

## A RENEWAL THEOREM FOR CURVED BOUNDARIES AND MOMENTS OF FIRST PASSAGE TIMES<sup>1</sup>

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Let  $X_1, X_2, \dots$  be i.i.d. with a finite positive mean  $\mu$  and a finite positive variance  $\sigma^2$  and let  $S_n = X_1 + \dots + X_n$ ,  $n \geq 1$ . Further, let  $0 \leq \alpha < 1$  and  $t_c$  be the first  $n \geq 1$  for which  $S_n > cn^\alpha$  and let  $W_c(a) = \sum_{n=1}^{\infty} P\{t_c > n, c(n+1)^\alpha - S_n < a\}$ . Under some additional conditions on the distribution of  $X_1$  we show that  $W_c$  converges weakly to a limit  $W$ , where  $W'(a) = \beta\mu^{-1}P\{S_k \geq (k+1)\alpha\mu - a, \text{ for all } k \geq 0\}$ , with  $\beta = 1/(1-\alpha)$ . We then find the asymptotic distribution of the excess  $R_c = S_{t_c} - ct_c^\alpha$  and show that  $R_c$  is asymptotically independent of  $t_c^* = (t_c - E(t_c))/E(t_c)^{1/2}$ , and we compute  $E(t_c)$  up to terms which are  $o(1)$  as  $c \rightarrow \infty$ .

**1. Introduction.** Let  $X_1, X_2, \dots$  denote independent and identically distributed random variables which have a finite positive mean  $\mu$ , let  $S_0 = 0$ , and let  $S_n = X_1 + \dots + X_n$  for  $n \geq 1$ . Further, let  $V$  be a positive function on  $(0, \infty)$  for which  $V(x) = o(x)$  as  $x \rightarrow \infty$ , and let

$$t_c = \inf \{n \geq 1 : S_n > cV(n)\}$$

for  $c > 0$ , where  $\inf \emptyset = \infty$ . Observe that  $t_c$  is finite w.p. 1. If  $V(x) = 1$  for all  $x > 0$ ,  $\Pr\{X_1 > 0\} = 1$ , and the distribution of  $X_1$  is nonarithmetic, then the renewal theorem (Feller (1966), page 347) asserts that  $U(c) = E\{t_c\}$  is finite for every  $c > 0$  and that

$$(1.1) \quad \lim U(c) - U(c - a) = a\mu^{-1}$$

as  $c \rightarrow \infty$  for every  $a > 0$ .

In this paper we will extend (1.1) to a wider class of boundaries  $V$  and to distributions which are not concentrated on  $(0, \infty)$ , although we will impose some additional conditions on the distribution of  $X_1$ .

We will assume throughout this paper that  $V$  is a positive, continuous, eventually concave function on  $(0, \infty)$  for which  $V(x) = x^\alpha L(x)$  for  $x > 0$ , where  $0 \leq \alpha < 1$ , and

$$(1.2) \quad \frac{L(x+y)}{L(x)} - 1 = o\left(\frac{1}{x}\right)$$

as  $x \rightarrow \infty$  for every  $y > 0$ . Equation (1.2) is more restrictive than requiring that  $L$  vary slowly at  $\infty$ , but includes all functions of the form  $L(x) = \log^k(1+x)$ ,

Received November 20, 1974; revised June 28, 1975.

<sup>1</sup> Research supported by the National Institute of Health under GM 20507-01.

AMS 1970 subject classifications. Primary 60F05; Secondary 60K05.

Key words and phrases. Renewal theorem, conditional distribution, excess over the boundary, first passage times.

$x > 0$ , with  $k \in R$ . Let

$$W_c(a) = \sum_{k=0}^{\infty} P\{t_c > k, cV(k+1) - S_k < a\}$$

for  $a > 0$  and  $c > 0$ . We observe that if  $V = 1$  and  $X_1 > 0$ , then  $W_c(a) = U(c) - U(c - a)$  for  $0 < a < c$ . In Section 3, we will show that if  $X_1$  has a finite positive variance, a finite third moment, and a suitably smooth distribution, then

$$(1.3) \quad \lim W_c(a) = \beta\mu^{-1} \int_0^a \phi(y) dy$$

for  $a > 0$  as  $c \rightarrow \infty$ , where  $\beta = 1/(1 - \alpha)$  and

$$(1.4) \quad \phi(y) = P\{S_k \geq (k+1)\alpha\mu - y, \text{ for all } k \geq 0\}$$

for  $0 \leq y < \infty$ . Of course, if  $V = 1$  and  $X_1 > 0$ , then (1.3) reduces to (1.1). If  $V = 1$ , then (1.3) is closely related to, but distinct from, a result of Spitzer (1960), Theorem 5.1 a.

Relation (1.3) has several interesting consequences. Let

$$R_c = S_{t_c} - cV(t_c), \quad c > 0,$$

denote the excess over the boundary. In Section 4 we will show that  $R_c$  has a limiting distribution  $H$  as  $c \rightarrow \infty$ , where  $H$  has density

$$h(y) = \beta\mu^{-1}P\{S_k \geq k\alpha\mu + y, \text{ for all } k \geq 1\}$$

for  $0 \leq y < \infty$ . Moreover, we will show that  $R_c$  is asymptotically independent of  $t_c^* = (t_c - E\{t_c\})/E\{t_c\}^{1/2}$ , which has an asymptotic normal distribution.

In Section 5 we use the asymptotic distribution of  $R_c$  to compute  $E\{t_c\}$  up to terms which are  $o(1)$  as  $c \rightarrow \infty$ . In fact, we show that if  $V(x) = x^\alpha$  for  $x > 0$ , then

$$E\{t_c\} = c^\beta\mu^{-\beta} + \beta\mu^{-1}\nu - \frac{1}{2}\alpha\beta^2\sigma^2\mu^{-2} + o(1)$$

as  $c \rightarrow \infty$ , where  $\sigma^2$  is the variance of  $X_1$  and  $\nu$  denotes the expectation of  $H$  (and is given explicitly in equation (4.5)).

The asymptotic properties of stopping times of the form  $t_c$  have been studied by Chow and Robbins (1963), Chow (1966), Siegmund (1967, 1968, and 1969), and Gut (1972 and 1974) under a variety of assumptions on  $V$  and the distribution of  $X_1, X_2, \dots$ .

**2. Preliminaries.** In this section we will suppose that  $X_1, X_2, \dots$  are i.i.d. random variables with a finite positive mean  $\mu$  and a finite positive variance  $\sigma^2$ . In addition, we will sometimes impose the following condition.

**CONDITION C.**  $X_1$  has a density  $f$  (with respect to Lebesgue measure) which is continuous a.e. (with respect to Lebesgue measure). Moreover, some power of the characteristic function of  $X_1$  is integrable.

