

CONTINUUM PERCOLATION AND EUCLIDEAN MINIMAL SPANNING TREES IN HIGH DIMENSIONS

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We prove that for continuum percolation in \mathbb{R}^d , parametrized by the mean number y of points connected to the origin, as $d \rightarrow \infty$ with y fixed the distribution of the number of points in the cluster at the origin converges to that of the total number of progeny of a branching process with a Poisson(y) offspring distribution. We also prove that for sufficiently large d the critical points for the existence of infinite occupied and vacant regions are distinct. Our results resolve conjectures made by Avram and Bertsimas in connection with their formula for the growth rate of the length of the Euclidean minimal spanning tree on n independent uniformly distributed points in d dimensions as $n \rightarrow \infty$.

1. Introduction. In the simplest Boolean model of continuum percolation, balls of unit diameter are centered at the points of a homogeneous d -dimensional Poisson process, and one studies the connected components of the union of the balls. Equivalently, one studies the components of the graph on the Poisson points obtained by connecting each pair of points separated by a distance less than 1. This and related models have been studied in the contexts of geometric probability, statistics and physics. For a survey, see Meester and Roy [14]; also Alexander [2], Grimmett [9], Hall [10] and Penrose [17].

For $y > 0$, let $\mathcal{P}_y = \mathcal{P}_y(d)$ be a homogeneous Poisson process of rate y/c_d on \mathbb{R}^d , with an extra point inserted at 0, where $c_d = \pi^{d/2}/\Gamma((d/2) + 1)$ is the volume of the ball of unit radius in dimension d . The added point at the origin represents a “typical point” of the Poisson process. Let the 1-graph on \mathcal{P}_y be the graph with vertices at each point of \mathcal{P}_y , and edges between each pair of points X, Y of \mathcal{P}_y such that $|X - Y| < 1$, where $|\cdot|$ is the Euclidean norm. The parameter y is the mean degree of the vertex at 0 in this graph.

Let $C_y = C_y(d)$ denote the set of points of \mathcal{P}_y in the component containing 0 of the 1-graph on \mathcal{P}_y ; this is analogous to the cluster at the origin in lattice percolation. Let $|C_y|$ denote the cardinality of C_y , a random variable taking values in $\{1, 2, 3, \dots\} \cup \{\infty\}$. Set

$$(1) \quad f_k^{(d)}(y) := P[|C_y(d)| = k], \quad k = 1, 2, 3, \dots, \infty.$$

Here the symbol $:=$ denotes definition. Informally, $f_k^{(d)}(y)$ is the probability that a “typical point” of the Poisson process $\mathcal{P}_y \setminus \{0\}$ lies in a component of size k of the 1-graph on $\mathcal{P}_y \setminus \{0\}$. Thus, $f_k^{(d)}(y)$ is a natural object of interest in

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continuum percolation; see, for example, [15], [3] and (11) below. For $k < \infty$, there is an explicit formula for $f_k^{(d)}(y)$ [see [15], Proposition 1(a), or [3], (7)], in terms of a rather complicated k -fold integral; there is no known closed-form formula for $f_\infty^{(d)}(y)$. In this paper we derive simple limiting expressions for these probabilities when the dimension becomes large.

Let T_y denote the total number of progeny from a single ancestor of a Galton–Watson branching process with Poisson(y) offspring distribution. In other words, set $T_y = \sum_{n=0}^{\infty} Z_n$, where (Z_n) is the Galton–Watson process with generating function $G_y(s) = \exp(y(s-1))$, with $Z_0 = 1$. Our main results are that $|C_y|$ is stochastically dominated by T_y , but that the (possibly defective) distribution of $|C_y(d)|$ converges weakly to that of T_y as $d \rightarrow \infty$.

A consequence of our results is the resolution of some conjectures made by Avram and Bertsimas [3], who were interested in the distributions of $|C_y|$ and of T_y in connection with the series expansions they developed for the growth rates, as n becomes large, of certain random minimal spanning trees on n points, which we describe in detail in Section 3. Our results help to clarify the relationship between these growth rates.

Another consequence of our results is a proof that, for the Boolean model described above in sufficiently high dimensions, there is a range of values of y for which unbounded occupied and vacant regions coexist, partially resolving a conjecture in [14]. This is explained in Section 4.

In Section 5 we describe a sequential construction of a set equivalent to C_y , and give a heuristic explanation of the ideas behind our proofs. In Sections 6 and 7 we finally prove the main results.

2. The main results. Let $(f_k(y), k \geq 1)$ denote the probability function of T_y ; that is, define

$$(2) \quad f_k(y) := P[T_y = k] = \frac{k^{k-2}}{(k-1)!} y^{k-1} e^{-ky}, \quad k = 1, 2, 3, \dots$$

The equality in (2) can be verified using the formula of Dwass [8] for the distribution of T_y ; see also Jagers [11], Theorem 2.11.2, and Section 3 below.

Our results may be compared with the conjectures in Avram and Bertsimas [3], page 129. Our $f_k^{(d)}(y)$ is written there as $f_k^{(E)}(y)$, while our $f_k(y)$ is written there as $f_k^{(T)}(y)$. Part (b) of our first result confirms Conjecture 1 of [3].

