

A REMARK ON THE STRONG LAW¹

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Let X_1, X_2, \dots be a sequence of random variables, each having conditional mean zero given the past. Let V_n be the conditional variance of X_n given the past. Let $S_n = X_1 + \dots + X_n$ and $T_n = V_1 + \dots + V_n$. For simplicity, suppose $\sum V_i = \infty$ a.e. For which non-decreasing function ϕ does $S_n/\phi(T_n)$ necessarily lead to 0 a.e. as n increases? It is necessary and sufficient that $\int_0^\infty 1/\phi(t)^2 dt < \infty$. That is, if the integral is finite, convergence to zero a.e. is guaranteed for all X_i and V_i satisfying the stated condition. If the integral is infinite, there is a sequence of independent, symmetric random variables X_i , each having variance 1, such that $S_n/\phi(n)$ oscillates between $\pm\infty$. The sufficiency is known, but a new proof is given.

Let (Ω, \mathcal{F}, P) be a probability triple, and let $\mathcal{F}_0 \subset \mathcal{F}_1 \subset \dots$ be sub σ -fields of \mathcal{F} . Let X_1, X_2, \dots be finite functions on Ω , such that

$$(1) \quad X_n \text{ is } \mathcal{F}_n\text{-measurable}$$

$$(2) \quad E\{X_{n+1} | \mathcal{F}_n\} = 0 \text{ a.e.}$$

This does not require X_{n+1} to be L^2 or even L^1 . It does require that a.e., given \mathcal{F}_n , the conditional distribution of X_{n+1} be integrable with mean 0. Let

$$(3) \quad V_n = E\{X_{n+1}^2 | \mathcal{F}_n\},$$

the conditional variance of X_{n+1} given \mathcal{F}_n . Let

$$(4) \quad S_n = X_1 + \dots + X_n \quad \text{and} \quad T_n = V_1 + \dots + V_n.$$

The main result of this note, to be proved later, is

(5) **THEOREM.** *Suppose (1)–(4). Let $A \geq 0$, and let ϕ be a positive, non-decreasing function on $[A, \infty)$, such that*

$$(6) \quad \int_A^\infty 1/\phi(t)^2 dt < \infty.$$

Then

$$\lim_{n \rightarrow \infty} S_n/\phi(T_n) = 0 \text{ a.e. on } \{\sum V_i = \infty\}.$$

This is known for L^2 variables; see [3] page 150. This result includes the usual strong law of large numbers for independent identically distributed L^2 variables: take $\phi(t) = t$. In a sense, (5) is sharp. Proved later will be

Received May 29, 1973; revised September 10, 1973.

¹ Research sponsored by the Air Force Office of Scientific Research, Air Force Systems Command USAF, under Grant AFOSR-71-2100. The United States Government is authorized to reproduce and distribute reports for governmental purposes notwithstanding any copyright notation hereon.

AMS 1970 subject classifications. 60F10, 60F15, 60G40, 60G45.

Key words and phrases. Strong law of large numbers, martingales, stopping times, gambling, dynamic programming.

(7) THEOREM. Let ϕ be a positive non-decreasing function on $[A, \infty)$, with $\int_A^\infty 1/\phi(t)^2 dt = \infty$. Then there is a sequence of independent, symmetric random variables, each having variance 1, such that

$$(8) \quad P\{\limsup_{n \rightarrow \infty} S_n/\phi(n) = \infty\} = 1.$$

The asymptotic behavior of $S_n/\phi(T_n)$ may well depend on the underlying process, when $\int_A^\infty 1/\phi(t)^2 dt = \infty$. For instance, suppose $\phi(x) = (x \log x)^{1/2}$, for large x . If the X_i are uniformly bounded, the law of the iterated logarithm shows $S_n/\phi(T_n) \rightarrow 0$ a.e. With this observation and (7), it is possible to construct a process satisfying (1)–(4), such that all V_i are 1, and

$$P\{S_n/(n \log n)^{1/2} \rightarrow 0\} = \frac{1}{2}.$$

The proof of (5) follows the pattern of [1]. Here is the main step. Fix one positive, measurable function ϕ on $[A, \infty)$, satisfying (6). Let

$$f(v) = \int_v^\infty 1/\phi(t)^2 dt$$

$$Q(x, v) = [x^2/\phi(v)^2] + f(v)$$

for $v \geq A$ and $-\infty < x < \infty$.

(9) LEMMA. Suppose X is a random variable, with $E(X) = 0$ and $E(X^2) = V$. Then

$$E\{Q(x + X, v + V)\} \leq Q(x, v),$$

for $v \geq A$ and $-\infty < x < \infty$.