To motivate the lemmas of the section, we observe that

$$(2.1) \quad W_c(b) - W_c(a) = \sum_{k=1}^{\infty} \int_{I_k} P\{t > k | S_k^* = z\} dG_k^*(z) + I_{[a,b)}(cV(1))$$

for  $0 \leq a < b < \infty$ , where

$$(2.2a) \quad S_k^* = \frac{S_k - k\mu}{\sigma k^{\frac{1}{2}}}$$

$$(2.2b) \quad I_k = I_k(a, b) = \left( y_k - \frac{b}{\sigma k^{\frac{1}{2}}}, y_k - \frac{a}{\sigma k^{\frac{1}{2}}} \right]$$

$$(2.2c) \quad y_k = \frac{cV(k+1) - k\mu}{\sigma k^{\frac{1}{2}}},$$

and  $G_k^*$  denotes the distribution function of  $S_k^*$  for  $k \geq 1$ . We observe that for a.e.  $z \in I_k[G_k^*]$ ,  $P\{t > k | S_k^* = z\} \leq \phi_{k,c}(b, z)$ , where

$$\phi_{k,c}(b, z) = P\{S_{kj} \leq cV(j) - jk^{-1}[cV(k+1) - b], 1 \leq j \leq k | S_k^* = z\}$$

with

$$S_{kj} = S_j - jk^{-1}S_k$$

for  $j = 1, \dots, k$ . Similarly,  $P\{t > k | S_k^* = z\} \geq \phi_{k,c}(a, z)$  for a.e.  $z \in I_k[G_k^*]$ . In this section we will derive an asymptotic expression for an appropriate version of  $\phi_{k,c}$ .

Let  $X_{ki} = X_i - k^{-1}S_k$ ,  $j = 1, \dots, k - 1$ . If Condition C is satisfied, then we may construct a regular conditional distribution for  $X_{k1}, \dots, X_{k(k-1)}$  given  $S_k^*$  as follows. By Condition C and the local limit theorem for densities (Feller (1966), pages 489-490),  $S_k^*$  has a density  $g_k^*$  which is continuous for  $k$  sufficiently large and converges to  $\phi$ , the standard normal density, uniformly on  $R$ . Let  $g_0$  denote any density on  $R$  and let

$$g_k(y_1, \dots, y_{k-1}, z) = \prod_{i=1}^{k-1} f(y_i + \mu + \sigma k^{-\frac{1}{2}}z) \sigma k^{\frac{1}{2}} f(\mu + \sigma k^{-\frac{1}{2}}z - \sum_{i=1}^{k-1} y_i)$$

and

$$g_k(y_1, \dots, y_{k-1} | z) = g_k(y_1, \dots, y_{k-1}, z) g_k^*(z)^{-1} I_{(0,\infty)}(g_k^*(z)) + \prod_{i=1}^{k-1} g_0(y_i) I_{\{0\}}(g_k^*(z))$$

for  $(y_1, \dots, y_{k-1}, z) \in R^k$ . Then

$$(2.3) \quad Q_k(z, B) = \int_B g_k(y_1, \dots, y_{k-1} | z) dy_1 \dots dy_{k-1},$$

$B \in \mathbf{B}(R^{k-1})$ ,  $z \in R$ , defines a regular conditional distribution for  $X_{k1}, \dots, X_{k(k-1)}$  given  $S_k^*$ . We will sometimes write  $Q_{k,z}$  for  $Q_k(z, \cdot)$ .

Let  $Y_{k1}, \dots, Y_{k(k-1)}$  denote the coordinate functions on  $R^{k-1}$ ,  $k \geq 2$ , and let  $\pi_{kj} = (Y_{k1}, \dots, Y_{kj})$  denote the projection from  $R^{k-1}$  onto  $R^j$ ,  $1 \leq j < k$ . Also, let  $F_1$  denote the distribution of  $X_1 - \mu$  and let  $F_1^j$  denote the product of  $j$  copies of  $F_1$ .

LEMMA 2.1. *If Condition C is satisfied and if  $I$  is any compact interval, then for any  $j \geq 1$ ,  $Q_{k,z} \circ \pi_{kj}^{-1}$  converges strongly to  $F_1^j$  uniformly with respect to  $z \in I$ . Moreover, letting  $T_{kj} = Y_{k1} + \dots + Y_{kj}$  for  $1 \leq j \leq k - 1$ ,*

$$\lim \int |T_{kj}| dQ_{k,z} = E|S_j - j\mu|$$

as  $k \rightarrow \infty$  uniformly with respect to  $z \in I$  for each  $j \geq 1$ .

PROOF. If  $I$  is given, then there is a  $k_0 = k_0(I)$  for which  $g_k^*$  is positive and continuous on  $I$  for  $k \geq k_0$ . It is then easily seen that for  $k \geq k_0$  and  $z \in I$ ,  $Q_{k,z} \circ \pi_{k,j}^{-1}$  has density

$$(2.4) \quad \begin{aligned} & h_{k,j}(y_1, \dots, y_j | z) \\ &= \left[ \left( \frac{k}{k-j} \right)^{\frac{1}{2}} \prod_{i=1}^j f(y_i + \mu + \sigma k^{-\frac{1}{2}} z) \right. \\ & \quad \left. \times g_{k-j}^* \left\{ \left( \frac{k-j}{k} \right)^{\frac{1}{2}} z - \sigma^{-1} (k-j)^{-\frac{1}{2}} \sum_{i=1}^j y_i \right\} \right] / g_k^*(z) \end{aligned}$$

with respect to Lebesgue measure. Thus if  $z_k, k \geq 1$ , is any sequence from  $I$ , then

$$\lim h_{k,j}(y_1, \dots, y_j | z_k) = \prod_{i=1}^j f(y_i + \mu)$$

for a.e.  $(y_1, \dots, y_j)$ . The first assertion of the lemma now follows from Scheffe's theorem (Lehmann (1959), page 351).

Let  $M_k$  and  $m_k$  denote the maximum of  $g_k^*$  and the minimum of  $g_k^*(z)$  for  $z \in I$ , respectively. Then  $h_{k,j}(y_1, \dots, y_j | z) \leq 2M_{k-j} m_k^{-1} \prod_{i=1}^j f(y_i + \mu + k^{-\frac{1}{2}} \sigma z)$  for  $y_1, \dots, y_j \in R^j, z \in I$ , and  $k$  sufficiently large, so

$$\int T_{k,j}^2 dQ_{k,z} \leq 2M_{k-j} m_k^{-1} j \sigma^2 \{1 + j k^{-1} z^2\}$$

is uniformly bounded for  $z \in I$  and  $k$  sufficiently large. The lemma now follows from the convergence of moments theorem (Loève (1963), page 184).

