# ON THE GENERALIZED MELLIN TRANSFORM OF A COMPLEX RANDOM VARIABLE AND ITS APPLICATIONS

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## 1. Introduction. The Mellin transform

$$h(s) = E[X^s]$$

of a real positive random variable X is a useful tool to treat products

$$(1.2) Y = A \cdot X_1 \cdot \cdot \cdot X_n$$

of independent positive random variables  $X_1, X_2, \dots, X_n, A$  being a positive constant. It can also be used to treat products of powers

$$(1.3) W = A \cdot X_1^{a_1} \cdot X_2^{a_1} \cdot \cdots \cdot X_n^{a_n}$$

where  $a_1, a_2, \dots, a_n$  are real (see [2], [3], [5], [6], [9]).

This Mellin transform is not as useful in cases for which  $X_k$  take both positive and negative values or complex values. W. M. Zolotariow [10] has given a tool to treat products of real (not necessary positive) random variables; this tool is not useful in cases when the factors are complex. P. Lévy has given a tool to treat products (1.2) of complex random variables (see [7]); this tool is not as useful for products (1.3) with  $a_k$  real.

In this paper a generalization of the Mellin transform (1.1) is given in such a way that it will be useful to treat products (1.3) where  $X_1$ ,  $X_2$ ,  $\cdots$ ,  $X_n$  are complex random variables for which  $P\{X_k = 0\} = 0$ , i.e. taking values in the set  $G^*$  of non-zero complex numbers, and  $a_k$  being real.

Under multiplication (1.2) the set  $G^*$  of non-zero complex numbers is an Abelian locally compact group isomorphic to the direct product  $\mathfrak{R} \times T$ , where  $\mathfrak{R}$  is the multiplicative group of positive real numbers, which is isomorphic to the additive group of real numbers, and T denotes the additive group of real numbers modulo  $2\pi$ . Given this structure of  $G^*$  the natural transform of a complex random variable  $Z = R \cdot e^{i\Phi}$  on  $G^*$  would be

$$(1.4) \quad h(t, n) = E[R^{it}e^{in\Phi}], \quad -\infty < t < +\infty, n = \cdots, -1, 0, 1, \cdots$$

(On this subject see [8], p. 141, [1], p. 73, [4], pp. 166-167).

The integral transform (1.4) does not suffice in cases where products (1.3) are treated with  $a_k$  being real but not necessary integer. In such a case it is more convenient to treat probability distributions on the set G being the Riemann surface of the function  $w = \log z$ . Under multiplication (1.3) the set G is isomorphic to the direct product  $G \times G$ , and that is why the natural transform of a probability distribution on G would be

(1.5) 
$$h(t, v) = E[R^{it}e^{iv\Phi}], \quad -\infty < t < +\infty, -\infty < v < +\infty.$$

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The way from a distribution on G to the corresponding distribution on  $G^*$  should be made by the suitable projection of the Riemann surface of the function  $w = \log z$  on the non-zero complex plane  $G^*$ .

In this paper we shall take the transform (1.5) for t and v complex, the case t and v real will be a particular one.

The generalized Mellin transform may also be used to obtain the distribution of the scalar product  $X_1X_2 + Y_1Y_2$  of two bivariate independent random vectors  $[X_1, Y_1]$ ,  $[X_2, Y_2]$  as well as the distribution of the determinant  $\begin{vmatrix} X_1, Y_1 \\ X_2, Y_2 \end{vmatrix}$  (see Section 4.4.).

2. The definition of the generalized Mellin transform of a complex random variable and its properties. Let us consider a bivariate random variable  $(R, \Phi)$  taking values  $(r, \varphi)$  on the half plane

$$(2.1) 0 < r < \infty, -\infty < \varphi < +\infty.$$

Denote  $\Phi^*$  the principal value of  $\Phi$ , i.e.

