

Invariance principle for a Brownian motion with large drift in a white noise environment

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ABSTRACT. This paper discusses an invariance principle for a Brownian motion with drift coefficient $\kappa/4$ in a white noise environment under the assumption that κ is large. Our method clarifies the relation between the environment-wise invariance principle discussed in [7] and the present result (the invariance principle in random environment).

Introduction

Let W be the space of continuous functions on \mathbf{R} vanishing at 0 that is equipped with the Wiener measure P . For an element $w \in W$ let us denote by w_κ the element of W defined by $w_\kappa(x) = w(x) - (\kappa x/2)$ where κ is a given positive constant. For $w \in W$, P_w denotes the probability measure on $\Omega = C[0, \infty)$ such that $\mathbf{X}_x = \{\omega(t), t \geq 0, P_w\}$ is a diffusion process with generator

$$\mathcal{L}_w = \frac{1}{2} e^{w_\kappa(x)} \frac{d}{dx} \left(e^{-w_\kappa(x)} \frac{d}{dx} \right)$$

starting at 0, where $\omega(t)$ is the value of a function $\omega \in \Omega$ at time t . We regard $\omega(t)$ as a process defined on the probability space $\{W \times \Omega, \mathcal{P}\}$ where $\mathcal{P}(d\omega d\omega) = P(dw)P_w(d\omega)$. Then symbolically

$$d\omega(t) = dB(t) + \frac{\kappa}{4} dt - \frac{1}{2} w'(\omega(t)) dt,$$

where $B(t)$ is a standard Brownian motion independent of the white noise $\{w'(x)\}$. We call the process $\mathbf{X} = \{\omega(t), t \geq 0, \mathcal{P}\}$ a Brownian motion with drift in a white noise environment; in [2] [6] [7] it is called a diffusion process in a Brownian environment with drift. The present authors obtained some limit theorems for \mathbf{X} in [2] (see [8] for further results; see also [6] for a brief survey on related problems), which are analogous to those of [3] and [5]; however, some problems remain open. The present paper is a continuation of [7] and

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discusses the central limit theorem, or more precisely speaking, invariance principle *in random environment* in the case $\kappa > 2$.

We set

$$M_x (= M(x)) = 2 \int_0^x dy \int_{-\infty}^y e^{w_\kappa(y) - w_\kappa(z)} dz, \quad x \in \mathbf{R},$$

$$\mu(t) = \text{the inverse function of } \{M_x, x \in \mathbf{R}\}, \quad t \in \mathbf{R},$$

$$T_x (= T(x)) = \inf\{t > 0 : \omega(t) > x\}, \quad x \geq 0.$$

We use also the following notation:

$$\bar{\omega}(t) = \max\{\omega(s) : 0 \leq s \leq t\}, \quad \underline{\omega}(t) = \inf\{\omega(s) : s \geq t\}, \quad (\omega \in \Omega),$$

$$\gamma = (\kappa - 1)/2, \quad m = 4/(\kappa - 1) = 2/\gamma,$$

$$A = 64(\kappa - 1)^{-2}(\kappa - 2)^{-1} = 16\gamma^{-2}(2\gamma - 1)^{-1},$$

$$B = 64(\kappa - 1)^{-3}(\kappa - 2)^{-1} = 8\gamma^{-3}(2\gamma - 1)^{-1},$$

$$C = A + B = 64\kappa(\kappa - 1)^{-3}(\kappa - 2)^{-1} = 8(2\gamma + 1)\gamma^{-3}(2\gamma - 1)^{-1}.$$

The following theorem was proved in [7].

THEOREM A (*Environment-wise invariance principle, see [7]*). *When $\kappa > 2$, we have the following:*

(i) *For almost all $w \in W$ with respect to P , the process*

$$\left\{ \frac{T_{\lambda x} - M_{\lambda x}}{\sqrt{A\lambda}}, x \geq 0, P_w \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$ (in the sense of convergence of probability measures on the Skorohod space).