It is easily seen that the equation

$$(2.5) \quad \mu \lambda = cV(\lambda)$$

has a unique solution  $\lambda = \lambda(c)$  for  $c$  sufficiently large, and it is known that  $t_c/\lambda \rightarrow 1$  w.p. 1 as  $c \rightarrow \infty$  (Siegmund (1967)).

In our next lemma we let  $T_{k,j} = Y_{k1} + \dots + Y_{kj}$  for  $1 \leq j < k$ .

LEMMA 2.2. *Suppose that Condition C is satisfied and let  $I$  be any compact interval. If  $k = k_c \rightarrow \infty$  as  $c \rightarrow \infty$  in such a manner that  $k \sim \lambda$ , then*

$$\lim_{m \rightarrow \infty} \limsup_{c \rightarrow \infty} Q_{k,z} \{ T_{k,j} \geq cV(j) - jk^{-1}[cV(k+1) - y] \text{ for some } j \leq k - m \} = 0$$

uniformly with respect to  $z \in I$  for any  $y \in R$ .

The proof of Lemma 2.2 is similar to that of Lemma 2.1, but longer and much more technical. We defer it to Section 6.

LEMMA 2.3. *Suppose that Condition C is satisfied and let  $I$  be any compact interval. Let*

$$\phi_{k,c}(y, z) = Q_{k,z} \{ T_{k,j} \leq cV(j) - jk^{-1}[cV(k+1) - y], 1 \leq j \leq k \}$$

for  $y \in R$  and  $z \in R$ . If  $k \rightarrow \infty$  as  $c \rightarrow \infty$  in such a manner that  $k \sim \lambda$ , then  $\lim \phi_{k,c}(y, z) = \phi(y)$  uniformly with respect to  $z \in I$  for every  $y \in R$ , where  $\phi(y) = P\{S_j \geq (j+1)\alpha\mu - y, \text{ for all } j \geq 0\}$ , as in (1.4).

**PROOF.** By Lemma 2.2,  $\phi_{k,c}(y, z) = \phi_{k,m}^*(y, z) + \varepsilon(c, m)$ , where  $\phi_{k,m}^*(y, z) = Q_{k,z}\{T_{kj} \leq cV(j) - jk^{-1}[cV(k+1) - y]\}$ , for  $k - m \leq j \leq k$  and  $\lim_{m \rightarrow \infty} \limsup_{c \rightarrow \infty} |\varepsilon(c, m)| = 0$ , uniformly on compacts (in  $z$ ) for each fixed  $y$ . By symmetry, we also have

$$\phi_{k,m}^*(y, z) = Q_{k,z}\{-T_{kj} \leq cV(k - j) - (k - j)k^{-1}[cV(k + 1) - y]\}, \text{ for } 0 \leq j \leq m\}.$$

Moreover,

$$cV(k - j) - (k - j)k^{-1}[cV(k + 1) - y] \rightarrow (1 - \alpha)\mu j - \alpha\mu$$

as  $c \rightarrow \infty$  for each fixed  $j$ . Thus, by Lemma 2.1,  $\phi_{k,m}^*(y, z) \rightarrow P\{S_j \geq (j + 1)\alpha\mu - y, \text{ for } 0 \leq j \leq m\}$  as  $c \rightarrow \infty$  for each fixed  $m$ . The lemma now follows by letting  $m \rightarrow \infty$ , since  $P\{S_j \leq (j + 1)\alpha\mu - y, \text{ for some } j \geq m\} \rightarrow 0$  as  $m \rightarrow \infty$ .

**3. A renewal theorem.** In this section we will suppose that  $X_1$  has a finite positive mean  $\mu$  and a finite positive variance  $\sigma^2$ . Let  $I$  be any compact interval and let

$$J_0 = \{k \geq 0 : y_k \in I\} \quad \text{and} \quad J_1 = \{k \geq 0 : y_k \notin I\},$$

where  $y_k$  is as in (2.2c). Further, let

$$(3.1) \quad W_c^i(a) = \sum_{k \in J_i} P\{t > k, cV(k + 1) - S_k < a\}, \quad i = 0, 1,$$

for  $a \geq 0$ , so that  $W_c = W_c^0 + W_c^1$ . The dependence of  $W_c^0$  and  $W_c^1$  on  $I$  will be suppressed in the notation.

**LEMMA 3.1.** *If Condition C is satisfied and if  $I$  is any finite interval, then*

$$\lim W_c^0(b) - W_c^0(a) = (\beta\mu^{-1} \int_a^b \phi(z) dz) (\int_I \phi(z) dz)$$

as  $c \rightarrow \infty$  for  $0 \leq a < b < \infty$ , where  $\phi$  denotes the standard normal density and  $\phi$  is as in (1.4).

**PROOF.** Let  $I_k$  be as in (2.2b). Then

$$(3.2) \quad W_c^0(b) - W_c^0(a) \leq \sum_{k \in J_0} \int_{I_k} \phi_{k,c}(b, z) g_k^*(z) dz,$$

where  $g_k^*$  denotes a density for  $S_k^*$  and  $\phi_{k,c}$  is as in Lemma 2.3. As  $c \rightarrow \infty$ ,  $k \sim \lambda$  uniformly with respect to  $k \in J_0$  and  $g_k^*(z) \rightarrow \phi(z)$  uniformly with respect to  $k \in J_0$  and  $z \in R$ , so that

$$\int_{I_k} \phi_{k,c}(b, z) g_k^*(z) dz \sim \phi(b)\sigma^{-1}k^{-\frac{1}{2}}(b - a)\phi(y_k)$$

uniformly with respect to  $k \in J_0$ . Since  $y_k - y_{k+1} \sim (1 - \alpha)\mu\sigma^{-1}k^{-\frac{1}{2}}$  uniformly with respect to  $k \in J_0$ , it now follows easily that

$$\limsup W_c^0(b) - W_c^0(a) \leq \beta\mu^{-1}(b - a)\phi(b) \int_I \phi(y) dy$$

as  $c \rightarrow \infty$ . A similar lower bound with  $\phi(a)$  replacing  $\phi(b)$  may be obtained for the  $\liminf$ , and the lemma then follows from the Riemann integrability of  $\phi$ .

LEMMA 3.2. *Suppose that Condition C is satisfied and that  $E|X_1^3| < \infty$ . Then*

$$\lim_{I \uparrow R} \limsup_{c \rightarrow \infty} W_c^1(b) - W_c^1(a) = 0$$

for  $0 \leq a < b < \infty$ . Moreover, if  $I$  is any finite interval, then there is a constant  $B = B_1$  for which  $W_c^1(y) \leq B(1 + y)$  for all  $y \geq 0$  and all  $c > 0$ .

