

## MAXIMA OF PARTIAL SUMS AND A MONOTONE REGRESSION ESTIMATOR

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Let  $\{t_k\}$  be a sequence of points in  $d$ -dimensional Euclidean space. Let  $\{X_k\}$  be a sequence of random variables with zero mean, i.i.d. or nearly so. If  $\mathcal{A}$  is a class of subsets of  $R^d$ , let

$$M_n(\omega) = \sup_{A \in \mathcal{A}} \sum_{\{k < n : t_k \in A\}} X_k(\omega).$$

$M_n$  is related to a commonly used estimator in monotone regression. Under various conditions on  $\mathcal{A}$  and the points  $\{t_k\}$ , we study the a.s. convergence to zero of  $M_n/n$  as  $n \rightarrow \infty$ .

**0. Introduction.** Let  $\{t_k\}$  be a sequence of points (not necessarily distinct) in  $d$ -dimensional Euclidean space. Let  $\{X_k\}$  be a sequence of random variables defined on a common probability space  $(\Omega, \mathcal{T}, P)$  and centered at their means; we think of  $X_k$  as being associated with the point  $t_k$  for  $k = 1, 2, \dots$ . Let  $F(y) = \sup_k P\{|X_k| \geq y\}$ ; for our purposes we will assume either that

(0.1) the  $\{X_k\}$  are independent with mean zero,  $F(y) \rightarrow 0$  as  $y \rightarrow \infty$ , and  $\int_0^\infty y |dF(y)| < \infty$ ;

or that

(0.2) the  $\{X_k\}$  form a stationary ergodic sequence with mean zero.

It is well known that the strong law of large numbers holds when either (0.1) or (0.2) is satisfied.

For  $A \subset R^d$  let

(0.3)  $S_n(A, \omega) \equiv \sum_{\{k < n : t_k \in A\}} X_k(\omega)$  where  $\sum_{\emptyset} = 0$ .

If  $\mathcal{A}$  is some collection of subsets of  $R^d$  we define, for  $\omega \in \Omega$ ,

(0.4)  $M_n(\omega) = \sup_{A \in \mathcal{A}} S_n(A, \omega)$ .

The question we consider in this note is:

(0.5) Under what conditions on the class  $\mathcal{A}$  and the sequence  $\{t_k\}$  does  $M_n/n \rightarrow 0$  a.s.?

This question arises in proving the (strong) consistency of a commonly used estimator in monotonic regression problems. For the regression motivation the

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reader is referred to Hanson, Pledger and Wright (1973) or to Wright (1979). In the former paper the question (0.4) is resolved under condition (0.1) when  $d = 1$  and  $\mathcal{Q}$  is the class  $u$  of *upper layers*, defined below (in the sequel the partial ordering  $<$  on  $R^d$  is always taken to be the coordinatewise ordering):

DEFINITION 0.1. Let  $U \subset R^d$ . If whenever  $s \in U$  and  $s < t$  it follows that  $t \in U$ ,  $U$  is called an *upper layer*. The complement  $L$  of an upper layer is called a *lower layer*; it clearly has the property that  $s \in L$  and  $s \geq t$  imply  $t \in L$ .

When  $d = 1$  the upper layers are simply intervals half-infinite to the right. There are two natural generalizations of this class of sets to the case  $d \geq 2$ : the class  $\mathcal{R}$  of "half-infinite rectangles"  $\{y : y \geq x\}$  for some  $x \in R^d$ , and the class  $u$  of upper layers, which is considerably larger. For statistical purposes, it is again the class of upper layers which is used to define an estimator with optimal (least-squares) fit (Brunk, Ewing, and Utz (1957); see also [4], page 403).

