ON BERRY-ESSEEN APPROXIMATION AND A FUNCTIONAL LIL FOR A CLASS OF DEPENDENT RANDOM FIELDS

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In this paper we derive a Berry-Esseen type approximation for a class of dependent random fields and use it to obtain a functional law of the iterated logarithm.

1. Introduction. In recent years there has been considerable interest in multiparameter stochastic collections or the so-called random fields. In this note we deal with stationary, dependent discrete-parameter random field. In [3] a concept of ϕ -mixing was introduced for such random fields and a functional central limit theorem was proved for them. Here we obtain a Berry-Esseen type approximation for such random fields and use it to prove a functional law of the iterated logarithm.

The set-up and the basic notation is as in [3]. Z^q is the set of all q-tuples of integers $(q \ge 1)$. We denote the points in Z^q by *i*, *n* etc. or sometimes explicitly by (i_1, i_2, \dots, i_q) , (n_1, n_2, \dots, n_q) etc. Let $\{\xi_n : n \in Z^q\}$ be a stationary, ϕ -mixing random field as defined in [3]. We denote the partial sums of this random field by S_n or S_{n_1,n_2,\dots,n_q} i.e.,

$$S_{n_1,n_2,\cdots,n_q} = \sum_{\imath_1=1}^{n_1} \sum_{\imath_2=1}^{n_2} \cdots \sum_{\imath_q=1}^{n_q} \hat{\xi}_{i_1,i_2,\cdots,i_q}$$

where $n_i \ge 1$. If some n_i are zero and others ≥ 1 then it is convenient to set $S_{n_1,n_2,\dots,n_q} = 0$.

Let $T^q = [0, 1]^q$ be the q-fold Cartesian product of the unit interval, and let D_q be the Skorohod function space on T^q . We use the uniform metric d on D_q i.e., if $x, y \in D_q$ then d(x, y) = $\sup_t |x(t) - y(t)|$.

A block B in T^q is a product of half-closed intervals i.e., a set of the form $\prod_{i=1}^{q} (s_i, t_i]$. If x is a function on T^q then x(B) denotes increment of x around B.

We will assume throughout that:

(1)
$$E(\xi_n)=0 ext{ and } E|\xi_n|^{2+\eta}<\infty ext{ for some } \eta>0$$
 .

We will also assume the following condition in [3] on the rate of ϕ -mixing:

$$(\ 2\) \qquad \qquad \sum_{q=1}^{\infty} r^{q-1} \phi^{1/2}(r) < \ \infty \ .$$

It is proved in [3] that under these conditions: $\lim_{n\to\infty} n^{-q} \operatorname{Var}(S_{n,n,\dots,n}) = \sigma^2(<\infty)$ where

$$\sigma^2 = \sum_{j \in \mathbb{Z}^q} E(\xi_0 \xi_j)$$
. (Here $\xi_0 = \xi_{0,0,\dots,0}$).

To avoid trivial complications we will assume $\sigma^2 > 0$.

We denote by K_{σ} the Strassen's set of continuous functions on T^{q} :

$$K_{\sigma} = \left\{x \colon x(t_1, t_2, \cdots, t_q) = \int_0^{t_1} \int_0^{t_2} \cdots \int_0^{t_q} y(u_1, u_2, \cdots, u_q) du_1 du_2 \cdots du_q
ight.$$

where $\int_0^1 \cdots \int_0^1 y^2(u_1, u_2, \cdots, u_q) du_1 \cdots du_q \leq \sigma^2
ight\}$.

Theorem 1 below is a Berry-Esseen type theorem dealing with the speed of convergence of (normalized) $S_{n,n,\dots,n}$ to normality. Theorem 2 is a functional LIL for these partial sums.

Denote by $(H_n: n \ge 1)$ the sequence of random functions in D_q defined by

$$H_n(t) = (2n^q \log \log n)^{-1/2} S_{[nt_1][nt_2], \dots, [nt_n]}$$

where $t = (t_1, t_2, \dots, t_q) \in T^q$ and $[\cdot]$ is the usual greatest-integer function.

