

Strong uniform consistency of recursive kernel density estimators*

By
Eiichi ISOGAI

1. Introduction

Let $f(x)$ be a (unknown) probability density function (p.d.f.) on the p -dimensional Euclidean space R^p with respect to Lebesgue measure. Based on a sequence X_1, X_2, \dots of independent identically distributed p -dimensional random vectors having the common p.d.f. $f(x)$, we wish to estimate the p.d.f. $f(x)$. Yamato [8] proposed recursive kernel estimators of the form

$$(1.1) \quad \begin{aligned} \tilde{f}_0(x) &\equiv 0 \\ \tilde{f}_n(x) &= \tilde{f}_{n-1}(x) + n^{-1} \{K_n(x, X_n) - \tilde{f}_{n-1}(x)\} \quad \text{for each } n \geq 1, \end{aligned}$$

where

$$(1.2) \quad K_n(x, y) = h_n^{-p} K((x-y)/h_n) \quad \text{for } x, y \in R^p \text{ and each } n \geq 1,$$

$\{h_n\}$ is a sequence of positive numbers and $K(x)$ is a real-valued Borel measurable function on R^p , on which certain properties were imposed. He showed the weak uniform consistency of these estimators as well as the weak pointwise consistency. DEVROYE [4] discussed several results related to the weak or the strong pointwise consistency of $\tilde{f}_n(x)$. DAVIES [3] showed the strong uniform consistency of $\tilde{f}_n(x)$ as well as the strong pointwise consistency.

In this paper we consider a class of recursive kernel estimators of the form

$$(1.3) \quad \begin{aligned} f_0(x) &\equiv K(x) \\ f_n(x) &= f_{n-1}(x) + a_n \{K_n(x, X_n) - f_{n-1}(x)\} \quad \text{for each } n \geq 1, \end{aligned}$$

or equivalently,

$$f_n(x) = \sum_{m=0}^n a_m \beta_{mn} K_m(x, X_m) \quad \text{for each } n \geq 0,$$

where $K_n(x, y)$ is defined by (1.2), $K_0(x, X_0) \equiv K(x)$, $\{a_n\}$ is a sequence of positive numbers satisfying

$$(1.4) \quad a_0 = 1, 0 < a_n \leq 1 \text{ for all } n \geq 1, \lim_{n \rightarrow \infty} a_n = 0 \text{ and } \sum_{n=1}^{\infty} a_n = \infty,$$

* Received July 15, 1981; revised Oct. 31, 1981.

and let

$$(1.5) \quad \beta_{mn} = \prod_{j=m+1}^n (1-a_j) \quad \text{if } n > m \geq 0 \\ = 1 \quad \text{if } n=m \geq 0.$$

We note that if $K(x)$ is a p.d.f., then $f_n(x)$ is a p.d.f. for each $n \geq 0$. It is easy to see that the class of our estimators contains both the estimators of YAMATO (1.1) and the estimators of ISOGAI [5] with $a_n = an^{-1}$ for $2^{-1} < a \leq 1$ and each $n \geq 1$. Putting $a_n = an^{-1}$ for $2^{-1} < a \leq 1$, the differences of properties between the estimators of (1.1) and (1.3) were discussed in ISOGAI [5].

In Section 2 we shall make some preparations for sections that follow. In Section 3 we shall show the weak or the strong pointwise consistency of $f_n(x)$. In Section 4 we shall show that $f_n(x)$ is strongly uniformly consistent and that $E[\sup_{x \in R^p} |f_n(x) - f(x)|^2]$ converges to 0 as n tends to infinity. We also consider a problem of estimating the mode θ of $f(x)$.

2. Preliminaries and auxiliary results

Let $K(x)$ in (1.2) be a bounded, integrable, real-valued Borel measurable function on R^p satisfying

$$(K 1) \quad \int K(x) dx = 1,$$

where all integrals with respect to Lebesgue measure are over R^p , unless otherwise specified. Let $\{h_n\}$ in (1.2) be a sequence of positive numbers satisfying $\lim_{n \rightarrow \infty} h_n = 0$. On this sequence $\{h_n\}$ we shall impose some of the following conditions:

$$(H 1) \quad a_n h_n^{-p} \xrightarrow{n} 0,$$

$$(H 2) \quad \sum_{n=1}^{\infty} a_n^2 h_n^{-p} < \infty,$$

$$(H 3) \quad h_1 \geq h_2 \geq h_3 \geq \dots,$$

$$(H 4) \quad h_n / h_{n+1} \xrightarrow{n} 1,$$

$$(H 5) \quad \sum_{n=1}^{\infty} (a_n h_n^{-p})^2 < \infty,$$

$$(H 6) \quad \sum_{n=1}^{\infty} \frac{A_n}{h_n^{2p-1}} \left| \frac{1}{h_{n+1}} - \frac{1}{h_n} \right| < \infty,$$

where a_n and β_{mn} are defined by (1.4) and (1.5) respectively, and let $A_n = \sum_{m=1}^n a_m^2 \beta_{mn}^2$ for each $n \geq 1$. Throughout this paper C_1, C_2, \dots denote positive constants, and let $\|g\|_{\infty} = \sup |g(x)|$ for any real-valued function g on R^p , where supremum is taken over R^p , unless otherwise specified.