(2.2) 
$$\Phi^* \equiv \Phi \mod 2\pi, \qquad -\pi < \Phi^* \leq \pi.$$

The bivariate random variable  $(R, \Phi^*)$  takes its values  $(r, \varphi^*)$  on the half strip

$$(2.3) 0 < r < \infty, \quad -\pi < \varphi^* \le \pi.$$

For a given distribution  $P\{R \leq r, \Phi \leq \varphi\}$  of  $(R, \Phi)$ , the distribution of  $(R, \Phi^*)$  may be easily found by the projecting formula

(2.4) 
$$P(R \le r, \Phi^* \le \varphi^*) = \sum_{k=-\infty}^{+\infty} P\{R \le r, k \cdot 2\pi - \pi < \Phi \le k \cdot 2\pi + \varphi^*\},$$
  
 $0 < r < \infty, -\pi < \varphi^* \le \pi.$ 

Denote

$$(2.5) Z = R \cdot e^{i\Phi}.$$

We obtain a complex random variable which does not meet the zero value. The function (2.5) is one-to-one for  $\Phi$  taking values  $-\pi < \varphi \leq \pi$ , but is not otherwise. Using the periodicity of the exponential function, we may also write (2.5) in form

$$(2.6) Z = R \cdot e^{i\Phi^*}.$$

This function is one-to-one.

Now we define the generalized Mellin transform of the complex random variable (2.5). It is given by the formula

$$(2.7) h(u,v) = E[R^{u}e^{iv\Phi}],$$

where u and v are complex variables. It is easy to see that the transform (2.7) is for u = it, the characteristic function of the bivariate random variable  $(\log R, \Phi)$ 

$$(2.8) \quad h(it, v) = E[R^{it}e^{iv\Phi}] = E[\exp (i(t \cdot \log R + v \cdot \Phi))] = \psi_{(\log R, \Phi)}(t, v).$$

From the known properties of characteristic functions it follows that the transform (2.7) is well defined in some pair of strips

$$(2.9) u_1 \leq \operatorname{Re} u \leq u_2, \quad v_1 \leq \operatorname{Im} v \leq v_2,$$

where  $u_1$ ,  $u_2$ ,  $v_1$ ,  $v_2$  are real and satisfy the conditions  $u_1 \le 0 \le u_2$ ,  $v_1 \le 0 \le v_2$ . It is enough to take the transform (2.7) for u = it, t and v being real. In such a case the inequalities (2.9) should be omitted. In this paper it is more convenient to take u and v complex.

It should be easily seen that h(u, 0) is the Mellin transform of the positive random variable R, and h(0, v) is the characteristic function of the random variable  $\Phi$ . The random variables R and  $\Phi$  are independent if and only if

$$(2.10) h(u, v) = h(u, 0) \cdot h(0, v).$$

From the properties of characteristic functions it follows that h(u, v) is continuous in the pair of strips (2.9), that h(u, v) defines the distribution of  $(R, \Phi)$  uniquely, and h(0, 0) = 1.

Further for t and v real there is

$$(2.11) |h(it, v)| \leq 1,$$

$$(2.12) h(-it, -v) = \overline{h(it, v)},$$

and h(it, v) is a positive definite function.

From the uniqueness property it follows that h(u, v) defines also the distribution of (X, Y), where

$$(2.13) X = R \cos \Phi, \quad Y = R \sin \Phi,$$

uniquely. The distribution of (X, Y) may be found by the suitable projecting. Let us consider n independent complex random variables

$$(2.14) Z_k = R_k \cdot e^{i\Phi_k}, k = 1, 2, \cdots, n,$$

satisfying the conditions  $P\{Z_k = 0\} = 0$ , and having their generalized Mellin transforms

$$(2.15) h_k(u,v) = E[R_k^{\ u}e^{iv\Phi_k}].$$

Let  $A = A_0 e^{i\varphi_0}$  be a free non-zero complex number, and let  $a_1$ ,  $a_2$ ,  $\cdots$ ,  $a_n$  be free real numbers. Denote

$$(2.16) W = A \cdot Z_1^{a_1} \cdot Z_2^{a_2} \cdot \cdots \cdot Z_n^{a_n}.$$

This product can also be written in form

(2.16') 
$$|W| = A_0 \cdot R_1^{a_1} \cdot R_2^{a_2} \cdot \cdots \cdot R_n^{a_n};$$