For  $d \geq 2$ , a simple example of Wright (1979) shows that some restriction on the  $\{t_k\}$  is essential when  $\mathcal{Q} = u$ , the upper layers, even when the  $\{X_k\}$  are i.i.d.: Take  $d = 2$  and let the points  $\{t_k\}$  be distinct and lie on the line  $y = -x$ . Given any  $\omega$ , there is an upper layer containing precisely the set of  $t_k, k \leq n$ , for which  $X_k \geq 0$ ; hence  $M_n(\omega) = \sum_{k=1}^n X_k^+(\omega)$  and  $M_n/n \rightarrow E(X_1^+)$  a.s. by the strong law of large numbers. We show in Section 2 that if the class  $\mathcal{Q}$  is taken to be  $\mathcal{R}$ , and (0.1) holds,  $M_n/n \rightarrow 0$  a.s. for *any* choice of points  $\{t_k\}$ . Although  $\mathcal{R}$  is not of statistical interest, this result is an exact analogue of that of Hanson, et al. for  $d = 1$  (and gives a different approach to that case).

The case of perhaps the greatest interest, for any value of  $d$ , is that in which the points  $\{t_k\}$  are a realization of an i.i.d. sequence  $\{T_k\}$  of  $R^d$ -valued random vectors. In the case considered here, the  $\{T_k\}$  are assumed to be defined on a probability space  $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$  and  $\{X_k\}$  and  $\{T_k\}$  are taken to be independent sequences on the product space  $\Omega \times \tilde{\Omega}$ . Our problem is then to determine conditions under which (0.6) below holds:

$$(0.6) \quad \text{For a.e. } \tilde{\omega}, \sup_{A \in \mathcal{Q}} 1/n \sum_{i=1}^n 1_A(T_i) X_i(\omega) \rightarrow 0 \quad \text{for a.e. } \omega.$$

Wright (1979) has shown that (0.6) holds under condition (0.1) for the class  $u$ , when the  $\{T_k\}$  are i.i.d. with a distribution having no singular (with respect to Lebesgue measure on  $R^d$ ) continuous part. In Section 1 we show that (0.6) holds under (0.1) or (0.2) for the class  $u$  whenever the sequence  $\{T_k\}$  is i.i.d., with a distribution whose continuous part does not charge the boundary of any upper layer; since all such boundaries are of Lebesgue measure zero (Brunk, et al. (1957)) this extends Wright's result. This condition appears more natural than Wright's and the proof (using a result of Steele (1978) on empirical discrepancies) is completely different.

In Section 3 we note briefly that when  $\{X_k\}$  satisfies (0.2) and the  $\{T_k\}$  form an i.i.d. sequence, (0.6) may be cast as a problem of identifying the (constant) limit of a subadditive process.

**1. Sufficient conditions for (0.6).** We suppose now that the sequence  $\{t_k\}$  is a realization of an i.i.d. sequence  $\{T_k\}$  of random vectors in  $R^d$ . Denote by  $\mu$  the probability measure of the  $T_k$ .

**THEOREM 1.1.** *Let  $u$  be the class of upper layers, and let  $\{X_k\}$  satisfy (0.1) or (0.2). If  $\{T_k\}$  is an i.i.d. sequence independent of  $\{X_k\}$ , with the property that the continuous part of its measure  $\mu$  does not charge the boundary of any upper layer, then (0.6) holds.*

**PROOF.** We note first that the  $X_k$  may be assumed to be uniformly bounded. For, given  $N > 0$ ,

$$(1.1) \quad \frac{M_n}{n} \leq \sup_{A \in u} \frac{1}{n} \sum_{i=1}^n 1_A(T_i) 1_{\{|X_i| > N\}} X_i + \sup_{A \in u} \frac{1}{n} \sum_{i=1}^n 1_A(T_i) 1_{\{|X_i| \leq N\}} X_i.$$

The first term on the right-hand side of (1.1) is bounded above by  $\frac{1}{n} \sum_{i=1}^n X_i + 1_{\{|X_i| > N\}}$  and it is easy to show that, given  $\epsilon > 0$ ,

$$(1.2) \quad \limsup_n \frac{1}{n} \sum_{i=1}^n X_i + 1_{\{|X_i| > N\}} < \epsilon \text{ a.s.} \quad \text{if } N > N_0.$$

Let  $Y_i = X_i 1_{\{|X_i| \leq N\}}$ ; the second term on the right-hand side of (1.1) is bounded above by

$$(1.3) \quad \sup_{A \in u} \frac{1}{n} \sum_{i=1}^n 1_A(T_i) \{Y_i - E(Y_i)\} + \frac{1}{n} \sum_{i=1}^n E(Y_i),$$

and the second term in (1.3) can be made arbitrarily small (since the  $X_i$  have mean zero) by choosing  $N$  sufficiently large. So it will suffice to assume that for all  $k$ ,  $|X_k| \leq N$  a.s.