2. Theorems and proofs.

THEOREM 1. Let Φ be the standard normal distribution function. Then under (1) and (2) there exists C > 0, $\alpha > 0$ such that

$$\sup_x |P\{\sigma^{-1}n^{-q/2}S_{n,n},...,n} < t\} - \varPhi(t)| < Cn^{-lpha}$$
 , for all n .

Proof. For simplicity suppose q = 2.

For given integers n, a = a(n) and b = b(n), let μ be the largest integer such that $\mu(a + b) \leq n$. Then subdivide the square $(0, n] \times$ (0, n] into blocks by taking the product of 2 copies of the partition $0 < a < a + b < 2a + b < \cdots < \mu(a + b) < n$. If $1 \leq m \leq \mu$, denote by I_{ma} the interval ((m - 1)(a + b), (m - 1)(a + b) + a], by I_{mb} the interval ((m - 1)(a + b) + a, m(a + b)] and $I_{(\mu+1)a} = (\mu(a + b), n]$. Set

$$\begin{aligned} \alpha_{m}(n) &= \sum_{j \in I_{m_{1}a} \times I_{m_{2}a}} \xi_{j} \quad (1, 1) \leq m \leq ((\mu + 1), \mu + 1) \\ \beta'_{m}(n) &= \sum_{j \in I_{m_{1}a} \times I_{m_{2}b}} \xi_{j} \quad (1, 1) \leq m \leq (\mu + \mu) \\ \beta''_{m}(n) &= \sum_{j \in I_{m_{1}a} \times I_{m_{2}b}} \xi_{j} \quad (1, 1) \leq m \leq (\mu + 1, \mu) \end{aligned}$$

$$\beta_{m}^{\prime\prime\prime}(n) = \sum_{j \in I_{m1b} \times I_{m2a}} \xi_{j} \quad (1, 1) \leq m \leq (\mu, \mu + 1)$$
$$u_{n} = \sum_{m_{i}=1}^{n+1} \alpha_{m}(n) \quad v_{n}^{\prime} = \sum_{m_{i}=1}^{n} \beta_{m}^{\prime\prime}(n) \quad v_{n}^{\prime\prime} = \sum_{m} \beta_{m}^{\prime\prime\prime}(n) \quad v_{n}^{\prime\prime\prime} = \sum_{m} \beta_{m}^{\prime\prime\prime}(n)$$

Then $S_{n,n} = u_n + v'_n + v''_n + v'''_n$.

Because of condition (2), by Proposition 1.1.20 of [6], we have $E(\gamma_m^2(n)) = \#(\gamma_m)(\alpha^2 + \rho_{\sharp(\gamma_m)})$ where γ_m^2 stands for one of the α_m or $\beta_m'', \beta_m''', \beta_m'''$ and $\#(\gamma_m)$ is the "size" of the block γ_m and $\rho_{\sharp(\gamma_m)} \to 0$ if $\#(\gamma_m) \to \infty$.

Furthermore, as in Theorem 1.1.22 of [6] we get:

(i) $Ev'_n^2 \leq [1 + 4\mu^2 \phi^{1/2}(a)][\mu^2 b^2(\alpha^2 + \rho_b)]$

(ii) $Ev_n^{\prime\prime 2} \leq [1 + 4(\mu + 1)^2 \phi^{1/2}(b)] [\mu^2 a b(\sigma^2 + \rho_{ab}) + \mu b(n - \mu(a + b))] (\sigma^2 + \rho_{n-\mu(a+b)})].$

(iii)
$$Ev_n^{\prime\prime\prime 2} \leq [1 + 4(\mu + 1)^2 \phi^{1/2}(b)] [\mu^2 a b(\alpha^2 + \rho_{ab}) + \mu b(n - \mu(a + b)) (\sigma^2 + \rho_{n-\mu(a+b)})].$$

For $(1, 1) \leq m \leq (\mu, \mu)$ define $\alpha'_m(n)$ to be independent random variables having the same law as $\alpha_{(1,1)}(n)$; then as in Theorem 1.1.22 of [6] we get:

$$(\mathrm{iv}) \qquad \left| P\left(\frac{u_n}{\sigma \sqrt{n^2}} < t \right) - P\left(\frac{\sum\limits_{m_i=1}^{\mu} \alpha'_m(n)}{\sigma \sqrt{n^2}} < t \right) \right| \leq (\mu + 1)^2 \phi(b) \; .$$