DEFINITION 2.1 Let $g(x)$ be a real-valued Borel measurable function on R^p . Then a point x is said to be a Lebesgue point of g if it holds that

$$\rho^{-p} \int_{S(x, \rho)} |g(y) - g(x)| dy \longrightarrow 0 \quad \text{as } \rho \rightarrow 0,$$

where $S(x, \rho)$ denotes a closed sphere in R^p centered at x with radius ρ .

REMARK 2.1. If $g(x)$ is integrable then almost every x is a Lebesgue point of g (see STEIN [7], p. 25). If x is a continuity point of g then x is a Lebesgue point of g .

The following lemma is a modification of Lemma 2 of DEVROYE [4].

LEMMA 2.1. Let $g(x)$ be a real-valued Borel measurable function on R^p , let $K(x)$ be a bounded, integrable, real-valued Borel measurable function on R^p , let $\{h_n\}$ be a sequence of positive numbers converging to zero, and let Condition $A(x, g, K)$ hold.

Condition $A(x, g, K)$. One of the following is true:

(i) x is a continuity point of g , g is integrable and

(K 2) $\|y\|^p |K(y)| \longrightarrow 0$ as $\|y\| \rightarrow \infty$,

where $\|\cdot\|$ denotes the Euclidean norm on R^p ,

(ii) x is a Lebesgue point of g and g is bounded on R^p ,

(iii) x is a Lebesgue point of g and K has compact support.

Then

$$\int h_n^{-p} K((x-y)/h_n) g(y) dy \longrightarrow g(x) \int K(y) dy.$$

PROOF. In case (i) the lemma holds by Theorem 2.1 of CACOULLOS [2]. Thus we need only show that in both cases (ii) and (iii)

$$\begin{aligned} & \left| \int h_n^{-p} K((x-y)/h_n) g(y) dy - g(x) \int K(y) dy \right| \\ & \leq \int h_n^{-p} |K((x-y)/h_n)| |g(y) - g(x)| dy \xrightarrow{n} 0. \end{aligned}$$

Let $U_n = \int h_n^{-p} |K((x-y)/h_n)| |g(y) - g(x)| dy$. Thus we need only show

$$(2.1) \quad \lim_{n \rightarrow \infty} U_n = 0.$$

For any $\rho > 0$

$$\begin{aligned} (2.2) \quad U_n &= \int_{S^c(x, \rho h_n)} h_n^{-p} |K((x-y)/h_n)| |g(y) - g(x)| dy \\ &+ \int_{S(x, \rho h_n)} h_n^{-p} |K((x-y)/h_n)| |g(y) - g(x)| dy \\ &= V_n + W_n, \text{ say,} \end{aligned}$$

where S^c denotes the complement of a set S . First we consider case (ii). Let any $\varepsilon > 0$

be fixed. Since $K(x)$ is integrable, there exists a positive constant ρ such that

$$\int_{S^c(0, \rho)} |K(y)| dy < \varepsilon. \quad \text{Such } \rho \text{ is fixed. It is easy to see that}$$

$$(2.3) \quad V_n \leq 2\|g\|_\infty \int_{S^c(0, \rho)} |K(y)| dy \leq 2\|g\|_\infty \varepsilon \quad \text{for all } n \geq 1.$$

Since x is a Lebesgue point of g and $\lim_{n \rightarrow \infty} h_n = 0$, we get

$$W_n \leq \|K\|_\infty \rho^p (\rho h_n)^{-p} \int_{S^c(0, \rho)} |g(y) - g(x)| dy \xrightarrow{n} 0,$$

which implies that

$$(2.4) \quad \lim_{n \rightarrow \infty} W_n = 0.$$

Combining (2.2), (2.3) and (2.4) we obtain (2.1). We next consider case (iii). Since $K(x)$ has compact support, there exists a positive constant ρ such that $K(y) = 0$ for all $y \in S^c(0, \rho)$, which implies that

$$(2.5) \quad V_n = \int_{S^c(0, \rho)} |K(y)| |g(x - h_n y) - g(x)| dy = 0 \quad \text{for all } n \geq 1.$$

Thus by (2.2), (2.4) and (2.5) we get (2.1). This completes the proof.

LEMMA 2.2. Condition (iii) in Condition A(x, g, K) is replaced by the following:

(iii)' x is a Lebesgue point of g , g is integrable and K has compact support.

