

**ON CERTAIN DISTRIBUTION PROBLEMS BASED ON  
POSITIVE DEFINITE QUADRATIC FUNCTIONS IN  
NORMAL VECTORS**

BY C. G. KHATRI

*Gujarat University*

**1. Introduction and summary.** Let  $X: p \times n$  be a matrix of random real variates such that the column vectors of  $X$  are independently and identically distributed as multivariate normals with zero mean vectors. Then a positive definite quadratic function in normal vectors is defined as  $XLX'$  where  $L$  is a symmetric positive definite (p.d.) matrix with real elements. In the analysis of variance, such functions appear. In the previous study, Khatri [14], [16], has established the necessary and sufficient conditions for the independence and the Wishartness of such functions. In this paper, we study the distribution of a positive definite quadratic function and the distribution of  $Y'(XLX')^{-1}Y$  where  $Y: p \times m$  is independently distributed of  $X$  and its columns are independently and identically distributed as multivariate normals with zero mean vectors. Moreover, we study the distribution of the characteristic (ch.) roots of  $(YY')(XLX')^{-1}$  and the similar related problems. When  $p = 1$ , the distribution of a p.d. quadratic function in normal variates (central or noncentral) has been studied by a number of people (see references).

In the study of the above and related topics in multivariate distribution theory, we are using zonal polynomials. A. T. James [10], [11], [12], [13], and Constantine [1], [2], have used them successfully and have given the final results in a very compact form, using hypergeometric functions  ${}_pF_q(S)$  in matrix arguments. These functions are defined by

$$(1) \quad {}_pF_q(a_1, \dots, a_p; b_1, \dots, b_q; Z) \\ = \sum_{k=0}^{\infty} \sum_{\kappa} [(a_1)_{\kappa} \dots (a_p)_{\kappa} / (b_1)_{\kappa} \dots (b_q)_{\kappa}] [C_{\kappa}(Z) / k!]$$

where  $C_{\kappa}(Z)$  is a symmetric homogeneous polynomial of degree  $k$  in the latent roots of  $Z$ , called zonal polynomials (for more detail study of zonal polynomials, see the references of A. T. James and Constantine),  $\kappa = (k_1, \dots, k_p)$ ,  $k_1 \geq k_2 \geq \dots \geq k_p \geq 0$ ,  $k_1 + k_2 + \dots + k_p = k$ ;  $a_1, \dots, a_p, b_1, \dots, b_q$  are real or complex constants, none of the  $b_j$  is an integer or half integer  $\leq \frac{1}{2}(m - 1)$  (otherwise some of the denominators in (1) will vanish),

$$(2) \quad (a)_{\kappa} = \prod_{j=1}^m (a - \frac{1}{2}(j - 1))_{k_j} = \Gamma_m(a, \kappa) / \Gamma_m(a), \\ (x)_n = x(x + 1) \dots (x + n - 1), (x)_0 = 1$$

and

$$(3) \quad \Gamma_m(a) = \pi^{\frac{1}{2}m(m-1)} \prod_{j=1}^m \Gamma(a - \frac{1}{2}(j - 1)) \\ \text{and } \Gamma_m(a, \kappa) = \pi^{\frac{1}{2}m(m-1)} \prod_{j=1}^m \Gamma(a + k_j - \frac{1}{2}(j - 1)).$$

---

Received 21 June 1965.



In (1),  $Z$  is a complex symmetric  $m \times m$  matrix, and it is assumed that  $p \leq q + 1$ , otherwise the series may converge for  $Z = 0$ . For  $p = q + 1$ , the series converge for  $\|Z\| < 1$ , where  $\|Z\|$  denote the maximum of the absolute value of ch. roots of  $Z$ . For  $p \leq q$ , the series converge for all  $Z$ . Similarly we define

$$(2b) \quad {}_pF_q^{(m)}(a_1, a_2, \dots, a_p; b_1, \dots, b_q; S, R) = \sum_{k=0}^{\infty} \sum_{\kappa} [(a_1)_{\kappa} \cdots (a_p)_{\kappa} / (b_1)_{\kappa} \cdots (b_q)_{\kappa}] [C_{\kappa}(S)C_{\kappa}(R) / C_{\kappa}(I_m)k!].$$

The Section 2 gives some results on integration with the help of zonal polynomials, the Section 3 derives the distributions based on p.d. quadratic functions, the Section 4 gives the moments of certain statistics arising in the study of multivariate distributions, and the Section 5 gives the results for complex multivariate Gaussian variates.

**2. Some results on integration.** We shall write  $X > X_0$  for  $X - X_0$  to be p.d.,  $O(m)$  for orthogonal group of  $m \times m$  orthogonal matrices and  $C_{\kappa}(X)$  for a zonal polynomial of degree  $k$ .  $R(Z)$  means the real part of  $Z$ . We shall denote

$$(4) \quad \Gamma_m(t, -\kappa) = \pi^{\frac{1}{2}m(m-1)} \prod_{j=1}^m \Gamma(t - k_j - \frac{1}{2}m + \frac{1}{2}j) \text{ and } R(t) > \frac{1}{2}(m - 1) + k_1.$$

LEMMA 1. Let  $S: m \times m$  and  $T: m \times m$  be symmetric matrices. Then

$$(5) \quad \int_{O(m)} C_{\kappa}(SHTH') dH = C_{\kappa}(S)C_{\kappa}(T) / C_{\kappa}(I_m).$$

(See A. T. James [10].)

LEMMA 2. Let  $Z: m \times m$  be a complex symmetric matrix whose real part is p.d. and let  $T: m \times m$  be an arbitrary complex symmetric matrix. Then

$$(6) \quad \int_{S>0} \exp(-\text{tr}ZS) |S|^{t-\frac{1}{2}(m+1)} C_{\kappa}(TS) dS = \Gamma_m(t, \kappa) |Z|^{-t} C_{\kappa}(TZ^{-1})$$

where  $\Gamma_m(t, \kappa)$  is defined in (3) and  $R(t) > \frac{1}{2}(m - 1)$ . (See Constantine [1].)

LEMMA 3. If  $R$  is any p.d.  $m \times m$  matrix, then

$$(7) \quad \int_0^I |S|^{t-\frac{1}{2}m-\frac{1}{2}} |I - S|^{u-\frac{1}{2}m-\frac{1}{2}} C_{\kappa}(RS) dS = \Gamma_m(t, \kappa) \Gamma_m(u) C_{\kappa}(R) / \Gamma_m(t + u, \kappa).$$

(See Constantine [1].)

