

ON THE EXISTENCE OF ABSOLUTE MOMENTS FOR THE EXTINCTION TIME OF A GALTON-WATSON PROCESS

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If $\{Z_n\}$ is a Galton-Watson process with mean one, and τ is the extinction time, it is shown that $EZ^{1+\alpha} < \infty$ implies $E\tau^\beta = \infty$ for $\beta > 1/\alpha$, $0 < \alpha < 1$. Conditions which imply $EZ^{1+\alpha} = \infty$ and $E\tau^\beta < \infty$ for $\beta < 1/\alpha$, $0 < \alpha < 1$ are given. Necessary and sufficient conditions for $EX^{m+\alpha} < \infty$ or $EX^m \log X < \infty$ are given in terms of the Laplace transform of a general nonnegative random variable X , $0 < \alpha < 1$, $m = 0, 1, \dots$.

1. Introduction. Let Z_n be the number of individuals in the n th generation of a reproducing system so that $\{Z_n\}$ is a Galton-Watson process, and let $\tau = \inf \{n \mid Z_n = 0\}$ be the extinction time for the process. It is well known that $\mu_1 = EZ_1 > 1$ iff $P(\tau < \infty) < 1$, and that if $\mu_1 < 1$ then $P(\tau > n) \leq \mu_1^n$. Thus, in the non-critical case, $\mu_1 \neq 1$, either $E\tau^\alpha < \infty$ for no $\alpha > 0$ or all $\alpha > 0$. There is a drastic difference in the critical case.

If $\mu_1 = 1$ (see Kesten, Ney, Spitzer [4]) then $P(\tau > n) \sim \sigma^2/2n$, when $\sigma^2 = \text{Var } Z_1 < \infty$, and in this case $E\tau^\alpha < \infty$ iff $0 < \alpha < 1$. It is pointed out in the above reference that $nP(\tau > n) \rightarrow 0$ if $\sigma^2 = \infty$. This fact is what led us to ask about higher moments of τ when Z_1 has worse behavior, and we have the following negative answer.

THEOREM 1. *Let $\mu_1 = 1$. If $EZ_1^{1+\alpha} < \infty$ then $E\tau^\beta = \infty$ for all $\beta > 1/\alpha$, $0 < \alpha < 1$. We have been unable to determine the validity of the next*

ASSERTION. *Let $\mu_1 = 1$. If $EZ_1^{1+\alpha} = \infty$ then $E\tau^\beta < \infty$ for all $\beta < 1/\alpha$, $0 < \alpha < 1$.*

However, if stronger restrictions than $EZ_1^{1+\alpha} = \infty$ are placed on the distribution of Z_1 results of the above type do hold. To make these precise we need additional notation.

Let $f(s) = \sum_0^\infty s^n P(Z_1 = n)$ be the probability generating function (pgf) of Z_1 , so that the pgf of Z_n is f_n , the n th functional iterate of f , i.e. $f_n = f \circ f_{n-1}$, $f_0 = \text{identity}$. Then $P(\tau > n) = 1 - f_n(0)$ and $\mu_1 = f'(1-)$.

From now on assume $\mu_1 = 1$.

We introduce the following conditions on f , where K is some finite positive constant.

$$A_\beta: f(s) - s \leq K(1-s)^{1+\beta} \quad \text{for } s_0 \leq s \leq 1, \quad \text{some } s_0 < 1.$$

$$B_\beta: f(s) - s \geq K(1-s)^{1+\beta} \quad \text{or } s_0 \leq s \leq 1, \quad \text{some } s_0 < 1.$$

It is easy to show that $EZ_1^{1+\beta} < \infty$ implies A_β , (see e.g. Loève [5] page 199).

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Further, A_β implies $EZ_1^{1+\alpha} < \infty$ for each $\alpha < \beta$, $0 < \beta < 1$, while B_β implies $EZ_1^{1+\beta} = \infty$, $0 < \beta < 1$ (see Remark 2 following Theorem 5 below). We know of no moment condition which implies B_β , but we do have the following sufficient condition (Theorem 7): If $P(Z_1 > x)$ is asymptotic to $x^{-1-\alpha}L(x)$, where L is slowly varying, then B_β holds for each $\beta > \alpha$.

