

# ON RANDOM SUMS OF RANDOM VECTORS

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**0. Summary.** To obtain the limit distribution of a sequence  $T_n$  of random vectors, the  $j$ th component of  $T_n$  being the sum of a random number  $N_n^{(j)}$  of  $j$ th components of independent, identically distributed chance vectors  $X_n$ , it is first necessary to treat the special case where the  $N_n^{(j)}$  are degenerate random variables. This is done in Section 2 and generalized to infinitely divisible limits in Section 4. The basic problem is treated in Section 3 and generalizations of theorems of Doeblin [3], Anscombe [1] and Rényi [7] are obtained.

**1. Preliminaries.** Underlying the succeeding sections will be a sequence

$$(1.1) \quad X_n = (X_{n1}, X_{n2}, \dots, X_{nm}), \quad n = 1, 2, \dots,$$

of independent, identically distributed random vectors defined on some probability space with  $P$  the probability measure thereupon. Of especial interest will be the corresponding sequence of partial sums

$$(1.2) \quad S_n = (S_n^{(1)}, S_n^{(2)}, \dots, S_n^{(m)}), \quad n = 1, 2, \dots,$$

where  $S_n^{(j)} = \sum_{i=1}^n X_{ij}$ ,  $1 \leq j \leq m$ . Further,

$$(1.3) \quad b(n); g(n); n = 1, 2, \dots \text{ (set } b(0) = 1)$$

will signify increasing sequences of positive numbers tending to infinity while

$$(1.4) \quad N_n = (N_n^{(1)}, N_n^{(2)}, \dots, N_n^{(m)}), \quad n = 1, 2, \dots,$$

will constitute a sequence of random vectors whose component random variables are positive integer-valued.

Apropos of terminology, to say that the random vector  $Y_n = (Y_{n1}, Y_{n2}, \dots, Y_{nm})$  converges in probability to a constant (vector)  $c = (c_1, c_2, \dots, c_m)$  signifies that each component random variable  $Y_{nj}$  converges in probability (denoted by  $\rightarrow_p$ ) to the corresponding scalar  $c_j$ ,  $1 \leq j \leq m$ . Convergence of cumulative distribution functions (cdf's) at continuity points of the limit cdf is denoted by  $\Rightarrow$ . If the original cdf's are attached to random vectors, the latter will be said to converge in distribution.

Throughout Sections 2 and 3, it will be supposed that there exists a cdf  $F(x_1, x_2, \dots, x_m)$  such that as  $n \rightarrow \infty$ ,

$$(1.5) \quad P\{S_n^{(1)} < x_1 b(n), \dots, S_n^{(m)} < x_m b(n)\} \Rightarrow F(x_1, \dots, x_m).$$

The characteristic function (cf) of the necessarily stable (see [5], p. 221) cdf  $F$  will invariably be denoted by  $\varphi$ .

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The following lemma is an analogue of and may be proved in the same fashion as the corresponding well-known one-dimensional result. The proof of the second lemma is likewise sufficiently simple to omit.

LEMMA 1. Let  $(X_{n1}, \dots, X_{nm}), (Y_{n1}, \dots, Y_{nm})$  and  $(Z_{n1}, \dots, Z_{nm}), n = 1, 2, \dots$  constitute three sequences of random vectors, the first converging in distribution and the latter two converging in probability to constants  $(c_1, \dots, c_m)$  and  $(d_1, \dots, d_m)$  respectively. Then the characteristic function of the random vector  $(X_{n1} \cdot Y_{n1} + Z_{n1}, \dots, X_{nm} \cdot Y_{nm} + Z_{nm})$  converges for all real  $(t_1, \dots, t_m)$  to  $\exp\{i \sum_{j=1}^m t_j d_j\} \cdot \varphi(c_1 t_1, \dots, c_m t_m)$  where  $\varphi$  is the limit cf of  $(X_{n1}, \dots, X_{nm})$ .

LEMMA 2. Let  $\phi(t_1, t_2, \dots, t_m), \phi_n(t_1, t_2, \dots, t_m), n \geq 1$  be sequence of cf's with  $\lim_{n \rightarrow \infty} \phi_n(t_1, t_2, \dots, t_m) = \phi(t_1, t_2, \dots, t_m)$  and  $\{d_{kj}, 1 \leq j \leq m, k = 1, 2, \dots\}$  constants such that  $\lim_{k \rightarrow \infty} d_{kj} = 0, 1 \leq j \leq m$ , and  $\{n_k\}$  a sequence of (not necessarily distinct) positive integers. Then,  $\lim_{k \rightarrow \infty} \varphi_{n_k}(d_{k1}t_1, \dots, d_{km}t_m) = 1$ , all real  $(t_1, \dots, t_m)$ .

**2. Some unbalanced limit theorems.** In this section, the effect of systematic variation of the number of elements in a component sum is examined. Recall that a function  $h(x)$  is slowly varying if  $\lim_{x \rightarrow \infty} h(cx)/h(x) = 1$  for all  $c > 0$ .