$$\arg W = \varphi_0 + a_1 \Phi_1 + a_2 \Phi_2 + \cdots + a_n \Phi_n.$$

Then the generalized Mellin transform of the random variable W is

$$h_{W}(u, v) = E[|W|^{u}e^{iv \operatorname{arg}W}]$$

$$= E[(A_{0}R_{1}^{a_{1}}R_{2}^{a_{2}}\cdots R_{n}^{a_{n}})^{u}$$

$$\cdot \exp[(iv(\varphi_{0} + a_{1}\Phi_{1} + a_{2}\Phi_{2} + \cdots + a_{n}\Phi_{n}))]$$

$$= A_{0}^{u} \cdot e^{i\varphi_{0}v} \cdot E[R_{1}^{a_{1}u}e^{iva_{1}\Phi_{1}}] \cdot E[R_{2}^{\alpha_{2}u}e^{iva_{2}\Phi_{2}}] \cdot \cdots \cdot E[R_{n}^{a_{n}u}e^{iva_{n}\Phi_{n}}]$$

$$= A_{0}^{u} \cdot e^{i\varphi_{0}v} \cdot h_{1}(a_{1}u, a_{1}v) \cdot h_{2}(a_{2}u, a_{2}v) \cdot \cdots \cdot h_{n}(a_{n}u, a_{n}v).$$

In particular cases we have

$$(2.18.1) h_{W}(u, v) = h_{Z}(au, av), for W = Z^{a};$$

(2.18.2) 
$$h_W(u, v) = h_Z(-u, -v),$$
 for  $W = 1/Z$ ;

(2.18.3) 
$$h_W(u, v) = h_{Z_1}(u, v) \cdot h_{Z_2}(u, v),$$
 for  $W = Z_1 \cdot Z_2$ ;

$$(2.18.4) h_{W}(u, v) = h_{Z_{1}}(u, v) \cdot h_{Z_{2}}(-u, -v), \text{for } W = Z_{1}/Z_{2}.$$

Further we have

$$(2.19) h_{\mathbf{w}}(u,v) = h_{\mathbf{z}}(u,-v), \text{for } W = \overline{Z}.$$

#### 3. Particular cases.

3.1. Let the complex random variable (2.5) have the uniform distribution on the arc  $r = \alpha$ ,  $\beta - \gamma < \varphi < \beta + \gamma$ . This distribution is defined by the density

(3.1.1) 
$$g(\varphi) = 1/(2\gamma), \quad \text{for } \beta - \gamma < \varphi < \beta + \gamma,$$
  
= 0, otherwise.

The generalized Mellin transform is in this case

$$(3.1.2) \quad h(u, v) = E[R^u e^{iv\Phi}] = \alpha^u \int_{\beta-\gamma}^{\beta+\gamma} e^{iv\varphi} (1/(2\gamma)) \, d\varphi = \alpha^u e^{i\beta v} [(\sin \gamma v)/\gamma v]$$

Taking  $\alpha = 1$ ,  $\beta = 0$ ,  $\gamma = \pi$ , we obtain the uniform distribution on the unity circle, having density

(3.1.3) 
$$g(\varphi) = 1/2\pi, \quad \text{for } -\pi < \varphi < +\pi,$$
$$= 0, \quad \text{otherwise,}$$

and the generalized Mellin transform

$$(3.1.4) h(u, v) = (\sin \pi v)/\pi v.$$

3.2. Let the complex random variable (2.5) have the uniform distribution inside the sector of the circle  $0 < r < \alpha, \beta - \gamma < \varphi < \beta + \gamma$ , where

$$(3.2.1) -\pi \leq \beta - \gamma < \beta + \gamma \leq \pi.$$

This distribution is given by the density

(3.2.2) 
$$f(x, y) = 1/\gamma \alpha^2$$
, for  $(0 < r < \alpha, \beta - \gamma < \varphi < \beta + \gamma)$ ,  
= 0, otherwise.