Then, under all assumptions of Lemma 2.1

$$\sup_{n \geq 1} \int h_n^{-p} |K((x-y)/h_n)| |g(y)| dy < \infty.$$

PROOF. If $g(x)$ is integrable we get that

$$\begin{aligned} U_n &\equiv \int h_n^{-p} |K((x-y)/h_n)| |g(y)| dy \\ &\leq \|K\|_\infty h_n^{-p} \int |g(y)| dy < \infty \quad \text{for all } n \geq 1. \end{aligned}$$

If $g(x)$ is bounded we have

$$U_n \leq \|g\|_\infty \int |K(y)| dy < \infty \quad \text{for all } n \geq 1.$$

Thus, under Condition A(x, g, K)

$$(2.6) \quad U_n < \infty \quad \text{for all } n \geq 1.$$

Replacing $g(x)$ and $K(x)$ of Lemma 2.1 by $|g(x)|$ and $|K(x)|$ respectively, and using Lemma 2.1, we get

$$U_n = \int h_n^{-p} |K((x-y)/h_n)| |g(y)| dy \xrightarrow{n} |g(x)| \int |K(y)| dy,$$

which, together with (2.6), yields the lemma. The proof is complete.

3. Weak or strong pointwise consistency

In this section we shall show the weak or the strong pointwise consistency of $f_n(x)$.

THEOREM 3.1. *Suppose that Condition $A(x, f, K)$ holds. If $\{h_n\}$ satisfies (H1), then*

$$E[(f_n(x) - f(x))^2] \xrightarrow{n} 0,$$

which implies that $f_n(x) \xrightarrow{n} f(x)$ in probability.

If (H2) is true, then

$$f_n(x) \xrightarrow{n} f(x) \quad \text{with probability one (w.p. 1)}$$

and

$$E[(f_n(x) - f(x))^2] \xrightarrow{n} 0.$$

PROOF. We note that

$$(3.1) \quad |f_n(x) - f(x)| \leq |f_n(x) - E[f_n(x)]| + |E[f_n(x)] - f(x)|$$

and

$$(3.2) \quad E[(f_n(x) - f(x))^2] = E[(f_n(x) - E[f_n(x)])^2] + (E[f_n(x)] - f(x))^2.$$

It follows from (1.4) and (1.5) that

$$a_m \beta_{mn} \geq 0 \text{ for all } m=1, \dots, n, \quad n=1, 2, \dots, \quad \lim_{n \rightarrow \infty} \beta_{mn} = 0 \text{ for each } m \geq 1,$$

$$\sum_{m=1}^n a_m \beta_{mn} \leq 1 \text{ for all } n \geq 1, \quad \sum_{m=1}^n a_m \beta_{mn} \xrightarrow{n} 1.$$

Thus from Lemma 2.1 and the Toeplitz lemma (see LOÈVE [6], p. 238), we get

$$(3.3) \quad E[f_n(x)] = \sum_{m=1}^n a_m \beta_{mn} E[K_m(x, X_m)] + \beta_{0n} K(x) \xrightarrow{n} f(x).$$

Since $f_n(x) - E[f_n(x)] = \sum_{m=1}^n a_m \beta_{mn} \{K_m(x, X_m) - E[K_m(x, X_m)]\}$,

$$\begin{aligned} & E[(f_n(x) - E[f_n(x)])^2] \\ & \leq \sum_{m=1}^n a_m^2 \beta_{mn}^2 E[K_m^2(x, X_m)] \\ & \leq \|K\|_\infty \sum_{m=1}^n a_m^2 \beta_{mn}^2 h_m^{-p} \int h_m^{-p} |K((x-y)/h_m)| f(y) dy. \end{aligned}$$

From Lemma 2.2 we get

$$\int h_m^{-p} |K((x-y)/h_m)| f(y) dy \leq C_1 \quad \text{for all } m \geq 1.$$

Hence

$$(3.4) \quad E[(f_n(x) - E[f_n(x)])^2] \leq C_2 \sum_{m=1}^n a_m^2 \beta_{mn}^2 h_m^{-p}$$

$$\leq C_2 \sum_{m=1}^n a_m \beta_{mn} (a_m h_m^{-p}),$$

because of the fact that $0 \leq \beta_{mn} \leq 1$. By (H1) and the Toeplitz lemma the last term of (3.4) approaches to 0 as n tends to infinity.

Hence

$$(3.5) \quad E[(f_n(x) - E[f_n(x)])^2] \xrightarrow{n} 0,$$

which, together with (3.2) and (3.3), yields the first assertion of the lemma. The proof of the second assertion parallels the proof of Theorem 3.1 of ISOGAI [5]. Since

$$E[(f_n(x) - E[f_n(x)])^2 | X_1, \dots, X_{n-1}]$$

$$\leq (1 - a_n)(f_{n-1}(x) - E[f_{n-1}(x)])^2$$

$$+ \|K\|_\infty a_n^2 h_n^{-p} \int h_n^{-p} |K((x-y)/h_n)| f(y) dy \quad \text{w.p.1,}$$

where $E[\cdot | \cdot]$ denotes a conditional expectation, we get, by Lemma 2.2,

$$E[(f_n(x) - E[f_n(x)])^2 | X_1, \dots, X_{n-1}]$$

$$\leq (1 - a_n)(f_{n-1}(x) - E[f_{n-1}(x)])^2 + C_3 a_n^2 h_n^{-p} \quad \text{w.p.1.}$$

Hence by (H2) and Proposition 2.4 of ISOGAI [5] we have (3.5) and

$$(3.6) \quad |f_n(x) - E[f_n(x)]| \xrightarrow{n} 0 \quad \text{w.p.1.}$$

Thus combining (3.1)~(3.3), (3.5) and (3.6), the second assertion is established. The proof of the lemma is complete.