(ii) *For almost all w , the process*

$$\left\{ \frac{\omega(\lambda t) - \mu(\lambda t)}{\sqrt{m^{-3}A\lambda}}, t \geq 0, P_w \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$ (in the sense of convergence of probability measures on $C[0, \infty)$). The same is true when $\omega(\lambda t)$ is replaced by either of $\bar{\omega}(\lambda t)$ and $\underline{\omega}(\lambda t)$.

Our main theorems are the following ($\kappa > 2$ is assumed throughout).

THEOREM 1. (i) *The process*

$$\left\{ \frac{M_{\lambda x} - \lambda mx}{\sqrt{B\lambda}}, x \in \mathbf{R}, P \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$.

(ii) *The process*

$$\left\{ \frac{\mu(\lambda t) - \lambda m^{-1}t}{\sqrt{m^{-3}B\lambda}}, t \in \mathbf{R}, P \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$.

THEOREM 2 (*Invariance principle in random environment*). (i) *The process*

$$\left\{ \frac{T_{\lambda x} - \lambda mx}{\sqrt{C\lambda}}, x \geq 0, \mathcal{P} \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$.

(ii) *The process*

$$\left\{ \frac{\omega(\lambda t) - \lambda m^{-1}t}{m^{-3}C\lambda}, t \geq 0, \mathcal{P} \right\}$$

converges in law to a Brownian motion as $\lambda \rightarrow \infty$. The same is true when $\omega(\lambda t)$ is replaced by either of $\bar{\omega}(\lambda t)$ and $\underline{\omega}(\lambda t)$.

As in [7] we introduce a one parameter family of measure preserving transformations $\theta_t, t \in \mathbf{R}$, on (W, P) defined by $(\theta_t w)(x) = w(x+t) - w(t), x \in \mathbf{R}$. Clearly $\theta_t \theta_s = \theta_{t+s}$ and $\{\theta_t\}$ is ergodic. Set

$$(0.1) \quad f_0(w) = \int_{-\infty}^0 e^{-w_\kappa(t)} dt.$$

Then $\theta_t f_0 \equiv f_0(\theta_t w) = \int_{-\infty}^t e^{w_\kappa(t) - w_\kappa(s)} ds$ and we have the following (see [7]):

$$(0.2) \quad E_w\{T_x\} = M_x = 2 \int_0^x \theta_y f_0 dy;$$

the first equality of (0.2) holds for $x \geq 0$ and the second one holds for $x \in \mathbf{R}$.

$$(0.3) \quad \text{Var}_w\{T_x\} = 8 \int_0^x \theta_y g dy \text{ for } x \geq 0 \quad (g(w) = \int_{-\infty}^0 e^{-w_\kappa(t)} (\theta_t f_0)^2 dt).$$

$$(0.4) \quad E\{f_0\} = \gamma^{-1}, E\{f_0^2\} = 2\gamma^{-1}(2\gamma - 1)^{-1}.$$

$$(0.5) \quad E\{\text{Var}_w(T_x)\} = Ax \text{ for } x \geq 0.$$

$$(0.6) \quad \text{Var}\{M_x\} = Bx + O(1), x \rightarrow \infty, (\text{Var} = \text{variance}).$$

It was also observed in [7] that

$$(0.7) \quad d\theta_t f_0 = \theta_t f_0 dw(t) - (\gamma \theta_t f_0 - 1) dt,$$

so $\theta_t f_0$ is a stationary diffusion process obtained as the unique stationary positive solution of the stochastic differential equation (0.7). Therefore

$$(0.8) \quad \theta_t f_0 - f_0 = \int_0^t \theta_s f_0 dw(s) - \lambda \int_0^t \theta_s f ds,$$

where $f = f_0 - \gamma^{-1}$.

1. Proof of Theorem 1

For the proof of Theorem 1 we need some lemmas.

LEMMA 1 ([7]). $t^{-1/2} \max\{\theta_s f_0 : |s| \leq t\} \rightarrow 0$ as $t \rightarrow \infty$.