THEOREM 1. Set  $b(n) = n^{1/\alpha} h(n), 0 < \alpha \leq 2$ , where  $h(n)$  is a slowly varying function and let  $X_n, S_n, b(n)$  be as described in (1.1), (1.2), (1.3). Suppose that (1.5) holds for some cdf  $F(x_1, x_2, \dots, x_m)$  whose cf is  $\phi(t_1, t_2, \dots, t_m)$  and let  $\{n_{kj}, 1 \leq j \leq m, k = 1, 2, \dots\}$  constitute  $m$  increasing sequences of positive integers for which

$$(2.0) \quad \lim_{k \rightarrow \infty} n_{k,j-1}/n_{k,j} = \bar{c}_j < \infty, \quad 2 \leq j \leq m.$$

(Without essential loss of generality  $\bar{c}_j, 2 \leq j \leq m$  will be supposed bounded by unity<sup>1</sup> and concomitantly  $n_{k,j-1}$  will be supposed at most equal to  $n_{kj}, 2 \leq j \leq m, k = 1, 2, \dots$ ). Define  $c_j = (\bar{c}_2 \cdot \bar{c}_3 \cdots \bar{c}_j)^{1/\alpha}$  and  $d_j^\alpha = 1 - \bar{c}_j, 2 \leq j \leq m$ . Then as  $k \rightarrow \infty$ ,

$$(2.1) \quad P\{S_{n_{k1}}^{(1)} < x_1 \cdot b(n_{k1}), \dots, S_{n_{km}}^{(m)} < x_m \cdot b(n_{km})\} \\ \Rightarrow G_\alpha(x_1, \dots, x_m; c_2, \dots, c_m)$$

where  $G_\alpha$  is a cdf with cf given by<sup>2</sup>

$$(2.2) \quad \phi(t_1, c_2 t_2, \dots, c_m t_m) \cdot \phi(0, d_2 t_2, d_2 c_3 t_3 / c_2, \dots, d_2 c_m t_m / c_2) \\ \cdots \phi(0, \dots, 0, d_{m-1} t_{m-1}, d_{m-1} c_m t_m / c_{m-1}) \cdot \phi(0, \dots, 0, d_m t_m).$$

PROOF. Note that  $0 < \bar{c}_j \leq 1, 2 \leq j \leq m$  is tantamount to  $1, c_2, c_3, \dots, c_m$  forming a non-increasing sequence. If, at the outset,  $\bar{c}_j$  is not bounded above by unity for all  $2 \leq j \leq m$ , this can be achieved by a permutation of the components of  $X_n$ , and we therefore suppose it to be true initially. Thus, for all sufficiently

<sup>1</sup> A method for obtaining the limit distribution when some  $\bar{c}_j > 1$  will be found in the "Prescription" following Theorem 1.

<sup>2</sup> 0/0 is to be interpreted as 0.

large  $k$ ,  $n_{k,j-1} < n_{kj}$  if  $c_j < c_{j-1}$  and otherwise (as has been supposed for convenience)  $n_{k,j-1} \leq n_{kj}$ .

As a consequence of (2.0), (1.3) and the slowly varying property,

$$\begin{aligned} \lim_{k \rightarrow \infty} [b(n_{k1})/b(n_{kj})] &= c_j, \quad 2 \leq j \leq m, \\ (2.3) \quad \lim_{k \rightarrow \infty} [b(n_{kj} - n_{k,j-1})/b(n_{kj})] &= d_j = (1 - c_j^\alpha / c_{j-1}^\alpha)^{1/\alpha} \\ &\quad \text{if } c_{j-1} > 0, \quad 2 \leq j \leq m, \\ &= d_j = 1 \quad \text{if } c_{j-1} = c_j = 0. \end{aligned}$$

Thus, denoting  $\sum_{i=h+1}^n X_{ij}$  by  $S_{h,n}^{(j)}$ , independence implies that

$$\begin{aligned} (2.4) \quad & E[\exp\{i \sum_{j=1}^m t_j S_{n_{kj}}^{(j)} / b(n_{kj})\}] \\ &= E[\exp\{i \sum_{j=1}^m t_j [b(n_{k1})/b(n_{kj})][S_{n_{k1}}^{(j)} / b(n_{k1})]\}] \\ &\quad \cdot E[\exp\{i \sum_{j=2}^m t_j [b(n_{k2} - n_{k1})/b(n_{kj})][S_{n_{k1}, n_{k2}}^{(j)} / b(n_{k2} - n_{k1})]\}] \\ &\quad \cdots E[\exp\{i t_m [b(n_{km} - n_{k,m-1})/b(n_{km})][S_{n_{k,m-1}, n_{km}}^{(m)} / b(n_{km} - n_{k,m-1})]\}]. \end{aligned}$$

If  $d_h > 0$ ,  $b(n_{kh} - n_{k,h-1})$  and therefore  $n_{kh} - n_{k,h-1}$  converges to infinity by (2.3). Consequently, by Lemma 1, the factor of (2.4) in which  $b(n_{kh} - n_{k,h-1})$  appears, approaches its correspondent in (2.2). On the other hand, if  $d_h = 0$ ,  $n_{kh} - n_{k,h-1}$  need not approach infinity. Nonetheless, applying Lemma 2, the limit of the appropriate factor in (2.4) will be unity, exactly the value assigned by  $\phi(0, 0, \dots, 0)$ . Thus, the limit of (2.4) exists and is given by (2.2).