THEOREM 2. *Let $0 < \beta < 1$. Condition B_β implies $EZ_1^{1+\beta} = \infty$ and $E\tau^\gamma < \infty$ for all $\gamma < 1/\beta$.*

Since $E\tau^\gamma < \infty$ iff $\sum n^{\gamma-1}(1-f_n(0)) < \infty$, $\gamma > 0$, Theorems 1 and 2 are corollaries of

THEOREM 3. *Let $\beta > 0$ and fix v_0 in $[0, 1)$. Then A_β implies $n^\gamma(1-f_n(v_0)) \rightarrow \infty$ for all $\gamma > 1/\beta$, and B_β implies $n^\gamma(1-f_n(v_0)) \rightarrow 0$ for all $\gamma < 1/\beta$.*

PROOF. Define $h(u) = 1-f(1-u)$; then $h_n(u) = 1-f_n(1-u)$ ($h_n = h_{n-1} \circ h$), and h is continuous, strictly increasing, concave downward and $0 < h(u) < u \leq 1$. Let $u_0 = 1-v_0$ and define $u_{n+1} = h(u_n) = 1-f_n(v_0)$. Further, define $a_n = (u_n - u_{n+1})u_n^{-1-\beta}$, $r_n = u_{n+1}/u_n \leq 1$ and $u_n^{-\beta} = \rho_1 + \dots + \rho_n$. Then

$$\rho_{n+1} = a_n(r_n)^{-\beta}(1-r_n^\beta)/(1-r_n).$$

Since $0 < a \leq (1-x^\beta)/(1-x) \leq b < \infty$ for $0 \leq x \leq 1$, B_β implies $\rho_{n+1} \geq K_1$ for some $K_1 > 0$ and all n . Under hypothesis A_β , $1 \geq r_n = 1 - a_n u_n^\beta \geq 1 - K u_n^\beta \uparrow 1$ so that $\rho_{n+1} \leq K_2$ for some $K_2 < \infty$ and all n .

The pattern for the above proof is found in the paper by Szekeres ([7] see Theorem 1c) which contains further information on the behavior of the iterates h_n .

Another interesting proof of at least part of Theorem 3 can be given by applying the A_β and B_β conditions to

$$\frac{1}{1-f_n(s_0)} - \frac{1}{1-s_0} = \sum_{k=0}^{n-1} \Delta(f_k(s_0))$$

where

$$\Delta(s) = \frac{1}{1-f(s)} - \frac{1}{1-s} = \frac{f(s)-s}{(1-s)^2} \cdot \frac{1-s}{1-f(s)}.$$

Assume that $0 < \beta < 1$. If $P_\gamma(Q_\gamma)$ is the proposition “ $n^\gamma(1-f_n(s_0)) \rightarrow \infty(0)$,” then under hypothesis A_β , P_γ implies P_α for all $\alpha > 1+(1-\beta)\gamma$, while under hypothesis B_β , Q_γ implies Q_α for all $\alpha < 1+(1-\beta)\gamma$. Iteration produces: A_β and P_γ imply P_α for all $\alpha > 1/\beta$; B_β and Q_γ imply Q_α for all $\alpha < 1/\beta$. Now Q_0 is well known, but we know of no way to get a P_γ except by using the arguments of the given proof.

While trying to get an insight into condition B_β and to prove the “assertion” above, we found necessary and sufficient conditions, in terms of Laplace transforms, for finiteness of non-integral absolute moments of general random variables (Theorem 5).

2. Moments and Laplace transforms. Let X be a nonnegative rv with df F and Laplace transform $\varphi(\lambda) = \int_0^\infty e^{-\lambda x} F(dx)$. Set $\mu_\alpha = EX^\alpha$ and introduce inductively the notation

$$F_0 := F, F_{n+1}(x) := \int_0^x \left[\frac{\mu_n}{n!} - F_n(y) \right] dy$$

for each n such that $\mu_n < \infty$.

The following theorem (perhaps well known) gives a representation for μ_α and φ in terms of F_n .