REMARK 3.1. The conditions (H1) and (H2) do not imply each other. Let

$$a_n = n^{-1},$$

$$h_n^p = n^{-1} \quad \text{if } n = 2^k \quad \text{for } k = 0, 1, 2, \dots$$

$$= n^{-\frac{1}{2}} \quad \text{if } n \neq 2^k,$$

whose sequences are given by DEVROYE [4]. These sequences satisfy (1.4) and (H2) but not (H1). Let

$$a_1 = 1, a_n = (n(\log n)^{\frac{1}{2}})^{-1} \quad \text{for } n \geq 2,$$

$$h_n^p = n^{-1}.$$

Then, these sequences satisfy (1.4) and (H1) but not (H2). Putting $a_n = an^{-1}$ with $0 < a \leq 1$ in (1.4), the conditions (H1) and (H2) coincide with (3) and (12) of DEVROYE [4] respectively. The following sequences satisfy (1.4), (H1) and (H2):

$$a_1 = 1, a_n = (n \log n)^{-1} \quad \text{for } n \geq 2, \quad h_n^p = n^{-1}.$$

4. Strong uniform consistency

In this section we shall show the strong uniform consistency of $f_n(x)$ and show that $E[\sup|f_n(x)-f(x)|^2]$ converges to 0 as n tends to infinity. Furthermore, the strong consistency of the mode estimator will be shown.

First we shall show the strong uniform consistency of $f_n(x)$ and the convergence of $E[\sup|f_n(x)-f(x)|^2]$ to 0. The method of proof is similar to that of DAVIES [3]. Let $k(t) = \int e^{it \cdot u} K(u) du$ for $t \in R^p$ be the Fourier transform of $K(u)$, where $t \cdot u = \sum_{j=1}^p t_j u_j$ for $t = (t_1, \dots, t_p)$ and $u = (u_1, \dots, u_p)$, and $i^2 = -1$. Also let $\varphi(t) = E[e^{it \cdot X_1}]$ for $t \in R^p$ be the characteristic function of the random vector X_1 . The next theorem is concerned with the strong uniform consistency of $f_n(x)$.

THEOREM 4.1. *In addition to the conditions (H3)~(H6), suppose that $K(x)$ is continuous on R^p and satisfies (K2) and*

$$(K 3) \quad \int |k(t)| dt < \infty,$$

where $|k(t)|$ is non-increasing on a ray $R(u) = \{v = qu; q > 0\}$ for each $u \neq 0 \in R^p$. (That is, $|k(v_1)| \geq |k(v_2)|$ for $v_1, v_2 \in R(u)$ with $\|v_1\| \leq \|v_2\|$). If $f(x)$ is uniformly continuous on R^p , then it holds that

$$\sup|f_n(x) - f(x)| \xrightarrow{n} 0 \quad \text{w.p.1}$$

and

$$E[\sup|f_n(x) - f(x)|^2] \xrightarrow{n} 0.$$

PROOF. It is easy to see that the theorem holds if we show that

$$(4.1) \quad \sup|E[f_n(x)] - f(x)| \xrightarrow{n} 0,$$

$$(4.2) \quad \sup|f_n(x) - E[f_n(x)]| \xrightarrow{n} 0 \quad \text{w.p.1}$$

and

$$(4.3) \quad E[(\sup|f_n(x) - E[f_n(x)]|)^2] \xrightarrow{n} 0,$$

because of the fact that $E[\sup|f_n(x) - f(x)|^2] = E[(\sup|f_n(x) - f(x)|)^2]$. First we shall show (4.1). Note that since $f(x)$ is uniformly continuous p.d.f., $f(x)$ is bounded. Since by Lemma 1 of YAMATO [8] $\sup|E[K_n(x, X_n)] - f(x)| \rightarrow 0$, using the Toeplitz lemma we get

$$\begin{aligned} & \sup|E[f_n(x)] - f(x)| \\ & \leq \sum_{m=1}^n \alpha_m \beta_{mn} (\sup|E[K_m(x, X_m)] - f(x)|) + \beta_{0n} (\|K\|_\infty + \|f\|_\infty) \xrightarrow{n} 0, \end{aligned}$$

which implies (4.1). Next we shall show (4.2) and (4.3). Since both $K(u)$ and $k(u)$ are

integrable and $K(u)$ is continuous on R^p , we have, by using the inversion theorem for the Fourier transform (see BOCHNER and CHANDRASEKHARAN [1], p. 66),

$$\begin{aligned} & K_m(x, X_m) - E[K_m(x, X_m)] \\ &= (2\pi)^{-p} \int [e^{iu \cdot X_m - \varphi(u)}] e^{-iu \cdot x} k(h_m u) du \quad \text{for all } x \in R^p \quad \text{w.p.1.} \end{aligned}$$