It will be necessary in Section 3 to stipulate the limiting distribution when the  $\bar{c}_j$ 's are positive but not necessarily bounded above by unity. To this end, we give the following

**PRESCRIPTION.** If, in Theorem 1,  $\lim_k n_{k,j}^{-1} n_{k,j-1} = \bar{c}_j'$ ,  $2 \leq j \leq m$  with the  $\bar{c}_j'$  positive but not all bounded above by unity, let  $\pi$  be a permutation of  $(1, 2, \dots, m)$  which when applied to the entries of  $(1, \bar{c}_2', \bar{c}_2' \bar{c}_3', \dots, \bar{c}_2' \bar{c}_3' \cdots \bar{c}_m')$  arranges these in non-increasing order. Denote by  $\pi(j)$  the image of  $j$ ,  $1 \leq j \leq m$  and set  $\tilde{X}_{i, \pi(j)} = X_{ij}$ ,  $\tilde{S}_n^{(j)} = \sum_{i=1}^n \tilde{X}_{i, \pi(j)}$  and  $\tilde{n}_{k, \pi(j)} = n_{k,j}$ . Then, clearly  $\lim_{k \rightarrow \infty} \tilde{n}_{k,j}^{-1} \tilde{n}_{k,j-1} = \bar{c}_j$  exists and  $\bar{c}_j \leq 1$ ,  $2 \leq j \leq m$ . Again, setting  $c_j = (\bar{c}_2 \cdots \bar{c}_j)^{1/\alpha}$ , the limiting joint cdf of  $(\tilde{S}_{\tilde{n}_{k1}}^{(1)}/b(\tilde{n}_{k1}), \dots, \tilde{S}_{\tilde{n}_{km}}^{(m)}/b(\tilde{n}_{km}))$  is given by (2.2). Now, in each factor of (2.2), apply  $\pi^{-1}$  to the arguments of the  $m$  entries of  $\phi$  and then relabel the  $t$ 's (only) to agree with the positions they occupy. The resulting cf will be the limit cf of  $(S_{n_{k1}}^{(1)}/b(n_{k1}), \dots, S_{n_{km}}^{(m)}/b(n_{km}))$ . Note that  $c_j = (c_{\pi^{-1}(j)}' / (c_{\pi^{-1}(1)}'))^{1/\alpha}$ ,  $1 \leq j \leq m$  where  $c_j' = (\bar{c}_2' \bar{c}_3' \cdots \bar{c}_j')^{1/\alpha}$ .

In Theorem 1,  $b(n)$  may vary with  $j$  if  $n_{k,j-1} < n_{kj}$  and

$$\lim_{k \rightarrow \infty} [b_j(n_{k,j-1})/b_j(n_{k,j})] = \bar{c}_j < \infty, \quad \lim_{k \rightarrow \infty} [b_j(n_{kj} - n_{k,j-1})/b_j(n_{kj})] = d_j.$$

**THEOREM 2.** Let  $X_n, S_n$  be as in (1.1), (1.2) with (1.5) holding for  $b(n) = n^{\frac{1}{\alpha}}$  and  $F(x_1, x_2, \dots, x_m) = \Phi_{\Sigma}(x_1, x_2, \dots, x_m)$ , where the cf of  $\Phi_{\Sigma}$  is  $\phi_{\Sigma}(t_1, t_2, \dots, t_m) = \exp\{-\frac{1}{2} \sum_{i,j=1}^m \sigma_{ij} t_i t_j\}$ . If  $\{n_{kj}, 1 \leq j \leq m, k = 1, 2, \dots\}$  are  $m$  increasing sequences of positive integers with  $\lim_{k \rightarrow \infty} (n_{k,j-1}/n_{kj}) = \bar{c}_j < \infty$ , where without

essential loss of generality, it will be supposed that  $\bar{c}_j \leq 1$ ,  $2 \leq j \leq m$ , then as  $k \rightarrow \infty$ ,

$$(2.5) \quad P\{S_{n_{k1}}^{(1)} < x_1 n_{k1}^{\frac{1}{2}}, \dots, S_{n_{km}}^{(m)} < x_m n_{km}^{\frac{1}{2}}\} \Rightarrow \Phi_{\Sigma_c}(x_1, \dots, x_m)$$

where, setting  $c_1 = 1$ , the cf of  $\Phi_{\Sigma_c}$  is<sup>3</sup>

$$(2.6) \quad \phi_{\Sigma_c}(t_1, t_2, \dots, t_m) = \exp\{-\frac{1}{2}[\sum_{i=1}^m \sigma_{ii} t_i^2 + 2 \sum_{i=1}^m \sum_{j>i} (c_j/c_i) \sigma_{ij} t_i t_j]\},$$

and  $c_j = \prod_{i=2}^j (\bar{c}_i)^{\frac{1}{2}}$ . Moreover, if  $\Sigma$  is positive definite, so is  $\Sigma_c$  for all  $1 \geq c_2 \geq c_3 \geq \dots \geq c_m \geq 0$ ; if  $\sigma_{ii} > 0$ ,  $1 \leq i \leq m$  but  $\Sigma$  is singular, then  $\Sigma_c$  is positive definite for  $1 > c_2 > c_3 > \dots > c_m > 0$ .