PROPOSITION 1. (a) *Let $y > 0$. Then, for all d , $|C_y|$ is stochastically dominated by T_y . That is,*

$$(3) \quad \sum_{j=1}^k f_j^{(d)}(y) \geq \sum_{j=1}^k f_j(y) \quad \text{for all } k \in \mathbb{N}.$$

(b) *For $y > 0$,*

$$(4) \quad \lim_{d \rightarrow \infty} f_k^{(d)}(y) = f_k(y) \quad \text{for all } k \in \mathbb{N}.$$

Our main result says that the limit in (4) can be taken inside the infinite sum over $k \geq 1$. Thus, the distribution of $C_y(d)$ converges to that of T_y as $d \rightarrow \infty$.

THEOREM 1. *Let $y > 0$. Let $t = \psi(y)$ be the smallest positive solution to the equation $t = \exp(y(t - 1))$. Then*

$$(5) \quad \lim_{d \rightarrow \infty} \sum_{j=1}^{\infty} f_j^{(d)}(y) = \sum_{j=1}^{\infty} f_j(y) = \psi(y).$$

The second equality in (5) is simply the classical extinction probability theorem for the Galton–Watson process.

Theorem 1 shows that as $d \rightarrow \infty$ the percolation probability $f_{\infty}^{(d)}$ converges to the survival probability $1 - \psi(y)$ for the Galton–Watson process with Poisson(y) offspring distribution. Since $\psi(1) = 1$ and $\psi(y) < 1$ for $y > 1$, it is easily deduced, from Proposition 1(a) and Theorem 1, that the critical value $y_c(d) := \inf\{y: f_{\infty}^{(y)} > 0\}$, above which percolation can occur, satisfies

$$(6) \quad \lim_{d \rightarrow \infty} y_c(d) = 1.$$

Equations (5) and (6) are continuum analogues of the results of Kesten [12] for percolation on the nearest-neighbor integer lattice \mathbb{Z}^d . The problem of evaluating $y_c(d)$ arises both physically (see Hall [10], Section 4.7) and in cluster analysis (see Penrose [17]); in the latter application, d may be arbitrarily large.

The next result confirms the second part of Conjecture 2 of [3], and is immediate from Proposition 1(b) and Theorem 1, along with a routine weak convergence argument.

COROLLARY 1. *For each $K \in \mathbb{N}$,*

$$(7) \quad \lim_{d \rightarrow \infty} \sum_{k \geq K} k^{-1} f_k^{(d)}(y) = \sum_{k \geq K} k^{-1} f_k(y).$$

Finally, we show among other things that the first part of Conjecture 2 of [3] is false.

PROPOSITION 2. *Let $d \in \mathbb{N}$. For all $y \geq 0$, $f_1^{(d)}(y) = f_1(y)$. Also*

$$(8) \quad \sum_{k \geq K} k^{-1} f_k^{(d)}(y) \geq \sum_{k \geq K} k^{-1} f_k(y), \quad K = 1, 2,$$

but, for some $y_1 > 0$,

$$(9) \quad \sum_{k \geq 3} k^{-1} f_k^{(d)}(y) < \sum_{k \geq 3} k^{-1} f_k(y), \quad y < y_1.$$

The quantity $\sum_{k \geq 1} k^{-1} f_k^{(d)}(y)$ has physical significance as the “free energy,” or mean number per unit volume of components of the 1-graph on $\mathcal{P}_y \setminus \{0\}$, divided by the intensity y/c_d ; see Hall [10], Section 4.6, and Penrose [15], Theorem 2. Equations (7) and (8) show that this quantity is bounded below by $E[T_y^{-1}]$, but converges to this lower bound as $d \rightarrow \infty$.

In applications, it is often natural to consider the connected components of the union of balls of radius $1/2$, centered at the points of \mathcal{P}_y . Let V_y denote the volume of the component including 0 , which is just the union of such balls centered at the points of C_y . Clearly, $V_y \leq 2^{-d} c_d |C_y|$; a further result is that the distribution of $(2^d/c_d)V_y$ converges to that of T_y , as $d \rightarrow \infty$. This can be proved using (4), (5) and the proof of Lemma 2 below; we omit the details.

3. The relation to minimal spanning trees. Suppose each edge $\{i, j\}$ of the complete graph on $\{1, 2, \dots, n\}$ is assigned a weight $D_{\{i, j\}} = D_{\{j, i\}}$. The minimal spanning tree (MST) is the connected graph τ on vertices $\{1, 2, \dots, n\}$ which minimizes $\sum_{\{i, j\} \in \tau} D_{\{i, j\}}$. The quantity thus minimized is the total length L_n of the MST. Here we are interested in two particular cases where the $D_{\{i, j\}}$ are *random*.

In the *d-dimensional Euclidean* model, X_1, X_2, \dots, X_n are independent and uniformly distributed over $[0, 1]^d$. We set $D_{\{i, j\}} = |X_i - X_j|$. For this model, the total length L_n is known to be almost surely asymptotic to $\beta n^{(d-1)/d}$, where $\beta = \beta(d)$ is a positive constant that has not been determined exactly. See Steele [18]; the method of that paper shows that the mean total length of the tree is also asymptotic to $\beta n^{(d-1)/d}$.