Hence

$$\begin{aligned} (4.4) \quad & \sup |f_n(x) - E[f_n(x)]| \\ & \leq (2\pi)^{-p} \int \left| \sum_{m=1}^n a_m \beta_{mn} [e^{iu \cdot X_m - \varphi(u)}] k(h_m u) \right| du \quad \text{w.p.1.} \end{aligned}$$

For simplicity we shall introduce the following notations:

$$\begin{aligned} S_n &= \{u \in R^p; k(h_n u) \neq 0\}, \\ \xi_{mn}(u) &= a_m \beta_{mn} [e^{iu \cdot X_m - \varphi(u)}] \quad \text{for } 1 \leq m \leq n, \\ g_{mn}(u) &= k(h_m u) / k(h_n u) \quad \text{if } u \in S_n \\ &= 0 \quad \text{if } u \in S_n^c \quad \text{for } 1 \leq m \leq n \end{aligned}$$

and

$$Z_{mn}(u) = \xi_{mn}(u) g_{mn}(u) \quad \text{for } 1 \leq m \leq n.$$

By the assumption on $|k(u)|$ and Condition (H3) we get

$$|k(h_m u)| \leq |k(h_n u)| \quad \text{for } u \in R^p \text{ and } 1 \leq m \leq n,$$

which implies that

$$(4.5) \quad S_n^c \subset S_m^c \quad \text{for } 1 \leq m \leq n$$

and

$$(4.6) \quad |g_{mn}(u)| \leq 1 \quad \text{for } 1 \leq m \leq n \text{ and } u \in R^p.$$

Hence by the definition of $g_{mn}(u)$ and (4.5) we have

$$(4.7) \quad \int \left| \sum_{m=1}^n \xi_{mn}(u) k(h_m u) \right| du = \int \left| \sum_{m=1}^n Z_{mn}(u) \right| |k(h_n u)| du \quad \text{w.p.1.}$$

In view of (4.4) and (4.7) we obtain, by the Schwarz inequality,

$$\begin{aligned} & \sup |f_n(x) - E[f_n(x)]| \\ & \leq (2\pi)^{-p} \left[\int |k(h_n u)| du \right]^{\frac{1}{2}} \left[\int \left| \sum_{m=1}^n Z_{mn}(u) \right|^2 |k(h_n u)| du \right]^{\frac{1}{2}} \\ & = (2\pi)^{-p} \left[\int |k(u)| du \right]^{\frac{1}{2}} \left[h_n^{-p} \int \left| \sum_{m=1}^n Z_{mn}(u) \right|^2 |k(h_n u)| du \right]^{\frac{1}{2}} \\ & = C_1 Y_n^{\frac{1}{2}} \quad \text{w.p.1,} \end{aligned}$$

where $Y_n = h_n^{-p} \int \left| \sum_{m=1}^n Z_{mn}(u) \right|^2 |k(h_n u)| du$.

Thus, in order to prove (4.2) and (4.3) it suffices to show that

$$(4.8) \quad Y_n \xrightarrow{n} 0 \quad \text{w.p.1}$$

and

$$(4.9) \quad E[Y_n] \xrightarrow{n} 0.$$

It is clear that

$$(4.10) \quad \begin{aligned} & Y_{n+1} \\ &= (1 - a_{n+1})^2 h_{n+1}^{-p} \int |\eta_n(u)|^2 |k(h_{n+1}u)| du \\ & \quad + a_{n+1}(1 - a_{n+1}) h_{n+1}^{-p} \left\{ \int \eta_n(u) [e^{-iu \cdot X_{n+1}} - \varphi(-u)] \overline{g_{n+1, n+1}(u)} |k(h_{n+1}u)| du \right. \\ & \quad \left. + \int \overline{\eta_n(u)} [e^{iu \cdot X_{n+1}} - \varphi(u)] g_{n+1, n+1}(u) |k(h_{n+1}u)| du \right\} \\ & \quad + a_{n+1}^2 h_{n+1}^{-p} \int |e^{iu \cdot X_{n+1}} - \varphi(u)|^2 |g_{n+1, n+1}(u)|^2 |k(h_{n+1}u)| du, \end{aligned}$$

where $\eta_n(u) = \sum_{m=1}^n \xi_{mn}(u) g_{m, n+1}(u)$ and \overline{b} denotes the conjugate of a complex number b . It follows from the independence of X_n 's that

$$E[e^{iu \cdot X_{n+1}} - \varphi(u) | X_1, \dots, X_n] = 0 \quad \text{w.p.1}$$

and

$$E[|e^{iu \cdot X_{n+1}} - \varphi(u)|^2 | X_1, \dots, X_n] = 1 - |\varphi(u)|^2 \leq 1 \quad \text{w.p.1.}$$