LEMMA 2 ([7]). For any positive constants c_1 ,

$$M_{t+u} - M_t = mu(1 + o(1)) + o(\sqrt{\lambda}), \quad |t| \leq c_1 \lambda, \quad u \in \mathbf{R},$$

where $o(1)$ represents a general term that tends to 0 as $\lambda \rightarrow \infty$ uniformly in (t, u) such that $|t| \leq c_1 \lambda$ and $u \in \mathbf{R}$, for almost all w ; $o(\sqrt{\lambda})$ is a term that can be expressed as $o(1)\sqrt{\lambda}$.

To prove (i) of Theorem 1 it is enough to consider $\int_0^x \theta_y f dy$ by virtue of (0.2). Making habitual use of t to indicate time we write

$$(1.1) \quad \frac{1}{\sqrt{\lambda}} \int_0^{\lambda t} \theta_s f ds = \frac{1}{\gamma \sqrt{\lambda}} \int_0^{\lambda t} \theta_s f_0 dw(s) - \frac{1}{\gamma \sqrt{\lambda}} (\theta_{\lambda t} f_0 - f_0).$$

By the ergodicity of $\{\theta_t\}$ and also by (0.4) we see that the quadratic variation of the stochastic integral term in (1.1) tends to $Bt/4$ as $\lambda \rightarrow \infty$, a.s., so the stochastic integral term itself converges in law to $\{(B/4)^{1/2} w(t), t \geq 0, P\}$ as $\lambda \rightarrow \infty$. The second term of the right hand side of (1.1) is negligible by Lemma 1. Therefore $X_\lambda^+ = \{\lambda^{-1/2} \int_0^{\lambda t} \theta_s f ds, t \geq 0\}$ converges in law to $\{(B/4)^{1/2} w(t), t \geq 0\}$, so does $X_\lambda^- = \{\lambda^{-1/2} \int_0^{-\lambda t} \theta_s f ds, t \geq 0\}$ because of the reversibility of the diffusion $\theta_t f$. Now the assertion (i) of Theorem 1 follows from the fact that X_λ^+ and X_λ^- are asymptotically independent as $\lambda \rightarrow \infty$.

To proceed let $\xi = \lambda m$ and put

$$(1.2) \quad \beta_\lambda(t) = (B\lambda)^{-1/2} (M_{\lambda t} - \lambda m t), \quad \tilde{\beta}_\lambda(t) = (m^{-3} B\xi)^{-1/2} (\mu(\xi t) - \xi m^{-1} t).$$

Then the assertion of (ii) of Theorem 1 follows immediately from the following Lemma.

LEMMA 3. For any $t_0 > 0$ and $\varepsilon > 0$

$$\lim_{\lambda \rightarrow \infty} P \left\{ \sup_{|t| \leq t_0} |\beta_\lambda(t) + \tilde{\beta}_\lambda(t)| > \varepsilon \right\} = 0.$$

PROOF. From the second equality of (1.2) we have $\lambda mt = M(\lambda t + \tilde{\beta}_\lambda(t)m^{-1}\sqrt{B\lambda})$ and hence

$$(B\lambda)^{-1/2}\{M(\lambda t + \tilde{\beta}_\lambda(t)m^{-1}\sqrt{B\lambda}) - M(\lambda t)\} = -\beta_\lambda(t),$$

so an application of Lemma 2 yields $\tilde{\beta}_\lambda(t)(1 + o(1)) + o(1) = -\beta_\lambda(t)$, where $o(1)$ is a term tending to 0 uniformly on each finite t -interval as $\lambda \rightarrow \infty$, a.s. This implies the lemma.

The following observation will lead to another proof of (ii) of Theorem 1. If $v(t)$ denotes the inverse function of $\int_0^x \theta_y f_0 dy$, then $v(t) = \mu(2t)$ and the derivative $v'(t)$, which equals to $1/\theta_{v(t)}f_0$, is a stationary diffusion process obtained from $\theta_t f_0$ by changing time and scale.