In the *independent* model with parameters $d > 0$ and $c > 0$, the variables $D_{\{i, j\}}$, $1 \leq i < j \leq n$, are mutually independent, positive random variables with common distribution function $(F(x), x \geq 0)$ satisfying

$$(10) \quad F(x) = cx^d(1 + o(1)) \quad \text{as } x \rightarrow 0.$$

In this model, the mean total length of the MST is again asymptotic to $\beta n^{(d-1)/d}$, but now with a different (known) constant which we denote $\beta = \beta^{(I)}(d, c)$. The superscript *I* stands for “independent.”

Progress in the evaluation of the Euclidean constant $\beta(d)$, and in unifying the two models, has been made by Avram and Bertsimas [3]. They derived closely related series expansions for the constants $\beta(d)$ and $\beta^{(I)}(d, c)$, which we now describe.

For $z > 0$, let $P_{k,n}^{(d)}(z)$ [respectively, $P_{k,n}^{(d,c)}(z)$] denote the probability that if one removes all edges with $D_{\{i, j\}} > z$ from the complete graph on n points, a specified point lies in a component with exactly k vertices, for the d -dimensional Euclidean model (respectively, the independent model with parameters d and c). Then, for fixed d ,

$$(11) \quad P_{k,n}^{(d)} \left[\left(\frac{y}{nc_d} \right)^{1/d} \right] \rightarrow f_k^{(d)}(y) \quad \text{as } n \rightarrow \infty,$$

where c_d is the volume of the unit ball and $f_k^{(d)}$ is given by (1). For fixed d and c ,

$$P_{k,n}^{(d,c)} \left[\left(\frac{y}{nc} \right)^{1/d} \right] \rightarrow f_k(y) \quad \text{as } n \rightarrow \infty,$$

with f_k given by (2). These limits are intuitively clear from the relevant Poisson limit theorems; for explicit proofs, see [3]. For convenience, we restate the main result in [3] [note that (1) in Theorem 1 of that paper has a factor of k^{-1} missing from the right-hand side].

THEOREM 2 (Avram and Bertsimas [3]). *In the Euclidean model in dimension $d \geq 2$,*

$$(12) \quad \beta(d) = \lim_{n \rightarrow \infty} \frac{E[L_n]}{n^{(d-1)/d}} = (dc_d^{1/d})^{-1} \int_0^\infty \sum_{k=1}^\infty (k^{-1} f_k^{(d)}(y)) y^{(1/d)-1} dy,$$

while in the independent model with parameters $d > 0$ and $c > 0$,

$$(13) \quad \beta^{(I)}(d, c) = \lim_{n \rightarrow \infty} \frac{E[L_n]}{n^{(d-1)/d}} = (dc^{1/d})^{-1} \int_0^\infty \sum_{k=1}^\infty (k^{-1} f_k(y)) y^{(1/d)-1} dy.$$

From this result and Proposition 2 above, it follows that $\beta(d) \geq \beta^{(I)}(d, c_d)$. This too was conjectured in [3]. The corollary (7) to Theorem 1 implies that, for each y , for large d the integrand in (12) is close to that in (13); it is possible that, as $d \rightarrow \infty$, $\beta(d)$ is asymptotic to $\beta^{(I)}(d, c_d)$. We do not have a proof, but see Bertsimas and van Ryzin [4] for an asymptotic expression for $\beta(d)$ as $d \rightarrow \infty$.

4. Coexistence in the Boolean model. In this section we consider the Boolean model, which we shall denote \mathcal{B}_y , given by the union of balls of diameter 1 centered at the points of $\mathcal{P}_y \setminus \{0\}$. We call the connected components of \mathcal{B}_y the *occupied regions*, and the components of $\mathbb{R}^d \setminus \mathcal{B}_y$ the *vacant regions*. Let \mathcal{O}_y (respectively, \mathcal{V}_y) denote the occupied region containing 0 (respectively, the vacant region containing 0), so that just one of \mathcal{O}_y and \mathcal{V}_y is nonempty. Define the following critical points, at which phase transitions occur for this model:

$$(14) \quad \begin{aligned} y_c &= y_c(d) = \inf \{ y : P[\mathcal{O}_y \text{ is unbounded}] > 0 \}, \\ y_c^* &= y_c^*(d) = \sup \{ y : P[\mathcal{V}_y \text{ is unbounded}] > 0 \}. \end{aligned}$$

It is easily seen that the definition of $y_c(d)$ in (14) is consistent with that given just before (6) above.

The number of infinite occupied regions is almost surely 1 if $y > y_c$, and a.s. 0 if $y < y_c$. The number of infinite vacant regions is a.s. 1 if $y < y_c^*$ and a.s. 0 if $y > y_c^*$. Also, $y_c = y_c^*$ for $d = 2$, and

$$(15) \quad y_c \leq y_c^* \quad \text{if } d \geq 3.$$

All these results can be found in [14], except for (15), which seems to be new.