Hence, taking conditional expectations of both sides of (4.10) and using (4.6) we have

$$(4.11) \quad \begin{aligned} & E[Y_{n+1} | X_1, \dots, X_n] \\ & \leq (1 - a_{n+1}) h_{n+1}^{-p} \int |\eta_n(u)|^2 |k(h_{n+1}u)| du + C_2 (a_{n+1} h_{n+1}^{-p})^2 \quad \text{w.p.1.} \end{aligned}$$

By the definition of $g_{mn}(u)$ and (4.5) we can prove that

$$\begin{aligned} & \int_{S_n} \left| \sum_{m=1}^n \xi_{mn}(u) g_{m, n+1}(u) \right|^2 |k(h_{n+1}u)| du \\ &= \int_{S_n} \left| \sum_{m=1}^n \xi_{mn}(u) g_{mn}(u) \right|^2 |g_{n, n+1}(u)| |k(h_n u)| du \quad \text{w.p.1} \end{aligned}$$

and

$$\begin{aligned} & \int_{S_n^c} \left| \sum_{m=1}^n \xi_{mn}(u) g_{m, n+1}(u) \right|^2 |k(h_{n+1}u)| du \\ &= \int_{S_n^c} \left| \sum_{m=1}^n \xi_{mn}(u) g_{mn}(u) \right|^2 |g_{n, n+1}(u)| |k(h_n u)| du = 0 \quad \text{w.p.1,} \end{aligned}$$

which, together with (4.6) and (4.11), yields that

$$\begin{aligned}
& E[Y_{n+1}|X_1, \dots, X_n] \\
(4.12) \quad & \leq (1-a_{n+1})h_n^p h_{n+1}^{-p} Y_n + C_2(a_{n+1}h_{n+1}^{-p})^2 \\
& \leq (1-a_{n+1})Y_n + h_n^p |h_{n+1}^{-p} - h_n^{-p}| Y_n + C_2(a_{n+1}h_{n+1}^{-p})^2 \quad \text{w.p.1.}
\end{aligned}$$

Since by (H4) $h_{n+1}^{-p} - h_n^{-p} \sim p h_n^{1-p} (h_{n+1}^{-1} - h_n^{-1})$ as $n \rightarrow \infty$, where " $a_n \sim b_n$ as $n \rightarrow \infty$ " means that $a_n/b_n \rightarrow 1$ as $n \rightarrow \infty$, we get

$$h_n^p |h_{n+1}^{-p} - h_n^{-p}| \leq C_3 h_n |h_{n+1}^{-1} - h_n^{-1}| \quad \text{for all } n \geq 1,$$

which, together with (4.12), yields that

$$\begin{aligned}
(4.13) \quad & E[Y_{n+1}|X_1, \dots, X_n] \\
& \leq (1-a_{n+1})Y_n + C_3 h_n |h_{n+1}^{-1} - h_n^{-1}| Y_n + C_2(a_{n+1}h_{n+1}^{-p})^2 \quad \text{w.p.1.}
\end{aligned}$$

It follows from (4.6) that

$$\begin{aligned}
E[|Z_{mn}(u)|^2] & \leq a_m^2 \beta_{mn}^2 E[|e^{iu \cdot X_m} - \varphi(u)|^2] \\
& \leq a_m^2 \beta_{mn}^2 \quad \text{for } 1 \leq m \leq n \text{ and } u \in R^p.
\end{aligned}$$

Thus, by the independence of $Z_{1n}(u), \dots, Z_{nn}(u)$ with $E[Z_{mn}(u)] = 0$ and Fubini's theorem we get

$$\begin{aligned}
E[Y_n] & = h_n^{-p} \sum_{m=1}^n \int E[|Z_{mn}(u)|^2] |k(h_n u)| du \\
& \leq C_3 h_n^{-2p} \sum_{m=1}^n a_m^2 \beta_{mn}^2,
\end{aligned}$$

which, together with Condition (H6), implies that

$$(4.14) \quad \sum_{n=1}^{\infty} h_n |h_{n+1}^{-1} - h_n^{-1}| E[Y_n] < \infty.$$

Thus, by the use of (4.13), (4.14), Condition (H5) and Proposition 2.4 of ISOGAI [5], we obtain (4.8) and (4.9). The proof is complete.

We shall present the strong uniform consistency of the estimators of ISOGAI [5] which contains the estimators of YAMATO [8] in a special case.