The corresponding density for  $(R, \Phi)$  is

(3.2.3) 
$$g(r, \varphi) = (1/\gamma \alpha^2) \cdot r$$
, for  $(0 < r < \alpha, \beta - \gamma < \varphi < \beta + \gamma)$ ,  
= 0, otherwise.

The integral transform is in this case

(3.2.4) 
$$h(u,v) = E[R^u e^{iv\Phi}] = \int_{\beta-\gamma}^{\beta+\gamma} d\varphi \int_0^\alpha e^{iv\varphi} r^u (1/\gamma \alpha^2) r \, dr$$
$$= [2/(2+u)] \alpha^u e^{i\beta v} [(\sin \gamma v)/\gamma v].$$

(The restriction (3.2.1.) may be omitted in Formulae (3.2.3), (3.2.4).)

Taking  $\alpha = 1$ ,  $\beta = 0$ ,  $\gamma = \pi$ , we obtain the uniform distribution inside the unity circle, having for (X, Y) the density

(3.2.5) 
$$f(x, y) = 1/\pi$$
, for  $x^2 + y^2 < 1$ ,  
= 0, otherwise.

The corresponding density of  $(R, \Phi)$  is

(3.2.6) 
$$g(r, \varphi) = (1/\pi)r$$
, for  $0 < r < 1, -\pi < \varphi < +\pi$ ,  
= 0, otherwise.

The corresponding integral transform is

(3.2.7) 
$$h(u, v) = [2/(2 + u)] \cdot [(\sin \pi v)/\pi v].$$

3.3. Let the complex random variable (2.5) have its distribution given by the density

(3.3.1) 
$$g(r,\varphi) = (1/2\gamma)g_0(r)$$
, for  $(0 < r < \infty, \beta - \gamma < \varphi < \beta + \gamma)$ ,  
= 0, otherwise.

The integral transform is in this case

(3.3.2) 
$$h(u, v) = E[R^u e^{iv\Phi}] = (1/2\gamma) \int_0^\infty r^u g_0(r) dr \int_{\beta-\gamma}^{\beta+\gamma} e^{iv\varphi} d\varphi.$$
 Denoting  $h_0(u) = \int_0^\infty r^u g_0(r) dr$ , we obtain

$$(3.3.3) h(u,v) = h_0(u)e^{i\beta v}[(\sin \gamma v)/\gamma v].$$

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(3.3.4) 
$$g_0(r) = [|q|a^{p/q}/\Gamma(p/q)]r^{p-1}e^{-ar^q}, q \neq 0, a > 0, p/q > 0,$$
 we obtain the corresponding integral transform

 $(3.3.5) h(u, v) = a^{-u/q} \left[\Gamma((p + u)/q)/\Gamma(p/q)\right] e^{i\beta v} \left[\sin \gamma v/\gamma v\right].$ 

(3.3.6) 
$$g_0(r) = [|q|\Gamma(a)/\Gamma(p/q)\Gamma(a - (p/q))][r^{p-1}/(1 + r^q)^a],$$
  
 $q \neq 0, p/q > 0, a - (p/q) > 0,$ 

we obtain the corresponding integral transform

(3.3.7) 
$$h(u, v) = e^{i\beta v} [(\sin \gamma v)/\gamma v] [\Gamma((p+u)/q)/\Gamma(p/q)] \cdot [\Gamma(a-(p+u)/q)/\Gamma(a-(p/q))].$$

3.4. Let the complex random variable (2.5) take its values on the sides of the angle  $\varphi = \alpha$ ,  $\varphi = \beta$ , and the densities of R on these sides are  $g_{\alpha}(r)$ ,  $g_{\beta}(r)$ ,  $0 < r < \infty$ . The integral transform is in this case

$$(3.4.1) \quad h(u, v) = E[R^{u}e^{iv\Phi}] = e^{i\alpha v} \int_{0}^{\infty} r^{u}g_{\alpha}(r) dr + e^{i\beta v} \int_{0}^{\infty} r^{u}g_{\beta}(r) dr.$$