The proof of Lemma 3.2 is somewhat similar to that of Lemma 3.1, but longer and more technical. We defer it to Section 6.

THEOREM 3.1. *Let  $X_1, X_2, \dots$  be i.i.d. with finite positive mean  $\mu$ , finite positive variance  $\sigma^2$ , and a finite third moment. Suppose also that Condition C is satisfied. Then*

$$\lim W_c(a) = \beta \mu^{-1} \int_0^a \phi(y) dy$$

for  $0 \leq a < \infty$  as  $c \rightarrow \infty$ . Here  $\beta = 1/(1 - \alpha)$  and  $\phi$  is as in (1.4).

PROOF. The theorem follows directly from Lemmas 3.1 and 3.2 by letting  $c \rightarrow \infty$  and  $I \uparrow R$  in that order.

It is clear from Theorem 3.1 that if  $z$  is a continuous function with compact support in  $[0, \infty)$ , then

$$(3.3) \quad \lim \int_0^\infty z(y) dW_c(y) = \beta \mu^{-1} \int_0^\infty z(y) \phi(y) dy$$

as  $c \rightarrow \infty$ . A more interesting class of functions is covered by the following theorem.

THEOREM 3.2. *Suppose that the hypotheses of Theorem 3.1 are satisfied. If  $z$  is a nonincreasing, nonnegative, integrable function on  $[0, \infty)$ , then (3.3) holds.*

PROOF. To establish Theorem 3.2, one integrates the left side of (3.3) by parts, applies the dominated convergence theorem, and then integrates the limit by parts. The dominating function is supplied by taking  $I = \emptyset$  in the second conclusion of Lemma 3.2.

**4. On the excess over the boundary.** In this section we will find and study the asymptotic distribution of the excess over the boundary

$$R_c = S_{t_c} - cV(t_c)$$

as  $c \rightarrow \infty$ . We denote the distribution function of  $R_c$  by  $H_c$ .

THEOREM 4.1. *If the hypotheses of Theorem 3.1 are satisfied, then  $H_c$  converges weakly to a limit  $H$  as  $c \rightarrow \infty$ , where  $H$  has density*

$$(4.1) \quad h(y) = \beta \mu^{-1} P\{S_j \geq j\alpha\mu + y, \text{ for all } j \geq 1\}.$$

Moreover, if  $E|X_1^{k+1}| < \infty$ , then the first  $k$  moments of  $H_c$  converge to those of  $H$ .

PROOF. Let  $F$  denote the distribution function of  $X_1$ . Then,

$$(4.2) \quad H_c(a) = \int_0^\infty [F(a + y) - F(y)] dW_c(y)$$

for  $0 \leq a < \infty$ . Moreover, for each  $a > 0$ , the integrand in (4.2) is the difference of two functions to which Theorem 3.2 applies, so that  $H_c(a)$  converges to

$$H(a) = \beta\mu^{-1} \int_0^\infty [F(a + y) - F(y)]\phi(y) dy$$

as  $c \rightarrow \infty$ . That  $H' = h$  follows from standard manipulations, but it is not clear that  $h$  is a density. To see this observe that

$$\int_0^\infty h(y) dy = \beta\mu^{-1}E\{M^+\},$$

where  $M = \min \{S_j' : j \geq 1\}$  with  $X_i' = X_i - \alpha\mu$  and  $S_j' = X_1' + \dots + X_j'$ . Now  $M = X_1' + \min \{M_1, 0\}$ , where  $M_1$  has the same distribution as  $M$ , so that  $E\{M\} = (1 - \alpha)\mu - E\{M^-\}$ —that is,  $E\{M^+\} = (1 - \alpha)\mu$ .

The proof of the second assertion in Theorem 4.1 is similar to that of the first.

We need the following result which is due to Siegmund (1968) in the case that  $V(x) = x^\alpha$  and has been extended by Gut (1974), pages 299–300.

**THEOREM 4.2.** *Let  $X_1, X_2, \dots$  be i.i.d. with finite positive mean  $\mu$  and finite positive variance  $\sigma^2$ . Then the distribution of*

$$t_c^* = \lambda^{-1}(t_c - \lambda)$$

*converges weakly to the normal distribution with mean 0 and variance  $\tau^2 = \beta^2\sigma^2\mu^{-2}$  as  $c \rightarrow \infty$ .*

**THEOREM 4.3.** *If the hypotheses of Theorem 3.1 are satisfied, then  $t_c^*$  and  $R_c$  are asymptotically independent. That is,*

$$(4.3) \quad \lim P\{t_c^* \in I, R_c \in J\} = (\tau^{-1} \int_I \phi(y\tau^{-1}) dy)(\int_J h(y) dy)$$

*as  $c \rightarrow \infty$  for all intervals  $I$  and  $J$ .*

**PROOF.** Let

$$s_c^* = \sigma^{-1}(t_c - 1)^{-1}[cV(t_c) - \mu(t_c - 1)];$$

then it follows easily from Taylor's theorem and Theorem 4.2 that  $s_c^* - \tau^{-1}t_c^* \rightarrow 0$  in probability as  $c \rightarrow \infty$ , so it will suffice to show that  $s_c^*$  and  $R_c$  are asymptotically independent. If  $I$  and  $J = [a, b]$  are finite intervals, then

$$P\{s_c^* \in I, R_c \in J\} = \int_0^\infty [F(b + y) - F(a + y)] dW_c^0(y),$$

where  $W_c^0$  is as in (3.1). It now follows easily from Lemma 3.1 and an argument similar to that given in the proof of Theorem 4.1 that

$$\lim P\{s_c^* \in I, R_c \in J\} = (\int_I \phi dy)(\int_J h dy).$$

The extension to unbounded intervals is routine.

We will now relate the asymptotic distribution  $H$  to the distribution  $F$  of  $X_1$ . Let  $X_i' = X_i - \alpha\mu$  and  $M = \min \{S_j' : j \geq 1\}$ , as above. Further, let  $F_\alpha$  and  $\phi_\alpha$  denote the distribution function and characteristic function of  $X_1'$  and let  $G$  and  $w$  denote the distribution function of  $M$  and the characteristic function of  $M^- = \min \{M, 0\}$ , respectively. Then  $h = \beta\mu^{-1}(1 - G)$  on  $[0, \infty)$ , and a result

of Spitzer (1960) asserts that

$$(4.4) \quad w(t) = \exp\left\{\sum_{k=1}^{\infty} k^{-1} \int_{-\infty}^0 (e^{itz} - 1) dF_{\alpha}^{*k}(x)\right\},$$

where \* denotes convolution. See also Feller (1966), page 576.

