

THE NONPARAMETRIC ORDERING: (1001) → (0110)

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Let $Z = (Z_1, Z_2, \dots, Z_N)$ be a random vector with $Z_i = 1(0)$ if the i th smallest in absolute value in a sample of N from the density $f(x)$ is positive (negative). Then

$$P(Z = z) = N! \int_{0 \leq y_1 \leq \dots \leq y_N} \prod_{i=1}^N [f^{1-z_i}(-y_i) f^{z_i}(y_i) dy_i].$$

In the case of normal slippage to the right (i.e., $f(x) = f(x, \mu)$ is $N(\mu, 1)$, $\mu > 0$), Savage [1] obtains a simple ordering of the 2^N possible values of Z for $N = 3$, namely:

$$111 \rightarrow 011 \rightarrow 101 \rightarrow 001 \rightarrow 110 \rightarrow 010 \rightarrow 100 \rightarrow 000,$$

where $\text{Prob}(Z = z) > \text{Prob}(Z = z')$ if and only if $z \rightarrow z'$.

For $N = 4$, Savage gives a partial ordering of the 2^4 possible values of Z . The following theorem orders two more values of Z .

THEOREM 1: *If X_1, X_2, X_3, X_4 are NID $(\mu, 1)$, with $\mu > 0$, then*

$$D = \text{Prob}[Z = (1001)] - \text{Prob}[Z = (0110)] > 0.$$

PROOF:

$$\begin{aligned} D &= \text{Prob}[Z = (1001)] - \text{Prob}[Z = (0110)] \\ &= 4!/(2\pi)^2 \int_{0 \leq y_1 \leq \dots \leq y_4} \exp(-\frac{1}{2}\Sigma y^2 - 2\mu^2) \{ \exp(\mu(y_4 - y_3 - y_2 + y_1)) \\ &\quad - \exp(-\mu(y_4 - y_3 - y_2 + y_1)) \} \prod dy_i. \\ D &= 4!e^{-2\mu^2}/(2\pi)^2 \int_{0 \leq y_1 \leq \dots \leq y_4} 2 \sinh \mu(y_4 - y_3 - y_2 + y_1) e^{-\Sigma y_i^2/2} \prod dy_i. \end{aligned}$$

Now make the transformation $y_i = \sum_{j=1}^i w_j$. The Jacobian is 1 and the region of integration becomes $0 \leq w_i \leq \infty, i = 1, \dots, 4$. Hence,

$$D = c \iiint\limits_0^\infty \sinh \mu(w_4 - w_2) \exp(-\Sigma(\Sigma w^2)/2) \prod dw_i,$$

where

$$c = 2 \cdot 4! e^{-2\mu^2} / (2\pi)^2.$$

Noting that the integral is positive for $w_4 > w_2$ and negative for $w_2 < w_4$, break up the region of integration into two parts in two different ways as follows

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$$D = \int_{w_1=0}^{\infty} \int_{w_3=0}^{\infty} \int_{w_2=0}^{\infty} \int_{w_4=w_2}^{\infty} + \int_{w_1=0}^{\infty} \int_{w_3=0}^{\infty} \int_{w_2=0}^{\infty} \int_{w_4=0}^{w_2} = (\text{say}) D_1^+ + D_1^-,$$

$$D = \int_{w_1=0}^{\infty} \int_{w_3=0}^{\infty} \int_{w_4=0}^{\infty} \int_{w_2=0}^{w_4} + \int_{w_1=0}^{\infty} \int_{w_3=0}^{\infty} \int_{w_4=0}^{\infty} \int_{w_2=w_4}^{\infty} = (\text{say}) D_2^+ + D_2^-.$$

Combining these two results, D may be written as $D = D_1^+ + D_2^-$. Setting $w_2 = s$ and $w_4 = t$ in D_1^+ and $w_2 = t, w_4 = s$ in D_2^- gives

$$D = c \int_0^{\infty} \int_0^{\infty} \int_0^{\infty} \int_s^{\infty} \sinh \mu(t - s) \exp(-\frac{1}{2}\{(w_1^2 + (w_1 + s)^2 + (w_1 + s + w_3)^2 + (w_1 + s + w_3 + t)^2)\}) dt ds dw_3 dw_1$$

$$- c \int_0^{\infty} \int_0^{\infty} \int_0^{\infty} \int_s^{\infty} \sinh \mu(t - s) \exp(-\frac{1}{2}\{(w_1^2 + (w_1 + t)^2 + (w_1 + t + w_3)^2 + (w_1 + t + w_3 + s)^2)\}) dt ds dw_3 dw_1,$$

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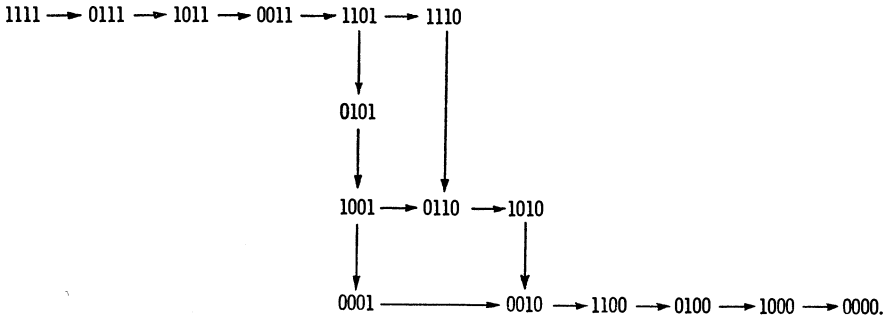
$$[\exp(-\frac{1}{2}\{(w_1 + s)^2 + (w_1 + s + w_3)^2\}) - \exp(-\frac{1}{2}\{(w_1 + t)^2 + (w_1 + t + w_3)^2\})] dt ds dw_3 dw_1.$$

Thus D is positive if the difference of the exponentials in the square brackets is positive, but this is obviously true since $t > s$ over the range of integration.

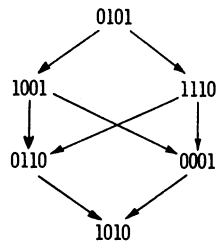
THEOREM 2: *If $X_1, \dots, X_N (N \geq 4)$ are NID $(\mu, 1)$ with $\mu > 0$, then $D = \text{Prob}(Z = z) - \text{Prob}(Z = z') > 0$ where z and z' are identical except that $z_1 = z'_2 = z'_3 = z_4 = 1$ and $z'_1 = z_2 = z_3 = z'_4 = 0$.*

PROOF: The proof of this Theorem follows from Theorem 1 in exactly the same way that Savage's Theorem 6.1 follows from Sobel's Theorem ([1], footnote p. 1024).

Using the results of Theorem 1 and Savage's Theorem 6.1, the partial ordering for $N = 4$ becomes, for normal slippage to the right,



Some preliminary Monte Carlo analysis suggests that the following orderings may be valid, although no analytic proofs are available:



REFERENCE

- [1] I. RICHARD SAVAGE, "Contributions to the theory of rank order statistics—the one-sample case," *Ann. Math. Stat.*, Vol. 30 (1959), pp. 1018–1023.