

## A NOTE ON A RENEWAL THEOREM

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1. Let  $X_i$  ( $i = 1, 2, \dots$ ) be independent random variables, having the finite mean values  $E(X_i) = m_i > 0$  ( $i = 1, 2, \dots$ ). When

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n m_i = m$$

exists, then it holds that

$$(1.1) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_{-\infty}^T dt \sum_{n=1}^{\infty} \Pr(t < S_n \leq t + h) = \frac{h}{m}$$

where

$$S_n = \sum_{i=1}^n X_i,$$

with some restrictions. This problem is a renewal theorem in a wide sense and have been treated by Kawata [2] and the author [1] under somewhat different conditions. In the following, we shall extend this theorem in a sense under the assumptions which have been considered in [2].

2. In the first place, we shall prepare the following lemmas.

LEMMA 1. *Let  $F(t)$  be a non-decreasing function,*

$$(2.1) \quad \int_{-\infty}^0 e^{-s_0 t} dF(t) < +\infty \quad \text{for some } s_0 > 0$$

and

$$(2.2) \quad \int_{-\infty}^{\infty} e^{-st} dF(t) \sim \frac{A}{s^\gamma} \quad \text{as } s \downarrow 0 \text{ for some } \gamma > 0.$$

Then

$$F(t) \sim \frac{At^\gamma}{\Gamma(\gamma + 1)} \quad \text{as } t \rightarrow \infty.$$

*Proof.* Since

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$$0 \leq \int_{-\infty}^0 e^{-st} dF(t) \leq \int_{-\infty}^0 e^{-s_0 t} dF(t) < +\infty \quad \text{for } 0 \leq s \leq s_0$$

by (2.1), we have

$$\int_0^{\infty} e^{-st} dF(t) \sim \frac{A}{s^r} \quad \text{for } s \downarrow 0.$$

Hence by a classical Tauberian theorem, it results that

$$F(t) \sim \frac{At^r}{\Gamma(r+1)} \quad \text{as } t \rightarrow \infty.$$

LEMMA 2. Let  $X_i$  ( $i=1, 2, \dots$ ) be independent random variables such that  $E(X_i) = m_i > 0$ . Suppose that the distribution function  $F_n(x)$  of  $X_n$  satisfies

$$(2.3) \quad \int_{-\infty}^0 e^{-s_0 t} dF_n(x) < +\infty \quad \text{for some } s_0 > 0$$

and further that both of

$$(2.4) \quad \lim_{A \rightarrow \infty} \int_A^{\infty} x dF_n(x) = 0$$

and

$$(2.5) \quad \lim_{A \rightarrow \infty} \int_{-\infty}^{-A} e^{-s_0 x} dF_n(x) = 0$$

hold uniformly with respect to  $n$ . If

$$(2.6) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n m_i = m > 0,$$

then

$$(2.7) \quad \lim_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} n^{\alpha} \varphi_n(s) = \frac{\alpha!}{m^{\alpha+1}} \quad \text{for } \alpha = 0, 1, 2, \dots,$$

where

$$\varphi_n(s) = \int_{-\infty}^{\infty} e^{-sx} d\sigma_n(x),$$

$\sigma_n(x)$  being the distribution function of  $S_n$ .

*Proof.* Let  $\varepsilon$  be any given positive number. Under the conditions of this lemma, there exist an  $N$  and an  $s_2 < s_0$  such that

$$(2.8) \quad e^{-ns(m+2\varepsilon)} \leq \varphi_n(s) \leq e^{-ns(m-2\varepsilon)} \quad \text{for } n > N \text{ and } 0 \leq s \leq s_2,$$

which has been given in the course of the proof of Lemma 2 in [2]. Hence we have

$$\begin{aligned} & s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \\ & < s^{\alpha+1} \sum_{n=1}^N (n+\alpha)(n+\alpha-1)\cdots(n+1)C \\ & \quad + s^{\alpha+1} \sum_{n=0}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)e^{-ns(m-2\varepsilon)} \\ & = s^{\alpha+1} \sum_{n=1}^N (n+\alpha)(n+\alpha-1)\cdots(n+1)C + \frac{s^{\alpha+1} \cdot \alpha!}{(1 - e^{-s(m-2\varepsilon)})^{\alpha+1}}, \quad \text{for } 0 < s \leq s_2 \end{aligned}$$

where

$$C = \sup_{\substack{0 < s \leq s_2 \\ n=1,2,\dots,N}} \varphi_n(s) \leq 1 + \max_{n=1,2,\dots,N} \int_{-\infty}^0 e^{-s_2 x} d\sigma_n(x) < +\infty.$$

