

AN ELEMENTARY PROOF OF ASYMPTOTIC NORMALITY OF
 LINEAR FUNCTIONS OF ORDER STATISTICS¹

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1. Introduction and summary. Let X_1, \dots, X_n be independent random variables, identically distributed with continuous distribution function F . Let $X_{(1)}, \dots, X_{(n)}$ denote the corresponding order statistics and F_n the empiric distribution function, which we assume to be right continuous. Jung [4] found the asymptotic mean and variance of linear functions of the form

$$T_n = n^{-1} \sum_{i=1}^n J(i/n)X_{(i)} = \int_{-\infty}^{\infty} xJ(F_n(x)) dF_n(x)$$

when the function J has four bounded derivatives. More recently, it has been shown ([1], [2]) that under suitable restrictions

$$(1.1) \quad \mathcal{L}\{n^{1/2}[T_n - \int_{-\infty}^{\infty} xJ(F(x)) dF(x)]\} \rightarrow N(0, \sigma^2)$$

where

$$(1.2) \quad \sigma^2 = 2 \int \int_{s < t} J(F(s))J(F(t))F(s)[1 - F(t)] ds dt.$$

Here $N(0, \sigma^2)$ denotes the normal distribution with mean zero and variance σ^2 , and (1.1) uses a standard notation for convergence in distribution.

The purpose of this note is to give a self-contained proof of (1.1) under the assumptions that the X_i have finite mean and that J' exists and is continuous and of bounded variation except at finitely many jumps of J . This theorem can be subsumed in a corrected version of Govindarajulu [2], where appeal is made to the results of [3]. The present proof is both more elementary and shorter. Chernoff, Gastwirth and Johns [1] do not require boundedness of J , but invoke compensating assumptions on F . Their methods of proof are quite different from those of [2] and of the present paper.

2. The result. Let G be any inverse of F . We remark here that for any choice of G all integrals below will exist wp 1.

THEOREM. (1.1) holds if $\sigma^2 < \infty$ and

(A) $E|X_1| = \int_0^1 |G(u)| du < \infty$

(B) J is continuous on $[0, 1]$ except for jump discontinuities at a_1, \dots, a_M , and J' is continuous and of bounded variation on $[0, 1] - \{a_1, \dots, a_M\}$.

PROOF. We will give the proof for the case in which J' is continuous and of bounded variation on $[0, 1]$. The treatment of the terms I_{1n} and I_{3n} below remains valid in the general case if all integrals are assumed to be over $[0, 1]$ —

Received 10 August 1966; revised 11 September 1967.

¹ Research supported in part by the Office of Naval Research, Contract NONR-401(50).
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$\{a_1, \dots, a_M\}$. Although the argument given here for I_{2n} does not generalize to the discontinuous case, it is not difficult to give a separate argument treating the contribution to I_{2n} by a jump in J' . Details may be found in [5].

Write $U_n(x)$ for the empiric distribution function of the uniform random variables $R_i = F(X_i)$ after n observations. Then the left side of (1.1) is $\mathcal{L}\{I_n\}$, where

$$I_n = n^{\frac{1}{2}}[\int_0^1 G(u)J(U_n(u)) dU_n(u) - \int_0^1 G(u)J(u) du].$$

By the mean value theorem, there exists a $0 < \theta < 1$ such that

$$J(U_n(u)) - J(u) = J'(V_n(u))(U_n(u) - u)$$

where $V_n(u) = \theta U_n(u) + (1 - \theta)u$. We may therefore write

$$\begin{aligned} I_n &= I_{1n} + I_{2n} + I_{3n}, \\ I_{1n} &= \int_0^1 G(u)J'(u)W_n(u) du + \int_0^1 G(u)J(u) dW_n(u), \\ I_{2n} &= \int_0^1 G(u)[J'(V_n(u)) - J'(u)]W_n(u) dU_n(u), \\ I_{3n} &= n^{-\frac{1}{2}} \int_0^1 G(u)J'(u)W_n(u) dW_n(u), \end{aligned}$$

