

**ON FUNCTIONS OF SEQUENCES OF INDEPENDENT CHANCE VECTORS  
WITH APPLICATIONS TO THE PROBLEM OF THE  
"RANDOM WALK" IN  $k$  DIMENSIONS**

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**1. Summary.** Consider a sequence  $\{x_i\}$  of independent chance vectors in  $k$  dimensions with identical distributions, and a sequence of mutually exclusive events  $S_1, S_2, \dots$ , such that  $S_i$  depends only on the first  $i$  vectors and  $\Sigma P(S_i) = 1$ . Let  $\varphi_i$  be a real or complex function of the first  $i$  vectors in the sequence satisfying conditions: (1)  $E(\varphi_i) = 0$  and (2)  $E(\varphi_j | X_1, \dots, X_i) = \varphi_i$  for  $j \geq i$ . Let  $\varphi = \varphi_i$  and  $n = i$  when  $S_i$  occurs. A general theorem is proved which gives the conditions  $\varphi_i$  must satisfy such that  $E\varphi = 0$ . This theorem generalizes some of the important results obtained by Wald for  $k = 1$ . A method is also given for obtaining the distribution of  $\varphi$  and  $n$  in the problem of the "random walk" in  $k$  dimensions for the case in which the components of the vector take on a finite number of integral values.

**2. A basic theorem.**

2.1 Let  $\{X_i\} = \{(X_{1i}, X_{2i}, \dots, X_{ki})\}$  be a sequence of independent  $k$ -dimensional chance variables with identical distributions. Let  $S_1, S_2, S_3, \dots$ , be mutually exclusive events such that (1)  $S_i$  depends only on  $X_1, X_2, \dots, X_i$ , and (2)  $\Sigma P(S_i) = 1$ . Let  $\varphi_i(X_1, X_2, \dots, X_i)$  be a sequence of real or complex variables satisfying the following two conditions:

*Condition 1:*  $E(\varphi_i) = 0$  for all  $i$ .

*Condition 2:*  $E(\varphi_j | X_1, X_2, \dots, X_i) = \varphi_i$  for all  $j \geq i$ , where  $E(\varphi_j | X_1, X_2, \dots, X_i)$  stands for the expected value of  $\varphi_j$  under the condition that  $X_1, X_2, \dots, X_i$  are held constant.<sup>1</sup> Define  $\varphi_i = \varphi$  and  $n = i$  when the event  $S_i$  occurs. We shall assume that  $E(n)$  is finite.

A problem of central importance in sequential theory may be formulated as follows: What conditions must  $\varphi_i$  satisfy so that  $E(\varphi)$  exists and equals zero? We shall prove the following:

**THEOREM 2.1.** *If there exists a function  $f(x_1, x_2, \dots, x_k) \geq 0$  such that (a)  $E[f(X_i)]$  is finite and (b)  $|\varphi_i| \leq \sum_{d=1}^i f(X_d)$  when  $n \geq i$ , then  $E(\varphi)$  exists and equals zero.*

Before proceeding to the proof, we consider two consequences of this theorem.

I. Assume that  $E(X_{ri}) = a_r$ . Let  $\varphi_i = \sum_{j=1}^i (X_{rj} - a_r)$ . It is easily verified that  $\varphi_i$  satisfies conditions 1 and 2. We set  $f(x_1, x_2, \dots, x_k) = |x_r$

<sup>1</sup> Chance variables  $\varphi_i$  satisfying condition 2 have been extensively studied by P. Levy [1] and J. L. Doob [2].

$- a_r |$ . Then Theorem 2.1 is applicable and we get  $E\varphi = 0$ . Now  $\varphi = W_r - na_r$  where  $W_r = \sum_{i=1}^n X_{ri}$ . Hence we have

$$(2.11) \quad E(W_r) = a_r E(n).$$

The relationship (2.11) has been proved for  $k = 1$  by Wald [3] and subsequently under somewhat more generalized conditions, by one of the authors [4].