For this computation, it is easy to show that the "end blocks" $\alpha_m(n)$ (with m_1 or m_2 equal to $\mu + 1$) become negligible for large n.

$$\begin{array}{l} (\mathrm{v}) \qquad \quad \left| P\left(\frac{\sum\limits_{m=1}^{\mu} \alpha'_{m}(n)}{E^{1/2} (\alpha'_{m}(n))^{2}} < \frac{t\sigma \sqrt{n^{2}}}{E^{1/2} (\sum \alpha'_{m})^{2}} \right) - \varPhi \left(\frac{t\sigma \sqrt{n^{2}}}{E^{1/2} (\sum \alpha'_{m}(n))^{2}} \right) \right| \\ & \leq \frac{c_{\delta} (\mu + 1)^{2} E |\alpha'_{(1,1)}|^{2+\delta}}{[(\mu + 1)^{2} E (\alpha^{2}_{(1,1)})]^{1+\delta/2}} \leq Ac_{\delta} (1 + \mu)^{-\delta} \end{array}$$

because by [4, Lemma 7]. $E(|\alpha_{_{(1,1)}}|^{_{2+\delta}}) \leq A(E(\alpha_{_{(1,1)}}'))^{_{1+\delta/2}}$

$$\begin{array}{l} (\mathrm{vi}) \quad \left| \left. \varPhi \left(\frac{\sqrt{n^2}}{\sqrt{(\mu+1)^2 a(\sigma^2+\rho_a)}} \cdot t \sigma \right) - \varPhi(t) \right| \\ \\ \leq & \frac{1}{2\pi e} \max \Big(1, \ \sqrt{\frac{(\mu+1)^2 a^2}{n^2}} \Big(1 + \frac{\rho_a}{\sigma^2} \Big) \Big| \ \sqrt{\frac{n^2 \sigma^2}{(\mu+1)^2 a^2 (\sigma^2+\rho_a)}} - 1 \Big) \\ \\ \\ = & \psi(n) \ . \end{array} \right.$$

From (i)-(vi) and using a similar argument as in Theorem 1.1.22 of [6], for $\tau > 0$

$$igg| \left| P \Big(rac{S_{n,n}}{\sigma \sqrt{n^2}} < t \Big) - arPsi(t)
ight| \leq (\mu + 1)^2 \phi(b) + A c_\delta(\mu + 1)^{-\delta} + \psi(n) \ + rac{ au}{\sigma \sqrt{n^2}} + rac{E {v'_n}^2}{ au^2/9} + rac{E {v''_n}^2}{ au^2/9} + rac{E {v''_n}^2}{ au^2/9} \,.$$

If we choose $a = [n^{\cdot 6}] b = [n^{\cdot 4}] \tau = [n^{1-\epsilon}] 0 < \epsilon < 1$ then $\mu = O(n^{\cdot 4})$ and $n - \mu(a + b) = O(n^{\cdot 6})$. Since condition (2) implies that $r^{q} \phi^{1/2}(r) \to 0$; then

$$\begin{split} (\mu + 1)^2 \phi(b) &= ((\mu + 1)^4 \phi(b))(\mu + 1)^{-2} = O(n^{-\cdot 8}) \ . \\ Ac_{\delta}(\mu + 1)^{-\delta} &= O(n^{-\delta(\cdot 4)}) = O(n^{-\cdot 4\delta}) \ . \\ \psi(n) &= O\Big(\Big|\sqrt{\frac{n^2 \sigma^2}{(\mu + 1)^2 a^2 (\sigma^2 + \rho_a)}} - 1\Big|\Big) = O\Big(\Big|\sqrt{\frac{n^2}{(\mu + 1)^2 a^2}} - 1\Big|\Big) \\ &= O\Big(\frac{b}{a}\Big) = O(n^{-\cdot 2}) \\ \frac{\tau}{\sigma \sqrt{n^2}} &= O(n^{-\epsilon}) \ . \\ \frac{Ev'_n^2}{\tau^2/9} &\leq (\text{constant}) \frac{\mu^2 b^2 (\sigma^2 + \rho_b)}{\tau^2/9} \text{ since } \mu^2 \phi^{1/2}(a) \longrightarrow 0 \ . \\ &= O\Big(\frac{n^{2-\cdot 4}}{n^{2-2\epsilon}}\Big) = O(n^{2\epsilon - \cdot 4}) \ . \end{split}$$