where $W_n(u) = n^{\frac{1}{2}}[U_n(u) - u]$. Since (A) implies that

$$\lim_{u \rightarrow 0^+} uG(u) = \lim_{u \rightarrow 1^-} (1 - u)G(u) = 0,$$

we have after integration by parts that wp 1

$$(2.1) \quad -I_{1n} = \int_0^1 J(u)W_n(u) dG(u).$$

Now $U_n(u)$ may be expressed as the arithmetic mean of n independent Bernoulli rv's, each distributed as $U_1(u)$. The right side of (2.1) therefore becomes

$$n^{-\frac{1}{2}} \sum_{i=1}^n H(R_i),$$

where $H(R_1) = \int J(u)[U_1(u) - u] dG(u)$ has mean zero and variance σ^2 . (Since the hypotheses of the theorem imply that

$$\int_0^1 \int_0^1 |J(u)| |J(t)| E[|U_1(u) - u| \cdot |U_1(t) - t|] dG(u) dG(t)$$

is finite, the interchange of expectation and integration required to obtain the variance σ^2 is valid.) Thus I_{1n} has the required asymptotic distribution by the central limit theorem.

Next note that

$$|I_{2n}| \leq \sup_{0 \leq u \leq 1} |J'(V_n(u)) - J'(u)| \cdot \sup_{0 \leq u \leq 1} |W_n(u)| \cdot \int_0^1 |G(u)| dU_n(u).$$

By the Glivenko-Cantelli theorem and uniform continuity of J' , the first factor on the right converges to zero wp 1. A well-known theorem of Kolmogorov asserts that $\sup |W_n(u)|$ has a limit in distribution, and by the strong law of large numbers

$$\int |G(u)| dU_n(u) = n^{-1} \sum_{i=1}^n |G(R_i)| \text{ converges to } \int |G(u)| du < \infty \text{ wp 1.}$$

Combining these results gives $I_{2n} \rightarrow 0$ in probability.

By writing I_{3n} as a sum of integrals over the continuity set of W_n and its complement we may obtain that wp 1

$$I_{3n} = \frac{1}{2}n^{-\frac{1}{2}} \int_0^1 G(u)J'(u) d[W_n(u)]^2 + \frac{1}{2}n^{-\frac{1}{2}} \sum_{i=1}^n G(R_i)J'(R_i).$$

The second term on the right clearly converges to zero wp 1. After integration by parts the first term becomes wp 1

$$-\frac{1}{2}n^{-\frac{1}{2}} \int_0^1 [W_n(u)]^2 d[G(u)J'(u)].$$

For brevity we will treat only the integral over $[\frac{1}{2}, 1]$, assuming that $\frac{1}{2}$ is a continuity point of G ; it will be clear how to treat the full integral. If $V(u)$ is the total variation of GJ' on $[\frac{1}{2}, u]$, we have that

$$\begin{aligned} E[n^{-\frac{1}{2}} \int_{\frac{1}{2}}^1 [W_n(u)]^2 d[G(u)J'(u)]] &\leq E[n^{-\frac{1}{2}} \int_{\frac{1}{2}}^1 [W_n(u)]^2 dV(u)] \\ &= n^{-\frac{1}{2}} \int_{\frac{1}{2}}^1 u(1-u) dV(u). \end{aligned}$$

To establish $I_{3n} \rightarrow 0$ in probability it is therefore sufficient to show

$$(2.2) \quad \int_{\frac{1}{2}}^1 u(1-u) dV(u) < \infty.$$

We need only consider the case $G(1) = \infty$, and may choose r so that $G(r) > |G(\frac{1}{2})|$. It is enough to treat the integral over $[r, 1]$. If $V_a^b(f)$ denotes the total variation of f on $[a, b]$, an obvious estimate is

$$\begin{aligned} V(u) &\leq \sup_{[r, u]} |G(t)| \cdot V_{\frac{1}{2}}^u(J') + \sup_{[r, u]} |J'(t)| \cdot V_{\frac{1}{2}}^u(G) \\ &= G(u) \cdot V_{\frac{1}{2}}^u(J') + \sup_{[r, u]} |J'(t)| \cdot [G(u) - G(\frac{1}{2})] \leq KG(u) \end{aligned}$$

for $u \geq r$ and some $K > 0$. Integrating $\int_r^1 u(1-u) dV(u)$ by parts and using this bound establishes (2.1) and concludes the proof.

3. Acknowledgment. I wish to thank Professors Gastwirth and Govindarajulu for providing copies of their papers [1] and [2].

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