II. Let  $t_1, t_2, \dots, t_k$  be any real or complex numbers for which  $Ee^{\sum_{r=1}^k t_r X_{ri}} = a$  is finite and  $|a| \geq 1$ . We assume that there exists a positive constant  $M$  such that

$$(2.12) \quad \left| \sum_{i=1}^m X_{ri} \right| \leq M, \quad r = 1, 2, \dots, k,$$

when  $n > m$ . Let

$$(2.13) \quad \varphi_i = a^{-i} e^{\sum_{j=1}^i \sum_{r=1}^k t_r X_{rj}} - 1$$

so that

$$(2.14) \quad \varphi = a^{-n} e^{\sum_{r=1}^k t_r W_r} - 1$$

where  $W_r$  is defined as above. It is easy to show that  $\varphi_i$  satisfies conditions 1 and 2. Now, in view of (2.12), when  $n \geq i$

$$(2.15) \quad |\varphi_i| \leq |a|^{-i} e^{M \sum_{r=1}^k |t_r|} e^{\sum_{r=1}^k \tau_r X_{ri}} + 1 \leq 1 + R e^{\sum_{r=1}^k \tau_r X_{ri}}$$

where  $\tau_j$  is the real part of  $t_j$  and  $R = e^{M \sum_{r=1}^k |t_r|}$  is a fixed positive constant. Then, letting

$$(2.16) \quad f(x_1, x_2, \dots, x_k) = 1 + R e^{\sum_{r=1}^k \tau_r X_{ri}}$$

we may apply Theorem 2.1 and obtain

$$(2.17) \quad E\left(a^{-n} e^{\sum_{r=1}^k t_r W_r}\right) = 1$$

which is a generalization of the Fundamental Identity proved by Wald [5] for the case  $k = 1$ .

PROOF OF THEOREM 2.1. Assume  $\varphi_i$  is real. Define chance variables  $N_m$  inductively as follows:  $N_0 = 0$ . Assuming  $N_0, \dots, N_m$  defined, define  $N_{m+1} = N_m + n(X_{N_m+1}, X_{N_m+2}, \dots)$ . Also let  $n_m = N_m - N_{m-1}$  and  $y_m = f(X_{N_{m-1}+1}) + \dots + f(X_{N_m})$ . It can be shown by induction that  $N_m$  is defined for all  $m$  with probability one, and that  $\{n_m\}, \{y_m\}$  are sequences of independent chance variables with identical distributions. Clearly  $n_1 = n$ .

The Strong Law of Large Numbers asserts that if  $z_1, z_2, \dots$  are independent chance variables with identical distribution, then  $\lim_{m \rightarrow \infty} \frac{z_1 + z_2 + \dots + z_m}{m} = c$  with probability one if and only if  $Ez_1$  exists and equals  $c$ .

It follows that, with probability one

$$(2.18) \quad \lim_{m \rightarrow \infty} \frac{f(X_1) + \cdots + f(X_m)}{m} = E[f(X_1)]$$

and

$$(2.19) \quad \lim_{m \rightarrow \infty} \frac{n_1 + \cdots + n_m}{m} = \lim_{m \rightarrow \infty} \frac{N_m}{m} = E(n).$$

Since  $\frac{y_1 + \cdots + y_m}{n_1 + \cdots + n_m} = \frac{y_1 + \cdots + y_m}{N_m}$  is a subsequence of  $\frac{f(X_1) + \cdots + f(X_m)}{m}$  we have with probability one,

$$(2.20) \quad \lim_{m \rightarrow \infty} \frac{y_1 + \cdots + y_m}{N_m} = E[f(X_1)]$$

so that

$$(2.21) \quad \lim_{m \rightarrow \infty} \frac{y_1 + \cdots + y_m}{m} = E[f(X_1)]E(n).$$

Consequently,  $E(y_1)$  exists and equals  $Ef(X_1)E(n)$ . Since  $|\varphi| \leq y_1$ ,  $E(\varphi)$  exists. Also using conditions (2) and (b) which were imposed on  $\varphi_i$  we have

$$(2.22) \quad \begin{aligned} \left| \int_{s_1 + \cdots + s_i} \varphi \, dp \right| &= \left| \sum_{j=1}^i \int_{s_j} \varphi_j \, dp \right| = \left| \sum_{j=1}^i \int_{s_j} \varphi_i \, dp \right| \\ &= \left| - \int_{n > i} \varphi_i \, dp \right| = \left| \sum_{i > j} \int_{s_j} \varphi_i \, dp \right| \\ &\leq \sum_{i > j} \int_{s_j} |\varphi_i| \, dp \leq \sum_{i > j} \int_{s_j} y_1 \, dp \end{aligned}$$

which approaches zero as  $i \rightarrow \infty$ . This completes the proof.