PROOF. To obtain (2.6), apply Theorem 1 with  $b(n) = n^{\frac{1}{2}}$ , noting that  $d_j^2 = 1 - c_j^2/c_{j-1}^2$ ,  $2 \leq j \leq m$ . Next, write  $a_j = \bar{c}_j$  and  $\Sigma_c = \Sigma(a_2, a_3, \dots, a_m)$ . Even if  $\Sigma$  is positive definite, from Theorem 1 it follows merely that  $\Sigma(a_2, a_3, \dots, a_m)$  is positive semi-definite. On the other hand, positive definiteness of  $\Sigma = \Sigma(1, 1, \dots, 1)$  readily implies that of  $\Sigma(a_2, a_3, \dots, a_m)$  for  $a_j = 0$  or 1,  $2 \leq j \leq m$ . Further, for any integer  $j$  in  $[2, m]$  and any real  $\lambda$  in  $(0, 1)$ ,

$$\begin{aligned} \Sigma(a_2, \dots, \lambda a_j' + (1 - \lambda)a_j'', \dots, a_m) \\ = \lambda \Sigma(a_2, \dots, a_j', \dots, a_m) + (1 - \lambda) \Sigma(a_2, \dots, a_j'', \dots, a_m). \end{aligned}$$

Since a convex linear combination of positive definite matrices is positive definite,  $\Sigma(a_2, a_3, \dots, a_m)$  is positive definite throughout the cube  $0 \leq a_j \leq 1$ ,  $2 \leq j \leq m$  which is tantamount to the statement immediately following (2.6). On the other hand, if  $\Sigma$  is singular but  $\sigma_{ii} > 0$ ,  $1 \leq i \leq m$ , the positive definiteness of  $\Sigma(0, 0, \dots, 0)$  together with the fact that a convex linear combination (with  $0 < \lambda < 1$ ) of a positive semi-definite and positive definite matrix is positive definite, insures the positive definiteness of  $\Sigma(a_2, a_3, \dots, a_m)$  at all interior points of the cube.

If  $m = 2$ ,  $n_{k1} = r_k$ ,  $n_{k2} = n_k$ ,  $\lim_{k \rightarrow \infty} r_k/n_k = c = 0$  and  $\sigma_{12} \neq 0$ , the random variables  $r_k^{-\frac{1}{2}} S_{r_k}^{(1)}$  and  $n_k^{-\frac{1}{2}} S_{n_k}^{(2)}$  are asymptotically independent, even though  $n^{-\frac{1}{2}} S_n^{(1)}$  and  $n^{-\frac{1}{2}} S_n^{(2)}$  are not.

**3. Random sums of random vectors.** The question of the limit distribution of the sum of a random number of random variables has sought and received attention in the literature. Let  $\{N_n, n = 1, 2, \dots\}$  denote a sequence of positive integer-valued random variables and  $\{X_i, i = 1, 2, \dots\}$  a sequence of independent identically distributed random variables. In [8] (see also [4]), independence of  $N_n$  and  $X_i$  was postulated, the (normalized)  $N_n$  were assumed to converge in distribution and the limit distribution of  $\sum_{i=1}^{N_n} X_i$  was determined. In more recent works, the independence assumption has been abandoned at the small cost of strengthening the convergence of the  $N_n$ . Anscombe, [1], considers the case of a randomly selected sequence (specializing to a sum) and postulates that  $N_n \rightarrow_P$  constant. In [7], Rényi demonstrates that if  $EX_i = 0$ ,  $EX_i^2 = 1$ ,

<sup>3</sup> 0/0 is to be interpreted as 0.

a central limit theorem continues to hold for  $N_n^{-1} \sum_{i=1}^{N_n} X_i$  provided that  $N_n/n \rightarrow_P Y$  where  $Y$  is a positive discrete random variable. The special case of  $Y = \text{constant}$  was implicitly dealt with by Doeblin [3] in his work on Markov Chains. Extension to the case of a positive (but non-discrete) random variable  $Y$ , conjectured in [7], was achieved by Blum, Hanson and J. Rosenblatt in [2] and Mogyorodi [6].

The following theorem generalizes some of the preceding to stable limits and higher dimensions. Analogous results for  $m = 1$  appear in [11].

**THEOREM 3.** Let  $X_n$ ,  $S_n$ ,  $b(n)$ ,  $g(n)$  and  $N_n$  be as described in (1.1), (1.2), (1.3), (1.4) and suppose that (1.5) obtains with  $b(n) = n^{1/\alpha} h(n)$ ,  $0 < \alpha \leq 2$ , where  $h(n)$  is a slowly varying function. Further, let  $a_j$  be positive constants (without loss of generality, take  $a_1 \leq a_2 \leq \dots \leq a_m$ ) such that

$$(3.1) \quad N_n^{(j)} / g(n) \rightarrow_P a_j, \quad 1 \leq j \leq m.$$

Then

$$(3.2) \quad P\{\bigcap_{j=1}^m [S_{N_n^{(j)}}^{(j)} < x_j (N_n^{(j)})^{1/\alpha} h(N_n^{(j)})]\} \Rightarrow G_\alpha(x_1, \dots, x_m; c_2, \dots, c_m)$$

where the cf corresponding to  $G_\alpha$  is given by (2.2) with  $c_j = (a_1/a_j)^{1/\alpha}$  and  $d_j = (1 - a_{j-1}/a_j)^{1/\alpha}$ ,  $2 \leq j \leq m$ .