LEMMA 4. Let  $Z$  be a complex symmetric matrix such that  $R(Z) > 0$ , and let  $T$  be an arbitrary complex symmetric matrix. Then.

$$(8) \quad \int_{S>0} \exp(-\text{tr}ZS) |S|^{t-\frac{1}{2}(m+1)} C_{\kappa}(TS^{-1}) dS = \Gamma_m(t, -\kappa) |Z|^{-t} C_{\kappa}(ZT),$$

where  $R(t) > \frac{1}{2}(m - 1) + k_1$  and  $\Gamma_m(t, -\kappa)$  is defined by (4).

PROOF. First, we shall prove the result for the special case  $Z = I_m$ , the  $m \times m$  identity matrix. Put

$$(9) \quad f(T) = \int_{S>0} \exp(-\text{tr}S) |S|^{t-\frac{1}{2}(m+1)} C_{\kappa}(TS^{-1}) dS.$$

Then  $f(T)$  is clearly a symmetric function of  $T$  (in fact, a homogeneous symmetric polynomial). Hence, making the transformation  $T \rightarrow H'TH$  and integrating  $H$  over  $O(m)$  with the help of (5), we have

$$(10) \quad f(T) = [f(I) / C_{\kappa}(I)] C_{\kappa}(T).$$

To find  $f(I)/C_\kappa(I)$ , let us assume that  $T$  is diagonal. Then Constantine [1] has showed that

$$C_\kappa(T) = d_{\kappa,\kappa} t_1^{k_1} t_2^{k_2} \dots t_m^{k_m} + \text{“lower terms”}$$

and

$$C_\kappa(T S^{-1}) = d_{\kappa,\kappa} \prod_{j=1}^m t_j^{k_j} |(S^{-1})_j|^{k_j - k_{j+1}} + \dots$$

with  $k_{m+1} = 0$  and  $(S^{-1})_j = (s^{uv})$ ,  $u, v, = 1, 2, \dots, j$ . Then using these results in (10) and comparing the coefficients of  $t_1^{k_1} \dots t_m^{k_m}$  from both sides, we get

$$(11) \quad f(I)/C_\kappa(I) = \int_{s>0} \exp(-\text{tr}S) |S|^{t-\frac{1}{2}(m+1)} \prod_{j=1}^m |(S^{-1})_j|^{k_j - k_{j+1}} dS.$$

Let  $S = VV'$  where  $V$  is an upper triangular matrix. The Jacobian of the transformation is  $2^m \prod_{j=1}^m v_{jj}^2$ , and  $|(S^{-1})_j| = \prod_{a=1}^j v_{aa}^{-2}$ . Then using these in (11), we get

$$(12) \quad \begin{aligned} f(I)/C_\kappa(I) &= \int \dots \int \exp\left(-\sum_{a=1}^m \sum_{j=a}^m v_{aj}^2\right) \prod_{j=1}^m (v_{jj}^2)^{t-k_j-\frac{1}{2}(m-j)-1} \\ &\quad \cdot \prod_{j=1}^m dv_{jj}^2 \prod_{a<j} dv_{aj} \\ &= \pi^{\frac{1}{2}(m-1)m} \prod_{j=1}^m \Gamma\left(t - k_j - \frac{1}{2}m + \frac{1}{2}j\right) = \Gamma_m(t, -\kappa), \end{aligned}$$

the range of integration being  $0 \leq v_{aa} \leq \infty, -\infty \leq v_{aj}(a < j) \leq \infty$ . For the general case, substitute  $Z^{\frac{1}{2}}SZ^{\frac{1}{2}}$  for  $S$  in  $f(T)$  with the Jacobian of the transformation  $|Z|^{\frac{1}{2}(m+1)}$

LEMMA 5. *If  $R$  is any arbitrary symmetric complex  $m \times m$  matrix, then*

$$(13) \quad \begin{aligned} \int_{s>0} |S|^{t-\frac{1}{2}(m+1)} |I + S|^{-t-u} C_\kappa(RS) dS \\ = \Gamma_m(t, \kappa) \Gamma_m(u, -\kappa) C_\kappa(R) / \Gamma_m(t + u) \end{aligned}$$

and

$$(14) \quad \begin{aligned} \int_{s>0} |S|^{t-\frac{1}{2}(m+1)} |I + S|^{-t-u} C_\kappa(RS^{-1}) dS \\ = \Gamma_m(t, -\kappa) \Gamma_m(u, \kappa) C_\kappa(R) / \Gamma_m(t + u). \end{aligned}$$

PROOF. By (6), we have for any p.d. matrix  $Z$

$$(15) \quad \int_{s>0} \exp(-\text{tr}ZS) |S|^{t-\frac{1}{2}(m+1)} C_\kappa(RS) |Z|^t dS = \Gamma_m(t, \kappa) C_\kappa(RZ^{-1}).$$

Multiplying both the sides of (15) by  $\exp(-\text{tr}Z) |Z|^{u-\frac{1}{2}(m+1)}$  and integrating over  $Z > 0$ , we get

$$(16) \quad \begin{aligned} \int_{s>0} \Gamma_m(t + u) |S|^{t-\frac{1}{2}(m+1)} |I + S|^{-t-u} C_\kappa(RS) dS \\ = \Gamma_m(t, \kappa) \int_{z>0} \exp(-\text{tr}Z) |Z|^{u-\frac{1}{2}m-\frac{1}{2}} C_\kappa(RZ^{-1}) dZ \end{aligned}$$

and now the use of (8) on the right side of (16) gives (13). If we transform  $S$  to  $S^{-1}$  in (13), then  $t$  and  $u$  will be interchanged and we shall get (14).

LEMMA 6. *Let  $R$  be a p.d. matrix. Then for  $t \geq m/2 + \frac{1}{2}k_1$ ,*

$$(17) \quad \begin{aligned} \int_0^I |S|^{t-\frac{1}{2}(m+1)} |I - S|^{u-\frac{1}{2}(m+1)} C_\kappa(RS^{-1}) dS \\ = \Gamma_m(t, -\kappa) \Gamma_m(u) C_\kappa(R) / \Gamma_m(t + u, -\kappa). \end{aligned}$$

PROOF. The left hand side of (17) is a symmetric function  $F(R)$  of  $R$ , so that, as in the proof of Lemma 4,

$$(18) \quad F(R) = [F(I)/C_\kappa(I)]C_\kappa(R).$$

For obtaining  $F(I)/C_\kappa(I)$ , we note that

$$\int_{S>0} \exp(-\text{tr}ZS) |S|^{t-\frac{1}{2}(m+1)} C_\kappa(Z^{-1}S^{-1}) dS |Z|^t = \Gamma_m(t, -\kappa) C_\kappa(I).$$

Multiplying the expression by  $|Z|^{u-\frac{1}{2}(m+1)} \exp(-\text{tr}Z)$ , integrating over  $Z > 0$  and transforming  $(S^{-1} + I)^{-1}$  to  $S$ , we get

$$\int_0^1 |S|^{t-\frac{1}{2}(m+1)} |I - S|^{u-\frac{1}{2}(m+1)} C_\kappa(S^{-1}) dS \Gamma_m(t + u, -\kappa) = \Gamma_m(u) \Gamma_m(t, -\kappa) C_\kappa(I),$$

which gives the expression for  $F(I)/C_\kappa(I)$ . The use of this result in (18) proves (17).