Denoting

$$(3.4.2) h_{\alpha}(u) = \int_0^{\infty} r^u h_{\alpha}(r) dr, h_{\beta}(u) = \int_0^{\infty} r^u g_{\beta}(r) dr,$$

we obtain

$$(3.4.3) h(u,v) = e^{i\alpha v} \cdot h_{\alpha}(u) + e^{i\beta v} \cdot h_{\beta}(u).$$

Taking

$$(3.4.4) \quad \alpha = 0, \quad \beta = \pi, \quad g_{\alpha}(r) = g_{\beta}(r) = (1/\sqrt{\pi})e^{-r^2}, \quad (0 < r < \infty),$$

we obtain Z as a real random variable taking positive values as well as negative; its integral transform is

(3.4.5) 
$$h(u, v) = (1 + e^{i\pi v}) \int_0^\infty r^u (1/\sqrt{\pi}) e^{-r^2} dr$$
  
=  $e^{i\frac{1}{2}\pi v} \cos \frac{1}{2}\pi v (1/\sqrt{\pi}) \Gamma(\frac{1}{2}(1 + u))$ .

3.5. Let the complex random variable (2.5) have R and  $\Phi$  independent and their distributions given by densities  $g_0(r)$  and  $g_1(\varphi)$  where

$$g_{1}(\varphi) = 0, \qquad \text{for } \varphi \leq \alpha - \beta, \varphi \geq \alpha + \beta,$$

$$(3.5.1) \qquad = [\beta + (\varphi - \alpha)]/\beta^{2}, \qquad \text{for } \alpha - \beta < \varphi < \alpha,$$

$$= [\beta - (\varphi - \alpha)]/\beta^{2}, \qquad \text{for } \alpha < \varphi < \alpha + \beta;$$

 $\Phi$  has in this case the triangular distribution. The integral transform is in this case

$$h(u,v) = E[R^{u}e^{iv\Phi}] = E[R^{u}] \cdot E[e^{iv\Phi}]$$

$$= \int_{0}^{\infty} r^{u}g_{0}(r) dr[\int_{\alpha-\beta}^{\alpha} e^{iv\varphi}[(\beta + (\varphi - \alpha))/\beta^{2}] d\varphi$$

$$+ \int_{\alpha}^{\alpha+\beta} e^{iv\varphi}[(\beta - (\varphi - \alpha))/\beta^{2}] d\varphi]$$

$$= h_{0}(u)e^{i\alpha v}[(\sin \frac{1}{2}\beta v)/\frac{1}{2}\beta v]^{2}$$

where  $h_0(u) = \int_0^\infty r^u g_0(r) dr$ .

Taking for instance,  $\alpha = 0$ ,  $\beta = 2\pi$ , we obtain

$$(3.5.3) g_1(\varphi) = 0, \text{for } \varphi \leq -2\pi, \varphi \geq 2\pi$$

$$= (2\pi + \varphi)/4\pi^2, \text{for } -2\pi < \varphi < 0$$

$$= (2\pi - \varphi)/4\pi^2, \text{for } 0 < \varphi < 2\pi$$

and

$$(3.5.4) h(u, v) = h_0(u) \cdot \left[ (\sin \pi v) / \pi v \right]^2.$$

Now let us project the Riemann surface of the function  $w = \log z$  ( $0 < r < \infty$ ,  $-2\pi < \varphi < 2\pi$ ) on the complex plane ( $0 < r < \infty$ ,  $-\pi < \varphi \le \pi$ ). It can be easy seen that the triangular distribution (3.5.1) on the interval ( $-2\pi$ ,  $2\pi$ ) becomes rectangular distribution on the interval ( $-\pi$ ,  $\pi$ ). That is why the distribution of  $Z^* = R \cdot e^{i\Phi^*}$ , where  $\Phi^*$  is the principal value of  $\Phi$ , is given by the integral transform

$$(3.5.5) h^*(u,v) = h_0(u) \cdot [(\sin \pi v)/\pi v].$$

# 4. Applications.

4.1. Let the complex random variable,

$$(4.1.1) Z = X + iY = R \cdot e^{i\Phi},$$

have the bivariate distribution given by the density

(4.1.2) 
$$f(x,y) = (1/\pi) \exp[-(x^2 + y^2)], -\infty < x < +\infty, -\infty < y < +\infty.$$