Thus

$$\varliminf_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \leq \frac{\alpha!}{(m-2\varepsilon)^{\alpha+1}}$$

and since  $\varepsilon$  is arbitrary, we get

$$(2.9) \quad \varliminf_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \leq \frac{\alpha!}{m^{\alpha+1}}.$$

On the other hand

$$\begin{aligned} & s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \\ & \geq s^{\alpha+1} \sum_{n=0}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)e^{-ns(m+2\varepsilon)} \\ & \quad - s^{\alpha+1} \sum_{n=0}^N (n+\alpha)(n+\alpha-1)\cdots(n+1) \\ & = \frac{s^{\alpha+1} \cdot \alpha!}{(1 - e^{-s(m+2\varepsilon)})^{\alpha+1}} - s^{\alpha+1} \sum_{n=0}^N (n+\alpha)(n+\alpha-1)\cdots(n+1) \quad \text{for } 0 < s \leq s_2. \end{aligned}$$

Hence

$$\varliminf_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \geq \frac{\alpha!}{(m+2\varepsilon)^{\alpha+1}}$$

from which it results that

$$(2.10) \quad \lim_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) \geq \frac{\alpha!}{m^{\alpha+1}}.$$

(2.9) and (2.10) show that

$$(2.11) \quad \lim_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} (n+\alpha)(n+\alpha-1)\cdots(n+1)\varphi_n(s) = \frac{\alpha!}{m^{\alpha+1}}.$$

From (2.11) with  $\alpha=0, 1, 2, \dots$ , we get (2.7).

LEMMA 3. *Let  $\{a_n\}$  be a sequence of real numbers. If, in addition to the conditions of Lemma 2, we assume that*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n a_i = a$$

*exists, then we have*

$$(2.12) \quad \lim_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} a_n n^\alpha \varphi_n(s) = \frac{a \cdot \alpha!}{m^{\alpha+1}} \quad \text{for } \alpha = 0, 1, 2, \dots.$$

*Proof.* Since, describing

$$\frac{1}{n} \sum_{i=1}^n a_i = a + \varepsilon_n,$$

we have

$$a_n = a + n\varepsilon_n - (n-1)\varepsilon_{n-1}, \quad \varepsilon_n \rightarrow 0 \quad (n \rightarrow \infty)$$

and

$$\sum_{n=1}^{\infty} a_n n^\alpha \varphi_n(s) = a \sum_{n=1}^{\infty} n^\alpha \varphi_n(s) + \sum_{n=1}^{\infty} n^\alpha \varphi_n(s) [n\varepsilon_n - (n-1)\varepsilon_{n-1}],$$

it suffices to prove

$$(2.13) \quad \lim_{s \downarrow 0} s^{\alpha+1} \sum_{n=1}^{\infty} n^\alpha \varphi_n(s) [n\varepsilon_n - (n-1)\varepsilon_{n-1}] = 0.$$

Noticing

$$\sum_{n=1}^{\infty} |n^\alpha \varphi_n(s) \cdot n\varepsilon_n| \leq \sup_{n=1,2,\dots} |\varepsilon_n| \cdot \sum_{n=1}^{\infty} n^{\alpha+1} \varphi_n(s) < +\infty,$$

we get

$$\begin{aligned} & \sum_{n=1}^{\infty} n^\alpha \varphi_n(s) [n\varepsilon_n - (n-1)\varepsilon_{n-1}] \\ &= \sum_{n=1}^{\infty} n\varepsilon_n [n^\alpha \varphi_n(s) - (n+1)^\alpha \varphi_{n+1}(s)] \end{aligned}$$