LEMMA 7. Let  $T$  be any arbitrary complex symmetric matrix. Then

$$(19) \quad \int_{S>0} \exp(-\text{tr}S) |S|^{t-\frac{1}{2}(m+1)} (\text{tr}S)^j C_\kappa(TS) dS \\ = \Gamma_m(t, \kappa) \Gamma(mt + j + k) C_\kappa(T) / \Gamma(mt + k)$$

while

$$(20) \quad \int_{S>0} \exp(-\text{tr}S) |S|^{t-\frac{1}{2}(m+1)} (\text{tr}S)^j C_\kappa(TS^{-1}) dS \\ = \Gamma_m(t, -\kappa) \Gamma(mt + j - k) C_\kappa(T) / \Gamma(mt - k).$$

PROOF. We shall only prove the result (19). We have by (6) for  $q < 1$ ,

$$(21) \quad \int_{S>0} \exp(-\text{tr}S(1 - q)) |S|^{t-\frac{1}{2}(m+1)} C_\kappa(TS) dS \\ = (1 - q)^{-tm-k} \Gamma_m(t, \kappa) C_\kappa(T).$$

Equating the coefficient of  $q^j/j!$  from both the sides of (21), we get (19). (20) can be proved in the same way.

LEMMA 8. Let  $B$  be symmetric matrix of order  $n \times n$  and let  $A$  be a p.d. symmetric  $p \times p$  matrix with  $n \geq p$ . Let  $X: p \times n$  be a real matrix. Then

$$(22) \quad \int_{XX'=S} \exp(\text{tr}AXBX') dX = (\pi)^{\frac{1}{2}pn} \{\Gamma_p(\frac{1}{2}n)\}^{-1} |S|^{\frac{1}{2}(n-p-1)} {}_0F_0^{(n)}(AS, B).$$

PROOF. Since  $\text{tr}(AXBX')$  can be written as the function of a symmetric matrix, it is easy to see from the results of A. T. James [12], [13], and Constantine [1] that

$$(23) \quad \exp(\text{tr}AXBX') = {}_0F_0(X'AXB).$$

Now let

$$(24) \quad g(B) = \int_{XX'=S} {}_0F_0(X'AXB) dX.$$

Since  $g(B)$  is a homogeneous symmetric function in  $B$ , by making the transformation  $B \rightarrow HBH'$  and integrating  $H$  over  $O(n)$  with the help of (5), we get

$$(25) \quad g(B) = \int_{XX'=S} {}_0F_0^{(n)}(AXX', B) dX \\ = \pi^{\frac{1}{2}pn} |S|^{\frac{1}{2}(n-p-1)} \{\Gamma_p(\frac{1}{2}n)\}^{-1} {}_0F_0^{(n)}(AS, B),$$

using Wishart's integral. This proves (22).

**3. Distributions related to the p.d. quadratic functions.**

**THEOREM 1.** Let  $X:p \times n$  be a real random matrix whose density function is

$$(26) \quad (2\pi)^{-\frac{1}{2}pn} |\Sigma|^{-\frac{1}{2}n} |B|^{-\frac{1}{2}p} \exp\left(-\frac{1}{2} \text{tr} \Sigma^{-1} X B^{-1} X'\right)$$

where  $\Sigma:p \times p$  and  $B:n \times n$  are p.d. Then the density function of  $S = X L X'$ ,  $L$  being a  $n \times n$  p.d. matrix, is

$$(27) \quad 2^{-\frac{1}{2}pn} \{\Gamma_p(\frac{1}{2}n)\}^{-1} |L B|^{-\frac{1}{2}p} |\Sigma|^{-\frac{1}{2}n} |S|^{\frac{1}{2}(n-p-1)} \cdot \exp\left(-\frac{1}{2} q^{-1} \text{tr} \Sigma^{-1} S\right) {}_0F_0^{(n)}\left(T, \frac{1}{2} q^{-1} \Sigma^{-1} S\right)$$

where  $q > 0$  and  $T = I_n - q L^{-\frac{1}{2}} B^{-1} L^{-\frac{1}{2}}$ .

**PROOF.** Let us use the transformation  $Y = X L^{\frac{1}{2}}$  in (26). Then the density of  $Y$  is

$$(2\pi)^{-\frac{1}{2}pn} |\Sigma|^{-\frac{1}{2}n} |L B|^{-\frac{1}{2}p} \exp\left(-\frac{1}{2} q^{-1} \text{tr} \Sigma^{-1} Y Y' + \frac{1}{2} q^{-1} \text{tr} \Sigma^{-1} Y T Y'\right)$$

and  $S = Y Y'$ . Now, the use of (22) gives (27).

**THEOREM 2.** The moment generating function of  $S$  defined in Theorem 1 is

$$(28) \quad E\{\exp(\text{tr} Z S)\} = |L B q^{-1}|^{-\frac{1}{2}p} |z|^{-\frac{1}{2}n} \{ {}_1F_0^{(n)}(\frac{1}{2}n; T, z^{-1}) = \prod_{j=1}^n |I_p - \phi_j z^{-1}|^{-\frac{1}{2}} \}$$

where  $\phi_j$ 's are the ch. roots of  $T$ ,  $z = I_p - 2q Z \Sigma$  and  $E$  stands for expectation.

**PROOF.** The first part follows from (27) with the help of (6). For the second part, we can write after a transformation  $X \rightarrow \Sigma^{\frac{1}{2}} X (L q)^{-\frac{1}{2}}$  in  $E \exp(\text{tr} Z X L X')$  as

$$(29) \quad E(\exp(\text{tr} Z S)) = (2\pi)^{-\frac{1}{2}pn} |L B q^{-1}|^{-\frac{1}{2}p} \int_X \exp\left(\frac{1}{2} \text{tr} X T X' - \frac{1}{2} \text{tr} X X'\right) dX.$$

Since  $X X'$  is invariant under post multiplication of  $X$  by an orthogonal matrix, we can consider  $T$  to be a diagonal matrix with  $\phi_j$ 's as diagonal elements. Let  $X = (\mathbf{x}_1, \dots, \mathbf{x}_n)$ . Then (29) can be rewritten as

$$E \exp(\text{tr} Z S) = |L B q^{-1}|^{-\frac{1}{2}p} \left\{ \prod_{j=1}^p \left\{ (2\pi)^{-\frac{1}{2}p} \int_{\mathbf{x}_j} \exp\left[-\frac{1}{2} \mathbf{x}_j'(z - \phi_j I_p) \mathbf{x}_j\right] d\mathbf{x}_j \right\} \right\}$$

and this gives the second part of (28).