If  $\varphi_j$  is a complex valued function, Theorem 2.1 still holds. For writing  $\varphi_j = g_j + ih_j$  then Condition 2 becomes  $E(g_p + ih_p | X_1, \cdots, X_j) = g_j + ih_j$  when  $p \geq j$ . Hence

$$(2.23) \quad E(g_p | X_1, \cdots, X_j) = g_j$$

and

$$(2.24) \quad E(h_p | X_1, \cdots, X_j) = h_j$$

when  $p \geq j$ . Since  $|g_j| \leq |\varphi_j|$  and  $|h_j| \leq |\varphi_j|$  and  $\varphi_j$  satisfies condition (b) we may apply Theorem 2.1 and get

$$(2.25) \quad Eg = E(h) = 0.$$

Hence  $E\varphi = 0$ .

**3. Applications to the problem of the random walk in  $k$  dimensions<sup>2</sup>**

3.1. *A theorem concerning decision points.* Let  $\{X_j\} = \{(X_{1j}, \dots, X_{kj})\}$  be a sequence of  $k$ -dimensional chance vectors with identical distributions. We assume that  $X_{ji}$  ( $j = 1, 2, \dots, k$ ), take on a finite number of integral values ranging from  $-r_j$  to  $m_j$  inclusive, where  $r_j$  and  $m_j$  are positive integers. We remark that any distribution can be approximated to any degree of accuracy by the distribution of a variate whose values are integral multiples of a constant  $d$ , which can be taken as the unit of measurement.

Let  $P_{u_1 u_2 \dots u_k}$  be the probability that  $X_i = (u_1, u_2, \dots, u_k)$ . We define  $W_{pi} = \sum_{j=1}^i X_{pj}$  and set  $U_i = (W_{1i}, W_{2i}, \dots, W_{ki})$ . Then  $\{U_i\}$  represents a sequence of points with integral coordinates in a  $k$ -dimensional space  $S_k = \{(y_1, y_2, \dots, y_k)\}$ . Let  $R$  be an arbitrary bounded region in  $S_k$ . We shall assume, without loss of generality, that the origin is an interior point of  $R$ . We now define a random variable  $n$  as the smallest subscript  $i$  of the sequence  $\{U_j\}$  for which  $W_i$  is either a boundary point or an exterior point of  $R$ . We set  $U_n = W = (W_1, W_2, \dots, W_k)$  and designate  $W$  as a decision point of  $R$ . Clearly, the number of decision points is finite.

The random variables  $n$  and  $W$  can be interpreted as follows: Consider a point  $Q$  which at the time  $t = 0$  is at the origin. At successive intervals of time  $t = 1, 2, \dots$ , the point  $Q$  moves with integral components in  $S_k$  the direction and distance of the motion being determined by chance. The point comes to rest as soon as, but not before it either reaches the boundary of  $R$  or falls outside of  $R$ . Let  $U_t$  be the co-ordinates of the point  $Q$  at time  $t$ . Then  $n$  represents the length of time it takes  $Q$  to come to rest, and  $W$  represents a possible resting point.<sup>3</sup>

We shall be concerned with the problem of finding the probability distribution of  $n$  and  $W$ . These will obviously depend on the shape of the region  $R$ . In what follows we shall restrict ourselves to the class of regions  $R$  which have the property that the intersection of any line parallel to the axes with  $R$  is an open interval. In view of the fact that  $W$  has integral coordinates, we can without any loss of generality, replace this class of regions by an equivalent class which are bounded by simple polygonal closed surfaces whose vertices have integral coordinates and whose sides are parallel to the planes  $y_j = 0$ . In the subsequent discussion we assume that the regions  $R$  are of this type.

Let

$$(3.10) \quad \text{l.u.b. } [y_i]_{(y_1, y_2, \dots, y_k) \in R}$$

<sup>2</sup> What follows is a generalization of a method previously employed by one of the authors [6] for the case  $k = 1$ .

<sup>3</sup> That  $Q$  will reach a resting point eventually can be asserted with probability one. See A. Wald [5], Lemma 1.

and

$$(3.11) \quad -b_i = \underset{y_i}{\text{g.l.b.}} [(y_1, y_2, \dots, y_k) \in R]$$

then  $a_i$  and  $b_i$  are positive integers.

