

SOME PROPERTIES OF THE EMPIRICAL DISTRIBUTION FUNCTION IN THE NON-i.i.d. CASE¹

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For $N = 1, 2, \dots$ let $X_{1N}, X_{2N}, \dots, X_{NN}$ be independent rv's having continuous df's $F_{1N}, F_{2N}, \dots, F_{NN}$. For the set $X_{1N}, X_{2N}, \dots, X_{NN}$, let us denote by $X_{1:N} \leq X_{2:N} \leq \dots \leq X_{N:N}$ the order statistics, by \mathbb{F}_N the empirical df and by \bar{F}_N the averaged df, i.e. $\bar{F}_N(x) = N^{-1} \sum_{n=1}^N F_{nN}(x)$ for $x \in (-\infty, \infty)$. It is shown that for each $\varepsilon > 0$ there exists a $0 < \beta(= \beta_\varepsilon) < 1$, independent of N , such that for $N = 1, 2, \dots$,

- (a) $P(\mathbb{F}_N(x) \leq \beta^{-1}\bar{F}_N(x), \text{ for } x \in (-\infty, \infty)) \geq 1 - \varepsilon,$
- (b) $P(\mathbb{F}_N(x) \geq \beta\bar{F}_N(x), \text{ for } x \in [X_{1:N}, \infty)) \geq 1 - \varepsilon.$

Moreover, these assertions hold uniformly in all continuous df's $F_{1N}, F_{2N}, \dots, F_{NN}$.

The theorem can be used to prove asymptotic normality of rank statistics and of linear combinations of functions of order statistics in the case where the sample elements are allowed to have different df's.

1. Notation and results. Let X_{nN}, F_{nN} , etc. be as in the abstract. Define \bar{F}_N^{-1} by $\bar{F}_N^{-1}(s) = \inf \{x: \bar{F}_N(x) \geq s\}$. Because \bar{F}_N is assumed to be continuous we have $\bar{F}_N(\bar{F}_N^{-1}(s)) = s$ for $s \in (0, 1)$.

THEOREM. For each $\varepsilon > 0$ there exists a $0 < \beta(= \beta_\varepsilon) < 1$, independent of N and of the continuous df's $F_{1N}, F_{2N}, \dots, F_{NN}$, such that for $N = 1, 2, \dots$,

- (a) $P(1 - \beta^{-1}(1 - \bar{F}_N(x)) \leq \mathbb{F}_N(x) \leq \beta^{-1}\bar{F}_N(x), \text{ for } x \in (-\infty, \infty)) \geq 1 - \varepsilon,$
- (b) $P(1 - \beta(1 - \bar{F}_N(x)) \geq \mathbb{F}_N(x) \geq \beta\bar{F}_N(x), \text{ for } x \in [X_{1:N}, \infty)) \geq 1 - \varepsilon.$

The theorem is useful for proving asymptotic normality of rank statistics and of linear combinations of functions of order statistics in the case where the sample elements are allowed to have different df's.

For the i.i.d. case these results are well known; statements appear in Shorack [2]. Note that our theorem fills the gap in extending the proof of Theorem 3 of Shorack [3], so that his Theorem 3 can now be claimed to hold without the assumption $b_1 = b_2 = 0$ (that is, for unbounded J).

Our basic tool is a result of Hoeffding [1]. Suppose that $Z_j, 1 \leq j \leq N$, are independent Bernoulli (p_j) rv's and suppose

$$0 < N^{-1} \sum_{j=1}^N p_j = \bar{p} < 1.$$

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LEMMA (Hoeffding). *If f is a strictly convex function defined on $(-\infty, \infty)$ then*

$$\mathcal{E}(f(\sum_{j=1}^N Z_j)) \leq \sum_{k=0}^N f(k) \binom{N}{k} \bar{p}^k (1 - \bar{p})^{N-k},$$

where equality holds if and only if $p_1 = p_2 = \dots = p_N = \bar{p}$.