**PROOF.** Set  $n_{kj} = [a_j g(k)]$  where  $[y]$  signifies the largest integer  $\leq y$  and observe that with  $b(n) = n^{1/\alpha} h(n)$ ,  $c_j = (a_1/a_j)^{1/\alpha}$ ,  $d_j = (1 - a_{j-1}/a_j)^{1/\alpha}$ , all the conditions of Theorem 1 are met. Consequently, writing  $q(n; a_j) = ([a_j g(n)])^{1/\alpha} h([a_j g(n)])$ ,

$$(3.3) \quad S_{[a_j g(n)]}^{(j)} / q(n; a_j), \quad 1 \leq j \leq m,$$

have the limiting cdf  $G_\alpha$  whose cf is exhibited in (2.2).

On the other hand, we have for  $1 \leq j \leq m$  (paralleling Rényi [7])

$$(3.4) \quad \begin{aligned} S_{N_n^{(j)}}^{(j)} / q(N_n^{(j)}; 1) &= S_{[a_j g(n)]}^{(j)} / q(n; a_j) \\ &+ [q(n; a_j) / q(N_n^{(j)}; 1)] [(S_{N_n^{(j)}}^{(j)} - S_{[a_j g(n)]}^{(j)}) / q(n; a_j)] \\ &+ [S_{[a_j g(n)]}^{(j)} / q(n; a_j)] [q(n; a_j) / q(N_n^{(j)}; 1) - 1]. \end{aligned}$$

In view of (3.1), (3.3) and the slowly varying property, the last factor of the right hand side converges in probability to zero for  $1 \leq j \leq m$ .

Once it is shown that

$$(3.5) \quad (S_{N_n^{(j)}}^{(j)} - S_{[a_j g(n)]}^{(j)}) / q(n; a_j) \rightarrow_P 0, \quad 1 \leq j \leq m,$$

the second factor of the right hand side of (3.4) will likewise converge in probability to zero for  $1 \leq j \leq m$ . Thus, according to Lemma 1 the limit distributions of the two random vectors whose  $j$ th components are respectively the first terms of each side of the equality (3.4) will be identical and the theorem will have been proved.

To validate (3.5), we modify somewhat an argument of Rényi (and Doeblin). Let  $\delta > 0$  and define

$$B_{nj}(\delta) = \{|N_n^{(j)} - a_j \cdot g(n)| < \delta g(n)\},$$

$$A_{nj}(\epsilon) = \{|S_{N_n^{(j)}}^{(j)} - S_{[a_j g(n)]}^{(j)}| > \epsilon g(n; a_j)\}.$$

Now, for any  $\epsilon > 0$ , omitting the component index  $j$ , denoting the complement of a set  $B$  by  $B'$ , and setting  $r_n = [\delta \cdot g(n)]$ ,

$$P\{|S_{N_n} - S_{[ag(n)]}| > \epsilon g(n; a)\} \leq P\{A_n(\epsilon) \cdot B_n(\delta)\} + P\{B_n'(\delta)\}$$

$$\leq P\{\max_{1 \leq i \leq r_n} |S_i - S_{[ag(n)]}| > \epsilon g(n; a)\} + P\{B_n'(\delta)\}$$

$$\leq 2P\{\max_{1 \leq i \leq r_n} |S_i| > \epsilon(a \cdot r_n / 2\delta)^{1/\alpha} \cdot h([a \cdot r_n / 2\delta])\} + P\{B_n'(\delta)\}.$$

Since ([9], [10]) the limiting distribution, say  $Q(\cdot)$ , of  $(r_n^{1/\alpha} h(r_n))^{-1} \cdot \max_{1 \leq i \leq r_n} |S_i|$  exists and is identical with that of  $\sup_{0 < \tau < 1} |X_\alpha(\tau)|$  where  $X_\alpha(\tau)$  is a stable process of index  $\alpha$ , the first term of the right hand side differs (for sufficiently large  $n$ ) by an arbitrarily small amount from  $2(1 - Q(\epsilon(a/2\delta)^{1/\alpha}))$ . But for sufficiently small positive  $\delta$ , the latter is, in turn, arbitrarily small while for any fixed  $\delta > 0$ , the second term  $P\{B_n'(\delta)\}$  is less than any preassigned quantity for all but a finite number of values of  $n$ . Thus, (3.5) and hence the theorem holds.

**COROLLARY 1.** Let  $X_n, S_n, N_n, g(n)$  and  $a_j$  be as described in (1.1), (1.2), (1.4), (3.1) and suppose that (1.5) obtains with  $F(x_1, x_2, \dots, x_m) = \Phi_2(x_1, x_2, \dots, x_m)$  and  $b(n) = n^{1/2}$ . Then (3.2) holds, the cf of  $G_\alpha$  being given by (2.6) with  $c_j = (a_1/a_j)^{1/2}$ ,  $2 \leq j \leq m$ .