Next we observe that

$$(1.4) \quad \sup_{A \in u} \frac{1}{n} \sum_{i=1}^n 1_A(T_i) X_i \leq \sup_{A \in u} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) X_i \right| + \frac{1}{n} \left| \sum_{i=1}^n X_i \right|.$$

Under conditions (0.1) or (0.2), the second term on the right-hand side of (1.4) tends a.s. to zero as  $n \rightarrow \infty$ ; the first term is (except for the presence of the  $X_i$ ) the empirical discrepancy for the class  $u$ . If  $\mathcal{L}$  denotes the class of lower layers, it was shown by Blum (1955) that

$$(1.5) \quad \sup_{A \in \mathcal{L}} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) \right| \rightarrow 0 \text{ a.s.} \quad \text{as } n \rightarrow \infty,$$

when  $\mu$  is absolutely continuous with respect to Lebesgue measure and the  $\{T_k\}$  are i.i.d.. Recently Steele (1978), drawing on a fundamental result of Vapnik and Chervonenkis (1971), showed that (1.5) holds for i.i.d.  $\{T_k\}$ , provided that the continuous part of  $\mu$  does not charge the boundary of any lower layer ([8, Corollary 7.2]); clearly the result then holds for upper layers as well.

Now define a sequence  $\{\bar{X}_k\}$  of random variables such that  $\bar{X}_k$  takes only a finite number  $\{c_1, c_2, \dots, c_m\}$  of values, and  $|X_k(\omega) - \bar{X}_k(\omega)| < \epsilon$  for all  $\omega \in \Omega$  and all  $k$  (where  $\epsilon > 0$  is prescribed). Clearly  $\{\bar{X}_k\}$  can be taken to satisfy (0.1) or (0.2).

Then

$$(1.6) \quad \sup_{A \in \mathfrak{U}} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) X_i \right| \\ \leq \sup_{A \in \mathfrak{U}} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) (X_i - \bar{X}_i) \right| + \sup_{A \in \mathfrak{U}} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) \bar{X}_i \right|.$$

The first term on the right-hand side of (1.6) is less than  $2\epsilon$ . Consider the second term as a function of  $\tilde{\omega}$ , for  $\omega$  fixed.

$$(1.7) \quad \sup_{A \in \mathfrak{U}} \frac{1}{n} \left| \sum_{i=1}^n (1_A(T_i) - \mu(A)) \bar{X}_i(\omega) \right| \\ \leq \sum_{k=1}^m \frac{|c_k|}{n} \sup_{A \in \mathfrak{U}} \left| \sum_{\{i < n : \bar{X}_i(\omega) = c_k\}} (1_A(T_i) - \mu(A)) \right|.$$

For each  $k = 1, 2, \dots, m$ ,  $\{i : \bar{X}_i(\omega) = c_k\}$  is a (finite or infinite) subsequence of integers, independent of  $\{T_k\}$ . The result of Steele quoted above then shows that the right-hand side of (1.7) converges to 0 as  $n \rightarrow \infty$ , for almost all  $\tilde{\omega} \in \tilde{\Omega}$ ; by (1.4) and (1.6), the proof is complete.

**2.  $\mathcal{Q} = \mathfrak{R}$  and  $\{t_k\}$  arbitrary.** We turn now to the case when  $\mathcal{Q} = \mathfrak{R}$ , the "half-infinite rectangles."