The density of the bivariate random variable  $(R, \Phi)$  is

(4.1.3) 
$$g(r,\varphi) = (1/\pi)re^{-r^2} \qquad 0 < r < \infty, -\pi < \varphi < \pi.$$

The integral transform of this random variable is (see Formulae (3.3.4) and (3.3.5))

(4.1.4) 
$$h(u, v) = [(\sin \pi v)/\pi v]\Gamma(1 + \frac{1}{2}u).$$

The reciprocal 1/Z has the integral transform (see Formula (2.18.2))

$$(4.1.5) h_1(u,v) = h(-u,-v) = [(\sin \pi v)/\pi v]\Gamma(1-\frac{1}{2}u).$$

The distribution of the modulus and the argument of the reciprocal is given by the corresponding to (4.1.5) density (see Formulae (3.3.4), (3.3.5)

(4.1.6) 
$$g_1(r,\varphi) = (1/\pi)r^{-3}e^{-r^{-2}} \qquad 0 < r < \infty, -\pi < \varphi < \pi.$$

Thus the distribution of the real and imaginary parts of the reciprocal is

(4.1.7) 
$$f_1(x, y) = (1/\pi)[1/(x^2 + y^2)^2]$$
  
  $\exp[-1/(x^2 + y^2)], > \infty < x < +\infty, -\infty < y < +\infty.$ 

4.2. Let us have two independent complex random variables  $Z_1$ ,  $Z_2$  having identical normal distributions given by density (4.1.2). We shall find the distribution of the quotient of these variables

$$(4.2.1) W = Z_1/Z_2.$$

The integral transform of W (see Formulae (2.18.4), (4.1.4)) is

$$(4.2.2) \quad h_1(u,v) = h(u,v) \cdot h(-u,-v) = \left[ (\sin \pi v) / \pi v \right]^2 \Gamma(1+\frac{1}{2}u) \Gamma(1-\frac{1}{2}u).$$

The density of the modulus and the argument of W is given by

$$(4.2.3) g_1(r,\varphi) = [2r/(1+r^2)^2] \cdot g_2(\varphi), 0 < r < \infty, -2\pi < \varphi < 2\pi$$

where  $g_2(\varphi)$  is given by Formula (3.5.3) (see also Formula (3.5.4)).

Projecting the Riemann surface of the function  $w = \log z$  ( $0 < r < \infty$ ,  $-2\pi < \varphi < 2\pi$ ) on the complex plane ( $0 < r < \infty$ ,  $-\pi < \varphi < \pi$ ) we see that the triangular distribution of the argument W becomes rectangular distribution on the interval  $(-\pi, +\pi)$ . That is why the density of the modulus and the principal value of the argument is

(4.2.4) 
$$g_1^*(r,\varphi) = (1/2\pi)[2r/(1+r^2)^2]$$
  
=  $(1/\pi)[r/(1+r^2)^2]$ ,  $0 < r < \infty, -\pi < \varphi < +\pi$ .

Hence the distribution of the real and imaginary parts of W is given by density (4.2.5)  $f_1(x, y)$ 

$$= (1/\pi)[1/(1+x^2+y^2)^2], -\infty < x < +\infty, -\infty < y < +\infty.$$

4.3. Let us have two independent real random variables  $Z_1$ ,  $Z_2$  having identical normal distribution given by density  $f(x) = (1/\pi)e^{-x^2}$ ,  $-\infty < x < +\infty$ . The integral transform of this random variable (see Formulae (3.4.4), (3.4.5)) is

$$(4.3.1) h(u,v) = e^{i\frac{1}{2}\pi v} \cos \frac{1}{2}\pi v (1/\sqrt{\pi}) \Gamma(\frac{1}{2}(1+u)).$$