**COROLLARY 2.** Let  $\{X_n\}$  be a sequence of independent identically distributed random variables and  $h(n)$  a slowly varying function with  $P\{S_n < xn^{1/\alpha}h(n)\} \Rightarrow G_\alpha(x)$ , a stable cdf of characteristic exponent  $\alpha$  ( $0 < \alpha \leq 2$ ). Further, let  $N_n$  be a sequence of positive integer-valued random variables,  $g(n)$  as in (1.3) and  $N_n/g(n) \rightarrow_P a > 0$ . Then  $P\{S_{N_n} < x(N_n)^{1/\alpha}h(N_n)\} \Rightarrow G_\alpha(x)$ .

**THEOREM 4.** Let  $X_n, S_n, g(n), b(n), N_n$  be as in (1.1)–(1.4) and suppose that (1.5) obtains with  $b(n) = n^{1/\alpha}h(n)$ ,  $0 < \alpha \leq 2$  with  $h$  slowly varying; let  $Y^{(j)}$  be positive discrete random variables such that  $N_n^{(j)}/g(n) \rightarrow_P Y^{(j)}$ ,  $1 \leq j \leq m$ . Then

$$P\{\bigcap_{j=1}^m [S_{N_n^{(j)}}^{(j)} < x(N_n^{(j)})^{1/\alpha}h(N_n^{(j)})]\}$$

$$\Rightarrow \sum_{y_2, \dots, y_m} G_\alpha(x_1, \dots, x_m; y_2, \dots, y_m) \cdot Q(y_2, \dots, y_m)$$

where  $Q(y_2, \dots, y_m) = P\{\bigcap_{j=2}^m [Y^{(j)}/Y^{(1)} = y_j]\}$  and  $G_\alpha$  is the cdf of (2.1) with  $c_j^\alpha = y_j^{-1}$ ,  $d_j^\alpha = 1 - (y_{j-1}/y_j)$  when  $1 \geq y_m \geq \dots \geq y_2 > 0$  and otherwise  $G_\alpha$  is the modification of the prescription following Theorem 1.

**THEOREM 5.** Let  $X_n, S_n, b(n), N_n$  be as in (1.1)–(1.4) and suppose that (1.5) obtains for non-singular  $F$ , where  $b(n) = n^{1/\alpha}h(n)$ ,  $0 < \alpha \leq 2$ , with  $h$  slowly varying; let  $Y^{(j)}$  be positive random variables such that  $N_n^{(j)}/n \rightarrow_P Y^{(j)}$ ,  $1 \leq j \leq$

$m$ . Then if  $R^+ = \{(y_2, \dots, y_m) : y_j > 0, 2 \leq j \leq m\}$ ,

$$(3.6) \quad P\{\bigcap_{j=1}^m [S_{N_n^{(j)}}^{(j)} < x_j (N_n^{(j)})^{1/\alpha} h(N_n^{(j)})]\} \\ \Rightarrow \int_{R^+} G_\alpha(x_1, \dots, x_m; y_2, \dots, y_m) dQ(y_2, \dots, y_m)$$

where  $Q(y_2, \dots, y_m) = P\{\bigcap_{j=2}^m [Y^{(j)}/Y^{(1)} < y_j]\}$  and  $G_\alpha$  is the cdf of (2.1) for  $1 \geq y_m \geq \dots \geq y_2 > 0$  and the modification of the Prescription otherwise. Again,  $y_j^{-1} = c_j^\alpha$ ,  $d_j^\alpha = 1 - y_{j-1}/y_j$ .

The proof of Theorem 5 is the  $m$ -dimensional analogue of [2], modified (as in Theorem 3) by the employment of a limiting stable process rather than Kolmogoroff's inequality. Similarly, the proof of Theorem 4 is essentially that of Rényi, again circumventing Kolmogoroff's inequality as in Theorem 3.

It is trivial to reformulate Rényi's basic result on mixing sequences in a manner suitable for use in Theorem 4:

LEMMA 3. Let  $A_{k_{1n}, k_{2n}, \dots, k_{mn}}$  be events of positive probability where the  $k_{jn}$  are non-negative integers with  $\lim_{n \rightarrow \infty} k_{jn} = \infty$ ,  $1 \leq j \leq m$ . Then a necessary and sufficient condition that  $\lim_{n \rightarrow \infty} P\{B \cdot A_{k_{1n}, \dots, k_{mn}}\} = \alpha P\{B\}$  for all events  $B$  (where  $0 < \alpha < 1$ ) is that  $\lim_{n \rightarrow \infty} P\{A_{k_{1n}, \dots, k_{mn}} \cdot A_{k_{1i}, \dots, k_{mi}}\} = \alpha \cdot P\{A_{k_{1i}, \dots, k_{mi}}\}$  and  $\lim_{n \rightarrow \infty} P\{A_{k_{1n}, \dots, k_{mn}}\} = \alpha$ , all  $i = 1, 2, \dots$ .