**THEOREM 2.1.** *Let  $\mathcal{Q} = \mathfrak{R}$  and suppose that  $\{X_k\}$  satisfies (0.1). Then for any sequence  $\{t_k\}$  of points,*

$$\frac{M_n}{n} \rightarrow 0 \text{ a.s.} \quad \text{as } n \rightarrow \infty.$$

**PROOF.** We give the proof for the case  $d = 2$  only; the extension to higher dimensions is straightforward. For each  $n$ , let

$$X'_n(\omega) = X_n(\omega) \quad \text{if } |X_n(\omega)| \leq n \\ = 0 \quad \text{if } |X_n(\omega)| > n.$$

Let  $X$  be a random variable defined on  $(\Omega, \mathfrak{F}, P)$  with

$$(2.1) \quad P\{|X| \geq n\} = F(n)$$

for every positive integer  $n$ , where  $F(y)$  is defined just above (0.1). Then

$$(2.2) \quad \sum_n P\{X_n \neq X'_n\} = \sum_n P\{|X_n| > n\} \leq \sum_n P\{|X| > n\} < \infty$$

by virtue of (0.1); so by Borel-Cantelli,  $\{X_n\}$  and  $\{X'_n\}$  are equivalent sequences, and it suffices to prove that

$$\frac{M'_n}{n} \rightarrow 0 \text{ a.s., where } M'_n = \sup_{A \in \mathfrak{R}} \sum_{\{k < n; t_k \in A\}} X'_k.$$

But clearly,

$$(2.3) \quad \frac{M'_n}{n} \leq \sup_{A \in \mathfrak{R}} \frac{1}{n} \sum_{\{k < n; t_k \in A\}} \{X'_k - E(X'_k)\} + \frac{1}{n} \sum_{k=1}^n |E(X'_k)|,$$

and the second term on the right-hand side of (2.3) tends to zero as  $n \rightarrow \infty$ , since

$E(X'_k) \rightarrow 0$  as  $k \rightarrow \infty$ . So it will be enough to prove the theorem in the case when  $|X_n| \leq n$  a.s.

Fix positive integers  $N$  and  $N_0$  with  $N_0 > N$ ; there is a (finite) rectangular box  $B$  in the plane which contains the points  $t_1, t_2, \dots, t_{N_0}$  (not assumed distinct). At each of these points draw a horizontal and a vertical line through  $B$  so that the box  $B$  is partitioned into a rectangular grid.

Now consider the rectangular parallelepiped  $V$  with the box  $B$  as its base and height  $N_0$ . Move the point  $t_k$  (lying in  $B$ ) vertically up so that its  $z$ -coordinate is  $k$ . Replicate the original grid on each plane  $z = k$  and draw a line through each point  $t_k$  parallel to the  $z$ -axis; this partitions  $V$  into a cubic grid  $G$ . Place the random variable  $X_k$  at the new point  $t_k$  in the grid, and place random variables which are identically zero at every other point of  $G$ .

Consider the upper right corner of  $B$  as the origin and label each point of the grid with three integral coordinates (counting  $x$  from right to left and  $y$  from top to bottom of  $B$ ). Given a point  $\mathbf{k}$  in the grid, let  $X_{\mathbf{k}}$  denote the (possibly zero) random variable at  $\mathbf{k}$ , and let

$$b_{\mathbf{k}} = \begin{cases} \text{the } z\text{-coordinate of } \mathbf{k} & \text{if this coordinate exceeds } N \\ = N & \text{otherwise.} \end{cases}$$

Then if  $A \in \mathcal{R}$  (in the original plane  $z = 0$ ), to form  $S_n(A)$  we sum over all  $\mathbf{k}$  with  $z$ -coordinate not greater than  $n$  whose projections on  $B$  lie in  $A$ . Therefore

$$(2.4) \quad \max_{N \leq n \leq N_0} \frac{M_n}{n} = \max_{N \leq n \leq N_0} \sup_A \frac{S_n(A)}{n} \leq \max_{\mathbf{k} \in G, N_0 \geq k_3 \geq N} (b_{\mathbf{k}})^{-1} \sum_{\mathbf{j} \leq \mathbf{k}} X_{\mathbf{j}} \\ = \max_{\mathbf{k} \in G, N_0 \geq k_3 \geq N} S_{\mathbf{k}} / b_{\mathbf{k}}.$$