COROLLARY 4.1. In (1.4) we put $a_n = a n^{-1}$ with $0 < a \leq 1$. Suppose that instead of Condition (H6) the sequence $\{h_n\}$ satisfies

$$(H7) \quad \sum_{n=1}^{\infty} \frac{d_n}{n^b h_n^{2p-1}} \left| \frac{1}{h_{n+1}} - \frac{1}{h_n} \right| < \infty,$$

where $b = \min(1, 2a)$, $d_1 = d_2 = 1$, and

$$\begin{aligned}
d_n & = \log n \quad \text{if } a = 2^{-1} \\
& = 1 \quad \text{otherwise} \quad \text{for } n = 3, 4, \dots
\end{aligned}$$

Then, under all conditions of Theorem 4.1 we obtain

$$\sup |f_n(x) - f(x)| \xrightarrow{n} 0 \quad w.p.1$$

and

$$E[\sup |f_n(x) - f(x)|^2] \xrightarrow{n} 0.$$

PROOF. In order to prove the corollary it suffices to verify Condition (H6). It is easy to see that $0 < \beta_{mn} \leq 2m^n n^{-a}$ for $1 \leq m \leq n$, which implies that

$$(4.15) \quad A_n \leq C_1 n^{-2a} \sum_{m=1}^n m^{2(a-1)} \quad \text{for each } n.$$

After some calculations we can show the following inequality:

$$(4.16) \quad n^{-2a} \sum_{m=1}^n m^{2(a-1)} \leq C_2 (d_n/n^b) \quad \text{for each } n.$$

Thus, combining (4.15), (4.16) and Condition (H7) we obtain Condition (H6). This completes the proof.

REMARK 4.1. If we put $a=1$ in (H7) then Condition (H7) coincides with Condition (11) of DAVIES [3]. Thus, Corollary 4.1 contains Theorem 2 of DAVIES [3] in a special case.

We shall give an example of $\{h_n\}$ satisfying all conditions of Corollary 4.1.

EXAMPLE 4.1.

Let $a_n = an^{-1}$ with $0 < a \leq 1$ and $h_n = n^{-r/b}$ with $0 < r < 1$. Then Conditions (H1)~(H5) and (H7) are fulfilled if we choose $0 < r < \min(2^{-1}, a)$ for fixed a with $0 < a \leq 1$.

Now, we consider a problem of estimating a mode θ of the continuous p.d.f. $f(x)$ defined by $f(\theta) = \max_{x \in R^p} f(x)$. We assume that θ is unique. If in Theorem 4.1 all conditions on $K(x)$ are fulfilled, there exists a random vector θ_n (called the sample mode) which satisfies

$$(4.17) \quad f_n(\theta_n) = \max_{x \in R^p} f_n(x) \quad \text{for each } n,$$

where $f_n(x)$ is defined by (1.3).

The following theorem is concerned with strong consistency of the sample mode θ_n .

THEOREM 4.2. *If all conditions of Theorem 4.1 or Corollary 4.1 are satisfied, then*

$$\|\theta_n - \theta\| \xrightarrow{n} 0 \quad w.p.1.$$

PROOF. Since the p.d.f. $f(x)$ is uniformly continuous on R^p and the mode θ is unique, for arbitrary $\varepsilon > 0$ there exists $\eta = \eta(\varepsilon) > 0$ such that

$$(4.18) \quad \|x - \theta\| \geq \varepsilon \text{ implies } |f(x) - f(\theta)| \geq \eta.$$

It is easily verified from (4.17) that

$$(4.19) \quad |f(\theta_n) - f(\theta)| \leq 2 \sup_x |f_n(x) - f(x)| \quad \text{for each } n.$$

Let any $\varepsilon > \delta_k > \tilde{\delta} > 0$ be fixed. By (4.18) and (4.19) we have

$$\begin{aligned} & P\{\sup_{m \geq n} \|\theta_m - \theta\| > \delta_k\} \\ & \leq P\{\sup_{m \geq n} \|\theta_m - \theta\| > \tilde{\delta}\} \\ & \leq P\{\sup_{m \geq n} \sup_x |f_m(x) - f(x)| \geq \eta(\tilde{\delta})/2\} \quad \text{for each } n, \end{aligned}$$

which implies that as $\delta_k \uparrow \varepsilon$

$$\begin{aligned} & P\{\sup_{m \geq n} \|\theta_m - \theta\| \geq \varepsilon\} \\ & \leq P\{\sup_{m \geq n} \sup_x |f_m(x) - f(x)| \geq \eta(\tilde{\delta})/2\} \quad \text{for each } n. \end{aligned}$$

By Theorem 4.1 or Corollary 4.1 the second term in the above inequality converges to zero as n tends to infinity. Thus

$$P\{\sup_{m \geq n} \|\theta_m - \theta\| \geq \varepsilon\} \xrightarrow{n} 0 \quad \text{for each } \varepsilon > 0,$$

which implies that $\|\theta_n - \theta\| \xrightarrow{n} 0$ w.p.1. This completes the proof.