**THEOREM 3.** Let  $X:p \times n$  and  $Y:p \times m$  be independently distributed, the density function of  $X$  be given by (26) and the density function of  $Y$  be given by

$$(2\pi)^{-\frac{1}{2}pm} |\Sigma|^{-\frac{1}{2}m} \exp\left(-\frac{1}{2} \text{tr} \Sigma^{-1} Y Y'\right).$$

Then, the density function of  $F = Y'(X L X')^{-1} Y$  if  $m \leq p \leq n$  is given by

$$(30) \quad \Gamma_p(\frac{1}{2}m + \frac{1}{2}n) \{\Gamma_p(\frac{1}{2}n) \Gamma_m(\frac{1}{2}p)\}^{-1} q^{\frac{1}{2}p(m+n)} |B L|^{-\frac{1}{2}p} |I_m + q F|^{-\frac{1}{2}(m+n)} \cdot |F|^{\frac{1}{2}(p-m-1)} {}_1F_0^{(p)}(\frac{1}{2}m + \frac{1}{2}n; T, R^*)$$

where  $q > 0$ ,  $R^* = \begin{pmatrix} (I_m + q F)^{-1} & 0 \\ 0 & I_{p-m} \end{pmatrix}$  and  $T = I_n - q(L B)^{-1}$ .

PROOF. Since  $F$  is invariant under the transformation  $X \rightarrow \Sigma^{\frac{1}{2}}X$  and  $Y \rightarrow \Sigma^{\frac{1}{2}}Y$ , we shall have  $\Sigma \rightarrow I$  and hence the joint density function of  $Z = S^{-\frac{1}{2}}Y:p \times m$  and  $S = XLX'$  with the help of (26) is

$$(31) \quad \{2^{\frac{1}{2}p(m+n)} \Gamma_p(\frac{1}{2}n) |LB|^{\frac{1}{2}p} \pi^{\frac{1}{2}pm}\}^{-1} |S|^{\frac{1}{2}(m+n-p-1)} \cdot \exp(-\frac{1}{2} \text{tr}(q^{-1}I_p + ZZ')S) {}_0F_0^{(n)}(T, \frac{1}{2}q^{-1}S).$$

Integrating  $S$  and noting

$$C_\kappa(I_p + qZZ')^{-1} = C_\kappa \begin{pmatrix} (I_m + qZ'Z)^{-1} & 0 \\ 0 & I_{p-m} \end{pmatrix} = C_\kappa(R^*), \quad (\text{say}),$$

we can write the density function of  $Z:p \times m$  as

$$(32) \quad \{\Gamma_p(\frac{1}{2}n) |LB|^{\frac{1}{2}p} \pi^{\frac{1}{2}pm}\}^{-1} \cdot \Gamma_p(\frac{1}{2}m + \frac{1}{2}n) |I_m q^{-1} + Z'Z|^{-\frac{1}{2}(m+n)} {}_1F_0^{(n)}(\frac{1}{2}m + \frac{1}{2}n, T, R^*).$$

Since  $p \geq m$ , we use the Wishart's integral and obtain the density function of  $F = Z'Z$ , which is given by (30).

THEOREM 4. Let  $X:p \times n$  and  $Y:p \times m$  be independently distributed, the density function of  $X$  be given by (26) and the density function of  $Y$  be given by

$$(2\pi)^{-\frac{1}{2}pm} |\Sigma_1|^{-\frac{1}{2}m} \exp(-\frac{1}{2} \text{tr} \Sigma_1^{-1}YY').$$

Then the density function of  $F_1 = X'(YY')^{-1}X$  when  $n \leq p \leq m$  is given by

$$(33) \quad c |\Omega|^{-\frac{1}{2}n} |B|^{-\frac{1}{2}p} |F_1|^{\frac{1}{2}(p-n-1)} \cdot |I_n + (qB)^{-1}F_1|^{-\frac{1}{2}(m+n)} {}_1F_0^{(p)}(\frac{1}{2}m + \frac{1}{2}n; \Omega^*, F_1(Bq + F_1)^{-1})$$

where  $\Omega^* = I_p - q\Omega^{-1}$ ,  $\Omega = \Sigma^{\frac{1}{2}}\Sigma_1^{-1}\Sigma^{\frac{1}{2}}$

and  $c = \Gamma_p(\frac{1}{2}m + \frac{1}{2}n) \{\Gamma_p(\frac{1}{2}m) \Gamma_n(\frac{1}{2}p)\}^{-1}$ ,  $q > 0$ .

PROOF. Since  $F_1$  is invariant under  $X \rightarrow \Sigma^{\frac{1}{2}}X$  and  $Y \rightarrow \Sigma^{\frac{1}{2}}Y$ . Then  $\Sigma \rightarrow I$ ,  $\Sigma_1^{-1} \rightarrow \Omega$ . Now in the joint density function of  $X$  and  $Y$  which is given by

$$(2\pi)^{-\frac{1}{2}p(m+n)} |\Omega|^{\frac{1}{2}m} |B|^{-\frac{1}{2}p} \exp(-\frac{1}{2} \text{tr} XB^{-1}X' - \frac{1}{2} \text{tr} \Omega YY'),$$

we transform  $X$  to  $Z$  by  $Z = (YY')^{-\frac{1}{2}}X$ , and then integrating with respect to  $Y$ , we get the density function of  $Z$  as

$$(34) \quad \pi^{-\frac{1}{2}pm} |\Omega|^{\frac{1}{2}m} |B|^{-\frac{1}{2}p} \Gamma_p(\frac{1}{2}m + \frac{1}{2}n) \{\Gamma_p(\frac{1}{2}m)\}^{-1} |\Omega + ZB^{-1}Z'|^{-\frac{1}{2}(m+n)}.$$

We note that  $F_1 = Z'Z$  and if  $\Omega^* = I_p - q\Omega^{-1}$ ,  $q > 0$ , then

$$|\Omega + ZB^{-1}Z'| = |\Omega| |B|^{-1} |B + Z'Zq^{-1}| |I_p - Z(Bq + Z'Z)^{-1}Z'\Omega^*|.$$

Now, integrating  $Z$  such that  $F_1 = Z'Z$  is fixed, we get the density function of  $F_1$  as

$$(35) \quad \pi^{-\frac{1}{2}pm} |\Omega|^{-\frac{1}{2}n} |B|^{-\frac{1}{2}p} \cdot \int_{F_1=Z'Z} \{|I_n + (Bq)^{-1}Z'Z| |I_p - Z(Bq + Z'Z)^{-1}Z'\Omega^*|\}^{-\frac{1}{2}(m+n)} dZ.$$

Since the integral in (35) is symmetric and homogeneous in  $(\Omega^*)$ , we get after using the method applied in the proof of Lemma 4 the density function of  $F_1$  as mentioned in (33).