**THEOREM 4.4.** *Let  $X_1, X_2, \dots$  be i.i.d. with finite positive mean  $\mu$  and let  $H$  be as in Theorem 4.1. Then the characteristic function of  $H$  is given by*

$$\hat{H}(t) = \beta\mu^{-1} \left( \frac{\phi_{\alpha}(t) - 1}{it} \right) w(t)$$

for  $t \neq 0$ , where  $w$  is as in (4.4). If, in addition,  $X_1$  has finite, positive variance  $\sigma^2$ , then the mean of  $H$  is

$$(4.5) \quad \nu = \frac{\sigma^2 + (1 - \alpha)^2\mu^2}{2(1 - \alpha)\mu} - \sum_{k=1}^{\infty} k^{-1} E\{(S_k - k\alpha\mu)^-\}.$$

The first assertion follows from the identity  $M = X_1' + \min\{M_1, 0\}$ , where  $M_1 = \min\{S'_{j+1} - S'_1 : j \geq 1\}$  has the same distribution as  $M$  and is independent of  $X_1'$ . The second then follows by differentiation. We omit the details.

**5. On the expectation of  $t_c$ .** In this section we will derive an asymptotic expression for the expectation of the first passage time  $t_c$ . We suppose throughout that  $X_1, X_2, \dots$  are i.i.d. with  $\mu > 0$  and  $0 < \sigma^2 < \infty$ . We will also suppose that  $V(x) = x^{\alpha}$ ,  $x > 0$ , in which case

$$\lambda = c^{\beta}\mu^{-\beta}$$

for  $c > 0$ . We will need to know when powers of  $t_c^{*2} = \lambda^{-1}(t_c - \lambda)$  are uniformly integrable.

**THEOREM 5.1.** *If  $E|X_1|^{\gamma} < \infty$  for some  $\gamma > \max\{4, \beta\}$ , then  $t_c^{*2}$  is uniformly integrable with respect to  $c > 0$ .*

The proof of Theorem 5.1 will be given in Section 7.

**THEOREM 5.2.** *Suppose that the hypotheses of Theorem 3.1 are satisfied, that  $V(x) = x^{\alpha}$ ,  $x > 0$ , and that  $E|X_1|^{\gamma} < \infty$  for some  $\gamma > \max\{4, \beta\}$ . Then*

$$E\{t_c\} = c^{\beta}\mu^{-\beta} + \beta\mu^{-1}\nu - \frac{1}{2}\alpha\beta^2\sigma^2\mu^{-2} + o(1)$$

as  $c \rightarrow \infty$ , where  $\nu$  is as in (4.5).

**PROOF.** By Wald's lemma we have  $\mu E\{t_c\} = cE\{t_c^{\alpha}\} + E\{R_c\}$ . This may also be written as

$$\begin{aligned} E\{t_c - \lambda\} &= \lambda^{1-\alpha}E\{t_c^{\alpha} - \lambda^{\alpha}\} + \mu^{-1}E\{R_c\} \\ &= \lambda^{1-\alpha}E\{\alpha\lambda^{\alpha-1}(t_c - \lambda) - \frac{1}{2}\alpha(1 - \alpha)\lambda_1^{\alpha-2}(t_c - \lambda)^2\} + \mu^{-1}E\{R_c\}, \end{aligned}$$

where  $|\lambda_1 - \lambda| \leq |t_c - \lambda|$ . Equivalently,

$$E\{t_c - \lambda\} = \beta\mu^{-1}E\{R_c\} - \frac{1}{2}\alpha E\{(\lambda \cdot \lambda_1^{-1})^{2-\alpha} t_c^{*2}\}.$$

Now  $E\{R_c\} \rightarrow \nu$ ,  $|\lambda/\lambda_1| \rightarrow 1$ , and  $t_c^{*2}$  is uniformly integrable, so it will suffice to



show that  $|\lambda/\lambda_1|$  is bounded. It is clear that  $|\lambda/\lambda_1|$  is bounded on  $\{t_c \geq \frac{1}{2}\lambda\}$ . Moreover, on  $\{t_c < \frac{1}{2}\lambda\}$ , we have

$$\frac{1}{2}\alpha(1 - \alpha)\lambda_1^{\alpha-2} = \{\lambda^\alpha - t_c^\alpha + \alpha\lambda^{\alpha-1}(t_c - \lambda)\}(t_c - \lambda)^{-2} \leq 4\lambda^{\alpha-2}.$$

**6. Proofs.** In this section we will present the proofs of Lemmas 2.2 and 3.2. We suppose throughout this section that  $X_1, X_2, \dots$  are i.i.d. with a positive mean  $\mu$  and a finite positive variance  $\sigma^2$ .

In order to prove Lemma 2.2, we need some auxiliary results. The first of these is an invariance principle for conditional distributions which may be of minor interest in its own right. In the notation of Section 2, let  $T_{kj} = Y_{k1} + \dots + Y_{kj}$ ,  $j < k$ , where  $Y_{ki}$  are the coordinate functions on  $R^{k-1}$ . Further, let  $Z_k$  be a continuous function on  $[0, 1]$  for which  $Z_k(0) = 0 = Z_k(1)$  and

$$Z_k(jk^{-1}) = \sigma^{-1}k^{-\frac{1}{2}}T_{kj}, \quad j = 1, \dots, k - 1,$$

and  $Z_k$  is linear on each of the intervals  $[(j - 1)k^{-1}, jk^{-1}]$ ,  $j = 1, \dots, k - 1$ . It is easily seen that  $Z_k$  is a measurable mapping from  $R^{k-1}$  into  $C[0, 1]$ , when both spaces are endowed with their Borel sigma algebras. Let  $Q_{k,z}$  be as in (2.3), let  $Q_0$  denote the distribution of a Brownian bridge in  $C[0, 1]$ , and let  $d$  denote the Prokhorov distance between probability measures on the Borel sets of  $C[0, 1]$  (Billingsley (1968), pages 237-238).

LEMMA 6.1. *If Condition C is satisfied and if  $I$  is any compact interval, then  $\lim d(Q_{k,z} \circ Z_k^{-1}, Q_0) = 0$  uniformly with respect to  $z \in I$  as  $k \rightarrow \infty$ .*

PROOF. It will suffice to show that if  $z_k \rightarrow z \in R$ , then  $Q_k^* = Q_{k,z_k} \circ Z_k^{-1}$  converges weakly to  $Q_0$ . That the finite dimensional distributions of  $Q_k^*$  converge (strongly) to those of  $Q_0$  follows from the local limit theorem for densities (Feller (1966), pages 489-490) by an argument which is similar to that given in Steck (1957).