$$rac{E v_n'^2}{ au^2/9} \leq rac{(ext{constant})}{ au^2/9} [\mu^2 a b (\sigma^2 +
ho_{ab}) + \mu b (n - \mu(a + b)) (\sigma^2 +
ho_{n-\mu(a+b)})] = O(n^{2arepsilon - 2}) \;.$$

Similarly

$$rac{E {v''_n}^2}{ au^2/9} = O(n^{2arepsilon-\cdot 2}) \; .$$

Then

$$ig| P\left(rac{S_{n,n}}{\sigma V n^2} < t
ight) - \varPhi(t) ig| \leq C n^{-lpha}$$

if we set: $arepsilon = lpha = rac{1}{15}$ whenever $\delta \geq rac{1}{6}$.
 $arepsilon = lpha = .48$ whenever $0 < \delta < rac{1}{6}$.

An analogous proof is valid for the q > 2, in that case take

$$arepsilon=lpha=rac{1}{15} ext{ if } \delta \geq rac{1}{3q} \ arepsilon=lpha=(.2q\delta) ext{ if not }.$$

REMARK. From the proof, it can be seen that a more general theorem can be obtained if we replace $S_{n,n,\dots,n}$ by S_n where $n' = (n\theta_1, n\theta_2, \dots, n\theta_q)$ $0 < \theta_i \leq 1$. Then we have

 $\sup_{n} |P\{\sigma^{-1}n^{-q/2}(\theta_1 \cdots \theta_q)^{-1/2}S_{n'} < t\} - \varPhi(t)| < Cn^{-\alpha} \, \forall n \, .$

In fact it is in this stronger form that we will use it in the proof of Theorem 2.

(a) $P\{\limsup_{n \to \infty} \sup d(H_n, K_n) = 0\} = 1,$

and

(b)
$$P\{\bigcap_{x \in K_{\sigma}} [\liminf_{n \to \infty} d(H_n, x) = 0]\} = 1.$$

Proof. We will give only a very brief sketch of the proof since the arguments used are fairly standard and can be found e.g., in Chover (1967) and Wichura (1973). Take q = 2 for simplicity and $\sigma = 1$ without loss of generality.

We begin by showing a kind of asymptotic equi-continuity in the following form: Let $B = \prod_i (s_i, t_i]$ be a block in T^q ; write $m(B) = \min_{1 \le i \le q} (t_i - s_i)$. Then

LEMMA 1. Given $\varepsilon > 0$, $\exists \delta > 0$ such that if B is any block with $m(B) < \delta$ then the event $\{|H_n(A)| > \varepsilon\}$ occurs only finitely often wp.1.

Proof. Standard arguments (using the triangle inequality) such as those appearing on pp.56-59 of Billingsley (1968) show that it suffices to prove the following: Given $\varepsilon > 0$, $\exists \delta > 0$ such that

$$\sum_{n} \left[P\{ \max_{\substack{1 \leq i \leq n \, \delta \ 1 \leq j \leq n}} |S_{i,j}| > arepsilon \sqrt{2n^2 \log \log n} \}
ight. \ + P\{ \max_{\substack{1 \leq i \leq n \ 1 \leq j \leq n \, \delta}} |S_{i,j}| > arepsilon \sqrt{2n^2 \log \log n} \}] < \infty \; .$$

But this can be proved in a straightforward manner using the maximal inequality developed on pp. 713-714 of [3], Theorem 1 above and the arguments in §3 of Chover (1967). We omit the details.

Let now *m* be a positive integer. Consider a partition of the unit square (T^2) into $m \times m$ squares with corners (i/m, j/m), $0 \leq i$, $j \leq m$. We enumerate these squares (blocks) arbitrarily as B_{im} , $1 \leq i \leq m^2$. Let $\gamma > 0$ be a small positive number and denote by $B_{im}^* = B_{im}^*(\gamma)$ the square which is concentric with B_{im} (and is contained in B_{im}) with each side being equal to $(1 - 2\gamma)/m$.