We make below few remarks or notes when  $m \geq p$  and  $n \geq p$  in the two theorems.

NOTE 1. In Theorem 4, when  $m \geq p$  and  $n \geq p$ , the density function of  $F_2 = (YY')^{-\frac{1}{2}}XX'(YY')^{-\frac{1}{2}}$  is given by

$$(36) \quad c_1 |B|^{-\frac{1}{2}p} |\Omega|^{-\frac{1}{2}n} |F_2|^{\frac{1}{2}(n-p-1)} \cdot |I_p + (q\Omega)^{-1}F_2|^{-\frac{1}{2}(m+n)} {}_1F_0^{(n)}(\frac{1}{2}m + \frac{1}{2}n; T, F_2(q\Omega + F_2)^{-1})$$

where  $T = I_n - qB^{-1}$ ,  $q > 0$ , and  $c_1 = \Gamma_p(\frac{1}{2}m + \frac{1}{2}n) \{ \Gamma_p(\frac{1}{2}m) \Gamma_p(\frac{1}{2}n) \}^{-1}$ .

NOTE 2. In Theorem 4, when  $m \geq p$  and  $n \geq p$ , the density function of  $F_3 = (XX')^{-\frac{1}{2}}YY'(XX')^{-\frac{1}{2}}$  can be obtained from (36) by the relation of transformation  $F_2 \rightarrow F_3^{-1}$ . This means that the density functions of  $F_3$  and  $F_4 = (YY')^{\frac{1}{2}}(XX')^{-1}(YY')^{\frac{1}{2}}$  are identical.

NOTE 3. When  $q \rightarrow \infty$  in (33) and (36), we get the density functions of  $F_1$  and  $F_2$  as

$$(37) \quad c |\Omega|^{-\frac{1}{2}n} |B|^{-\frac{1}{2}p} |F_1|^{\frac{1}{2}(p-n-1)} {}_1F_0^{(p)}(\frac{1}{2}m + \frac{1}{2}n; -\Omega^{-1}, F_1B^{-1})$$

and

$$(38) \quad c_1 |\Omega|^{-\frac{1}{2}n} |B|^{-\frac{1}{2}p} |F_2|^{\frac{1}{2}(n-p-1)} {}_1F_0^{(n)}(\frac{1}{2}m + \frac{1}{2}n; -B^{-1}, F_2\Omega^{-1})$$

where  $c$  and  $c_1$  are constants defined in (33) and (36) respectively. From the expressions (37) and (38), we can obtain the density functions of the ch. roots of  $F_2$  and  $F_3$ . We may also note that when  $B = I_n$  in (33), we can explicitly write down the density function of the ch. roots of  $F_2$  when  $m \geq p$  and  $n \leq p$  which will converge rapidly by choosing  $q$ , while that of  $F_3$  can be obtained from (36) when  $n \geq p$  and  $m \geq p$  and  $\Omega = I_p$ , and the density function will depend on  $B$ . Those derivable from (37) and (38) do not require such conditions.

**4. Moments of certain statistics.** (a) Let  $X: p \times n$  be distributed as normal whose density function is given by (26). Then, for any symmetric matrix  $Z$ , and for  $l_j$ 's being the ch. roots of  $LB$ , we have by Theorem 2,

$$(39) \quad E \exp(\text{tr } ZXLX') = \prod_{j=1}^n |I_p - 2l_j Z \Sigma|^{-\frac{1}{2}} \\ = \sum_{k=0}^{\infty} \sum_{\kappa} (\frac{1}{2}n)_{\kappa} C_{\kappa}(LB) C_{\kappa}(Z \Sigma) 2^k / k! C_{\kappa}(I_n).$$

Now, we note that the density function of  $S = X' LX$  given in (27) can be re-written as

$$(40) \quad 2^{-\frac{1}{2}pn} \{ \Gamma_p(\frac{1}{2}n) \}^{-1} |\Sigma|^{-\frac{1}{2}n} |Q|^{\frac{1}{2}p} \int_{O(n)} |S|^{\frac{1}{2}(n-p-1)} \exp(-\frac{1}{2} \text{tr } \Sigma^{-\frac{1}{2}} H_1 Q H_1' \Sigma^{-\frac{1}{2}} S) dH$$

where  $H' = (H_1' H_2')$  is an  $n \times n$  orthogonal matrix with  $H_1$  and  $H_2$  of dimension  $p \times n$  and  $(n-p) \times n$ , respectively, and  $Q^{-1} = L^{\frac{1}{2}} B L^{\frac{1}{2}}$ . Using this in finding  $E \exp(\text{tr } ZS)$  and interchanging the integration signs, we get

$$(41) \quad E \exp(\operatorname{tr} ZXLX') = |Q|^{\frac{1}{2}p} \int_{o(n)} |H_1QH_1'|^{-\frac{1}{2}n} \cdot |I_p - 2(\Sigma^{\frac{1}{2}}Z\Sigma^{\frac{1}{2}})(H_1QH_1')^{-1}|^{-\frac{1}{2}n} dH = \sum_{k=0}^{\infty} \sum_{\kappa} (\frac{1}{2}n)_{\kappa} 2^k (k!)^{-1} |Q|^{\frac{1}{2}p} \cdot \int_{o(n)} |H_1QH_1'|^{-\frac{1}{2}n} C_{\kappa}\{(\Sigma^{\frac{1}{2}}Z\Sigma^{\frac{1}{2}})(H_1QH_1')^{-1}\} dH.$$

Since this is homogeneous and symmetric function in  $\Sigma^{\frac{1}{2}}Z\Sigma^{\frac{1}{2}}$ , we have as is the proof of Lemma 4,

$$(42) \quad E \exp(\operatorname{tr} ZXLX') = \sum_{k=0}^{\infty} \sum_{\kappa} (\frac{1}{2}n)_{\kappa} 2^k (k!)^{-1} C_{\kappa}(Z\Sigma)\{C_{\kappa}(I_p)\}^{-1} |Q|^{\frac{1}{2}p} \int_{o(n)} |H_1QH_1'|^{-\frac{1}{2}n} C_{\kappa}(H_1QH_1')^{-1} dH.$$