We shall find the distribution of the quotient (4.2.1). Its integral transform is (see Formulae (2.18.4), (4.3.1))

$$(4.3.2) \quad h_1(u, v) = h(u, v) \cdot h(-u, -v) = \cos^2 \frac{1}{2} \pi v \cdot (1/\pi) \Gamma(\frac{1}{2}(1+u)) \Gamma(\frac{1}{2}(1-u)).$$

Hence we see that the modulus and the argument of the quotient W are independent, the argument being distributed according to the characteristic function  $h_1(0, v) = \cos^2 \frac{1}{2}\pi v$ .

Then we see that the argument of W takes values  $-\pi$ , 0,  $+\pi$  with probabilities  $\frac{1}{4}$ ,  $\frac{1}{2}$ ,  $\frac{1}{4}$  respectively. Projecting the half line  $\varphi = -\pi$  on the half line  $\varphi = \pi$  we see that the principal argument of W takes values 0 and  $\pi$  with probabilities  $\frac{1}{2}$ ,  $\frac{1}{2}$ .

The modulus of W is distributed according to the Mellin transform

$$(4.3.3) h_1(u,0) = (1/\pi)\Gamma(\frac{1}{2}(1+u))\Gamma(\frac{1}{2}(1-u)) = 1/\cos\frac{1}{2}\pi u.$$

That is why the distribution of W is on the real line, it is given by the density

$$(4.3.4) g_0(r) = (1/\pi) \cdot [1/(1+r^2)],$$

$$g_{\pi}(r) = (1/\pi) \cdot [1/(1+r^2)], \quad 0 < r < \infty.$$

From the Formula (4.3.4) it follows that W is distributed on the whole real line according to the Cauchy law,  $f_1(x) = (1/\pi) \cdot [1/(1+x^2)], -\infty < x < +\infty$ .

4.4. Let us have two independent real bivariate random vectors  $Q_1 = [X_1, Y_1]$ ,

 $Q_2 = [X_2, Y_2]$ . We shall find the distribution of the scalar product  $U = X_1 \cdot X_2 + Y_1 \cdot Y_2$ , and the determinant  $V = \begin{vmatrix} X_1, & Y_1 \\ X_2, & Y_2 \end{vmatrix}$ .

Denote

$$(4.4.1) Z_1 = X_1 + iY_1, Z_2 = X_2 + iY_2, W = U + iV.$$

The complex random variables  $Z_1$ ,  $Z_2$  are connected with U, V by the formula

$$(4.4.2) W = U + iV = \overline{Z_1} \cdot Z_2.$$

Taking the integral transform of (4.4.2) we obtain

$$(4.4.3) h_{\mathbf{w}}(u, v) = h_{\mathbf{z}_1}(u, -v) \cdot h_{\mathbf{z}_2}(u, v),$$

(see Formulae (2.18.3), (2.19)).

Let, for example,  $Q_1$  be normal according to the density (4.1.2) and  $Q_2$  be reciprocal normal according to the density (4.1.7). Then  $h_{Z_1}(u, v)$  and  $h_{Z_2}(u, v)$  are given by Formulae (4.1.4) and (4.1.5).

Using Formula (4.4.3) we obtain

$$(4.4.4) h_{W}(u, v) = h_{Z_{1}}(u, -v) \cdot h_{Z_{2}}(u, v)$$
$$= [(\sin \pi v)/\pi v]\Gamma(1 + \frac{1}{2}u)\Gamma(1 - \frac{1}{2}u).$$

The density of W corresponding to (4.4.4) is

$$(4.4.5)$$
  $f_{\mathbf{w}}(x, y)$ 

$$= (1/\pi)[1/(1+x^2+y^2)^2], -\infty < x < +\infty, -\infty < y < +\infty.$$

(see Section 4.2). From this density we see that U and V have in this case both the same marginal distributions given by density

$$(4.4.6) f_{v}(x) = f_{v}(x) = 1/[2(1+x^{2})^{3/2}], -\infty < x < +\infty.$$

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