

A COUNTEREXAMPLE ON TRANSLATION INVARIANT ESTIMATORS¹

BY ERIK N. TORGENSEN

University of Oslo

It seems to be generally known that the proof of the continuity part of Theorem 1 in Hodges' and Lehmann's paper (1963) is incorrect. The fact that the theorem is incorrect is—perhaps—not so well known. We show this by constructing independent real random variables X_1, \dots, X_n , each having the same non-atomic symmetric distribution, and an odd translation invariant estimator $h(X_1, \dots, X_n)$ such that $P(h(X_1, \dots, X_n) = 0) > 0$. h may be chosen symmetric provided $n \geq 3$.

The construction. Define first three odd measurable functions ϕ , ψ and g from $] -\infty, +\infty[$ to $] -\infty, +\infty[$ such that

$$\begin{aligned} \phi(x) &= 0.0x^{(1)}000x^{(2)}000x^{(3)} \dots, \\ \psi(x) &= 0.000x^{(1)}000x^{(2)}000x^{(3)} \dots \quad \text{and} \\ g(x) &= 0.000x^{(4)}000x^{(8)}000x^{(12)} \dots \end{aligned}$$

when $x \in [0, 1[$ and $0, x^{(1)}x^{(2)}x^{(3)} \dots$ is the dyadic expansion of x which has infinitely many zeroes. Note that $x_1 + g(x_2 - x_1) = 0$ when $x_1 \in \psi] -1, 0]$ and $x_2 \in \phi]0, 1[$. If $n = 2$, then we may proceed as follows. Define h as the map $(x_1, x_2) \rightarrow x_1 + g(x_2 - x_1)$. Choose independent random variables U_1 and U_2 , each being uniform on $] -1, +1[$. Then put $X_i = \phi(U_i)$ or $\psi(U_i)$ with equal probability $\frac{1}{2}$, — the selections being independent and independent of the U 's. It is easily seen that the distribution is non-atomic and symmetric and $P(h(X_1, X_2) = 0) \geq P(X_1 \in \psi] -1, 0])P(\frac{1}{4} \leq X_2 \in \phi]0, 1[) > 0$. In this inequality $h(X_1, X_2)$ could be replaced by $h(X_1, X_2) \wedge h(X_2, X_1)$ —the latter being symmetric but not odd. (The only odd translation invariant and symmetric estimator is $(X_1 + X_2)/2$.)

If $n \geq 3$, then counterexamples with odd translation invariant and symmetric estimators of the form $h(X_1, \dots, X_n) = \text{median}(X_1, \dots, X_n) + g((\wedge X_i + \vee X_i)/2 - \text{median}(X_1, \dots, X_n))$ may be constructed. It suffices to do this for $n = 3$ —the general case being very similar.

Choose independent random variables U_1, U_2 and U_3 , each being uniform on $] -1, +1[$. Let $U_i^{(j)}$ denote the j th binary digit in the dyadic expansion of $|U_i|$. Then put

Received December 3, 1970.

¹ Work done while the author had grants from Norges Almen vitenskapelige Forskningsråd and U.S. Army Research Grant DA-31-124-AROD-548, DA-ARO-D-31-124-G-816.

$X_i = \alpha(U_i) = [-0. U_i^{(1)} 0000000 U_i^{(2)} 0000000 U_i^{(3)} \dots + 2] \operatorname{sgn} U_i$ with probability $\frac{1}{3}$,

$X_i = \beta(U_i) = [0.000 U_i^{(1)} 000 U_i^{(2)} 000 U_i^{(3)} \dots] \operatorname{sgn} U_i$ with probability $\frac{1}{3}$,

$X_i = \gamma(U_i) = [0.0000 U_i^{(1)} 0000000 U_i^{(2)} 0000000 U_i^{(3)} \dots + 2] \operatorname{sgn} U_i$ with probability $\frac{1}{3}$,

the selections being independent and independent of the U 's.

The event “ $X_1 = \alpha(U_1)$ with $U_1 < 0$,
 $X_2 = \beta(U_2)$ with $U_2 < 0$ and
 $X_3 = \gamma(U_3)$ with $U_3 > 0$ ”

has positive probability and implies the event

$$\begin{aligned} & \text{“ } X_1 \leq X_2 \leq X_3, \\ & X_2 \in \psi] - 1, 0] \text{ and } (X_1 + X_3)/2 \in \phi [0, 1[\text{”} \end{aligned}$$

which in turn implies that $h(X_1, X_2, X_3) = 0$.

It would be interesting to know if there are singular probability measures P on the line such that $P \times P(\{(x_1, x_2): h(x_1, x_2) = 0\}) = 0$ for all translation invariant estimators h .

REFERENCE

- [1] HODGES, J. L. and LEHMANN, E. L. (1963). Estimates of location based on rank tests. *Ann. Math. Statist.* **34** 598–611.