To show that  $Q_k^*$ ,  $k \geq 1$ , is tight, it will suffice to show that for every  $\epsilon > 0$ , there is an  $n = n_\epsilon$  for which

$$(6.1) \quad \sup_{k \geq n_\epsilon} Q_{k,z_k}\{|T_{kj}| > \epsilon k^{\frac{1}{2}}, \text{ for some } j \leq kd\} = o(\delta)$$

as  $\delta \rightarrow 0$  (Billingsley (1968), page 56). Given  $\epsilon > 0$ , let  $\delta > 0$  and let  $m = [k\delta]$  be the greatest integer which is less than or equal to  $k\delta$ . Further, let  $A_k$  be the set of  $y \in R^m$  for which  $|T_{m+1,j}| > \epsilon k^{\frac{1}{2}}$  for some  $j \leq m$ , so that the left side of (6.1) is simply  $Q_{k,z_k} \circ \pi_{km}^{-1}(A_k)$ . If  $h_{km}$  is as in (2.4), then, as in the proof of Lemma 2.1, there is a constant  $B$  for which

$$h_{km}(y_1, \dots, y_m | z_k) \leq B \prod_{i=1}^m f(y_i + \mu + \sigma k^{-\frac{1}{2}}z_k)$$

for all  $y = (y_1, \dots, y_m) \in R^m$  and all sufficiently large  $k$ . It then follows easily that

$$\begin{aligned} Q_{k,z_k} \circ \pi_{km}^{-1}(A_k) &\leq BP\{|S_j - j\mu - j\sigma k^{-\frac{1}{2}}z_k| > \epsilon k^{\frac{1}{2}}, \text{ for some } j \leq m\} \\ &\leq BP\{|S_j - j\mu| > \frac{1}{2}\epsilon k^{\frac{1}{2}}, \text{ for some } j \leq m\} \end{aligned}$$

for  $k$  sufficiently large and  $\delta$  sufficiently small. Relation (6.1) now follows from standard arguments (Billingsley (1968), page 69).

LEMMA 6.2. *Suppose that Condition C is satisfied. Let  $F_j \subset \mathbf{B}(R^{k-1})$  be the smallest sigma algebra with respect to which  $Y_{ki}$ ,  $i = j, \dots, k-1$ , are measurable. If  $I$  is any compact interval, then  $\{j^{-1}T_{kj}, F_j, j = 1, \dots, k-1\}$  is a reverse martingale on the probability space  $(R^{k-1}, \mathbf{B}(R^{k-1}), Q_{k,z})$  for all  $z \in I$  for  $k$  sufficiently large.*

Lemma 6.2 follows easily from considerations of symmetry.

We will now prove Lemma 2.2, which asserts that if Condition C is satisfied, and if  $k = k_c \sim \lambda$  as  $c \rightarrow \infty$ , then  $\lim_{m \rightarrow \infty} \limsup_{c \rightarrow \infty} Q_{k,z}\{T_{kj} \geq cV(j) - jk^{-1}[cV(k+1) - y], \text{ for some } j \leq k - m\} = 0$  uniformly in  $z \in I$  for any compact interval  $I$  and any  $y \in R$ . We divide the range  $j \leq k - m$  into  $j \leq k\delta$ ,  $k\delta < j < k(1 - \varepsilon)$ , and  $k(1 - \varepsilon) \leq j \leq k - m$ , where  $0 < \delta < \varepsilon < \frac{1}{2}$ . Since  $v(j) = V(j) - jk^{-1}V(k+1)$  is concave on  $k\delta \leq j \leq k(1 - \varepsilon)$  for  $c$  sufficiently large, we have  $\min\{v(j) : k\delta \leq j \leq k(1 - \varepsilon)\} \geq \min\{v(k\delta), v(k(1 - \varepsilon))\} = m(c)$ , say, for  $c$  sufficiently large. It is easily seen that  $m(c) \sim v(k\delta) \sim V(k)[\delta^\alpha - \delta]$  as  $c \rightarrow \infty$  for  $\delta$  sufficiently small. In particular,  $ck^{-1}m(c) \rightarrow \infty$  as  $c \rightarrow \infty$ , so that

$$(6.2) \quad \lim Q_{k,z}\{T_{kj} \geq cv(j), \text{ for some } j, k\delta \leq j \leq k(1 - \varepsilon)\} = 0$$

uniformly with respect to  $z \in I$  as  $c \rightarrow \infty$  by Lemma 6.1.

From (1.2) it is easily seen that  $j^{-1}V(j)$  is nonincreasing in  $j$  for  $j$  sufficiently large. Let  $u(j) = j^{-1}V(j) - k^{-1}V(k)$ ; then  $\min\{u(j) : j \leq k\delta\} \sim u(k\delta) \sim k^{-1}V(k)[\delta^{\alpha-1} - 1]$  as  $c \rightarrow \infty$ . Thus, for  $c$  sufficiently large and  $\delta$  sufficiently small

$$(6.3) \quad \begin{aligned} Q_{k,z}\{T_{kj} \geq cV(j) - jk^{-1}[cV(k+1) - y], \text{ for some } j \leq k\delta\} \\ \leq Q_{k,z}\{j^{-1}T_{kj} \geq \frac{1}{2}cV(k)\delta^{\alpha-1}k^{-1}, \text{ for some } j \leq k\delta\} \\ \leq \frac{2k\delta^{1-\alpha}}{cV(k)} \int |y|h_{k1}(y) dy \leq 4\mu^{-1}\delta^{1-\alpha}E|X_1 - \mu| \end{aligned}$$

by Lemmas 2.1 and 6.2 and the martingale inequality.

Finally, if  $\varepsilon$  is sufficiently small and  $m$  is sufficiently large, then

$$cV(k-j) - (k-j)k^{-1}[cV(k+1) - y] \geq \frac{1}{2}j\mu(1 - \alpha)$$

for  $m \leq j \leq k\varepsilon$  and  $c$  sufficiently large. It now follows from symmetry, Lemma 6.2, and the martingale inequality that

$$(6.4) \quad \begin{aligned} Q_{k,z}\{T_{kj} \geq cV(j) - jk^{-1}[cV(k+1) - y], \\ \text{for some } j, k(1 - \varepsilon) \leq j \leq k - m\} \\ \leq Q_{k,z}\{-T_{kj} \geq \frac{1}{2}j\mu(1 - \alpha), \text{ for some } j, m \leq j \leq k\varepsilon\} \\ \leq 2\beta\mu^{-1} \int |m^{-1}T_{km}| dQ_{k,z} \end{aligned}$$

for all  $z \in I$ , if  $c$  is sufficiently large (and  $m$  is sufficiently large and  $\varepsilon$  is sufficiently small). By Lemma 2.1, the right side of (6.4) tends to zero as  $c \rightarrow \infty$  and

$m \rightarrow \infty$  (in that order). Thus, Lemma 2.2 follows from (6.2), (6.3), and (6.4) by letting  $c \rightarrow \infty$ ,  $\delta \rightarrow 0$ , and  $m \rightarrow \infty$  (in that order).