2. The proof of Theorem 2

We give the proof of the part (i). Taking an arbitrary positive sequence $\{\lambda_n, n = 1, 2, \dots\}$ tending to ∞ , we denote by $P^{(n)}$ the probability law of the process $\{\lambda_n^{-1/2}(T_{\lambda_n x} - \lambda_n mx), x \geq 0, \mathcal{P}\}$. Note that $P^{(n)}$ is a probability measure on the Skorohod space $D = D[0, \infty)$. For the proof of the part (i) it is enough to show that $P^{(n)}$ converges to the probability law of the process $\{\sqrt{C}w(x), x \geq 0, P\}$ as $n \rightarrow \infty$. We first prove that the sequence $\{P^{(n)}, n \geq 1\}$ is tight. If $Q_w^{(n)}$ denotes the probability law of the process $\{\lambda_n^{-1/2}(T_{\lambda_n x} - M_{\lambda_n x}), x \geq 0, P_w\}$, then $Q_w^{(n)} \rightarrow Q_1$ (P -a.s.) by Theorem A and hence $Q^{(n)} = \int Q_w^{(n)} P(dw)$ also converges to Q_1 as $n \rightarrow \infty$ where Q_1 is the probability law (on D) of the process $\{\sqrt{A}w(x), x \geq 0, P\}$. All the convergence here is to be understood as the convergence of probability measures on D . Therefore for any $\varepsilon > 0$ there exists a compact set $K_1 \subset D$ such that $Q^{(n)}(K_1^c) < \varepsilon^2$ for all $n \geq 1$. We then have $P\{L_n\} < \varepsilon$ where $L_n = \{w : Q_w^{(n)}(K_1^c) > \varepsilon\} = \{w : Q_w^{(n)}(K_1) \leq 1 - \varepsilon\}$. We also introduce, for each fixed w , an element $\varphi_n(w)$ of Ω which is defined to be the function $\lambda_n^{1/2}(M_{\lambda_n x} - \lambda_n mx)$ of x . Then $P \circ \varphi_n^{-1} \rightarrow Q_2$ as $n \rightarrow \infty$ by Theorem 1 where Q_2 is the probability law (on Ω) of the process $\{\sqrt{B}w(x), x \geq 0, P\}$. We can thus find a compact set $K_2 \subset \Omega$ such that $P \circ \varphi_n^{-1}(K_2^c) < \varepsilon$ for all $n \geq 1$. We now put $K = \{w_1 + w_2 : w_1 \in K_1, w_2 \in K_2\}$. Then K is a compact subset of D . Since

$$\frac{T_{\lambda_n x} - \lambda_n mx}{\sqrt{\lambda_n}} = \frac{T_{\lambda_n x} - M_{\lambda_n x}}{\sqrt{\lambda_n}} + \frac{M_{\lambda_n x} - \lambda_n mx}{\sqrt{\lambda_n}},$$

we have

$$\begin{aligned}
P^{(n)}(K) &= \iint 1_K(w_1 + \varphi_n(w)) Q_w^{(n)}(dw_1) P(dw) \\
&\geq \iint 1_{K_1}(w_1) 1_{K_2}(\varphi_n(w)) Q_w^{(n)}(dw_1) P(dw) \\
&= \int_{\varphi_n^{-1}(K_2)} Q_w^{(n)}(K_1) P(dw) \\
&\geq \int_{\varphi_n^{-1}(K_2) \cap L_n^c} (1 - \varepsilon) P(dw) \geq (1 - \varepsilon)(1 - 2\varepsilon),
\end{aligned}$$

which proves that $\{P^{(n)}, n \geq 1\}$ is tight. Therefore, for the proof of Theorem 2 it is enough to show that

$$(2.1) \quad \lim_{n \rightarrow \infty} \int f(w) P^{(n)}(dw) = \iint f(w_1 + w_2) Q_1(dw_1) Q_2(dw_2),$$