We sketch a proof of (15) for $d = 3$ (the proof for other values of d is similar). Let T denote the plane $\{(x_1, x_2, x_3) \in \mathbb{R}^3: x_3 = 0\}$. Then $\mathcal{B}_n \cap T$ is a Boolean model consisting of a collection of disks, centered at the points of a two-dimensional Poisson process of rate y/c_3 , and with independent identically distributed radii whose common distribution is that of $((1/4) - U^2)^{1/2}$, where U is uniform on the interval $(-1/2, 1/2)$. If (15) were false, then for y in the interval (y_c^*, y_c) , both $\mathcal{C}_n \cap T$ and $\mathcal{V}_n \cap T$ would a.s. have no infinite component. But this cannot happen because of known results for the two-dimensional Boolean model ([14], Theorem 4.4).

It is conjectured in [14], Chapter 4, that the inequality (15) is strict, which would imply that for an interval of values of y there coexist a.s. an unbounded occupied region *and* an unbounded vacant region. This is believed to be true, both because of the outcome of simulation studies and because of analogous known results for lattice percolation [6]. We can confirm that the conjecture is true, at least for sufficiently large values of d .

THEOREM 3. (a) *Let p_c be the critical probability for Bernoulli site percolation on the nearest-neighbor lattice \mathbb{Z}^2 . Then, for $d \geq 2$,*

$$(16) \quad y_c^*(d) \geq 2^{d-2} \log\left(\frac{1}{p_c}\right) \pi^{1/2} \left(\frac{\Gamma((d+1)/2)}{\Gamma((d+2)/2)}\right).$$

Consequently, there exists $d_0 \in \mathbb{N}$ such that $y_c^ > y_c$ for $d \geq d_0$.*

PROOF. For each $(i, j) \in \mathbb{Z}^2$, let $\Gamma_{i,j} \subset \mathbb{R}^d$ be the cube of side 1 centered at $(i, j, 0, \dots, 0)$. For each $(i, j) \in \mathbb{Z}^2$, we define two perpendicular cylinders $U_{i,j}$ and $V_{i,j}$ contained in $\Gamma_{i,j}$. Let $U_{0,0}$ (respectively $V_{0,0}$) be the cylinder of radius $\frac{1}{2}$ and height 1, with its axis given by the line segment from $(-\frac{1}{2}, 0, 0, \dots, 0)$ to $(\frac{1}{2}, 0, 0, \dots, 0)$ [respectively, from $(0, -\frac{1}{2}, 0, \dots, 0)$ to $(0, \frac{1}{2}, 0, \dots, 0)$]. In other words, set

$$U_{0,0} = \left\{ x = (x_1, \dots, x_d) \in \Gamma_{0,0}: \sum_{i=2}^d x_i^2 \leq \frac{1}{4} \right\},$$

$$V_{0,0} = \left\{ x = (x_1, \dots, x_d) \in \Gamma_{0,0}: x_1^2 + \sum_{i=3}^d x_i^2 \leq \frac{1}{4} \right\}.$$

Using the notation of (17) below, let $U_{i,j} = (i, j, 0, \dots, 0) + U_{0,0}$, and let $V_{i,j} = (i, j, 0, \dots, 0) + V_{0,0}$.

We say $(i, j) \in \mathbb{Z}^2$ is *good* if $U_{i,j} \cup V_{i,j}$ contains no points of $\mathcal{P}_y \setminus \{0\}$. If (i, j) is good, then the straight-line segment from $(i - \frac{1}{2}, j, 0, \dots, 0)$ to $(i + \frac{1}{2}, j, 0, \dots, 0)$ is contained in a vacant region, as is the line segment from $(i, j - \frac{1}{2}, 0, \dots, 0)$ to $(i, j + \frac{1}{2}, 0, \dots, 0)$. Therefore, if there is an infinite connected set of good sites (i, j) in \mathbb{Z}^2 , there is an infinite connected vacant region.

Since the volume of $U_{i,j} \cup V_{i,j}$ is less than $2(2^{-(d-1)}c_{d-1})$,

$$P[(i, j) \text{ is good}] \geq \exp(-2^{2-d}(c_{d-1}/c_d)y).$$

Therefore, by a comparison with independent Bernoulli site percolation on \mathbb{Z}^2 ,

$$y \leq y_c^* \quad \text{if } \exp(-2^{2-d}(c_{d-1}/c_d)y) > p_c,$$

so that $y_c^* \geq 2^{d-2} \log(1/p_c)(c_d/c_{d-1})$, which yields (16).

It follows from (16) that $y_c^*(d) \rightarrow \infty$ as $d \rightarrow \infty$. A comparison with (6) shows that $y_c(d) < y_c^*(d)$ for sufficiently large d . \square

5. Sequential construction of the cluster at 0. In this section we describe a random algorithm for generating the cluster C_y . The algorithm also applies to the following more general setting. A bounded open set $B_1 \subset \mathbb{R}^d$ is specified, symmetric in the sense that $-x \in B_1$ for all $x \in B_1$. Let the B_1 -graph on the points of \mathcal{P}_y be given by putting edges between every $X, Y \in \mathcal{P}_y$ with $X - Y \in B_1$. Let C_{B_1} be the component of this graph including the point 0, and let $|C_{B_1}|$ be the number of vertices of this component.