Now (42) and (39) must be identical in  $Z$  and hence comparing the coefficients of  $C_{\kappa}(Z\Sigma)$ , we must have

$$(43) \quad |Q|^{\frac{1}{2}p} \int_{o(n)} |H_1QH_1'|^{-\frac{1}{2}n} C_{\kappa}(H_1QH_1')^{-1} dH = C_{\kappa}(Q^{-1})C_{\kappa}(I_p)/C_{\kappa}(I_n),$$

where  $H' = (H_1' H_2')$  is as defined in (40). Now with the help of (43), using the density function of  $S$  as given in (40), we get

$$(44) \quad EC_{\kappa}(ZS) = 2^k (\frac{1}{2}n)_{\kappa} C_{\kappa}(LB)C_{\kappa}(Z\Sigma)/C_{\kappa}(I_n).$$

The answer was given by Constantine [1] when  $LB = I_n$ . Further, when  $LB = I_n$ , then  $S = X' LX$  is distributed as Wishart whose density function is  $W(S; \frac{1}{2}n; \Sigma)$ . Hence using (8), we have

$$(45) \quad EC_{\kappa}(ZS^{-1}) = \Gamma_p(\frac{1}{2}n, -\kappa)\{\Gamma_p(\frac{1}{2}n)\}^{-1} 2^{-k} C_{\kappa}(Z\Sigma^{-1}), \text{ if } n > (p-1) + k_1.$$

(b) When the distribution of  $S = XLX'$  is given by (27), then it is easy to show that

$$(46) \quad E |S|^k = q^{p(\frac{1}{2}n+k)} 2^{pk} \Gamma_p(\frac{1}{2}n+k)\{\Gamma_p(\frac{1}{2}n)\}^{-1} |LB|^{-\frac{1}{2}p} |\Sigma|^k {}_1F_0^{(n)}(\frac{1}{2}n+k; T, I_p).$$

Now when  $LB = I_n$ , we get the well-known result

$$E |S|^k = (2^{pk}) \Gamma_p(\frac{1}{2}n+k) |\Sigma|^k / \Gamma_p(\frac{1}{2}n).$$

In (36) under the condition  $\Omega = I_p$  or  $\Sigma_1 = \Sigma_2$ , we shall choose  $q = 1$  in order to obtain the  $j$ th moment of  $|S|/|YY' + S|$  and that of  $|YY'|/|S + YY'|$  are given by

$$(47) \quad \begin{aligned} E\{|S|/|S + YY'|\}^j &= \Gamma_p(\frac{1}{2}n)/\Gamma_p(\frac{1}{2}n+j) \\ &= E\{|YY'|/|S + YY'|\}^j \Gamma_p(\frac{1}{2}m)/\Gamma_p(\frac{1}{2}m+j) \\ &= \Gamma_p(\frac{1}{2}n + \frac{1}{2}m)\{\Gamma_p(\frac{1}{2}m + \frac{1}{2}n + j)\}^{-1} {}_2F_1(\frac{1}{2}p, j; \frac{1}{2}m + \frac{1}{2}n + j; I_n - BL). \end{aligned}$$

When  $BL = I_n$ , we get the well-known results for the moments of the likelihood ratio statistics. Hence the distribution of  $|YY'|/|S + YY'|$  is obtained from that of  $|S|/|S + YY'|$  by interchanging  $n$  and  $m$ .

(c) If the density function of  $V: p \times p$  is given by

$$(48) \quad \text{constant } |V|^{t-\frac{1}{2}(m+1)} |I_p + V|^{-t-u},$$



then from (13), (14), (7), and (17), we have

$$\begin{aligned}
 & EC_{\kappa}(ZV) && \text{if } u \geq p + k_1, \\
 &= (t)_{\kappa} \Gamma_p(u, -\kappa) C_{\kappa}(Z) / \Gamma_p(u) \\
 & EC_{\kappa}(ZV^{-1}) && \text{if } t \geq p + k_1, \\
 &= (u)_{\kappa} \Gamma_p(t, -\kappa) C_{\kappa}(Z) / \Gamma_p(t) \\
 (49) \quad & EC_{\kappa}(Z(V^{-1} + I_p)^{-1}) \\
 &= (t)_{\kappa} C_{\kappa}(Z) / (t + u)_{\kappa}, \\
 & EC_{\kappa}(Z(V + I_p)^{-1}) && \text{and} \\
 &= (t)_{\kappa} C_{\kappa}(Z) / (t + u)_{\kappa}, \\
 & EC_{\kappa}(Z(V^{-1} + I_p)) \\
 &= \Gamma_p(t, -\kappa) \Gamma_p(t + u) C_{\kappa}(Z) / \Gamma_p(t + u, -\kappa) \Gamma_p(t) \text{ if } t \geq p + k_1.
 \end{aligned}$$

From these, we can easily write down the moments for  $(\text{tr } V)$ ,  $(\text{tr } V^{-1})$ ,  $\text{tr } (V^{-1} + I_p)^{-1}$ ,  $\text{tr } (V + I_p)^{-1}$  and  $(\text{tr } V^{-1} + p)$  with some conditions.

**5. Corresponding results for complex Gaussian variates.**

(5.1) In this section, we shall state the above results for complex Gaussian distributions studied by Wooding [29], Goodman [3], [4], James [13] and Khatri [15], [17], [18]. We shall denote

$$\begin{aligned}
 (50) \quad & \tilde{\Gamma}_m(a) = \pi^{\frac{1}{2}m(m-1)} \prod_{i=1}^m \Gamma(a - i + 1), \quad \tilde{\Gamma}_m(a, \kappa) = \pi^{\frac{1}{2}m(m-1)} \\
 & \prod_{i=1}^m \Gamma(a + k_i - i + 1), \\
 & \tilde{\Gamma}_m(a, -\kappa) = \pi^{\frac{1}{2}m(m-1)} \prod_{i=1}^m \Gamma(a - m - k_i + i) \text{ and} \\
 & [a]_{\kappa} = \tilde{\Gamma}_m(a, \kappa) / \tilde{\Gamma}_m(a) = \prod_{i=1}^m (a - i + 1)_{k_i}.
 \end{aligned}$$

The corresponding hypergeometric functions are defined as

$$(51) \quad {}_p\tilde{F}_q^{(m)}(a_1, \dots, a_p; b_1, \dots, b_q; A, B) = \sum_{k=0}^{\infty} \sum_{\kappa} \frac{[a_1]_{\kappa} \dots [a_p]_{\kappa} \tilde{C}_{\kappa}(A) \tilde{C}_{\kappa}(B)}{[b_1]_{\kappa} \dots [b_q]_{\kappa} \tilde{C}_{\kappa}(I_m) k!}.$$

When  $B = I$ , it is denoted as  ${}_p\tilde{F}_q^{(m)}(a_1, \dots, a_p; b_1, \dots, b_q; A)$ , and  $\tilde{C}_{\kappa}(A)$  is a zonal polynomial of a Hermitian matrix  $A$  and is the symmetric functions of ch. roots of  $A$ .