for any function f of the form

$$f(w) = \exp \left\{ \sqrt{-1} \sum_{j=1}^k \alpha_j w(t_j) \right\},$$

where $\alpha_j \in \mathbf{R}$ and $t_j \geq 0$, $1 \leq j \leq k$. For such an f the left hand side of (2.1) equals

$$\begin{aligned}
&\lim_{n \rightarrow \infty} \iint f(w_1 + \varphi_n(w)) Q_w^{(n)}(dw_1) P(dw) \\
&= \lim_{n \rightarrow \infty} \int_W \left\{ \int_D f(w_1) Q_w^{(n)}(dw_1) \right\} f(\varphi_n(w)) P(dw) \\
&= \int f(w_1) Q_1(dw_1) \int f(w_2) Q_2(dw_2)
\end{aligned}$$

which also equals the right hand side of (2.1). This completes the proof of (i) of Theorem 2.

The part (ii) of Theorem 2 can be proved in a way similar to the above by making use of Theorem A and Theorem 1.

3. Supplement to the proof of (i) of Theorem A

The proof of Theorem A was given in [7]; however, some details in the proof of the part (i) were omitted. It will be worth supplementing them.

The proof of Theorem A given in [7] proceeds as follows. Let $\tau_k =$

$T_k - T_{k-1}$, $\tilde{\tau}_k = \tau_k - E_w\{\tau_k\}$, $k \geq 1$. Then it was proved that, for almost all w , $\{\tilde{\tau}_k, k \geq 1, P_w\}$ is a sequence of independent random variables satisfying the Lindeberg condition. Therefore the central limit theorem holds for T_n with respect to P_w , for almost all w . Note that $E_w\{T_n\} = M_n$ and $\text{Var}_w\{T_n\} \sim An$ as $n \rightarrow \infty$ (P-a.s.). Now the rest of the proof, whose detail was omitted in [7], is given as follows.

Let $t_{nk} = \text{Var}_w\{\tau_k\}/\text{Var}_w\{T_n\}$, $\zeta_{nk} = (An)^{-1/2}\{T_k - M_k\}$, $1 \leq k \leq n$, and $t_{n0} = \zeta_{n0} = 0$. For each fixed w we construct a piece-wise linear function $\xi_n(x)$, $0 \leq x \leq 1$, with vertexes $(\sum_{j=0}^k t_{nj}, \zeta_{nk})$, $0 \leq k \leq n$. We regard $\{\xi_n(x), 0 \leq x \leq 1, P_w\}$ as a process with time parameter x . Then by Theorem 3.1 of [5], for almost all w , the process $\{\xi_n(x), 0 \leq x \leq 1, P_w\}$ converges in law to a Brownian motion as $n \rightarrow \infty$. We now modify $\xi_n(x)$ slightly, namely, we consider a piece-wise linear function $\eta_n(x)$ with vertexes $(k/n, \zeta_{nk})$, $0 \leq k \leq n$. Then $\eta_n(x)$ can be represented as $\eta_n(x) = \xi_n(\varphi_n(x))$ where $\varphi_n(x)$ is the piece-wise linear function with vertexes $(k/n, \sum_{j=0}^k t_{nj})$, $0 \leq k \leq n$. On the other hand it is easy to see that, for each fixed x , $\varphi_n \rightarrow x$ as $n \rightarrow \infty$ for almost all w . This combined with the fact that φ is increasing implies that $\varphi_n \rightarrow x$ uniformly as $n \rightarrow \infty$ (P-a.s.). Therefore

$$(3.1) \quad \text{the process } \{\eta_n(x), 0 \leq x \leq 1, P_w\} \text{ converges in law to a Brownian motion as } n \rightarrow \infty \text{ for almost all } w.$$

