

MONOTONICITY OF AN INTEGRAL OF M. KLAŠS¹

BY JAMES REEDS

University of California, Berkeley

For each value of β , $0 < \beta < 2$, the integral

$$\int_{-\infty}^{\infty} \{1 - \exp(-x^{-2} \sin^2 tx)\} |t|^{-1-\beta} dt$$

decreases monotonically as a function of x , $x > 0$. This result is useful in approximating the absolute β th moment of the sum of zero mean i.i.d. random variables.

Let $0 < \beta < 2$; define $g: \mathbb{R} \rightarrow \mathbb{R}$ by

$$g(x) = \int_{-\infty}^{\infty} \{1 - \exp(-x^{-2} \sin^2 tx)\} |t|^{-1-\beta} dt \quad \text{for } x \neq 0$$

and by

$$g(0) = \lim_{x \rightarrow 0} g(x) = \int_{-\infty}^{\infty} (1 - e^{-t^2}) |t|^{-1-\beta} dt.$$

It is easy to check that g is continuously differentiable for $x \neq 0$.

The following result was conjectured by M. Klašs (1978):

THEOREM. $g'(x) \leq 0$ for all $x > 0$.

REMARKS. Actually, Klašs worked with the integral

$$\int_{-\infty}^{\infty} \left\{ 1 - \exp \frac{\cos tx - 1}{x^2} \right\} |t|^{-1-\beta} dt;$$

the double angle formula $\cos 2\theta - 1 = -2 \sin^2 \theta$ shows that his integral is our $2^{-\beta/2} g(2^{-1/2} x)$. He gave a proof of the theorem in the special case $\beta = 1$ and verified it by computer for many other values of β . He used the present result to derive high precision bounds on $E|S_n|^\beta$, where S_n is the sum of i.i.d. random variables.

PROOF. It is clear from the definition of g as an integral that $0 < g(x) < g(0)$ for $x > 0$, and that $\lim_{x \rightarrow \infty} g(x) = 0$. Hence $g'(x) < 0$ for certain values of x arbitrarily close to 0 and for certain other values of x arbitrarily large. A Laplace transform argument below will show that the set of x for which $g'(x) \leq 0$ is an interval; taken together these observations certainly imply that $g'(x) \leq 0$ for all $x > 0$.

The Laplace transform argument is easiest done under the change of variables $s = x^{-2}$. Let

$$(*) \quad k(s) = \int_{-\infty}^{\infty} (1 - e^{-s \sin^2 t}) |t|^{-(1+\beta)} dt;$$

Received March 13, 1978; revised July 25, 1978.

¹This research was prepared with the support of National Science Foundation Grant No. MCS75-10376-A01.

AMS 1970 subject classifications. Primary 60G50; secondary 44A10, 26A48.

Key words and phrases. Laplace transform, total positivity, variation diminishing.

then $k(s) = g(s^{-\frac{1}{2}})s^{\beta/2}$ so if $x > 0$, $g'(x) \leq 0$ if and only if $h(s) = s^{\beta/2}(d/ds)k(s)s^{-\beta/2} \geq 0$.

The claim is that the set of $s > 0$ such that $h(s) \geq 0$ forms an interval. Actually, more is true: $h(s)$ has at most one sign change in $(0, \infty)$. This follows because $h(s)$ is (a constant multiple of) the Laplace transform of a certain function $f_Z - (\beta/2)f_{UZ}$ defined below, which itself has at most one sign change in $(0, \infty)$. (By the “variation diminishing property of the Laplace transform” (Karlin, 1968) Laplace transforming cannot increase the number of sign changes of a function.) The gist of the proof consists in exhibiting $f_Z - (\beta/2)f_{UZ}$, showing that its Laplace transform is a constant multiple of the function $h(s)$, and finally, showing that $f_Z - (\beta/2)f_{UZ}$ has at most one sign change.

By definition

$$h(s) = s^{\beta/2} \frac{d}{ds} (k(s)s^{-\beta/2}) = k'(s) - \frac{\beta}{2s} k(s).$$

It is clear from (*) that $k(0) = 0$, so

$$\begin{aligned} h(s) &= k'(s) - \frac{\beta}{2s} \int_0^s k'(\sigma) d\sigma \\ &= k'(s) - \frac{\beta}{2} Ek'(sU), \end{aligned}$$

where U is a random variable uniformly distributed on $[0, 1]$. Examination of (*) shows that $k'(s)$ is a Laplace transform:

$$k'(s) = \int_{-\infty}^{\infty} \sin^2 t |t|^{-(1+\beta)} e^{-s \sin^2 t} dt.$$

Let T be a random variable (independent of U) with density function $(1/c)\sin^2 t |t|^{-(1+\beta)}$, where $c = \int_{-\infty}^{\infty} \sin^2 t |t|^{-(1+\beta)} dt$. Then $k'(s)$ is just c times the Laplace transform of $Z = \sin^2 T$, that is,

$$k'(s) = c \cdot E(e^{-sZ}).$$

Let f_Z be the density function of the random variable Z and let f_{UZ} be the density function of the random variable UZ . $f_Z(t)$ and $f_{UZ}(t)$ both vanish if t is outside the range $[0, 1]$. Then

$$\begin{aligned} h(s) &= c \left(Ee^{-sZ} - \frac{\beta}{2} Ee^{-sUZ} \right) \\ &= c \int_0^1 \left(f_Z(z) - \frac{\beta}{2} f_{UZ}(z) \right) e^{-sz} dz, \end{aligned}$$

as promised.