By the Hajek-Renyi inequality established in [7], for any  $\lambda > 0$ :

$$(2.5) \quad P \left\{ \max_{N \leq n \leq N_0} \frac{M_n}{n} \geq \lambda \right\} \leq P \left\{ \max_{N \leq k_3 \leq N_0} |S_{\mathbf{k}}| / b_{\mathbf{k}} \geq \lambda \right\} \\ \leq \frac{c}{\lambda^2} \left\{ \frac{1}{N^2} \sum_{k=1}^N \sigma^2(X_k) + \sum_{k=N+1}^{N_0} \frac{\sigma^2(X_k)}{b_k^2} \right\},$$

where  $c$  is a constant depending only on the dimension  $d$ . But we have truncated the  $\{X_k\}$  so that by standard arguments (cf. Chung (1974), page 126) the series  $\sum_k \frac{\sigma^2(X_k)}{k^2}$  converges; by choosing  $N$  large enough we can therefore get

$$(2.6) \quad P \left\{ \sup_{n \geq N} \frac{M_n}{n} \geq \lambda \right\} < \varepsilon.$$

for any prescribed  $\varepsilon > 0$ , proving Theorem 2.1.

**3. Subadditive processes.** Assume now that  $\{T_k\}$  is a stationary sequence in  $R^d$ , with measure  $\mu$ ; we assume also that  $\{X_k\}$  is a stationary sequence with finite mean, independent of  $\{T_k\}$ .

Under these conditions it is not difficult to verify that, for any class  $\mathcal{Q}$ , the process

$$Y_{mn} \equiv \sup_{A \in \mathcal{Q}} \sum_{i=m+1}^n 1_A(T_i) X_i$$

is a *subadditive process* in the sense of Kingman (1968). By Kingman's ergodic theorem for subadditive processes,

$$\sup_{A \in \mathcal{Q}} \frac{1}{n} \sum_{i=1}^n 1_A(T_i) X_i \rightarrow C(\mathcal{Q}, \mu) \quad \text{a.s.} \quad \text{as } n \rightarrow \infty;$$

under certain conditions (including, but not limited to, the case when both  $\{X_k\}$  and  $\{T_k\}$  are mixing, and  $\mathcal{Q}$  and  $\mu$  are arbitrary), the limit will be a constant. The problem then reduces to finding the value of  $C(\mathcal{Q}, \mu)$ .

When  $\{X_k\}$  is ergodic and  $\{T_k\}$  is i.i.d., we have seen that if  $\mathcal{Q} = u$  and  $\mu$  lives (and is continuous) on the line  $y = x$ , then  $C(\mathcal{Q}, \mu) = E(X_1^+)$ , and that if  $\mathcal{Q} = u$  and the continuous part of  $\mu$  does not charge the boundary of any upper layer,  $C(\mathcal{Q}, \mu) = 0$ . If  $\mathcal{Q}$  is the class of convex sets and the continuous part of  $\mu$  does not charge the boundary of any convex set, it again turns out that  $C(\mathcal{Q}, \mu) = 0$ ; this follows as in Theorem 1.1 from a result of Ranga Rao (1962) for the empirical discrepancy. It would be of interest to know what values between 0 and  $E(X_1^+)$  (if any) can be taken by  $C(\mathcal{Q}, \mu)$  for other interesting choices of  $\mathcal{Q}$  and  $\mu$ .

The formulation above in terms of subadditive processes introduces an apparent simplification of the solution; for an ergodic subadditive process  $z_{mn}$ , convergence a.s. of  $z_{0n}/n$  to zero is equivalent to convergence to zero of  $E(z_{0n}/n)$  (see [5]). The simplification is largely illusory, however, for the latter verification does not seem substantially simpler.

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