For $x \in \mathbb{R}^d$, define the translate

$$(17) \quad x + B_1 := \{x + w : w \in B_1\}.$$

To generate C_{B_1} sequentially, start with the set $\{0\}$. Add the points of \mathcal{P}_y in B_1 , say X_1, \dots, X_N . Next add the points of \mathcal{P}_y in $(X_1 + B_1) \setminus B_1$ (a Poisson process), and then the points in $(X_2 + B_1) \setminus (B_1 \cup (X_1 + B_1))$, and so on. At each stage we add the points of \mathcal{P}_y in a region that is disjoint from those already considered, so by the independence properties of \mathcal{P}_y , at each stage we add a new Poisson process on the relevant region.

The proofs of the results in Section 2 are based on this procedure. We describe it more formally by the following algorithm, which generates a sequence (F_n) of random finite subsets of \mathbb{R}^d , with $F_1 \subset F_2 \subset F_3 \subset \dots$, and a second increasing sequence $G_1 \subset G_2 \subset G_3 \subset \dots$, with $G_i \subset F_i$ for each i . Set $A_i := \bigcup_{x \in G_i} (x + B_1)$, which represents the region of space already considered before stage i (the construction of sets F_{i+1} and G_{i+1}), and F_i represents the set of points of \mathcal{P}_y in A_i .

Initially set $F_1 = \{0\}$ and $G_1 = A_1 = \emptyset$, the empty set. Then at each successive stage $i = 1, 2, 3, \dots$ of the algorithm, perform the following two operations:

1. Select $x_i \in F_i \setminus G_i$, and set $G_{i+1} = G_i \cup \{x_i\}$.
2. Writing $|B_1|$ for the Lebesgue measure of B_1 , let N_i be a Poisson variable with mean $(y/c_d)|B_1|$, and let N_i points be placed independently and uniformly in $x_i + B_1$. These points are to be viewed as the *offspring* of x_i . Throw away those offspring that lie in A_i . Let F_{i+1} be the union of the remaining set of offspring [a Poisson process of rate y/c_d on $(x_i + B_1) \setminus A_i$] with F_i .

The algorithm is terminated at stage i if $G_i = F_i$. In this case, set $F_j = F_i$ and $G_j = G_i$ for all $j > i$. In any event, set $F_\infty = \bigcup_{i=1}^\infty F_i$ and $|F_\infty| = \text{card}(F_\infty)$.

In step 1 of the i th stage, the choice of x_i from possibly several candidates is to be given by a deterministic rule, based on the positions and family trees of the candidates. Any such rule will be said to specify a *version* of the algorithm.

Intuitively, the Poisson process on $(x_i + B_1) \setminus A_i$ may be viewed as the set of points of \mathcal{P}_y in that set, so that the result of stage i is to add the points of \mathcal{P}_y in $(x_i + B_1) \setminus A_i$. Thus, the random set F_∞ should have the same distribution as C_{B_1} (when finite) or a subset of C_{B_1} (when infinite). Therefore, the following result is clear, given sufficient intuition about the Poisson process. A formal proof (not given here) can be effected by a discretization argument; see [13].

LEMMA 1. For $k = 0, 1, 2, 3, \dots$ and $k = \infty$, $P[|F_\infty| = k] = P[|C_{B_1}| = k]$.

We shall also require the following modification of the above algorithm. Specify some open set B_2 with $B_2 \supset B_1$. In step 2 of the algorithm, set $A'_i := \bigcup_{x \in F_i \setminus \{x_i\}} (x + B_2)$ (so that $A'_i \supset A_i$), and throw away all offspring lying in A'_i (rather than A_i). Let F'_∞ be the set of points produced by this modified algorithm. This has the distribution of a subset of C_{B_1} . Indeed, there is a coupling in which F'_∞ is a subset of the set F_∞ produced by a version of the original algorithm, and therefore

$$(18) \quad P[|F'_\infty| = \infty] \leq P[|C_{B_1}| = \infty].$$

Heuristics. Here is a rough guide to the ideas behind the proofs of Proposition 1 and Theorem 1. Take B to be the unit ball in the above algorithm, so that the Poisson variables N_i each have mean y . Assume we use a version of the algorithm in which x_i is selected from the earliest available “generation” of points of F_i . Then an appropriate subsequence of the sequence F_1, F_2, \dots would form a simple branching random walk (BRW), if it were not for the necessary throwing away of points, a phenomenon we denote “interference.”

The path traced out in \mathbb{R}^d by the above BRW consists of edges which in high dimensions are likely to be (1) of length only just smaller than 1, and (2) mutually orthogonal. Therefore, the probability of any point being discarded before generation k vanishes as $d \rightarrow \infty$. For example, each point is likely to be at a distance approximately $\sqrt{2}$ from its grandparent; hence, interference due to grandparents is unlikely since $\sqrt{2} > 1$. This argument will give us Proposition 1.