$$\begin{aligned}
&= \sum_{n=1}^{\infty} n \varepsilon_n \varphi_n(s) [n^\alpha - (n+1)^\alpha f_{n+1}(s)] \\
&= \sum_{n=1}^{\infty} n \varepsilon_n \varphi_n(s) \cdot n^\alpha (1 - f_{n+1}(s)) + \sum_{n=1}^{\infty} n \varepsilon_n \varphi_n(s) (n^\alpha - (n+1)^\alpha) f_{n+1}(s) \\
&= I_1 + I_2,
\end{aligned}$$

say, where

$$f_n(s) = \int_{-\infty}^{\infty} e^{-sx} dF_n(s), \quad 0 \leq s \leq s_0.$$

Now, for any given positive number  $\varepsilon$ , we choose an integer  $N$  such that

$$|\varepsilon_n| < \varepsilon \quad \text{for } n > N.$$

Since there are positive constants  $C_1$  and  $s_2$  such that  $m_n < C_1$  for  $n = 1, 2, \dots$  and

$$f_n(s) = 1 - sm_n + s\eta_n,$$

where

$$|\eta_n| < \varepsilon(C_1 + 2) \equiv C_2 \quad \text{for } 0 \leq s \leq s_2,$$

uniformly with respect to  $n$ , which have been proved in the course of the proof of Lemma 2 in [2], we have

$$\begin{aligned}
|s^{\alpha+1} \cdot I_1| &< s^{\alpha+2} \cdot (C_1 + C_2) \cdot C \sum_{n=1}^N n^{\alpha+1} |\varepsilon_n| + \varepsilon(C_1 + C_2) \cdot s^{\alpha+2} \sum_{n=N+1}^{\infty} n^{\alpha+1} \varphi_n(s) \\
&\quad \text{for } 0 < s \leq s_2,
\end{aligned}$$

where

$$C = \sup_{\substack{0 < s \leq s_2 \\ n=1, 2, \dots, N}} \varphi_n(s) < +\infty,$$

and hence

$$\overline{\lim}_{s \downarrow 0} |s^{\alpha+1} \cdot I_1| \leq \varepsilon (C_1 + C_2) \frac{(\alpha+1)!}{m^{\alpha+2}}$$

and since  $\varepsilon$  is arbitrary, we get

$$(2.14) \quad \lim_{s \downarrow 0} s^{\alpha+1} \cdot I_1 = 0.$$

On the other hand,

$$|s^{\alpha+1} \cdot I_2| < s^{\alpha+1} [1 + (C_1 + C_2)s] \cdot C \sum_{n=1}^N \alpha(n+1)^\alpha \cdot |\varepsilon_n|$$

$$+ [1 + (C_1 + C_2)s] \varepsilon \cdot s^{\alpha+1} \sum_{n=N+1}^{\infty} \left( \alpha n^{\alpha} + \binom{\alpha}{2} n^{\alpha-1} + \dots + n \right) \varphi_n(s)$$

for  $0 < s \leq s_2$ .

Hence

$$\overline{\lim}_{s \downarrow 0} |s^{\alpha+1} \cdot I_2| \leq \alpha \cdot \varepsilon \frac{\alpha!}{m^{\alpha+1}},$$

from which it results that

$$(2.15) \quad \lim_{s \downarrow 0} s^{\alpha+1} \cdot I_2 = 0.$$

(2.14) and (2.15) give (2.12).

**3. THEOREM 1.** *If, in addition to the conditions of Lemma 3, we assume that  $a_n \geq 0$  ( $n = 1, 2, \dots$ ), then we have*

$$(3.1) \quad \lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n \cdot n^{\alpha} \Pr(S_n \leq t) = \frac{a}{(\alpha+1)m^{\alpha+1}} \quad \text{for } \alpha = 0, 1, 2, \dots$$

*Proof.* Describing

$$H_N(t) = \sum_{n=1}^N a_n n^{\alpha} \Pr(S_n \leq t),$$

we have

$$\int_{-\infty}^{\infty} e^{-st} dH_N(t) = \sum_{n=1}^N a_n \cdot n^{\alpha} \int_{-\infty}^{\infty} e^{-st} d\sigma_n(t) = \sum_{n=1}^N a_n n^{\alpha} \varphi_n(s)$$

and so

$$(3.2) \quad \lim_{N \rightarrow \infty} \int_{-\infty}^{\infty} e^{-st} dH_N(t)$$

exists and is equal to

$$\sum_{n=1}^{\infty} a_n n^{\alpha} \varphi_n(s) \quad \text{for } 0 < s \leq s_2.$$