In particular this lemma together with Chebyshev's inequality implies that for $n > N\bar{p}$,

$$\begin{aligned} &P(\sum_{j=1}^N Z_j \geq n) \\ &\leq P(|\sum_{j=1}^N Z_j - N\bar{p}| \geq n - N\bar{p}) \\ (1.1) \quad &\leq (n - N\bar{p})^{-4} \mathcal{E}(\sum_{j=1}^N Z_j - N\bar{p})^4 \\ &\leq (n - N\bar{p})^{-4} \sum_{k=0}^N (k - N\bar{p})^4 \binom{N}{k} \bar{p}^k (1 - \bar{p})^{N-k} \\ &= (n - N\bar{p})^{-4} \{(\bar{p}(1 - \bar{p}))^2 (3N^2 - 6N) + N\bar{p}(1 - \bar{p})\} \\ &\leq (n - N\bar{p})^{-4} \min \{ (3N^2\bar{p}^2 + N\bar{p}), (3N^2(1 - \bar{p})^2 + N(1 - \bar{p})) \}. \end{aligned}$$

2. Proof of the theorem. (a) For all $N \geq 1$ and all $0 < \beta < 1$ we have

$$\begin{aligned} &P(\mathbb{F}_N(x) \leq \beta^{-1} \bar{F}_N(x), \text{ for } x \in (-\infty, \infty)) \\ (2.1) \quad &= P(\mathbb{F}_N(X_{n:N}) \leq \beta^{-1} \bar{F}_N(X_{n:N}), \text{ for } n = 1, 2, \dots, N) \\ &\geq 1 - \sum_{n=1}^N P(\bar{F}_N(X_{n:N}) < \beta n N^{-1}) = 1 - \sum_{n=1}^N P(\sum_{j=1}^N Z_j \geq n), \end{aligned}$$

where $Z_j, 1 \leq j \leq N$, are independent Bernoulli (p_j) rv's with

$$p_j = F_{jN}(\bar{F}_N^{-1}(\beta n N^{-1})).$$

From (2.1) and (1.1) with $\bar{p} = N^{-1} \sum_{j=1}^N F_{jN}(\bar{F}_N^{-1}(\beta n N^{-1})) = \beta n N^{-1}$ it is now immediate that for $n = 1, 2, \dots, N$,

$$P(\bar{F}_N(X_{n:N}) < \beta n N^{-1}) \leq \beta(1 - \beta)^{-4} (3\beta n^{-2} + n^{-3}) \leq 4\beta(1 - \beta)^{-4} n^{-2},$$

so that

$$(2.2) \quad \sum_{n=1}^N P(\bar{F}_N(X_{n:N}) < \beta n N^{-1}) \leq M_1 \beta (1 - \beta)^{-4} \rightarrow 0 \quad \text{as } \beta \rightarrow 0,$$

where M_1 is a finite constant, independent of N and of the df's $F_{1N}, F_{2N}, \dots, F_{NN}$. Assertion (a) in the theorem now follows from (2.2), (2.1) and the symmetric result that comes from replacing X_{nN} by $-X_{nN}$.

(b) For all $N \geq 1$ and all $0 < \beta < 1$ we have (using $(n - 1)/\beta N > 1$ for $n > [\beta N] + 1$)

$$\begin{aligned} &P(\mathbb{F}_N(x) \geq \beta \bar{F}_N(x), \text{ for } x \in [X_{1:N}, \infty)) \\ &\geq P(\bigcap_{n=2}^N [\mathbb{F}_N(X_{n:N}) \geq \beta \bar{F}_N(X_{n:N})]) \\ (2.3) \quad &\geq 1 - \sum_{n=2}^N P(\bar{F}_N(X_{n:N}) > \beta^{-1}(n - 1)N^{-1}) \\ &= 1 - \sum_{n=2}^{[\beta N] + 1} P(\bar{F}_N(X_{n:N}) > \beta^{-1}(n - 1)N^{-1}), \\ &= 1 - \sum_{n=2}^{[\beta N] + 1} P(\sum_{j=1}^N Z_j \geq N - n + 1), \end{aligned}$$

where $Z_j, 1 \leq j \leq N$, are independent Bernoulli (p_j) rv's with

$$p_j = 1 - F_{jN}(\bar{F}_N^{-1}(\beta^{-1}(n - 1)N^{-1})).$$

From (2.3) and (1.1) with now $\bar{p} = 1 - \beta^{-1}(n-1)N^{-1}$ it is immediate again that for $n = 2, 3, \dots, [\beta N] + 1$,

$$P(\bar{F}_N(X_{n:N}) > \beta^{-1}(n-1)N^{-1}) \leq 4\beta^2(1-\beta)^{-4}(n-1)^{-2},$$

so that

$$(2.4) \quad \sum_{n=2}^{[\beta N]+1} P(\bar{F}_N(X_{n:N}) > \beta^{-1}(n-1)N^{-1}) \leq M_2 \beta^2 (1-\beta)^{-4} \rightarrow 0$$

as $\beta \rightarrow 0$,

where M_2 is a finite constant, independent of N and of the df's $F_{1N}, F_{2N}, \dots, F_{NN}$. Assertion (b) in the theorem now follows from (2.4), (2.3) and symmetry.

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