(5.2) We shall denote by  $dU$  the invariant measure on the unitary group  $U(n)$  normalized to make the total measure unity.

$$(52) \quad \int_{U(n)} \tilde{C}_{\kappa}(AUB\tilde{U}') dU = \tilde{C}_{\kappa}(A) \tilde{C}_{\kappa}(B) / \tilde{C}_{\kappa}(I_n).$$

$$(53) \quad \int_{\text{tr } A > 0} \exp(-\text{tr } A) |A|^{a-m} \tilde{C}_{\kappa}(AB) dA = \tilde{\Gamma}_m(a, \kappa) C_{\kappa}(B),$$

and

$$(54) \quad \int_{\bar{A} \rightarrow A > 0} \exp(-\text{tr } A) |A|^{\alpha-m} \tilde{C}_\kappa(BA^{-1}) dA = \tilde{\Gamma}_m(a, -\kappa) \tilde{C}_\kappa(B).$$

(5.3) Let  $F: p \times p$  be a Hermitian p.d. matrix, and its density function is

$$\tilde{\Gamma}_p(t+u) \{ \tilde{\Gamma}_p(t) \tilde{\Gamma}_p(u) \}^{-1} |F|^{t-p} |I + F|^{-t-u}.$$

Then,

$$\begin{aligned} E\tilde{C}_\kappa(ZF) &= [t]_\kappa \tilde{\Gamma}_p(u, -\kappa) \tilde{C}_\kappa(Z) / \tilde{\Gamma}_p(u) && \text{if } u \geq p + k_1, \\ E\tilde{C}_\kappa(ZF^{-1}) &= [u]_\kappa \tilde{\Gamma}_p(t, -\kappa) \tilde{C}_\kappa(Z) / \tilde{\Gamma}_p(t) && \text{if } t \geq p + k_1, \\ (55) \quad E\tilde{C}_\kappa(Z(F^{-1} + I_p)^{-1}) &= [t]_\kappa \tilde{C}_\kappa(Z) / [t + u]_\kappa, \\ E\tilde{C}_\kappa(Z(F + I_p)^{-1}) &= [u]_\kappa \tilde{C}_\kappa(Z) / [t + u]_\kappa, && \text{and} \\ E\tilde{C}_\kappa(ZF^{-1} + Z) &= \tilde{\Gamma}_p(t, -\kappa) \tilde{\Gamma}_p(t+u) \tilde{C}_\kappa(Z) / \{ \tilde{\Gamma}_p(t+u, -\kappa) \tilde{\Gamma}_p(t) \} \\ &&& \text{if } t \geq p + k_1. \end{aligned}$$

(5.4) If the complex random matrix  $X: p \times n$  is distributed as Gaussian whose density function is given by

$$(56) \quad \pi^{-pn} |\Sigma|^{-n} |B|^{-p} \exp(-\text{tr } \Sigma^{-1}XB^{-1}\bar{X}')$$

where  $\Sigma$  and  $B$  are Hermitian p.d., then the density function of  $XL\bar{X}' = S$  ( $L$  being a Hermitian p.d. matrix) is given by

$$(57) \quad (\tilde{\Gamma}_p(n) |LB|^p |\Sigma|^n)^{-1} |S|^{n-p} \exp(-q^{-1} \text{tr } \Sigma^{-1}S) {}_0\tilde{F}_0^{(n)}(T, q^{-1}\Sigma^{-1}S)$$

where  $q > 0$  and  $T = I_n - qL^{-1}B^{-1}L^{-1}$ . Moreover

$$(58) \quad E\tilde{C}_\kappa(ZS) = [n]_\kappa \tilde{C}_\kappa(LB) \tilde{C}_\kappa(Z\Sigma) / \tilde{C}_\kappa(I_n).$$

When  $LB = I$ , then  $E\tilde{C}_\kappa(ZS^{-1})$  can be written down with the help of (54).

(5.5) Let the density function of  $X: p \times n$  be given by (56), the density function of  $Y: p \times m$  be  $\pi^{-pm} |\Sigma|^{-m} \exp(-\text{tr } \Sigma^{-1}Y\bar{Y}')$  and  $X$  and  $Y$  be independent. Then for  $m \leq p \leq n$ , the density function of  $F = \bar{Y}'(XL\bar{X}')^{-1}Y$  is given by

$$(59) \quad \tilde{\Gamma}_p(m+n) \{ \tilde{\Gamma}_p(n) \tilde{\Gamma}_m(p) \}^{-1} \cdot |BL|^{-p} |F|^{p-m} |I_m q^{-1} + F|^{-m-n} {}_1\tilde{F}_0^{(n)}(m+n; T, R^*)$$

where  $q > 0$ ,  $T = I_n - q(LB)^{-1}$  and  $R^* = \begin{pmatrix} (I_m + qF)^{-1} & 0 \\ 0 & I_{p-m} \end{pmatrix}$ .

(5.6) Let the density function of  $X: p \times n$  be given by (56),  $X$  and  $Y: p \times m$  be independent and the density function of  $Y$  be given by

$$(60) \quad \pi^{-pm} |\Sigma_1|^{-m} \exp(-\text{tr } \Sigma_1^{-1}Y\bar{Y}').$$

Then for  $n \leq p \leq m$ , the density function of  $F_1 = \bar{X}'(Y\bar{Y}')^{-1}X$  is given by

$$(61) \quad \tilde{\Gamma}_p(m+n) \{ \tilde{\Gamma}_p(m) \tilde{\Gamma}_n(p) \}^{-1} |B|^{-p} |\Omega|^{-n} |F_1|^{p-n} \cdot |I_n + (qB)^{-1}F_1|^{-m-n} {}_1\tilde{F}_0^{(p)}(m+n; \Omega^*, F_1(Bq + F_1)^{-1})$$

where  $q > 0$ ,  $\Omega^* = I_p - q\Omega^{-1}$  and  $\Omega = \Sigma^{\frac{1}{2}}\Sigma_1^{-1}\Sigma^{\frac{1}{2}}$ .

