SOME OPTIMUM WEIGHING DESIGNS

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1. Introduction and summary. Suppose we are given N objects to be weighed in N weighings with a chemical balance having no bias. Let

 $x_{ij} = 1$ if the jth object is placed in the left pan in the ith weighing,

=-1 if the *j*th object is placed in the right pan in the *i*th weighing,

= 0 if the jth object is not weighed in the ith weighing.

The Nth order matrix $X = (x_{ij})$ is known as the design matrix. Also let y_i be the result recorded in the *i*th weighing, ϵ_i be the error in this result and w_j be the true weight of the *j*th object, so that we have the N equations

$$x_{i1}w_1 + x_{i2}w_2 + \cdots + x_{iN}w_N = y_i + \epsilon_i, \quad i = 1, \dots, N.$$

We assume X to be a non-singular matrix. The method of Least Squares or theory of Linear Estimation gives the estimated weights (\hat{w}_i) by the equation

$$\hat{w} = (X'X)^{-1}X'Y,$$

where Y is the column vector of the observations and \hat{w} is the column vector of the estimated weights.

If σ^2 is the variance of each weighing, then

Var
$$(\hat{w}) = (X'X)^{-1}\sigma^2 = (c_{ij})\sigma^2$$
,

where (c_{ij}) is the inverse matrix of (X'X).

An expository article reviewing the work done in weighing designs is given by Banerjee [2].

Kishen [4] treats the reciprocal of the increase in variance resulting from the adoption of any design other than the most efficient design, with mean variance σ^2/N , as the efficiency of the design. This efficiency can be measured by

$$1/\sum_{i=1}^N c_{ii}.$$

Mood [5] gives an alternative definition for the best weighing design. In his view the best weighing design should give the smallest confidence region in the $\hat{w}_i(i=1,\dots,N)$ space for the estimates of the weights. Hence a design will be called best if the determinant of the matrix (c_{ij}) is minimised.

In this paper we follow Kishen's definition in obtaining the best weighing designs.

Hotelling [3] proved that for the best weighing design $c_{ii} = 1/N$ and $c_{ij} = 0$ ($i \neq j$). The weighing designs for which $c_{ii} = 1/N$ and $c_{ij} = 0$ are best in the sense of both Kishen and Mood. Later Mood proved that the above property is

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satisfied by Hadamard matrices. Plackett and Burman [6] have constructed Hadamard matrices, H_N , up to and including N=100, excepting N=92. It may be remarked here that a necessary condition for the existence of H_N is $N\equiv 0\pmod 4$, with the exception of N=2. It is not known whether this condition is sufficient or not.

In this paper, the best weighing designs are obtained in the cases (i) N is odd and (ii) $N \equiv 2 \pmod{4}$ subject to the conditions:

- i) The variances of the estimated weights are equal; and
- ii) The estimated weights are equally correlated.

The 2nd condition here is the same as that of Banerjee [1].

2. Some theorems relating to the best weighing designs. With the conditions made above (X'X) matrix takes the form

(2.1)
$$\begin{bmatrix} r & \lambda & \cdots & \lambda \\ \lambda & r & \cdots & \lambda \\ \vdots & & & \\ \lambda & \lambda & \cdots & r \end{bmatrix}.$$

Now

$$\det (X'X) = \{\det (X)\}^2,$$

= $(r - \lambda)^{N-1} \{r + \lambda(N-1)\}.$

Therefore,

(2.2)
$$\det (X) = \pm (r - \lambda)^{(N-1)/2} \{r + \lambda(N-1)\}^{\frac{1}{2}}.$$

The det (X) is real and not equal to zero. Hence we have

$$(2.3) r > \lambda$$

and

$$(2.4) r + \lambda(N-1) > 0.$$

Relation (2.4) holds good when λ is non negative and $\lambda = -1$. In the latter case r = N.

Therefore, in this paper, we consider only the values of r and λ satisfying

$$r > \lambda \ge 0$$
, or $r = N$, $\lambda = -1$.

When the matrix (X'X) is of the form (2.1), the variance of the estimated weight is

(2.5)
$$\frac{\{r+\lambda(N-2)\}\sigma^2}{(r-\lambda)\{r+\lambda(N-1)\}}.$$

Therefore, the efficiency of the weighing design is

(2.6)
$$\frac{(r-\lambda)\{r+\lambda(N-1)\}}{N\{r+\lambda(N-2)\}} = f(r,\lambda), \text{ say.}$$

LEMMA 2.1.

- i) Let r = N. Then λ cannot be even (including zero) when N is odd and λ cannot be odd (including -1) when N is even.
- ii) Let r = N 1. Then λ cannot be even (including zero) when N is odd and λ cannot be odd when N is even.

Proof. Let x_i and x_j be any two column vectors of the design matrix X.