All that remains to be proven is that the function $f_Z - (\beta/2)f_{UZ}$ has at most one sign change in $(0, 1)$.

To see this define the random variable T' to be the mod π residue of T , so $P(0 \leq T' < \pi) = 1$ and $P(\exists n \in \mathbb{Z} \text{ such that } T' = T + n\pi) = 1$. The density function of T' is equal to 0 for $t < 0$ and $t > \pi$, and for $0 < t < \pi$ is given by

$(1/c)\sin^2 t \phi(t)$ where

$$\phi(t) = \sum_{n=-\infty}^{\infty} |t - n\pi|^{-1-\beta}.$$

Calculus, change of variables, and the fact that $\phi(t) = \phi(\pi - t)$, yields

$$A \left(f_Z(z) - \frac{\beta}{2} f_{UZ}(z) \right) = \tan t \phi(t) - \beta \int_0^{\pi/2} \phi(u) du = l(t), \quad \text{say,}$$

where $A > 0$ is some suitable constant and $t = \arcsin z^{1/2}$ for $0 < z < 1$, i.e., for $0 < t < \pi/2$. We prove $l'(t) \geq 0$ on $(0, \pi/2]$; this shows $f_Z - \frac{\beta}{2} f_{UZ}$ is monotone on $(0, 1]$.

Differentiating, we obtain

$$l'(t) = \sec^2 t \phi(t) + \tan t \phi'(t) + \beta \phi(t);$$

we show this is positive by arguing term by term in the summation defining ϕ . Since

$$\begin{aligned} \phi'(t) &= \sum_{n=-\infty}^{\infty} \frac{-(1+\beta)}{|t - \pi n|^{2+\beta}} \cdot \operatorname{sgn}(t - \pi n) \\ &= -(1+\beta) \sum_{n=-\infty}^{\infty} \frac{1}{|t - \pi n|^{1+\beta}} \frac{1}{t - \pi n}, \end{aligned}$$

the n th term in $l'(t)$ is

$$\frac{1}{|t - \pi n|^{1+\beta}} \left\{ \sec^2 t + \beta - (1+\beta) \frac{\tan t}{t - \pi n} \right\}.$$

We argue that the quantity in braces is ≥ 0 for each t in $(0, \pi/2]$ and each integer n .

Since $\sec^2 t$ and $\tan t$ are periodic with period π , it suffices to show that

$$(**) \quad \sec^2 w + \beta \geq (1+\beta) \frac{\tan w}{w}$$

for all w of the form $w = t - \pi n$, where n is an integer and $0 < w \leq \pi/2$.

By periodicity, and since $\tan w \geq 0$ for all our w 's, it suffices to check the "worst" case $0 < w \leq \pi/2$. Further, $\tan w \geq w$ for each such w , so the function $\sec^2 w + \beta - (1+\beta)(\tan w/w)$ decreases as β increases. Thus if the inequality $(**)$ is verified for $\beta = 2$ it is automatically true for all $\beta < 2$. When $\beta = 2$, $(**)$ reduces to

$$w(1 + 2 \cos^2 w) \geq 3 \cos w \sin w$$

for all $w \in (0, \pi/2]$; the double angle formulae imply that this is equivalent to

$$\lambda(\theta) = \theta(2 + \cos \theta) - 3 \sin \theta \geq 0$$

for all $\theta \in (0, \pi]$. Now,

$$\lambda'(\theta) = 2 - 2 \cos \theta - \theta \sin \theta,$$

$$\lambda''(\theta) = \sin \theta - \theta \cos \theta,$$

and

$$\lambda'''(\theta) = \theta \sin \theta \geq 0 \text{ in } (0, \pi].$$

Further, $\lambda(0) = \lambda'(0) = \lambda''(0) = 0$, so by Taylor's theorem, if $0 < \theta \leq \pi$,

$$\lambda(\theta) = \int_0^\theta \lambda'''(t) \frac{(\theta - t)^2}{2} dt \geq 0. \quad \square$$

Acknowledgement. I am grateful to Mike Klass for having introduced this problem to me and for helpful suggestions in writing this note.

REFERENCES

- [1] KARLIN, S. (1968). *Total Positivity*. Stanford Univ. Press.
- [2] KLASS, M. (1980). Precision bounds for the relative error in the approximation of $E|S_n|$ and extensions. *Ann. Probability* 8 350-367.

DEPARTMENT OF STATISTICS
UNIVERSITY OF CALIFORNIA
BERKELEY, CALIFORNIA 94720