The proof of Theorem 1 is harder since one is required to consider infinitely many generations. We lay out a grid of squares on \mathbb{R}^2 , and project \mathbb{R}^d onto \mathbb{R}^2 in such a way that the projected BRW has successive steps which are approximately standard bivariate normals. Assuming $y > 1$, one can fix k and m such that, with high probability for the projected BRW starting from m particles in one of the squares, there are m particles of generation k in each of the neighboring squares. We run a sequence of BRW's indexed by the squares of the grid, starting with the square at the origin, each BRW running for k generations. The zeroth generation of the BRW for square 0 is the single point at 0, and for each subsequent square the zeroth generation

of the corresponding BRW consists (if possible) of m points lying “above” that square (in terms of the projection), taken from the k th generation of one of the previous BRW’s. The whole collection of BRW’s is a sort of pruned BRW starting from the origin.

We say a square is “occupied” if the BRW for the corresponding step of the algorithm produces at least m particles at the k th generation lying above each neighboring square. The probability that the square 0 is occupied is close to $1 - \psi(y)$, while other squares have probability close to 1 of being occupied.

To make this algorithm really correspond to a version of the algorithm above for producing F_∞ , we must throw some of the particles of the BRW’s away. The key to the proof is the fact that the probability of interference is small at each step, and therefore, even with interference taken into account, the probability that the algorithm produces an infinite path from 0 is close to $1 - \psi(y)$ by a comparison with (oriented) site percolation of occupied squares on the grid.

At each step the BRW runs for only a fixed finite number of generations. By the same argument as for Proposition 1, for large d the chances of interference due to any particular previous step are small. Finally, the chance of interference due to far-away steps can be made 0, if we truncate the (two-dimensional normal) distribution of the steps of the projected random walk at some large range ρ ; indeed, the projected BRW starting from a given square can then only spread within a range ρk from that square, and so cannot interfere with any BRW from a distance greater than $2\rho k$ from that square. If ρ is sufficiently large, the truncation need not invalidate any of the earlier parts of the argument.

6. Proof of propositions.

PROOF OF PROPOSITION 1. In this section and the next, let B be the unit ball $\{x \in \mathbb{R}^d: |x| < 1\}$. Let $(Z_n^d, n = 0, 1, 2, \dots)$ be the following BRW in \mathbb{R}^d . Here Z_n^d is a random subset of \mathbb{R}^d . Each element (or “particle”) of Z_n^d is replaced in Z_{n+1}^d by a Poisson number of offspring with mean y , and the offspring of a particle at x are uniformly distributed over $x + B$. Set Z_0 to be a single particle at 0. The distribution of the total progeny $\sum_{n=0}^{\infty} \text{card}(Z_n^d)$ is that of T_y .

Let the particles of this BRW be ordered as follows; all members of an earlier generation precede all members of a later generation, and the members of any particular generation are ranked in order of increasing modulus. Modify the BRW by throwing away any particle X that lies in any translate of B centered at a particle of the BRW that precedes the parent of X in the above ordering. Also throw away all descendants of any particle thus discarded.

It is not hard to see that this construction is equivalent to a version of the algorithm described in Section 5 (with $B_1 = B$). The throwing away of branches of the BRW corresponds to the discarding of offspring in A_i at stage i of that algorithm. The construction shows by an explicit coupling that $|F_\infty|$ given by the algorithm is stochastically dominated by T_y , and therefore, by Lemma 1, so is $|C_y|$.

To prove part (b), that $f_k^{(d)}(y) \rightarrow f_k(y)$ as $d \rightarrow \infty$, it suffices to prove the following lemma.

LEMMA 2. *Let $y \geq 0$ and $k \in \mathbb{N}$. Then $\lim_{d \rightarrow \infty} (\sum_{j=1}^k f_j^{(d)}(y) - \sum_{j=1}^k f_j(y)) = 0$.*

PROOF. The quantity inside the limit is the probability that, in the above construction, (i) the BRW has total progeny greater than k , and (ii) after throwing away some of the branches in the prescribed manner, we are left with at most k particles. This event is contained in the complement of the event, denoted E_k , that none of the first k particles in the above ordering of the BRW is thrown away.

Let W_i denote the position of the i th particle of the BRW (in the given ordering). Let E'_i denote the event that (i) none of the offspring of W_i lies in the union of the unit balls centered at its predecessors W_1, \dots, W_{i-1} ; (ii) no two of its offspring are separated by a distance less than 1; and (iii) each of its offspring lies outside the Euclidean ball of radius $3/4$ centered at W_i .

Set $F_k = \bigcap_{i=1}^k E'_i$. Clearly, $E_k \supset F_{k-1}$. The next lemma shows that $P[F_1] \rightarrow 1$ as $d \rightarrow \infty$, and that, for each $k > 1$, $P[F_k | F_{k-1}] \rightarrow 1$ as $d \rightarrow \infty$; therefore, $P[E_k] \rightarrow 1$, which completes the proof.