Since we know by (2.8) that there exists a constant  $C_3$  such that

$$\varphi_n(s) \leq C_3 \quad \text{for } n = 1, 2, \dots \text{ and } 0 < s \leq s_2,$$

we have

$$\int_{-\infty}^{\infty} e^{-st} dH_N(t) \leq C_3 \sum_{n=1}^N a_n n^{\alpha}$$

and

$$\int_{-\infty}^{-A} e^{-st} dH_N(t) \leq C_3 \sum_{n=1}^N a_n n^\alpha \quad \text{for any positive } A.$$

Taking  $\sigma$  less than  $s_2$  and  $s = s_2$ ,

$$\begin{aligned} C_3 \sum_{n=1}^N a_n \cdot n^\alpha &\geq \int_{-\infty}^{-A} e^{-(s_2-\sigma)t} \cdot e^{-\sigma t} dH_N(t) \\ &\geq e^{(s_2-\sigma)A} \int_{-\infty}^{-A} e^{-\sigma t} dH_N(t) \\ &\geq e^{(s_2-\sigma)A} \int_{-B}^{-A} e^{-\sigma t} dH_N(t), \end{aligned}$$

which gives

$$\int_{-B}^{-A} e^{-\sigma t} dH_N(t) \leq e^{-(s_2-\sigma)A} C_3 \cdot \sum_{n=1}^N a_n n^\alpha.$$

Letting  $\sigma \rightarrow 0$  and  $B \rightarrow \infty$ , we get

$$H_N(-A) \leq e^{-s_2 A} C_3 \cdot \sum_{n=1}^N a_n n^\alpha.$$

Therefore if  $0 < s \leq s_3 < s_2$ ,

$$\lim_{t \rightarrow -\infty} e^{-st} H_N(t) = 0.$$

Thus we get by partial integration that for  $0 < s \leq s_3$

$$(3.3) \quad \int_{-\infty}^{\infty} e^{-st} dH_N(t) = s \int_{-\infty}^{\infty} e^{-st} H_N(t) dt.$$

Since  $H_N(t)$  increases as  $N \rightarrow \infty$  and tends to a non-decreasing function  $H(t)$ , the existence of the limit (3.2) and (3.3) show that

$$(3.4) \quad \lim_{N \rightarrow \infty} \int_{-\infty}^{\infty} e^{-st} H_N(t) dt = \int_{-\infty}^{\infty} e^{-st} H(t) dt$$

exists for  $0 < s \leq s_3$ .  $H(t)$  is equal to

$$\sum_{n=1}^{\infty} a_n n^\alpha \Pr(S_n \leq t).$$

The existence of the right side integral shows that for  $0 < s \leq s_4 \leq s_3$

$$H(t) = o(e^{st}) \quad \text{for } |t| \rightarrow \infty.$$

Hence

$$s \int_{-\infty}^{\infty} e^{-st} H(t) dt = \int_{-\infty}^{\infty} e^{-st} dH(t)$$

exists and is equal to

$$\sum_{n=1}^{\infty} a_n n^{\alpha} \varphi_n(s) \quad \text{for } 0 < s \leq s_4$$

by (3.2), (3.3) and (3.4), so that we have by Lemma 3

$$\int_{-\infty}^{\infty} e^{-st} dH(t) \sim \frac{a \cdot \alpha!}{m^{\alpha+1} s^{\alpha+1}} \quad \text{as } s \downarrow 0.$$

Then, by Lemma 1, we get

$$\lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} H(t) = \lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(S_n \leq t) = \frac{a}{(\alpha+1)m^{\alpha+1}}.$$

**COROLLARY 1.** *If, in addition to the conditions of Lemma 3, we assume that  $a_n$  ( $n=1, 2, \dots$ ) are bounded from below, we have*

$$\lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(S_n \leq t) = \frac{a}{(\alpha+1)m^{\alpha+1}} \quad \text{for } \alpha = 0, 1, 2, \dots.$$