However, it is not quite so straightforward to verify that the condition (and hence conclusion) of the lemma holds for

$$A_{k_{1n}, \dots, k_{mn}} = \{S_{k_{1n}}^{(1)}/b(k_{1n}) < x_1, \dots, S_{k_{mn}}^{(m)}/b(k_{mn}) < x_m\}, \quad n = 1, 2, \dots$$

Let  $I_A$  be the indicator (set characteristic function) of the event  $A$  and set  $k_{0i} = \max_{1 \leq j \leq m} k_{ji}$  and  $J_i = \{j : k_{ji} < k_{0i}\}$ . Then denoting conditional probabilities and expectations by  $P\{\cdot | \cdot\}$  and  $E\{\cdot | \cdot\}$ ,

$$\begin{aligned} \lim_{n \rightarrow \infty} P\{\bigcap_{j=1}^m [S_{k_{jn}}^{(j)} < x_j b(k_{jn})] | I_{[S_{k_{ji}}^{(j)} < x_j b(k_{ji})]}, 1 \leq j \leq m; S_{k_{0i}}^{(j)}, j \in J_i\} \\ = \lim_{n \rightarrow \infty} P\{\bigcap_{j=1}^m [S_{k_{jn}}^{(j)} - S_{k_{0i}}^{(j)} < x_j b(k_{jn})] | I_{[S_{k_{ji}}^{(j)} < x_j b(k_{ji})]}, 1 \leq j \leq m; \\ S_{k_{0i}}^{(j)}, j \in J_i\} \\ = \lim_{n \rightarrow \infty} P\{\bigcap_{j=1}^m [(S_{k_{jn}}^{(j)} - S_{k_{0i}}^{(j)})/b(k_{jn} - k_{0i}) < x_j b(k_{jn})/b(k_{jn} - k_{0i})]\} \\ = G_\alpha(x_1, x_2, \dots, x_m) \end{aligned}$$

if  $b(n) = n^{1/\alpha} h(n)$  with  $h$  slowly varying and the hypotheses of Theorem 1 hold. From the relation

$$\begin{aligned} P\{\bigcap_{j=1}^m [S_{k_{jn}}^{(j)} < x_j b(k_{jn})] | I_{[S_{k_{ji}}^{(j)} < x_j b(k_{ji})]}, 1 \leq j \leq m\} \\ = E[P\{\bigcap_{j=1}^m [S_{k_{jn}}^{(j)} < x_j b(k_{jn})] | I_{[S_{k_{ji}}^{(j)} < x_j b(k_{ji})]}, 1 \leq j \leq m; S_{k_{0i}}^{(j)}, j \in J_i\} \\ \cdot I_{[S_{k_{ji}}^{(j)} < x_j b(k_{ji})]}, 1 \leq j \leq m] \end{aligned}$$

and the dominated convergence theorem, it follows that

$$\lim_{n \rightarrow \infty} P\{A_{k_{1n}, \dots, k_{mn}} A_{k_{1i}, \dots, k_{mi}}\} = G(x_1, \dots, x_m)$$

and consequently  $\lim_{n \rightarrow \infty} P\{A_{k_{1n}, \dots, k_{mn}} \cdot B\} = G(x_1, \dots, x_m)P\{B\}$  for  $0 < G(x_1, \dots, x_m) < 1$ .

#### 4. Double sequences. Let

$$(4.1) \quad \begin{array}{c} X_{11} \\ X_{21}, X_{22} \\ \vdots \\ X_{n1}, X_{n2}, \dots, X_{nn} \\ \vdots \end{array}$$

constitute a double sequence of row-wise independent, identically distributed random vectors, say  $X_{ni} = (X_{ni}^{(1)}, X_{ni}^{(2)}, \dots, X_{ni}^{(m)})$ ,  $1 \leq i \leq n$ ,  $n = 1, 2, \dots$ . Consider analogously the row sum vectors

$$(4.2) \quad S_n = (S_n^{(1)}, S_n^{(2)}, \dots, S_n^{(m)}), \quad n = 1, 2, \dots,$$

where  $S_n^{(j)} = \sum_{i=1}^n X_{ni}^{(j)}$  and suppose that for  $b(n)$  as in (1.3), there exists a (necessarily infinitely divisible) cdf  $F(x_1, x_2, \dots, x_m)$  with cf  $\phi(t_1, t_2, \dots, t_m)$  such that

$$(4.3) \quad P\{S_n^{(1)} < x_1 b(n), \dots, S_n^{(m)} < x_m b(n)\} \Rightarrow F(x_1, x_2, \dots, x_m).$$

Then, we have

**THEOREM 6.** Let  $X_{ni}$ ,  $1 \leq i \leq n$ ,  $S_n$ ,  $b(n)$ ,  $F(x_1, x_2, \dots, x_m)$  be as described in (4.1), (4.2), (1.3) and (4.3). Further, let  $\{k_{jn}, 1 \leq j \leq m, n = 1, 2, \dots\}$  represent  $m$  increasing sequences of positive integers such that

$$(4.4) \quad \begin{array}{ll} \text{(i)} & k_{j,n} \leq n, \quad \lim_{n \rightarrow \infty} k_{j,n}/n = \delta_j \text{ where } 0 < \delta_1 < \dots < \delta_m \leq 1, \\ \text{(ii)} & \lim_{n \rightarrow \infty} b(n)/b(k_{jn}) = \beta_j, \quad 1 \leq j \leq m. \end{array}$$