THEOREM 4. Assume $\mu_m < \infty$, m a nonnegative integer. Then for $n = 0, 1, \dots, m$

(a) $F_n(x) \uparrow \mu_n/n!$ as $x \uparrow \infty$;

(b)
$$\mu_{n+\alpha} = \frac{\Gamma(n+\alpha+1)}{\Gamma(n+\alpha-k)} \int_0^\infty x^{n+\alpha-k-1} \left[\frac{\mu_k}{k!} - F_k(x) \right] dx,$$

$k = 0, 1, \dots, n, \alpha \geq 0, n+\alpha-k > 0$, in the sense that if one side is finite so is the other;

(c)
$$(-1/\lambda)^{n+1} \left[\varphi(\lambda) - \sum_0^n \mu_k (-\lambda)^k / k! \right] = \int_0^\infty e^{-\lambda x} \left[\frac{\mu_n}{n!} - F_n(x) \right] dx.$$

The proof uses induction and integration by parts and is omitted. See Feller [2] for the first step.

For some reason Harkness and Shantaram [3] have been led to study the df's $n!F_n/\mu_n = G_n$ and they give a version of (c) for characteristic functions. Professor James Hannan has pointed out (oral communication) that an induction shows that

$$1 - G_n(x) = \int_0^\infty \left[\left(1 - \frac{x}{y} \right)^+ \right]^n H_n(dy), \quad x \geq 0$$

where $H_n(y) = \int_0^y u^n F(du) / \mu_n$, so that G_n is the mixture of the minimums of n independent rv's with uniform distribution on $(0, y)$, where H_n is the mixing distribution.

In what follows we assume $\mu_m < \infty$ and define

$$\eta_{m,\alpha}(\lambda) = \lambda^{-m-1-\alpha} \left| \varphi(\lambda) - \sum_0^m \mu_k (-\lambda)^k / k! \right|.$$

THEOREM 5. (a) For $0 < \alpha < 1, \mu_{m+\alpha} < \infty$ iff for some (and thus all) $c > 0 \int_0^c \eta_{m,\alpha} < \infty$, and then (whether finite or not)

$$\mu_{m+\alpha} = \frac{\Gamma(m+\alpha+1)}{\Gamma(1-\alpha)\Gamma(\alpha)} \int_0^\infty \eta_{m,\alpha}.$$

(b) $v_m := EX^m \log X < \infty$ iff for some (and all) $c > 0 \int_0^c \eta_{m,0} < \infty$ and then (whether finite or not)

$$v_m = m! \int_0^\infty \eta_{m,0}.$$

PROOF. Since $\eta_{m,\alpha}(\lambda) \sim \lambda^{-1-\alpha}$ as $\lambda \rightarrow \infty$, $\alpha > -1$, part (a) is a simple consequence of Theorem 4 and Fubini’s theorem. To prove part (b) note that

$$\int_0^c \eta_{m,0} = \int_0^\infty \frac{1 - e^{-xc}}{x} \left[\frac{\mu_m}{m!} - F_m(x) \right] dx,$$

so if it is finite for some c it is finite for $c = \infty$. To complete the proof, use the following formula which follows by induction and integration by parts:

$$v_m := \int_0^\infty x^m \log x F(dx) = \int_0^\infty L^{(k+1)}(x) \left[\frac{\mu_k}{k!} - F_k(x) \right] dx,$$

where $L(x) = x^m \log x$, so that $L^{(k)}(x) = (m!/(m-k)!) x^{m-k} (\log x + a_k)$ $k = 0, \dots, m$, and $L^{(m+1)}(x) = m!/x$.

REMARKS. (1) The above extends the results of Athreya [1] who shows that $EX \log X < \infty$ iff $\int_0^c \eta_{1,0} < \infty$ some $c > 0$.

(2) When X is integer valued (identify X and Z_1 and all corresponding notation) $f(e^{-\lambda}) = \varphi(\lambda)$ and the change of variable $s = e^{-\lambda}$ shows that

$$\eta_{1,\alpha-1}(\lambda) = \frac{\varphi(\lambda) - (1-\lambda)}{\lambda^{1+\alpha}} \sim \frac{f(s) - s}{(1-s)^{1+\alpha}} \text{ as } \lambda \rightarrow 0, s \rightarrow 1.$$

Hence the B_α condition is equivalent to $\eta_{1,\alpha-1}(\lambda) \geq K > 0$, $0 \leq \lambda \leq \lambda_0$. It is now clear that A_β implies $EZ_1^{1+\alpha} < \infty$, $0 \leq \alpha < \beta < 1$ and that B_β implies $EZ_1^{1+\beta} = \infty$.