To prove Lemma 3.2 we will also need some preparation. We will use the fact that if  $E|X_1^3| < \infty$ , then there is a constant  $B$  for which

$$(6.5) \quad |g_k^*(z) - \phi(z)| \leq Bk^{-\frac{1}{2}}(1 + |z^2|)^{-1}$$

for all  $z \in R$  for  $k$  sufficiently large, where  $g_k^*$  denotes the density of  $S_k^*$ . See Petrov (1964). In particular, it follows from (6.5) that  $g^* = \sup \{g_k^* : k \geq 1\}$  is integrable.

LEMMA 6.3. *Let  $y_k = \sigma^{-1}k^{-\frac{1}{2}}[cV(k + 1) - k\mu]$ ,  $k \geq 1$ , as in (2.2c). Then there is a sequence  $\varepsilon_k$ ,  $k \geq 1$ , for which  $\varepsilon_k \rightarrow 0$  as  $k \rightarrow \infty$  and*

$$y_k - y_{k+1} \geq \frac{1}{(k + 1)^{\frac{1}{2}}} \left\{ \frac{\mu}{2\sigma} - \left(\alpha - \frac{1}{2} - 2\varepsilon_k\right) \frac{cV(k + 1)}{\sigma(k + 1)} \right\} \quad k \geq 1.$$

PROOF.

$$\begin{aligned} y_k - y_{k+1} &= c\sigma^{-1}\{k^{-\frac{1}{2}}(k + 1)^\alpha - (k + 1)^{-\frac{1}{2}}(k + 2)^\alpha\}L(k + 1) \\ &\quad + c\sigma^{-1}(k + 1)^{-\frac{1}{2}}(k + 2)^\alpha\{L(k + 1) - L(k + 2)\} \\ &\quad + \mu\sigma^{-1}\{(k + 1)^{\frac{1}{2}} - k^{\frac{1}{2}}\} \\ &= \Delta_1 + \Delta_2 + \Delta_3, \end{aligned} \quad \text{say.}$$

Simple convexity arguments show that  $-\Delta_1 \leq c\sigma^{-1}V(k + 1)(k + 1)^{\alpha-\frac{3}{2}}[\alpha - \frac{1}{2}]$  and that  $\Delta_3 \geq \frac{1}{2}\mu\sigma^{-1}(k + 1)^{-\frac{1}{2}}$  for  $k \geq 1$ . By (1.2) there is a sequence  $\varepsilon_k$ ,  $k \geq 1$  for which  $\varepsilon_k \rightarrow 0$  as  $k \rightarrow \infty$  and

$$|\Delta_2| \leq c\sigma^{-1}(k + 1)^{-\frac{3}{2}}(k + 2)^\alpha L(k + 1)\varepsilon_k \leq 2c\sigma^{-1}(k + 1)^{-\frac{3}{2}}V(k + 1)\varepsilon_k$$

for  $k \geq 1$ . Lemma 6.3 follows easily.

We will now give the proof of Lemma 3.2, which asserts that if Condition C is satisfied and if  $E|X_1^3|$  is finite, then  $\lim_{I \uparrow R} \limsup_{c \rightarrow \infty} W_c^1(b) - W_c^1(a) = 0$  for  $0 \leq a < b < \infty$  and that  $W_c^1(y) \leq B(1 + y)$  for all  $y \geq 0$  and all  $c > 0$  for some constant  $B$  (see (3.1)). Let

$$\left(\alpha - \frac{1}{2}\right)^+ < \gamma < \frac{1}{2} \quad \text{and} \quad 2\delta = 1 - \gamma^{-1}\left(\alpha - \frac{1}{2}\right)^+.$$

Then, it will suffice to prove the first assertion of the lemma in the special case that  $b - a \leq \delta\mu/4\sigma$ , for the general case may be reduced to a finite sum of terms to which the special case applies.

Let  $J_{11} = \{k \in J_1 : k \leq 2\gamma\lambda\}$  and  $J_{12} = J_1 - J_{11}$ . It follows easily from the fact that  $k^{-1}V(k)$  is eventually nonincreasing that  $cV(k) \leq k\mu/2\gamma$  for  $k \geq 2\gamma\lambda$  and  $c$  sufficiently large. Thus

$$y_k - y_{k+1} \geq \frac{\mu}{2\sigma} \left[1 - \gamma^{-1}\left(\alpha - \frac{1}{2} - 2\varepsilon_k\right)\right](k + 1)^{-\frac{1}{2}} \geq \frac{\delta\mu}{2\sigma k^{\frac{1}{2}}}$$

for  $k \geq 2\gamma\lambda$  and  $c$  sufficiently large by Lemma 6.3. Thus for  $c$  sufficiently large

and  $b - a \leq \delta\mu/4\sigma$ , the intervals  $I_k = (y_k - \sigma^{-1}k^{-\frac{1}{2}}b, y_k - \sigma^{-1}k^{-\frac{1}{2}}a]$ ,  $k \in J_{12}$ , are disjoint. Consequently (with the obvious notational conventions)

$$(6.6) \quad W_c^{12}(b) - W_c^{12}(a) \leq \sum_{k \in J_{12}} \int_{I_k} g_k^*(z) dz \leq \int_{(I', \pm 1)} g^*(z) dz,$$

which is independent of  $a, b$ , and  $c$  and tends to zero as  $I \uparrow R$ .

To estimate  $W^{11}$  let  $k_0$  be so large that  $k^{-1}V(k)$  is nonincreasing for  $k \geq k_0$ . If  $\alpha < \alpha' < 1$ , it is then the case that  $ck^{-1}V(k) \geq \mu(2\gamma)^{\alpha'-1}$  for  $k_0 \leq k \leq 2\gamma\lambda + 1$  for  $c$  sufficiently large. Consequently,

$$cV(k + 1) - k\mu \geq k\mu\{(2\gamma)^{\alpha'-1} - 1\} = k\mu\gamma^*, \quad \text{say,}$$

for  $k_0 \leq k \leq 2\gamma\lambda + 1$  for  $c$  sufficiently large. Given  $b > 0$ , let

$$k_1 = k_1(b) = \max \left\{ \frac{2b}{\mu\gamma^*}, k_0 \right\}.$$

Then for  $k \geq k_1(b)$ , we have

$$(6.7) \quad W^{11}(b) \leq \sum_{j \leq 2\gamma\lambda} P\{S_j > cV(j + 1) - b\} \\ \leq \sum_{j \leq k} P\{S_j > cV(j + 1) - b\} + \sum_{j > k} \left( \frac{2}{j\mu\gamma^*} \right)^3 E|S_j - j\mu|^3$$

which tends to zero as  $c \rightarrow \infty$  and  $k \rightarrow \infty$  (in that order). Moreover, letting  $k = k_1(b)$  in (6.7), we find that  $W^{11}(b) \leq A(1 + b)$  for  $b \geq 0$ , where  $A$  is independent of  $b$  and  $c$ . Thus (6.6) and (6.7) combine to establish Lemma 3.2.

