The following theorem can be proved by making use of the relations in the proof of Theorem 3. 1.

**THEOREM 3. 2.** *Assume the conditions (3. 1) to (3. 9). Then, under all the conditions of Lemma 2.3 except for (2. 3. 7) and (2. 3. 9), we have*

$$(3. 2. 1) \quad X_n - \theta_n \rightarrow 0 \quad \text{a.s.} \quad \text{as } n \rightarrow \infty$$

and, with probability one,  $X_n - \theta_n > 0$  only finitely many time.

**PROOF.** By Lemma 2. 3 we obtain (3. 2. 1). Let  $Z_n = X_n - \theta_n$ . It follows from (3. 1. 4), (3. 1. 8), (3. 1. 10) and (3. 1. 14) that

$$Z_{n+1} = A \sum_{m=3}^n \beta_{mn} m^{-1} v_m - AD_n + o((n^{-1} \log_2 n)^{1/2}) \quad \text{a.s.} \quad \text{for all } n \geq 3,$$

which, together with (3. 3), (3. 8), (3. 9) and (3. 1. 15), implies

$$\begin{aligned} & P\{Z_n > 0 \text{ i.o.}\} \\ & \leq P\{\limsup_{n \rightarrow \infty} (\sigma^{-1}(2A\alpha_1(n)-1)^{1/2} (2n^{-1}\log_2 n)^{-1/2} \sum_{m=3}^n \beta_{mn} m^{-1} v_m \\ & \quad \geq \limsup_{n \rightarrow \infty} (\sigma^{-1}(2A\alpha_1(n)-1)^{1/2} (2n^{-1}\log_2 n)^{-1/2} D_n)\} \\ & = P\{1 \geq \sigma^{-1}(2A\alpha_1-1)^{1/2} D\} = 0. \end{aligned}$$

Thus  $P\{Z_n > 0 \text{ i.o.}\} = 0$ , which concludes that, with probability one,  $X_n - \theta_n > 0$  occur only finitely many times. This completes the proof.

We shall give an example of  $\{b_n\}$  satisfying (3. 8) and (3. 9).

**EXAMPLE.**

Let

$$b_1 = b_2 = 0$$



$$b_n = D(2n^{-1} \log_2 n)^{1/2} \quad \text{for all } n \geq 3,$$

where  $D$  is any positive number such that  $D > 3\sigma(2A\alpha_1 - 1)^{1/2}/2$  for  $\sigma$  and  $\alpha$  in (3.3) and (3.7).

We shall, now, consider two cases where ANBAR [1] and DUPAČ [2] treated. First we shall consider the case of ANBAR [1] in which it holds that

$$M_n(x) = M(x) \quad \text{for all } x \in R \text{ and all } n \geq 1$$

$$\theta_n = \theta \quad \text{for all } n \geq 1 \text{ with } \theta \text{ being the unique root of } M(x) = 0.$$

Theorem 3.2 gives

COROLLARY 3.3 *Theorem 2 of ANBAR [1] holds.*

Next, we shall consider the case of DUPAČ [2] in which  $Y(x)$  is a random variable for each  $x \in R$  with  $Y(x)$  being a Borel measurable function on  $R$ , the expectation of  $Y(x)$ ,  $M(x) = E[Y(x)]$ , exists and is a Borel measurable function on  $R$ , and

$$M_n(x) = M(x - \theta_n + \theta_1) \quad \text{for all } n \geq 1$$

where  $\theta_1$  is the unique root of  $M(x) = 0$  and  $\{\theta_n\}$  is a sequence of real numbers, so  $\theta_n$  is the unique root of the equation  $M_n(x) = 0$ .

COROLLARY 3.4. *Suppose the following conditions:*

$$(3.4.1) \quad (x - \theta_1)M(x) \geq 0 \quad \text{for all } x \in R,$$

$$(3.4.2) \quad K_1|x - \theta_1| \leq |M(x)| \leq K_2|x - \theta_1| \quad \text{for all } x \in R \text{ and some positive constants } K_1 \text{ and } K_2,$$

$$(3.4.3) \quad M(x) = \alpha_1(x - \theta_1) + \alpha_2(x - \theta_1)^2 + o((x - \theta_1)^2)$$

with  $\alpha_1 > 0$  and  $\alpha_2$  being real numbers,

there exist finite positive constants  $\eta$  and  $\sigma^2$  such that

$$(3.4.4) \quad \sup_{x \in R} E[|V(x)|^{2+\eta}] < \infty$$

and

$$(3.4.5) \quad E[V^2(x)] \rightarrow \sigma^2 \quad \text{as } x \rightarrow \theta_1 \text{ with } V(x) = M(x) - Y(x),$$

$$(3.4.6) \quad |\theta_n - \theta_{n+1}| = o(n^{-3/2}(\log_2 n)^{1/2}),$$

$$(3.4.7) \quad a_n = A n^{-1}$$

where  $A$  is an arbitrary real number with  $2AK_1 > 1$ . Assume that  $\{b_n\}$  is a sequence of nonnegative numbers satisfying (3.1.1) and

$$D_n \geq D(2n^{-1} \log_2 n)^{1/2} + o((n^{-1} \log_2 n)^{1/2})$$

where  $D_n$  is defined as (3.1.5) and  $D$  is any positive number with  $D > \sigma(2AK_1 - 1)^{-1/2}$ . Then Theorem 3.2 holds.

PROOF. In Theorem 3.1 set  $\alpha_1(n)=\alpha_1$ ,  $\alpha_2(n)=\alpha_2$ ,  $\eta_n=\eta$ ,  $V_n(x)=V(x-\theta_n+\theta_1)$  for all  $n \geq 1$ . Then all the conditions (3.1) to (3.8) are satisfied. From (3.4.2) and (3.4.3) we get

$$D > \sigma(2AK_1-1)^{-1/2} \geq \sigma(2A\alpha_1-1)^{-1/2},$$

by which (3.9) is fulfilled. The relation (3.4.1) implies the relation (2.3.1). In Lemma 2.3 set  $K_{n1}=K_1$  and  $K_{n2}=K_2$  for all  $n \geq 1$ . By the Hölder inequality we get

$$E[(M_n(x)-Y_n(x))^2] \leq \{E[|V(x-\theta_n+\theta_1)|^{2+\eta}]\}^{2/(2+\eta)}.$$

Set  $L_n(x) = \{E[|V(x-\theta_n+\theta_1)|^{2+\eta}]\}^{2/(2+\eta)}$  in (2.3.5). Since  $l_n \leq \{\sup_{x \in R} E[|V(x)|^{2+\eta}]\}^{2/(2+\eta)}$ , we have (2.3.6) from (3.4.4). Another conditions of Lemm 2.3 are all satisfied. Thus, since all the conditions of Theorem 3.2 are fulfilled, Theorem 3.2 holds. This completes the proof.

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