LEMMA 3. *Suppose $\mathbf{X}(d)$ and $\mathbf{Y}(d)$ are independent and uniformly distributed on B . Then*

$$(19) \quad \lim_{d \rightarrow \infty} P[|\mathbf{X}(d)| > 3/4] = 1,$$

$$(20) \quad \lim_{d \rightarrow \infty} (\sup\{P[|\mathbf{X}(d) - x| \leq 1]: x \in \mathbb{R}^d, |x| \geq 3/4\}) = 0$$

and

$$(21) \quad \lim_{d \rightarrow \infty} P[|\mathbf{X}(d) - \mathbf{Y}(d)| \leq 1] = 0.$$

PROOF. The first conclusion (19) is trivial. To prove (20), note that

$$|\mathbf{X}(d) - x|^2 = |x|^2 + |\mathbf{X}(d)|^2 - 2|\mathbf{X}(d) \cdot x|.$$

By (19), it suffices to prove that $\mathbf{X}(d) \cdot x$ converges to 0 in probability, uniformly on $\{x: (3/4) \leq |x| \leq 2\}$. Write $\mathbf{X}(d)$ in coordinates as $(X^1(d), X^2(d), \dots, X^d(d))$. By symmetry, $\mathbf{X}(d) \cdot x$ has the same distribution as $|x|X^1(d)$. Again, by symmetry, $E[|X^1(d)|^2] \leq 1/d$, so $X^1(d)$ converges to 0 in L^2 , hence in probability.

Finally, (21) follows at once from (19) and (20). \square

PROOF OF PROPOSITION 2. Clearly, $|C_y| = 1$ if and only if there are no points of \mathcal{P}_y in the unit ball centered at 0. Therefore, $f_1^{(d)}(y) = \exp(-y) = f_1(y)$. The inequality

$$\sum_{k \geq K} k^{-1} f_k^{(d)}(y) \geq \sum_{k \geq K} k^{-1} f_k(y)$$

holds for $K = 1$ by the stochastic domination in part (a) of Proposition 1. It then holds also for $K = 2$ because $f_1^{(d)}(y) = f_1(y)$. However, we now show that, for fixed d and small enough y , the inequality is false for $K = 3$.

As $y \rightarrow 0$, $f_3(y)$ and $f_3^{(d)}(y)$ are both $O(y^2)$, whereas $\sum_{k>3} f_k(y)$ and $\sum_{k>3} f_k^{(d)}(y)$ are $O(y^3)$. Therefore, it suffices to show that there are constants $c_1 < c_2$ such that, as $y \rightarrow 0$,

$$(22) \quad \sum_{k \geq 3} f_k^{(d)}(y) \sim c_1 y^2,$$

while

$$(23) \quad \sum_{k \geq 3} f_k(y) \sim c_2 y^2.$$

Let N_1 and N_2 be independent Poisson(y) variables, where N_1 represents the size of the first generation and N_2 represents the number of offspring of the (only) member of the first generation in the case where $N_1 = 1$. Then

$$\begin{aligned} \sum_{k \geq 3} f_k(y) &= P[N_1 \geq 2] + P[N_1 = 1, N_2 \geq 1] \\ &= (e^{-y} y^2 / 2) + e^{-2y} y^2 + o(y^2), \end{aligned}$$

so that (23) holds with $c_2 = 3/2$.

Set $\gamma := P[|\mathbf{X} + \mathbf{Y}| > 1]$, where \mathbf{X} and \mathbf{Y} are independent and uniform on the unit ball. Consider the construction of F_∞ via a BRW, as described in the previous section. In this construction, no members of generation 1 will be thrown away (as their parent has no predecessor). However, given that generations 1 and 2 of the BRW are of size 1, the particle in generation 2 is discarded with probability $1 - \gamma$. Therefore,

$$\sum_{k \geq 3} f_k^{(d)}(y) = P[N_1 \geq 2] + \gamma P[N_1 = 1] P[N_2 = 1] + o(y^2),$$

so that (22) holds with $c_1 = (1/2) + \gamma$. \square

7. Proof of Theorem 1. Define the linear mapping $L: \mathbb{R}^d \rightarrow \mathbb{R}^2$ by

$$(24) \quad L(x_1, x_2, \dots, x_d) = d^{1/2}(x_1, x_2).$$

The next result is an extension of the well-known fact that a single coordinate of a uniformly distributed variable on a high-dimensional sphere is approximately normal.

LEMMA 4. *Suppose $\mathbf{X}(d) = (X_1(d), \dots, X_d(d))$ is uniformly distributed over $\{x \in \mathbb{R}^d: |x| < 1\}$. Then the two-dimensional random vector $L(\mathbf{X}(d))$ converges in distribution to the bivariate normal distribution, written $N(0, I)$, with mean 0 and the identity matrix I as covariance matrix.*

PROOF. Let $Z_i, i \geq 1$, be independent $N(0, 1)$ random variables. For each d , set $Y_i(d) = d^{-1/2}Z_i$, set $\mathbf{Y}(d) = (Y_1(d), Y_2(d), \dots, Y_d(d))$ and $R_d = |\mathbf{Y}(d)|$. Let U_d be an independent random variable, distributed over $(0, 1)$ with density $dy^{d-1}, 0 < y < 1$. By the strong law of large numbers, $R_d \rightarrow 1$ as $d \rightarrow \infty$, so that U_d/R_d converges to 1 in probability.