If x is a function on T^2 we denote by $\pi_m x$ the function on T^2 defined by

$$(\pi_m x)(t_1, t_2) = \int_0^{t_1} \int_0^{t_2} \sum_{i=1}^{m_2} m^2 x(B_{im}) I_{B_{im}}(u_1, u_2) du_1 du_2$$

where I_{B} stands for the indicator of the block B.

Lemmas 2 and 3 below follow easily from the arguments used in proving Corollaries 1 and 2 in Chover (1967). Lemma 4 is immediate from Lemma 1.

LEMMA 2. Given $\varepsilon > 0$, $\exists m \text{ such that}$

 $P\{d(\pi_m H_n, H_n) > \varepsilon \text{ only finitely often } (in n)\} = 1$.

LEMMA 3. Given $\varepsilon > 0$, $\exists c > 1$ such that, wp. 1, $\max_{c^n \leq m \leq c^{n+1}} d(H_m, H_{[c^n]} > \varepsilon$ for only finitely many n.

LEMMA 4. Given $\varepsilon > 0$, $\exists \gamma > 0$ such that for each m and i $(1 \leq i \leq m^2)$,

$$P\{|H_n(B_{im}) - H_n(B_{im}^*)| > \varepsilon \text{ for only finitely many } n\} = 1$$
 .

We now proceed to prove (a) of the theorem. Let $\{\theta_i: 1 \leq i \leq m^2\}$ be real numbers such that $\sum_{i=1}^{m^2} \theta_i^2 = 1$. To prove (a) it suffices to show that for each m,

 $Pig\{\sum_{i=1}^{m^2} heta_i(\pi_m H_n)(B_{im}) < (1+arepsilon) ext{ for all large } n ig\} = 1 \;.$

In view of the preceding lemmas it thus suffices to prove (with c>1 sufficiently close to 1 and $\gamma>0$ sufficiently small)

$$\sum\limits_{n=1}^{\infty} P\Big\{m\sum\limits_{i=1}^{m^2} heta_i H_{[\mathfrak{c}^n]}(B^*_{i\mathfrak{m}}) > (1+arepsilon)\Big\} < \infty \; .$$

But the proof of this is essentially the same as given in §4 of Chover (1967). The only complication here is that the m^2 random variables $\{H_{[c^n]}(B_{im}^*): 1 \leq i \leq m^2\}$ are not independent. But there is enough separation among these and it suffices to apply Lemma 1.1.5 in Iosifescu and Theodorescu (1969).

To prove (b) take $x \in K$ with $\int_{0}^{1} \int_{0}^{1} (\partial^{2}x/\partial t_{1}\partial t_{2})^{2} dt_{1} dt_{2} < 1$. We need to show that $\forall \varepsilon > 0$, $P(\liminf d(H_{n}, x) < \varepsilon) = 1$. Again in view of the preceding lemmas and the arguments in Sec.5 of Chover (1967) it is enough to prove for sufficiently small $\delta > 0$, $\gamma > 0$

 $P(\lim_{n\to\infty}\sup F_n)=1$ where

$$F_n = \{ |H_{[\mathfrak{c}^n]}(B^*_{i\mathfrak{m}}) - x(B_{i\mathfrak{m}})| < \delta \ , \ \text{ all } i, \ 1 \leq i \leq m^2 \} \ .$$

[It might be noted here that (35) in [2] is insufficient; it should be strengthened to $P(\lim_{r\to\infty} \sup \bigcap_{\nu} C_r^{(\nu)}) = 1$.] Now if the probability of F_n in computed on the assumption that the m^2 random variables $\{H_{[c^n]}(B_{im}^*): 1 \leq i \leq m^2\}$ are independent then the error committed is

at most $m^2\phi_{[\nu e^n]}$ which forms a term of a convergent series in n. Hence using part (a) of the lemma on page 142 of [5] it is enough to show $\sum_n P(F_n) = \infty$. But given Theorem 1 this follows from computations which are standard in the proof of Strassen's theorem. This completes the proof of Theorem 2.

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