

A NOTE ON MINIMUM DISCRIMINATION INFORMATION¹

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This note contains a simple proof of the minimum discrimination information theorem in Kullback (1959), pp. 36–39 and an affirmative answer to a suggestion in a personal communication from Dr. I. J. Good that the theorem could be applied even to random elements of a Banach space.

Let X be a space of points x , \mathcal{S} a σ -field of sets of X , and P_2 a probability measure on \mathcal{S} . Let $T(x)$ be a real valued \mathcal{S} -measurable function such that

$$(1) \quad M_2 = \int_X \exp(T(x)) dP_2 < \infty$$

and let the probability measure P^* be defined by

$$(2) \quad P^*(A) = \int_A (\exp(T(x))/M_2) dP_2, \quad \text{for } A \in \mathcal{S}.$$

Suppose that $T(x)$ is P^* -integrable, and let

$$(3) \quad \theta = \int_X T(x) dP^*.$$

Now let P_1 be an arbitrary probability measure on \mathcal{S} . If $P_1 \ll P_2$ define

$$(4) \quad I(P_1, P_2) = \int_X [\log(dP_1/dP_2)] dP_1,$$

otherwise define $I(P_1, P_2) = \infty$. It is clear that $I(P^*, P_2)$ is finite; in fact, from (2) and (3),

$$(5) \quad I(P^*, P_2) = \theta - \log M_2.$$

THEOREM. *If P_1 is a probability measure on \mathcal{S} such that T is P_1 -integrable and*

$$(6) \quad \int_X T(x) dP_1 = \theta,$$

then

$$(7) \quad I(P_1, P_2) \geq I(P^*, P_2) = \theta - \log M_2$$

with equality if and only if $P_1 = P^$ on \mathcal{S} .*

PROOF. If $I(P_1, P_2) = \infty$ there is nothing to prove. Suppose then that $I(P_1, P_2) < \infty$. In this case $P_1 \ll P_2$, and we write $f(x) = dP_1/dP_2$. Then

$$(8) \quad I(P_1, P_2) = \int_X f(x) \log f(x) dP_2$$

and

$$(9) \quad I(P^*, P_2) = \int_X f^*(x) \log f^*(x) dP_2$$

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where $f^* = dP^*/dP_2 = e^T/M_2$ by (2). In view of (3), (6) and f^* as just defined, we have that

$$(10) \quad \int_X f(x) \log f^*(x) dP_2 = \int_X f^*(x) \log f^*(x) dP_2.$$

We see from (8), (9) and (10) that it suffices to show that the left-hand side of (10) cannot exceed the right-hand side of (8), that is, that

$$(11) \quad \int_X f(x) \log [f^*(x)/f(x)] dP_2 \leq 0.$$

Since $\log z \leq z - 1$ for all $z \geq 0$ with equality only for $z = 1$, and since

$$(12) \quad \int_X f dP_2 = \int_X f^* dP_2 = 1,$$

it follows that (11) holds with equality only if $P_1 \equiv P^*$.

The foregoing applies to *any* probability space (X, S, P_2) and any statistic T . Suppose now that

- (a) X is a real separable Banach-space;
- (b) X^* is the dual space of X consisting of all the continuous linear functionals $x^*(x)$ on X ;
- (c) S is the σ -algebra generated by the continuous linear functionals on X ;
- (d) m_1, m_2 are elements of X ;
- (e) P_2 is a probability measure defined on S with the mean value m_2 defined via a Pettis integral, that is, $x^*(x)$ is P_2 -integrable for each x^* and m_2 is a (necessarily unique) element in X such that $x^*(m_2) = \int_X x^*(x) dP_2$ for all x^* in X^* .

We write $m_2 = E_2(x)$ [cf. Grenander (1963), p. 128, Mourier (1953), p. 164; (1956), p. 231, Pettis (1938)]. Let x^* be a fixed continuous linear functional and take $T(x) \equiv x^*(x)$ and write M_2 in (1) as $M_2(x^*)$. Let P_1 be a probability measure and m_1 an element of X such that

$$(13) \quad \int_X y^*(x) dP_1 = y^*(m_1) = \int_X (y^*(x) \exp x^*(x)/M_2(x^*)) dP_2$$

for all continuous linear functionals y^* in X^* , that is, $m_1 = E_1(x)$. Then (13) holds in particular when $y^* = x^*$, that is, P_1 satisfies (6) with $\theta = x^*(m_1)$ defined by (3), and we have (7) with $\theta - \log M_2 = x^*(m_1) - \log M_2(x^*)$.

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