- i) When r = N, $x_i'x_j$ will have N terms each term being either +1 or -1. Since $x_i'x_j = \lambda$, amongst the N terms $\{N |\lambda|\}$ terms sum to zero. Hence N and $|\lambda|$ should either be odd or even and the statement follows.
- ii) When r = N 1, $x_i'x_j$ will have N terms each term being +1 or -1 or 0. Since $x_i'x_j = \lambda$, amongst the N terms $(N \lambda)$ terms sum to zero. If N is odd and λ is even $(N \lambda)$ will become odd and the $(N \lambda)$ terms cannot sum to zero unless there is a single zero term. x_i and x_j will contribute a single zero term to $x_i'x_j$ when and only when the zeros of x_i and x_j are in the same row. This is also the case for any two columns of X. Hence, if N is odd and λ is even, we get a row of zeros in X, and in this case det (X) = 0, contrary to our assumption.

Therefore, λ cannot be even when N is odd.

Similarly we can show that λ cannot be odd when N is even.

Theorem 2.1. When N is odd the best weighing design X is that for which

$$X'X = \begin{bmatrix} N & 1 & \cdots & 1 \\ 1 & N & \cdots & 1 \\ \vdots & & & \\ 1 & 1 & \cdots & N \end{bmatrix}.$$

Proof.

$$f(N,1) - f(r,\lambda)$$

$$= \frac{2N-1}{2N} - \frac{(r-\lambda)\{r+\lambda(N-1)\}}{N\{r+\lambda(N-2)\}}, \quad \text{or } r=N, \lambda = -1.$$

$$(2.7) = \frac{(2N-1)(N-2)\lambda + (2N-1)r - 2r^2 + 2r\lambda - 2(r-\lambda)(N-1)\lambda}{2N\{r+\lambda(N-2)\}},$$

$$= \frac{(N-1)\lambda(2N-2r+2\lambda-1) + (2N-2r-1)(r-\lambda)}{2N\{r+\lambda(N-2)\}} > 0,$$

when r < N.

(2.8)
$$f(N,1) - f(r,\lambda) = \frac{\lambda(\lambda - 1)(N - 2) + N(\lambda^2 - 1)}{2N\{r + \lambda(N - 2)\}},$$

when r = N. (2.8) is again greater than zero for all values of λ excepting zero in which case it is less than zero. But Lemma 2.1 proves that λ cannot be zero since N is odd. Hence the best weighing design in this case has efficiency f(N, 1).

The theorem is thus proved.

THEOREM 2.2. When $N \equiv 2 \pmod{4}$ the best weighing design X is that for which

$$X'X = \text{diag}\{(N-1), (N-1), \dots, N \text{ terms}\}.$$

PROOF.

$$f(N-1,0) - f(r,\lambda) = \frac{(N-1)}{N} - \frac{(r-\lambda)\{r+\lambda(N-1)\}}{N\{r+\lambda(N-2)\}},$$

$$r, \lambda \text{ are positive, } r > \lambda, \text{ or } r = N, \lambda = -1.$$

$$= \frac{r(N-1) + (N-1)(N-2)\lambda - r(r-\lambda)}{-(N-1)(r-\lambda)\lambda},$$

$$= \frac{-(N-1)(r-\lambda)\lambda}{N\{r+\lambda(N-2)\}},$$

$$= \frac{r(N-r+\lambda-1) + \lambda(N-1)}{N\{r+\lambda(N-2)\}} > 0,$$

when r < N.

(2.9)
$$f(N-1,0) - f(r,\lambda) = \frac{N(\lambda-1) + (N-1)(\lambda-2)\lambda}{N\{r+\lambda(N-2)\}},$$

when r = N. (2.9) is greater than zero when $\lambda \ge 2$ and $\lambda = -1$ and it is less than zero when $\lambda = 0$ or 1.

Lemma 2.1 proves that $\lambda=1$ cannot exist and it is known that the Hadamard matrix cannot exist in this case and λ cannot be zero.

Therefore, the best weighing design in this case has efficiency f(N-1, 0). The theorem is hence proved.

3. P_N matrices.

Definition 3.1. A P_N matrix is an Nth order matrix with elements +1 and -1 such that

$$P'_{N}P_{N} = (N-1)I_{N} + E_{NN}$$

where I_N is the identity matrix of order N and E_{NN} is an Nth order matrix with positive unit elements everywhere.

It is obvious from theorem 2.1 above that the P_N matrix is the best weighing design whenever it exists and N must be odd.

Theorem 3.1. A necessary condition for the existence of P_N is that

$$N = \frac{d^2 + 1}{2}$$

where d is an odd integer.