We finally prove that the process $\{(A\lambda)^{-1/2}(T_{\lambda x} - M_{\lambda x}), x \in [0, 1], P_w\}$ converges in law to a Brownian motion as $\lambda \rightarrow \infty$ for almost all w ; the time interval $[0, 1]$ can be replaced by an arbitrary interval $[0, t_0]$ with a minor modification of the proof. Given $x \in (0, 1]$ and an integer $n \geq 1$ we take the integer k such that $(k-1)/n < x \leq k/n$. Then $T_{nx} - M_{nx} > T_{k-1} - M_k > \sqrt{An}\eta_n(x) - \tau_k - m_k$ where $m_k = M_k - M_{k-1}$. Similarly $T_{nx} - M_{nx} < \sqrt{An}\eta_n(x) + \tau_k + m_k$ and hence

$$\sqrt{An}\eta_n(x) - (\tau_k + m_k) < T_{nx} - M_{nx} < \sqrt{An}\eta_n(x) + (\tau_k + m_k).$$

This implies that for $x \in [0, 1]$

$$(3.2) \quad \sqrt{An}\eta_n(x) - (\hat{\tau}_n + \hat{m}_n) < T_{nx} - M_{nx} < \sqrt{An}\eta_n(x) + (\hat{\tau}_n + \hat{m}_n),$$

where $\hat{\tau}_n = \max\{\tau_k : 1 \leq k \leq n\}$ and $\hat{m}_n = \max\{m_k : 1 \leq k \leq n\}$. Next, given $\lambda > 0$ let $n = n(\lambda)$ be the integer such that $n-1 < \lambda \leq n$. Then $T_{(n-1)x} - M_{\lambda x} < T_{\lambda x} - M_{\lambda x} \leq T_{nx} - M_{\lambda x}$, which combined with (3.2) implies

$$\begin{aligned} & \sqrt{A(n-1)}\eta_{n-1}(x) - (\hat{\tau}_{n-1} + \hat{m}_{n-1}) - (M_{\lambda x} - M_{(n-1)x}) \\ & < T_{\lambda x} - M_{\lambda x} < \sqrt{An}\eta_n(x) + (\hat{\tau}_n + \hat{m}_n) + (M_{nx} - M_{\lambda x}). \end{aligned}$$

Since $M_{\lambda x} - M_{(n-1)x}$ and $M_{nx} - M_{\lambda x}$ are dominated by $2\hat{m}_n$, we obtain

$$(3.3) \quad \begin{aligned} & \left(\frac{n-1}{\lambda}\right)^{1/2} \eta_{n-1}(x) - (A\lambda)^{-1/2}(\hat{\tau}_n + 3\hat{m}_n) < (A\lambda)^{-1/2}(T_{\lambda x} - M_{\lambda x}) \\ & < \left(\frac{n}{\lambda}\right)^{1/2} \eta_n(x) + (A\lambda)^{-1/2}(\hat{\tau}_n + 3\hat{m}_n). \end{aligned}$$

On the other hand we can prove that for almost all w

$$(3.4) \quad P_w \left\{ \lim_{n \rightarrow \infty} \hat{\tau}_n / \sqrt{n} = 0 \right\} = 1,$$

$$(3.5) \quad \lim_{n \rightarrow \infty} \hat{m}_n / \sqrt{n} = 0.$$

In fact, it is easy to see that $\{\tau_k, k \geq 1, \mathcal{P}\}$ is stationary and ergodic. Since τ_k^2 is integrable we have $n^{-1} \sum_{k=1}^n \tau_k^2 \rightarrow \text{const.}$ as $n \rightarrow \infty$ (\mathcal{P} -a.s.) and hence $n^{-1} \tau_n^2 \rightarrow 0$, namely, $n^{-1/2} \tau_n \rightarrow 0$ (\mathcal{P} -a.s.). This implies $n^{-1/2} \hat{\tau}_n \rightarrow 0$ (\mathcal{P} -a.s.) and hence (3.4). (3.5) can be proved in a similar manner. By virtue of (3.1), (3.4) and (3.5) the processes of the leftmost and rightmost hands of (3.3) converge in law to a Brownian motion as $\lambda \rightarrow \infty$. Therefore (3.3) implies the assertion for $(A\lambda)^{-1/2}(T_{\lambda x} - M_{\lambda x})$ that we wanted to prove.

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