We now prove the following:

LEMMA 3.1. *For the given sequence of chance vectors  $\{X_i\}$  and the given region  $R$ , the number of possible decision points  $N_R$  is given by*

$$(3.12) \quad N_R = \prod_{j=1}^k (a_j + b_j + r_j + m_j - 1) - \prod_{j=1}^k (a_j + b_j - 1).$$

PROOF: We shall first prove this theorem for a rectangular region  $R = R_1$  where  $R_1$  is defined by  $-b_i < y_i < a_i$ , ( $i = 1, 2, \dots, k$ ) and then generalize the proof to any region of the class specified.

Let  $R_2$  be a closed rectangular region defined by  $-(b_i + r_i - 1) \leq y_i \leq (a_i + m_i - 1)$ . Then  $R_2 \supseteq R_1$ . Let  $S = R_2 - R_1$ . It is clear that every integral point of  $S$  is a possible decision point. Moreover, no point exterior to  $R_2$  is a possible decision point. For assume, for example, that there exists a point  $W = (W_1, W_2, \dots, W_k)$  which is an exterior point of  $R_2$ . Then at least one of its coordinates, say  $W_j$ , has the property that  $W_j > a_j + m_j - 1$  or  $W_j < -(b_j + r_j - 1)$ . But since  $-(b_j - 1) \leq W_{j,n-1} \leq a_j - 1$ , it must follow that  $X_{j,n}$  took on a value greater than  $m_j$  or less than  $-r_j$  which is contrary to assumption. Now the total number of integral points contained in  $R_1$  is  $\prod_{j=1}^k (a_j + b_j + r_j + m_j - 1)$  and the total number of integral points in  $R_1$

which by assumption are not decision points, is  $\prod_{j=1}^k (a_j + b_j - 1)$ . Hence the Lemma is proved if  $R$  is a rectangular region.

Now, let  $R$  be any polygonal region of the type specified and let  $R_1$  be the corresponding rectangular region. Consider two randomly moving points  $Q$  and  $Q_1$ , each having coordinates  $W_t$  at time  $t$ . Let the decision points for  $Q$  be defined in terms of  $R$  and the decision points of  $Q_1$  in terms of  $R_1$ . We shall prove that the number of decision points for  $Q$  and  $Q_1$  are the same.

By assumption, every line parallel to the axes intersects  $R$  in an open interval. Moreover  $R_1 \supseteq R$ . Hence the sum of the areas of the segments which compose the boundary of  $R$  must equal the area of the boundary of  $R_1$ . The same must be true for the total number of integral points on the boundaries of the two regions. Thus, the theorem is true for  $r_j = m_j = 1$ , ( $j = 1, 2, \dots, k$ ). We assume that the theorem is true for  $r_j = r'_j$  and  $m_j = m'_j$  and prove that it must hold for  $m_u = m'_u + 1$  for a fixed but arbitrary  $u$ . Now it is obvious that if the range of  $X_{u,i}$  is increased by unity in the positive direction, the point  $Q$  can move an extra unit in the positive direction parallel to the  $y_u$  axes. Thus, the total number of additional decision points that  $Q$  gains by the unit increase in the range of  $X_{u,i}$  is identical with the total number that  $Q_1$  gains. This proves the theorem.

It is clear that the smallest rectangular region which includes all the decision points of  $W$  is  $R_2$ . We now prove the following:

**THEOREM 3.1.** *For any polygonal region  $R$  of the class previously specified, and for any random sequence  $\{X_i\}$  in which  $X_i$  takes on a finite number of integral values, the number of points in the rectangular region  $R_2$  which are not decision points is always equal to  $\prod_{j=1}^k (a_j + b_j - 1)$  where  $a_j + b_j$  are the dimensions of the rectangular region  $R_1$ .*

**PROOF:** This Theorem follows from Lemma 3.1 and the fact that the total number of integral points in  $R_2$  is  $\prod_{j=1}^k (a_j + b_j + r_j + m_j - 1)$ .