We close this section with a presentation of sufficient conditions on $K(x)$ which are required in Theorem 4.1.

PROPOSITION 4.1. *Let $K_j(x)$ for $1 \leq j \leq p$ be bounded, integrable, continuous functions on the real line R satisfying*

$$(4.20) \quad |x|^p |K_j(x)| \longrightarrow 0 \quad \text{as } |x| \rightarrow \infty$$

and

$$\int_R K_j(x) dx = 1.$$

Suppose that for each $j=1, \dots, p$

$$(4.21) \quad \int_R |k_j(u)| du < \infty,$$

where $|k_j(u)| = \left| \int_R e^{iux} K_j(x) dx \right|$ is non-decreasing for $u < 0$ and non-increasing for $u \geq 0$.

Let $K(x) = \prod_{j=1}^p K_j(x_j)$ for $x = (x_1, \dots, x_p)$. Then $K(x)$ satisfies all conditions of Theorem 4.1.

PROOF. It is easy to see that $K(x)$ is a bounded, integrable, continuous function on R^p satisfying (K1). We shall verify Condition (K2). For $x = (x_1, \dots, x_p)$ let $j(x)$ be the smallest integer j such that $\max \{|x_i|; 1 \leq i \leq p\} = |x_j|$. Then we get

$$(4.22) \quad \|x\|^2 \leq p |x_{j(x)}|^2.$$

Let $C = \max \left\{ \prod_{\substack{i=1 \\ i \neq q}}^p \sup_{y \in R} |K_i(y)|; 1 \leq q \leq p \right\}$. By the boundedness of $K_j(y)$'s, C is finite. It

follows from (4.22) that

$$(4.23) \quad \|x\|^p |K(x)| \leq p^{p/2} C |x_{j(x)}|^p |K_{j(x)}(x_{j(x)})|.$$

Thus, by (4.20) and (4.23) we have Condition (K2), since $|x_{j(x)}| \rightarrow \infty$ as $\|x\| \rightarrow \infty$ by (4.22). From (4.21) and the fact that

$$(4.24) \quad |k(t)| = \prod_{j=1}^p |k_j(t_j)| \quad \text{for } t = (t_1, \dots, t_p),$$

it holds that $\int |k(t)| dt < \infty$, that is, Condition (K3) holds. Let $s = (s_1, \dots, s_p)$ and $t = (t_1, \dots, t_p)$ in a ray $R(u)$ for $u \neq 0 \in R^p$ with $\|s\| \leq \|t\|$ be fixed. Then we can write s and t as $s = q_1 u$ and $t = q_2 u$ with $0 < q_1 \leq q_2$, respectively. Hence, for each $j = 1, \dots, p$ we get $|k_j(s_j)| \geq |k_j(t_j)|$ by the property of $|k_j(y)|$. Thus by (4.24) we have $|k(s)| \geq |k(t)|$, that is, $|k(t)|$ is non-increasing on the ray $R(u)$. This completes the proof.

EXAMPLE 4.2.

The following functions satisfy all conditions of Proposition 4.1: For $x = (x_1, \dots, x_p)$

- (i) $K(x) = 2^{-p} \exp(-\sum_{j=1}^p |x_j|)$ for $p \geq 1$
- (ii) $K(x) = (2\pi)^{-p/2} \exp(-\|x\|^2/2)$ for $p \geq 1$
- (iii) $K(x) = (1/\pi)(1+x^2)^{-1}$ for $p=1$
- (iv) $K(x) = (1/2\pi)(\sin(x/2)/(x/2))^2$ for $p=1$.

References

- [1] BOCHNER, S., and CHANDRASEKHARAN, K. *Fourier transforms*. Princeton University Press, 1949.
- [2] CACOULLOS, T. *Estimation of a multivariate density*. Ann. Inst. Statist. Math., 18 (1966), 179-189.
- [3] DAVIES, H. I. *Strong consistency of a sequential estimator of a probability density function*. Bull. Math. Statist., 15 (1973), No. 3-4, 49-54.
- [4] DEVROYE, L. P. *On the pointwise and the integral convergence of recursive kernel estimates of probability densities*. Util. Math., 15 (1979), 113-128.
- [5] ISOGAI, E. *Strong consistency and optimality of a sequential density estimator*. Bull. Math. Statist., 19 (1980), No. 1-2, 55-69.
- [6] LOÈVE, M. *Probability theory*. 3rd Edition. D. Van Nostrand, Princeton, 1963.
- [7] STEIN, E. M. *Singular integrals and differentiability properties of functions*. Princeton University Press, Princeton, New Jersey, 1970.
- [8] YAMATO, H. *Sequential estimation of a continuous probability density function and mode*. Bull. Math. Statist., 14 (1971). No 1-2, 1-12.

Department of Mathematics
Faculty of Science
NIIGATA UNIVERSITY
Niigata, 950-21, Japan