(5.7) Let  $X:p \times n$  and  $Y:p \times m$  be independently distributed and their respective density functions be given by (56) and (60). Then for  $n \geq p$ ,  $m \geq p$ , the density function of  $F_2 = (Y\bar{Y}')^{-\frac{1}{2}}(X\bar{X}')(Y\bar{Y}')^{-\frac{1}{2}}$  is given by

$$(62) \quad \bar{\Gamma}_p(m+n) \{\bar{\Gamma}_p(n)\bar{\Gamma}_p(m)\}^{-1} |\Omega|^{-n} |B|^{-p} |F_2|^{n-p} \\ \cdot |I_p + (q\Omega)^{-1}F_2|^{-m-n} {}_1\bar{F}_0^{(n)}(m+n; T, F_2(q\Omega + F_2)^{-1})$$

where  $\Omega = \Sigma^{\frac{1}{2}}\Sigma_1^{-1}\Sigma^{\frac{1}{2}}$  and  $T = I_n - q(LB)^{-1}$ . The density functions of  $F_2^{-1}$  and  $F_3 = (X\bar{X}')^{-\frac{1}{2}}(Y\bar{Y}')(X\bar{X}')^{-\frac{1}{2}}$  are identical. Further, we can obtain the density function of the ch. roots of  $F_2$  in the form of (62) provided  $\Omega = I_p$  or  $\Sigma_1 = \Sigma$ ; otherwise, we can obtain the density function of the ch. roots of  $F_2$  in the form obtained by taking  $q \rightarrow \infty$  in (62). Further, if  $S = XL\bar{X}'$ , we have

$$(63) \quad E(|Y\bar{Y}'|/|S + Y\bar{Y}'|)^j \bar{\Gamma}_p(m)/\bar{\Gamma}_p(m+j) \\ = E(|S|/|S + Y\bar{Y}'|)^j \bar{\Gamma}_p(n)/\bar{\Gamma}_p(n+j) \\ = \bar{\Gamma}_p(m+n) \{\bar{\Gamma}_p(m+n+j)\}^{-1} {}_2\bar{F}_1(p, j; m+n+j; I_n - BL).$$

Hence the distribution of  $|Y\bar{Y}'|/|S + Y\bar{Y}'|$  can be obtained from that of  $|S|/|S + Y\bar{Y}'|$  by interchanging  $n$  and  $m$ .

#### REFERENCES

- [1] CONSTANTINE, A. G. (1963). Some noncentral distribution problems in multivariate analysis. *Ann. Math. Statist.* **34** 1270-1285.
- [2] CONSTANTINE, A. G. and JAMES, A. T. (1958). On the general canonical correlation distribution. *Ann. Math. Statist.* **29** 1146-1166.
- [3] GOODMAN, N. R. (1963a). Statistical analysis based on a certain multivariate complex Gaussian distribution (an introduction). *Ann. Math. Statist.* **34** 152-177.
- [4] GOODMAN, N. R. (1963b). The distribution of the determinant of a complex Wishart distributed matrix. *Ann. Math. Statist.* **34** 178-180.
- [5] GRAD, ARTHUR and SOLOMON, HERBERT (1955). Distribution of quadratic forms and some applications. *Ann. Math. Statist.* **26** 464-477.
- [6] GURLAND, J. (1948). Inversion formulae for the distribution of ratios. *Ann. Math. Statist.* **19** 228-237.
- [7] GURLAND, J. (1955). Distribution of definite and of indefinite quadratic forms. *Ann. Math. Statist.* **26** 122-127. Corrections in *Ann. Math. Statist.* **33** (1962) 813.
- [8] GURLAND, J. (1957). Quadratic forms in normally distributed variables. *Sankhyā* **17** 37-50.
- [9] JAMES, A. T. (1955). A generating function for averages over the orthogonal group. *Proc. Roy. Soc. London Ser. A.* **229** 367-375.
- [10] JAMES, A. T. (1960). The distribution of the latent roots of the covariance matrix. *Ann. Math. Statist.* **31** 151-158.
- [11] JAMES, ALAN T. (1961). The distribution of noncentral means with known covariance. *Ann. Math. Statist.* **32** 874-882.
- [12] JAMES, ALAN T. (1961). Zonal polynomials of the real positive definite symmetric matrices. *Ann. of Math.* **74** 456-469.
- [13] JAMES, ALAN T. (1964). Distribution of matrix variates and latent roots derived from normal samples. *Ann. Math. Statist.* **35** 475-501.
- [14] KHATRI, C. G. (1959). On conditions for the forms of the type  $XAX'$  to be distributed independently or to obey Wishart distribution. *Calcutta Statist. Assoc. Bull.* **8** 162-168.

- [15] KHATRI, C. G. (1961). Cumulants and higher order uncorrelation of certain functions of normal variates. *Calcutta Statist. Assoc. Bull.* **10** 93-98.
- [16] KHATRI, C. G. (1962). Conditions for Wishartness and independence of second degree polynomials in a normal vector. *Ann. Math. Statist.* **33** 1002-1007.
- [17] KHATRI, C. G. (1965a). Classical statistical analysis based on a certain multivariate complex Gaussian distribution. *Ann. Math. Statist.* **36** 98-114.
- [18] KHATRI, C. G. (1965b). A test for reality of a covariance matrix in a certain complex Gaussian distribution. *Ann. Math. Statist.* **36** 115-119.
- [19] MARSAGLIA, GEORGE (1960). Tables of the distribution of quadratic forms of ranks two and three. Mathematics Research Laboratory, Boeing Scientific Research Laboratories, Seattle, Report No. D1-82-0015-1.
- [20] OKAMOTO, MASASHI (1960). An inequality for the weighted sum of  $\chi^2$  variates. *Bull. Math. Statist.* **9** 69-70.
- [21] PACHARES, JAMES (1955). Note on the distribution of a definite quadratic form. *Ann. Math. Statist.* **26** 728-731.
- [22] ROBBINS, HERBERT (1948). The distribution of a definite quadratic form. *Ann. Math. Statist.* **19** 266-270.
- [23] ROBBINS, HERBERT AND PITMAN, E. J. G. (1949). Application of the method of mixtures to quadratic forms in normal variables. *Ann. Math. Statist.* **20** 552-560.
- [24] RUBEN, HAROLD (1960). Probability content of regions under spherical normal distributions, I. *Ann. Math. Statist.* **31** 598-618.
- [25] RUBEN, HAROLD (1962). Probability content of regions under spherical normal distributions, IV: The distribution of homogeneous and non-homogeneous quadratic functions of normal variables. *Ann. Math. Statist.* **33** 542-570.
- [26] SHAH, B. K. (1963). Distribution of definite and of indefinite quadratic forms from a non-central normal distribution. *Ann. Math. Statist.* **34** 186-190.
- [27] SHAH, B. K. AND KHATRI, C. G. (1961). Distribution of a definite quadratic form for non-central normal variates. *Ann. Math. Statist.* **32** 883-887. Correction note in *Ann. Math. Statist.* **34** (1963) 673.
- [28] SOLOMON, HERBERT (1961). On the distribution of quadratic forms in normal variates. *Proc. 4th Berkeley Symp.* **1** Univ. of California Press, Berkeley.
- [29] WOODING, R. A. (1956). The multivariate distribution of complex normal variates. *Biometrika* **43** 212-215.