**7. Uniform integrability.** In this section we will give the proof of Theorem 5.1. We suppose throughout that  $X_1, X_2, \dots$  are i.i.d. with positive mean and finite positive variance and that  $V(x) = x^\alpha, x > 0$ .

**LEMMA 7.1.** *If either  $\alpha < \frac{1}{2}$  or  $E|X_1|^\gamma < \infty$  for some  $\gamma > \beta = 1/(1 - \alpha)$ , then  $\lim \lambda P\{t_c \leq \alpha\lambda\} = 0$  as  $c \rightarrow \infty$ .*

**PROOF.** We use the easily verified fact that

$$P\{\max_{k \leq n} (S_k - k\mu) > y\} \leq KP\{S_n - n\mu > y\}$$

for  $n \geq 1$  and  $y > 0$ , where  $K^{-1} = \inf \{P\{S_k - k\mu > 0\} : k \geq 1\} > 0$  (cf. Loève (1963), pages 247–248). Let  $n \geq 2$  be an integer and let  $L = L(n, c)$  be an integer for which  $n^{L-1} < \alpha\lambda \leq n^L$ . Further, let  $\delta = 1 - \alpha^{1-\alpha}$ . Then

$$(7.1) \quad P\{t_c \leq \alpha\lambda\} \leq \sum_{i=1}^L P\{S_k - k\mu > \delta cn^{\alpha(i-1)}, \text{ for some } k \leq n^i\} \\ \leq K \sum_{i=1}^L P\{S_{n^i} - n^i\mu > \delta cn^{\alpha(i-1)}\}.$$

Let  $A_k$  be the event that  $|S_k^*| > \delta ck^{-\frac{1}{2}}$  and let

$$r(c) = \max_{k \leq \lambda} \int_{A_k} |S_k^*|^2 dP.$$

If  $\alpha < \frac{1}{2}$ , then  $r(c) \rightarrow 0$  as  $c \rightarrow \infty$  by the uniform integrability of  $|S_k^*|^2, k \geq 1$ , and the right side of (7.1) does not exceed

$$K(\delta c)^{-2} \sum_{i=1}^L n^{i-2\alpha(i-1)} \cdot r(c) = c^{-2}r(c)O\{n^{(1-2\alpha)L}\} = o(\lambda^{-1})$$

as  $c \rightarrow \infty$ . Similarly, if  $\alpha \geq \frac{1}{2}$  and if  $E|X_1|^\gamma$  is finite for some  $\gamma > \beta$ , then the right side of (7.1) does not exceed

$$K(\delta c)^{-\gamma} \sum_{i=1}^L n^{-\alpha\gamma(i-1)} E|S_{n^i} - n^i\mu|^\gamma = O\{Lc^{-\gamma}\} = o(\lambda^{-1})$$

as  $c \rightarrow \infty$ .

LEMMA 7.2. *If  $E|X_1|^\gamma < \infty$  for some  $\gamma > 4$ , then  $\lim_{\{t_c > 4\lambda\}} t_c^2 dP = 0$  as  $c \rightarrow \infty$ .*

PROOF. The integral in question does not exceed

$$32\lambda^2 P\{t_c > 4\lambda\} + \sum_{j \geq 4\lambda} 8jP\{t_c > j\}.$$

Letting  $\delta = 1 - 4^{\alpha-1}$ , we find that  $P\{t_c > j\} \leq P\{S_j - j\mu \leq -j\delta\mu\} \leq Bj^{-\frac{1}{2}}$  for all  $j \geq 4\lambda$  for some constant  $B$ . The result follows easily.

We will now prove Theorem 5.1, which asserts that if  $V(x) = x^\alpha$  for  $x > 0$  and if  $E|X_1|^\gamma$  is finite for some  $\gamma > \max\{4, \beta\}$ , then  $t_c^{**}$ ,  $c > 0$ , are uniformly integrable. By Lemmas 7.1 and 7.2, it will suffice to show that there is a function  $J$  for which  $yJ(y)$  is integrable over  $(0, \infty)$  and

$$P\{|t_c^*| > y, \alpha\lambda < t_c < 4\lambda\} \leq J(y)$$

for all  $y > 0$  and all  $c > 0$ . Let

$$J_0(y) = \sup_{j \geq 1} P\{|S_j^*| > y\}$$

for  $y > 0$ . It is easily seen that  $yJ_0(y)$  is integrable. Given  $y > 0$ , let  $n$  be the greatest integer which is  $< \lambda - y\lambda^{\frac{1}{2}}$ . Then since  $cx^\alpha - \mu x$  is decreasing  $x \geq \alpha^\beta \lambda$ , we have

$$(7.2) \quad P\{t_c > \alpha\lambda, t_c^* < -y\} \leq P\{S_j - j\mu > cj^\alpha - j\mu, \text{ for some } j \leq n\} \\ \leq KP\{S_n - n\mu > cn^\alpha - n\mu\}$$

for  $n \geq \alpha\lambda$ ; and, of course, the left side of (7.2) is zero if  $n < \alpha\lambda$ . Let  $h(x) = cx^\alpha - \mu x$  for  $x > 0$ . Then  $h$  is concave and  $h(\lambda) = 0$ , so that

$$h(n) \geq h'(n)(n - \lambda) \geq \mu[1 - \alpha\lambda^{1-\alpha}n^{\alpha-1}]y\lambda^{\frac{1}{2}} \geq \delta\mu y n^{\frac{1}{2}},$$

where  $\delta = 1 - \alpha^\alpha > 0$ . It follows that

$$P\{t_c > \alpha\lambda, t_c^* < -y\} \leq KJ_0(\delta\mu y)$$

for  $y > 0$ . A similar, somewhat simpler, argument will show that  $P\{t_c \leq 4\lambda, t_c^* > y\} \leq J_0[\frac{1}{2}(1 - \alpha)y]$  for  $y > 0$  to complete the proof of Theorem 5.1.

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