We now state the Abelian–Tauberian theorems of Laplace transforms in the form we need. Recall that a function $L: (0, \infty) \rightarrow (0, \infty)$ is *slowly varying* iff $\lim_{t \rightarrow \infty} L(xt)/L(t) = 1$ for all $x > 0$.

THEOREM 6. Let X be a nonnegative rv with df F and Laplace transformation φ . Assume that $0 < \alpha < 1$ and that m is the nonnegative integer such that $\mu_m < \infty = \mu_{m+1}$. The following are equivalent (where $k = 0, 1, \dots, m$, and $x = 1/\lambda \rightarrow \infty$):

$$(a_k) \frac{\mu_k}{k!} - F_k(x) \sim \frac{\Gamma(m+\alpha-k)}{\Gamma(m+\alpha)} x^{-m-\alpha+k} L(x)$$

$$(a_{m+1}) F_{m+1}(x) \sim \frac{\Gamma(\alpha)}{(1-\alpha)\Gamma(m+\alpha)} x^{1-\alpha} L(x)$$

$$(b_k) \eta_{k,\alpha-1}(\lambda) \sim \frac{\Gamma(1-\alpha)\Gamma(\alpha)}{\Gamma(m+\alpha)} \lambda^{m-k} L(1/\lambda).$$

PROOF. The equivalence of the (a_k) ’s is given in Feller [2], Theorem VIII.9.1 and the lemma to Theorem XIII.5.4 (which holds for $0 < |\rho| < \infty$). The equivalence of (a_m) and (b_m) is the familiar Abelian–Tauberian theorem, see Feller [2], Theorem XIII.5.4. The equivalence of the (b_k) ’s is trivial.

This gives our final

THEOREM 7. If $P(Z_1 > x) \sim x^{-1-\alpha}L(x)$, where L is slowly varying, then f satisfies condition B_β for all $\beta > \alpha$, whence $E\tau^\gamma < \infty$ for all $\gamma < 1/\alpha$ and $EZ_1^\beta = \infty$ for all $\beta > \alpha$.

PROOF. For $s = e^{-\lambda}$, as $\lambda \rightarrow 0$

$$\frac{f(s)-s}{(1-s)^{1+\alpha}} \sim \eta_{1,\alpha-1}(\lambda) \sim L\left(\frac{1}{\lambda}\right).$$

By the representation of $L(x) = a(x) \exp \int_1^x \varepsilon(y)/y \, dy$ where $a(x) \rightarrow c \in (0, \infty)$ and $\varepsilon(y) \rightarrow 0$ (see Feller [2], corollary to Theorem VIII.9.1) we see that $x^\varepsilon L(x) \rightarrow \infty$ if $\varepsilon > 0$, whence $B_{\alpha+\varepsilon}$ holds.

REMARKS. It was hoped that $\int_0^\infty \eta_{1,\alpha-1} = \infty$ (iff $EZ_1^{1+\alpha} = \infty$) would shed light on the "assertion" at the beginning of this note, but it has failed to do so as yet.

Seneta [6] did show that

$$E\tau < \infty \quad \text{iff} \quad \int_0^1 \frac{1-u}{f(u)-u} \, du < \infty,$$

but his approach does not seem to work for non-integral moments.

REFERENCES

- [1] ATHREYA, K. (1969). On the supercritical one dimensional age dependent branching processes. *Ann. Math. Statist.* **40** 743–763.
- [2] FELLER, W. (1966). *An Introduction to Probability Theory and Its Applications* **2**. Wiley, New York.
- [3] HARKNESS, W. and SHANTARAM, R. (1969). Convergence of a sequence of transformations of distribution functions. *Pacific J. Math.* **31** 403–415.
- [4] KESTEN, H., NEY P. and SPITZER, F. (1966). The Galton–Watson process with mean one and finite variance. *Theor. Probability Appl.* **11** 513–540. (English translation.)
- [5] LOÈVE, M. (1963). *Probability Theory*, 3rd ed. Van Nostrand, Princeton.
- [6] SENETA, E. (1967). The Galton–Watson process with mean one. *J. Appl. Probability* **4** 489–495.
- [7] SZEKERES, G. (1960). Regular iteration of real and complex functions. *Acta. Math.* **100** 203–258.