Then

$$(4.5) \quad P\{S_{k_{1n}}^{(1)}/b(k_{1n}) < x_1, \dots, S_{k_{mn}}^{(m)}/b(k_{mn}) < x_m\} \Rightarrow G(x_1, x_2, \dots, x_m)$$

where the Fourier transform of the cdf  $G$  is given by

$$(4.6) \quad \phi^{\delta_1}(\beta_1 t_1, \beta_2 t_2, \dots, \beta_m t_m) \cdot \phi^{\delta_2 - \delta_1}(0, \beta_2 t_2, \dots, \beta_m t_m) \dots \phi^{\delta_m - \delta_{m-1}}(0, 0, \dots, 0, \beta_m t_m).$$

**PROOF.** Let  $\phi_n(t_1, t_2, \dots, t_m)$  denote the cf of  $X_{n1}$ . The hypothesis (4.3) implies that for all real  $t = (t_1, t_2, \dots, t_m)$ ,

$$\lim_{n \rightarrow \infty} \phi_n^n(t_1/b(n), \dots, t_m/b(n)) = \phi(t_1, t_2, \dots, t_m).$$

Since  $\phi$  is infinitely divisible and hence non-vanishing for real  $t$ , inside any fixed but arbitrary cube  $C = \{|t_i| \leq T, 1 \leq i \leq m\}$  the cf's  $\phi_n$  are non-vanishing for real  $t$  for all sufficiently large  $n$ . Thus, setting, for all  $n$ ,  $k_{0n} \equiv 0 = \delta_0$ ,



$$(4.7) \quad \lim_{n \rightarrow \infty} \phi^{k_{jn} - k_{j-1,n}}(t_1/b(n), \dots, t_m/b(n)) = \phi^{\delta_j - \delta_{j-1}}(t_1, t_2, \dots, t_m)$$

uniformly in  $C$  for  $1 \leq j \leq m$ .

Denote  $S_n^{(j)}$  by  $S^{(j)}(n)$  and  $\sum_{i=k_{un}+1}^{k_{u+1,n}} X_{ni}^{(j)}$  by  $S^{(j)}(k_{un}, k_{u+1,n})$ . For all sufficiently large  $n$ , (4.4)(i) implies

$$\begin{aligned} E[\exp \{i \sum_{j=1}^m t_j S^{(j)}(k_{jn})/b(k_{jn})\}] \\ = E[\exp \{i \sum_{j=1}^m t_j [b(n)/b(k_{jn})][S^{(j)}(k_{1n})/b(n)]\}] \\ \cdot E[\exp \{i \sum_{j=2}^m t_j [b(n)/b(k_{jn})][S^{(j)}(k_{1n}, k_{2n})/b(n)]\}] \\ \dots E[\exp \{i t_m [b(n)/b(k_{m,n})][S^{(m)}(k_{m-1,n}, k_{mn})/b(n)]\}]. \end{aligned}$$

Consequently, employing (4.7) and (ii) of (4.4)

$$\begin{aligned} \lim_{n \rightarrow \infty} E[\exp \{i \sum_{j=1}^m t_j S^{(j)}(k_{jn})/b(k_{jn})\}] \\ = \phi^{\delta_1}(\beta_1 t_1, \beta_2 t_2, \dots, \beta_m t_m) \cdot \phi^{\delta_2 - \delta_1}(0, \beta_2 t_2, \dots, \beta_m t_m) \\ \dots \phi^{\delta_m - \delta_{m-1}}(0, \dots, 0, \beta_m t_m). \end{aligned}$$

It is clear that extensions to the case of randomly chosen indices can be given under suitable assumptions.

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#### REFERENCES

- [1] ANSCOMBE, F. J. (1952). Large-sample theory of sequential estimation. *Proc. Cambridge Philos. Soc.* **48** 600-607.
- [2] BLUM, J. R., HANSON, D. L., and ROSENBLATT, J. I. (1963). On the central limit theorem for the sum of a random number of independent random variables. *Z. Wahrscheinlichkeitstheorie und Verw. Gebiete* **1** 389-393.
- [3] DOEBLIN, W. (1938). Sur deux problemes de M. Kolmogoroff concernant les chaines denombrables. *Bull. Soc. Math. France* 210-220.
- [4] DOBRUSIN, R. L. (1955). Lemma on the limit of compound random functions (in Russian). *Uspehi Mat. Nauk* **10** 157-159.
- [5] LÉVY, P. (1937). *Theorie de l'Addition des Variables Aleatoires*. Gautier-Villars, Paris.
- [6] MOGYORODI, J. (1962). A central limit theorem for the sum of a random number of independent random variables. *Publ. Math. Inst. Hungar. Acad. Sci.* **7** 409-424.
- [7] RÉNYI, A. (1960). On the central limit theorem for the sum of a random number of independent random variables. *Acta Math. Acad. Sci. Hungar.* **11** 97-102.
- [8] ROBBINS, H. (1948). Asymptotic distribution of the sum of a random number of random variables. *Bull. Amer. Math. Soc.* **54** 1151-1161.
- [9] SKOROHOD, A. V. (1957). Limit theorems for stochastic processes with independent increments. *Teor. Veroyatnost. i Primenen* **2** 145-177.
- [10] UDAGAWA, M. (1952). Asymptotic properties of some functionals of random variables. Reports of Statistical Application Research, Union of Japanese Scientists and Engineers.
- [11] WITTENBURG, H. (1964). Limiting distributions of random sums of random variables. *Z. Wahrscheinlichkeitstheorie und Verw. Gebiete* **3** 7-18.