PROOF. Clearly,

$$E\{Q(x + X, v + V)\} = x^2/\phi(v + V)^2 + V/\phi(v + V)^2 + f(v + V),$$

and $x^2/\phi(v + V)^2 \leq x^2/\phi(v)^2$. Next,

$$f(v) - f(v + V) = \int_{v+V}^{v+V} 1/\phi(t)^2 dt$$

$$\geq V/\phi(v + V)^2;$$

so

$$V/\phi(v + V)^2 + f(v + V) \leq f(v). \quad \square$$

(10) PROPOSITION. Suppose (1)–(4) and (6). Let τ be a stopping time: $\{\tau = n\} \in \mathcal{F}_n$ for all n . For $v \geq A$ and $-\infty < x < \infty$,

$$E\{(x + S_\tau)^2/\phi(v + T_\tau)^2\} \leq Q(x, v).$$

PROOF. This follows from (9), using Theorem (22) of [1]. \square

(11) COROLLARY. Suppose (1)–(4) and (6). Let $a > 0$. Then

$$P\{|S_n| > a\phi(v + T_n) \text{ for some } n\} \leq Q(0, v)/a^2.$$

PROOF. This follows from (10), using Chebychev and Lemma (30) of [1]. \square

(12) COROLLARY. Suppose (1)–(4) and (6). Let $k = 1, 2, \dots$ and $a > 0$. Then

$$P\{|X_{k+1} + \dots + X_{k+n}| > a\phi(T_n) \text{ for some } n | \mathcal{F}_k\} \leq Q(0, T_k)/a^2.$$

PROOF. Condition on \mathcal{F}_k , and use (11). \square

I believe estimates (9)—(12) are new.

PROOF OF (5). Fix $a > 0$. Let $D = \{\sum V_i = \infty\}$. Let

$$D_k(a) = \{|X_{k+1} + \cdots + X_{k+n}| \leq a\phi(T_n) \text{ for all } n\}.$$

Since $\phi(\infty) = \infty$,

$$(13) \quad \limsup |S_n|/\phi(T_n) \leq a \quad \text{on } D \cap D_k(a).$$

Fix $\varepsilon > 0$. Choose $v \geq A$ so large that $Q(0, v)/a^2 \leq \varepsilon$. Then choose a positive integer k so large that

$$P\{D \text{ and } T_k \geq v\} \geq (1 - \varepsilon)P\{D\}.$$

But

$$\begin{aligned} P\{D_k(a) \mid \mathcal{F}_k\} &\geq 1 - Q(0, T_k)/a^2 \\ &\geq 1 - \varepsilon \quad \text{on } \{T_k \geq v\}. \end{aligned}$$

So

$$P\{D_k(a) \text{ and } T_k \geq v\} \geq (1 - \varepsilon)P\{T_k \geq v\}.$$

Now

$$P\{D \text{ but not } D_k(a)\} \leq \varepsilon P\{D\} + \varepsilon P\{T_k \geq v\} \leq 2\varepsilon.$$

This, together with (13), shows $\limsup |S_n|/\phi(T_n) \leq a$ on all of D apart from a set of measure 2ε . Since ε was arbitrary, $\limsup_{n \rightarrow \infty} |S_n|/\phi(T_n) \leq a$ a.e. on D . Now let a tend to zero through a sequence. \square

The next argument follows[2].

PROOF OF (7). Choose $\varepsilon_n > 0$ with $\varepsilon_n \rightarrow 0$ so slowly that

$$\varepsilon_n^2/\phi(n)^2 < 1 \quad \text{for all } n = 1, 2, \dots \quad \text{and} \quad \sum_n \varepsilon_n^2/\phi(n)^2 = \infty.$$

Make X_n take the values $\pm \phi(n)/\varepsilon_n$ with chance $\frac{1}{2}\varepsilon_n^2/\phi(n)^2$ each, and the value 0 with remaining chance $1 - \varepsilon_n^2/\phi(n)^2$; make the X_n independent. So $E(X_n) = 0$ and $\text{Var } X_n = 1$. By way of contradiction, suppose

$$P\{\limsup S_n/\phi(n) < \infty\} > 0.$$

By the Kolmogorov 0-1 Law,

$$P\{\limsup S_n/\phi(n) < \infty\} = 1.$$

By symmetry,

$$P\{\liminf S_n/\phi(n) > -\infty\} = 1.$$

So

$$P\{\liminf S_{n-1}/\phi(n-1) > -\infty\} = 1.$$

Since ϕ is non-decreasing,

$$P\{\liminf S_{n-1}/\phi(n) > -\infty\} = 1.$$

So

$$P\{\limsup X_n/\phi(n) < \infty\} = 1.$$

But $\sum \varepsilon_n^2/\phi(n)^2 = \infty$, so $X_n = \phi(n)/\varepsilon_n$ for infinitely many n , and $X_n/\phi(n) = 1/\varepsilon_n \rightarrow \infty$ for these n .

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