3.2. *The distribution of  $W$ .* Let  $\psi(t_1, \dots, t_k)$  be the joint generating function of  $X_{ui}$ , ( $u = 1, 2, \dots, k$ ), and  $\varphi(t_1, \dots, t_k)$  the joint generating function of  $W_j$  ( $j = 1, 2, \dots, k$ ). Then

$$(3.21) \quad \psi(t_1, \dots, t_k) = \sum_{u=-r_1}^{m_1} \dots \sum_{u_k=-r_k}^{m_k} P_{u_1 \dots u_k} t_1^{u_1} \dots t_k^{u_k}$$

$$(3.22) \quad \phi(t_1, \dots, t_k) = \sum_{v_1=-(b_1+r_1-1)}^{a_1+m_1-1} \dots \sum_{v_k=-(b_k+r_k-1)}^{a_k+m_k-1} \xi_{v_1 \dots v_k} t_1^{v_1} \dots t_k^{v_k}$$

where  $\xi_{v_1 \dots v_k}$  is the probability that  $W = (v_1, \dots, v_k)$ . In terms of the generating function  $\psi$  the Fundamental Identity (3.17) states that

$$(3.23) \quad Et_1^{w_1} \dots t_k^{w_k} [\psi(t_1, \dots, t_k)]^{-n} = 1$$

for all  $t_1, \dots, t_k$  for which  $|\psi(t_1, \dots, t_k)| \geq 1$ . Hence, it follows that for  $t_1, \dots, t_k$  for which  $\psi(t_1, \dots, t_k) = 1$ ,  $\varphi(t_1, \dots, t_k) = 1$ . Let

$$(3.24) \quad f(t_1, \dots, t_k) = t_1^{r_1} \dots t_k^{r_k} [\psi(t_1, \dots, t_k) - 1]$$

and

$$(3.25) \quad g(t_1, \dots, t_k) = t_1^{b_1+r_1-1} \dots t_k^{b_k+r_k-1} [\varphi(t_1, \dots, t_k) - 1].$$

Then  $f(t_1, \dots, t_k)$  is a polynomial of degree  $r_j + m_j$  in  $t_j$  and  $g(t_1, \dots, t_k)$  is a polynomial of degree  $(a_j + b_j + r_j + m_j - 2)$  in  $t_j$ .

We shall assume that  $f(t_1, \dots, t_k)$  is an irreducible polynomial. Then, since  $g(t_1, \dots, t_k)$  vanishes for all values of  $t_1, \dots, t_k$  for which  $f(t_1, \dots, t_k)$  vanishes, it follows<sup>4</sup> that  $f$  is a factor of  $g$ . That is

$$(3.26) \quad g(t_1, \dots, t_k) = f(t_1, \dots, t_k) \sum_{s_1=0}^{a_1+b_1-1} \dots \sum_{s_k=0}^{a_k+b_k-1} C_{s_1 \dots s_k} t_1^{s_1} \dots t_k^{s_k}$$

where the  $C_{s_1, \dots, s_k}$  are unknown. Equating coefficients on both sides of (3.26) we get

<sup>4</sup> See, for example, Bôcher [7], Theorem 7, Chapter 16.

$$(3.27) \quad \xi_{v_1-b_1-r_1+1 \dots v_k-b_k-r_k+1} = \sum_{u_1=0}^{v_1} \dots \sum_{u_k=0}^{v_k} (P_{u_1-r_1 \dots u_k-r_k} - \delta_{u_j r_j}) C_{v_1-r_1 \dots v_k-r_k} + \prod_{j=1}^k \delta_{v_j, b_j+r_j-1}$$

where  $\delta_{i,j}$  is the Kronecker delta. But by Theorem 3.1,  $\prod_{j=1}^k (a_j + b_j - 1)$  of the  $\xi_{v_1 \dots v_k}$  in  $\varphi(t_1, \dots, t_k)$  are zero since they correspond to values of  $W$  which are non-decision points. Hence  $\prod_{j=1}^k (a_j + b_j - 1)$  terms in (3.27) are zero with the exception of the term  $\xi_{b_1+r_1-1 \dots b_k+r_k-1}$  (corresponding to the non-decision point (0, 0)) which is  $-1$ . Hence, we have the required number of equations to solve for the unknown  $C$ 's and consequently for the  $\xi$ 's provided the determinant of the coefficients is different from zero.

As an illustration, let  $R = R_1$ , then the  $C$ 's are obtained by solving the set of linear equations

$$(3.28) \quad \sum_{u_1=0}^{v_1} \dots \sum_{u_k=0}^{v_k} \left( \prod_{j=1}^k \delta_{u_j r_j} - P_{u_1-r_1 \dots u_k-r_k} \right) C_{v_1-r_1 \dots v_k-r_k} = \prod_{j=1}^k \delta_{v_j, b_j+r_j-1}$$

where  $v_j$  takes on all integral values from  $r_j$  to  $a_j + b_j + r_j - 2$  inclusive.