PROOF. det $\{P'_N P_N\} = \{\det P_N\}^2 = (2N-1)(N-1)^{N-1}$. Therefore det $(P_N) = (2N-1)^{\frac{1}{2}}(N-1)^{(N-1)/2}$. Also since P_N is a matrix with integral elements det (P_N) is an integer.

Hence (2N-1) should be a perfect square. Let

$$2N - 1 = d^2$$

$$N = \frac{d^2 + 1}{2}$$

Since N is an integer, d is an odd integer and thus the theorem is proved. Theorem 3.2. If a Balanced Incomplete Block Design exists with parameters

$$v^* = b^* = N$$
, $r^* = k^* = (N \pm d)/2$, $\lambda^* = (N \pm 2d + 1)/4$,

then, by changing the zeros into -1's in the incidence matrix of the incomplete block design, we get a P_N matrix.

Proof. Let the column vectors of the incidence matrix after 0's are changed to -1 be p_1 , p_2 , \cdots , p_N .

The negative contribution to $p'_i p_j = 2(r^* - \lambda^*) = (N-1)/2$ $(i, j = 1, 2, \dots, N; i \neq j)$.

Therefore, the positive contribution to $p'_i p_j = (N+1)/2$. Hence $p'_i p_j = 1$. Thus the theorem is established.

4. S_N matrices. Williamson [7] proved that when

$$N=p^h+1,$$

where p is an odd prime and h is a positive integer such that $p^h \equiv 1 \pmod{4}$, then a symmetric matrix S_N exists such that

$$S_N'S_N = (N-1)I_N,$$

where I_N is the Nth order identity matrix. In that case the S_N matrix can be taken as our best weighing design. The construction of the S_N matrices is based on Galois Fields and the Legendre function ζ , and it is discussed in detail in [7].

5. Numerical examples. Now we construct some designs that belong to the P_N and S_N series. Among the designs given below, P_5 is proved to be the best design by Mood. A similar type of S_6 is constructed by Banerjee [1] intuitively.

$$P_5 = \begin{bmatrix} -1 & 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & 1 & 1 \\ 1 & 1 & -1 & 1 & 1 \\ 1 & 1 & 1 & -1 & 1 \\ 1 & 1 & 1 & 1 & -1 \end{bmatrix}.$$

Variance of each estimated weight = $2\sigma^2/9$. Covariance of each pair of estimated weights = $-\sigma^2/36$. Efficiency = 9/10.

$$S_6 = \begin{bmatrix} 0 & 1 & 1 & 1 & 1 & 1 \\ 1 & 0 & 1 & -1 & -1 & 1 \\ 1 & 1 & 0 & 1 & -1 & -1 \\ 1 & -1 & 1 & 0 & 1 & -1 \\ 1 & -1 & -1 & 1 & 0 & 1 \\ 1 & 1 & -1 & -1 & 1 & 0 \end{bmatrix}.$$

Variance of each estimated weight = $\sigma^2/5$. Covariance of each pair of estimated weights = 0. Efficiency = 5/6.

Variance of each estimated weight $= \sigma^2/9$. Covariance of each pair of estimated weights = 0. Efficiency = 9/10.

Variance of each estimated weight = $2\sigma^2/25$. Covariance of each pair of estimated weights = $-\sigma^2/300$. Efficiency = 25/26.

Variance of each estimated weight = $\sigma^2/13$. Covariance of each pair of estimated weights = 0. Efficiency = 13/14.

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Variance of each estimated weight = $\sigma^2/17$. Covariance of each pair of estimated weights = 0. Efficiency = 17/18.

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Variance of each estimated weight = $2\sigma^2/49$. Covariance of each pair of estimated weights = $-\sigma^2/1176$. Efficiency = 49/50.

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REFERENCES

- [1] K. S. Banerjee, "Some contributions to Hotelling's weighing designs," Sankhyā, Vol. 10 (1950), pp. 371-382.
- [2] K. S. Banerjee, "Weighing designs," Calcutta Stat. Assn. Bull., Vol. 3 (1950-51), pp. 64-76.
- [3] H. Hotelling, "Some improvements in weighing and other experimental techniques," *Ann. Math. Stat.*, Vol. 15 (1944), pp. 297-306.
- [4] K. Kishen, "On the design of experiments for weighing and making other types of measurements," Ann. Math. Stat., Vol. 16 (1945), pp. 294-300.
- [5] A. M. Moop, "On Hotelling's weighing problem," Ann. Math. Stat., Vol. 17 (1946), pp. 432-446.
- [6] R. L. Plackett and J. P. Burman, "Designs of optimum multifactorial experiments," Biometrika, Vol. 33 (1946), pp. 305-325.
- [7] J. WILLIAMSON, "Hadamard's determinant theorem and the sun of four squares," Duke Math. Jour., Vol. 11 (1944), pp. 65-82.