*Proof.* Since we can take a constant  $c \geq 0$  such that

$$a_i + c \geq 0 \quad \text{for } i = 1, 2, \dots,$$

we have

$$\frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(S_n \leq t) = \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} (a_n + c) n^{\alpha} \Pr(S_n \leq t) - \frac{c}{t^{\alpha+1}} \sum_{n=1}^{\infty} n^{\alpha} \Pr(S_n \leq t)$$

which converges to

$$\frac{a+c}{(\alpha+1)m^{\alpha+1}} - \frac{c}{(\alpha+1)m^{\alpha+1}} = \frac{a}{(\alpha+1)m^{\alpha+1}} \quad (t \rightarrow \infty).$$

**COROLLARY 2.** *If, in addition to the conditions Lemma 2, we assume that  $\{a_n\}$  is a sequence of real numbers and  $a_n$  can be expressed as the difference of  $b_n$  and  $c_n$  for  $n=1, 2, \dots$ , where  $b_n$  and  $c_n$  are bounded from below and have the properties that*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n b_i = b \quad \text{and} \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n c_i = c$$

*exist, we have*



$$\lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(S_n \leq t) = \frac{a}{(\alpha+1)m^{\alpha+1}} \quad \text{for } \alpha = 0, 1, 2, \dots,$$

where  $a = b - c$ .

4. Taking  $\sigma_n(t)$  as  $a_n \Pr(S_n \leq t)$  in the course of the proof of Theorem 2 in [1], we get the following

**THEOREM 2.** *Under the conditions of Theorem 1, we have*

$$(4.1) \quad \lim_{T \rightarrow \infty} \frac{1}{T^{\alpha+1}} \int_{-\infty}^T dt \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(t < S_n \leq t+h) = \frac{ah}{(\alpha+1)m^{\alpha+1}} \\ \text{for } \alpha = 0, 1, 2, \dots$$

**COROLLARY 3.** *Under the conditions of Corollary 2, we have*

$$\lim_{T \rightarrow \infty} \frac{1}{T^{\alpha+1}} \int_{-\infty}^T dt \sum_{n=1}^{\infty} a_n n^{\alpha} \Pr(t < S_n \leq t+h) = \frac{ah}{(\alpha+1)m^{\alpha+1}} \\ \text{for } \alpha = 0, 1, 2, \dots$$

The proof follows easily from Theorem 2.

**REMARK.** If  $\alpha \geq 0$  may be not necessarily an integer, we have in place of (3.1) and (4.1) that

$$\lim_{t \rightarrow \infty} \frac{1}{t^{\alpha+1}} \sum_{n=1}^{\infty} a_n \cdot \frac{\Gamma(n+\alpha+1)}{n!} \Pr(S_n \leq t) = \frac{a}{(\alpha+1)m^{\alpha+1}},$$

and

$$\lim_{T \rightarrow \infty} \frac{1}{T^{\alpha+1}} \int_{-\infty}^T dt \sum_{n=1}^{\infty} a_n \cdot \frac{\Gamma(n+\alpha+1)}{n!} \Pr(t < S_n \leq t+h) = \frac{ah}{(\alpha+1)m^{\alpha+1}}.$$

**COROLLARY 4.** *Let  $A_i$  ( $i=1, 2, \dots$ ) be independent identically distributed random variables, having finite variances. Then, under the conditions of Lemma 2, we have*

$$(4.2) \quad \lim_{T \rightarrow \infty} \frac{1}{T^{\alpha+1}} \int_{-\infty}^T dt \sum_{n=1}^{\infty} A_n n^{\alpha} \Pr(t < S_n \leq t+h) = \frac{ah}{(\alpha+1)m^{\alpha+1}}$$

with probability 1, where  $a = E\{X_i\}$ .

*Proof.* Since the sequences  $\{\max(A_i, 0)\}$  and  $\{\max(-A_i, 0)\}$  of random variables obey respectively the the strong law of large numbers and  $A_i = \max(A_i, 0) - \max(-A_i, 0)$ , this theorem follows from Corollary 3.

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## REFERENCES

- [ 1 ] HATORI, H., Some theorems in an extended renewal theory, II. Kōdai Math. Sem. Rep. 12 (1960), 21-27.
- [ 2 ] KAWATA, T., A renewal theorem, Journ. Math. Soc. Japan 8 (1956), 118-126.

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