3.3. *The distribution of n.* For any random variable  $U$ , let  $E_{v_1 \dots v_k} U$  stand for the expected value of  $U$  under the restriction that  $W = (v_1, v_2, \dots, v_k)$ . Let  $\varphi_1(t_1, \dots, t_k; \tau)$  be the joint generating function of  $W_1, W_2, \dots, W_k$ , and  $n$ . Then

$$(3.31) \quad \varphi_1(t_1, \dots, t_k; \tau) = \sum_{u_1} \dots \sum_{u_k} \xi_{u_1 \dots u_k} t_1^{u_1} \dots t_k^{u_k} E_{u_1 \dots u_k} \tau^n.$$

Let

$$(3.32) \quad \psi_1(t_1, \dots, t_k; \tau) = \tau \psi(t_1, t_2, \dots, t_k) - 1$$

where  $\psi(t_1, \dots, t_k)$  is the joint generating function of  $X_{1i}, \dots, X_{ki}$  and is given by (3.21) and let

$$(3.33) \quad \psi_2(t_1, \dots, t_k; \tau) = \varphi_1(t_1, \dots, t_k; \tau) - 1.$$

Then, if we fix  $\tau$  so that  $|\tau| \leq 1$ , we see by (3.23) that for all values of  $t_1, \dots, t_k$  for which  $\psi_1$  vanishes,  $\psi_2$  also vanishes. Let

$$(3.34) \quad f_1(t_1, \dots, t_k; \tau) = t_1^{r_1} \dots t_k^{r_k} \psi(t_1, \dots, t_k; \tau)$$

and

$$(3.35) \quad f_2(t_1, \dots, t_k; \tau) = t_1^{b_1+r_1-1} \dots t_k^{b_k+r_k-1} \psi_2(t_1, \dots, t_k; \tau).$$

Then for  $\tau$  fixed,  $f_1$  is a polynomial of degree  $r_j + m_j$  in  $t_j$  and  $f_2$  is a polynomial of degree  $a_j + b_j + r_j + m_j - 2$  in  $t_j$ . Since  $f_2$  vanishes for all values of  $t_1, \dots, t_k$  for which  $f_1$  vanishes then if  $f_1$  is irreducible,  $f_1$  will be a factor of  $f_2$ . That is  $f_2$  can then be written as

$$(3.36) \quad f_2(t_1, \dots, t_k; \tau) = f_1(t_1, \dots, t_k; \tau) \sum_{v_1=1}^{a_1+b_1-2} \cdots \sum_{v_k=1}^{a_k+b_k-2} d_{v_1 \cdots v_k} t_1^{v_1} \cdots t_k^{v_k}.$$

The rest of the argument is identical with that employed in section 3.3. The unknowns in the present case, however, are  $\xi_{v_1 \cdots v_k} E_{v_1 \cdots v_k} \tau^n$ . When  $\xi_{v_1 \cdots v_k} E_{v_1 \cdots v_k} \tau^n$  is expanded in a power series in  $\tau$ , the coefficient of  $\tau^m$  is the probability that  $W = (v_1, \dots, v_k)$  in exactly  $m$  steps. We shall, therefore, examine the validity of the expansion of the above function in the neighborhood of  $\tau = 0$ .

Let us first consider the rectangular region  $R = R_1$ . In this case the  $d$ 's are obtained from the equations

$$(3.37) \quad \sum_{u_1=1}^{v_1} \cdots \sum_{u_k=1}^{v_k} \left( \prod_{j=1}^k \delta_{u_j, r_j} - \tau P_{u_1-r_1 \cdots u_k-r_k} \right) d_{v_1-r_1 \cdots v_k-r_k} = \prod_{j=1}^k \delta_{v_j, b_j+r_j-1},$$

$$(v_j = r_j, \quad r_j + 1, \dots, \quad a_j + b_j + r_j - 2),$$

so that  $\xi_{v_1 \cdots v_k} E_{v_1 \cdots v_k} \tau^n$  will be given as a ratio of two polynomials in  $\tau$  the denominator of which will be the determinant of the coefficients of (3.37). But this determinant equals unity when  $\tau = 0$ . Hence the validity of the expansion is established for a rectangular region.

If  $R$  is not a rectangle, the value of the determinant of the equations in  $d$  will still be unity. This follows from the fact that the number of non-decision points in  $R_2$  is precisely the same as the number of non-decision points contained in  $R_1$ , hence by rearranging of the equations they can be made to assume the form (3.37).

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