Set $\mathbf{X}(d) = (U_d/R_d)\mathbf{Y}(d)$. Then $\mathbf{X}(d)$ is uniformly distributed over the unit ball. Since $L(\mathbf{Y}(d))$ has the $N(0, I)$ distribution, it follows that $L(\mathbf{X}(d)) = (U_d/R_d)L(\mathbf{Y}(d)) \rightarrow N(0, I)$ in distribution (see Billingsley [5], page 28, exercise 1). \square

PROOF OF THEOREM 1. It follows from (3) in Proposition 1 that

$$\liminf_{d \rightarrow \infty} \sum_{j=1}^{\infty} f_j^{(d)}(y) \geq \psi(y).$$

Therefore, it suffices to show that, for any y with $\psi(y) < 1$ (i.e., $y > 1$) and any $\varepsilon > 0$,

$$(25) \quad \limsup_{d \rightarrow \infty} \sum_{j=1}^{\infty} f_j^{(d)}(y) \leq \psi(y) + 3\varepsilon.$$

With $\mathbf{X}(d)$ as in Lemma 4 and with $\mathbf{Z} \sim N(0, I)$, the two-dimensional standard normal, for $\rho > 0$ define $y_\rho(d)$ and $y_\rho(\infty)$ by

$$(26) \quad P[|L(\mathbf{X}(d))| < \rho] = y_\rho(d)/y, \quad P[|\mathbf{Z}| < \rho] = y_\rho(\infty)/y.$$

We shall show that ρ can be chosen so that, for large d , with probability close to $1 - \psi(y)$, there is an infinite path from 0 in the 1-graph on \mathcal{S}_y , consisting of Poisson points Y_i satisfying $|L(Y_{i+1} - Y_i)| < \rho$. In other words, we consider the B^ρ -graph on \mathcal{S}_y , where we set

$$(27) \quad B^\rho := \{x \in \mathbb{R}^d: |x| < 1 \text{ and } |L(x)| < \rho\}.$$

We shall show there exists ρ such that, for large d ,

$$(28) \quad P[C_{B^\rho} \text{ is infinite}] \geq 1 - \psi(y) - 2\varepsilon,$$

and (25) will then follow at once.

The argument to prove (28) is similar to that in Penrose [16] and elsewhere, a comparison with oriented percolation on the lattice $\mathcal{L} := \{(i, j) \in \mathbb{Z}^2: i \geq 0, |j| \leq i, (i + j)/2 \in \mathbb{Z}\}$, with oriented edges from (i, j) to $(i + 1, j \pm 1)$. We can (and do) choose $\delta \in (0, \varepsilon/3)$ such that, for oriented site percolation on \mathcal{L} with parameter $p \geq 1 - 5\delta$, the probability exceeds $1 - \varepsilon$ that there is an infinite path from 0 of occupied sites in \mathcal{L} . See Durrett [7]. For $(i, j) \in \mathcal{L}$, let $A_{i,j}$ be the closed square of side 1 in \mathbb{R}^2 centered at (i, j) .

Let $(Z_n^d, n = 0, 1, 2, \dots)$ be the d -dimensional branching random walk (BRW) described in Section 6, in which a particle at x has a $\text{Poisson}(y)$ number of offspring, uniformly distributed on the unit ball $x + B$. Let $Z_n^{d,\rho}$ be the BRW of the same form, except that the number of offspring is Poisson with mean $y_\rho(d)$, and the offspring are placed uniformly on $x + B^\rho$.

Let Z_n^∞ be the BRW in \mathbb{R}^2 with the $\text{Poisson}(y)$ offspring distribution, whose steps have the $N(0, I)$ distribution. Let $Z_n^{\infty,\rho}$ be the BRW in \mathbb{R}^2 with the $\text{Poisson}(y_\rho(\infty))$ offspring distribution, whose steps have the distribution of an $N(0, I)$ variable conditioned to lie in the disk of radius ρ . For $A \subset \mathbb{R}^2$, let $Z_n^\infty(A)$ denote the number of particles of the n th generation of the BRW (Z_n^∞) in A , and likewise for $Z_n^{\infty,\rho}$, Z_n^d and $Z_n^{d,\rho}$.

Let $y > 1$. By the proof of Lemma 2 of [16] (but see the remark below), there exist $m > 0$ and $k > 0$ such that, for sufficiently large d ,

$$(29) \quad P[Z_k^\infty(A_{1,1}) \geq m \text{ and } Z_k^\infty(A_{1,-1}) \geq m] > 1 - \delta \quad \text{if } |Z_0^\infty(A_{0,0})| \geq m.$$

The proof of (3.3) of [16] uses the multivariate local limit theorem, but this is not required here because the individual steps of the BRW (Z_n^∞) are already normal, and the distribution of the sum of independent normal variables is known exactly. This was pointed out to the author by R. Meester.

By the proof of Lemma 3 of [16] (again, a slight simplification is possible here), there exists $k_1 > 0$ such that, for sufficiently large d ,

$$(30) \quad P[Z_{k_1}^\infty(A_{1,1}) \geq m \text{ and } Z_{k_1}^\infty(A_{1,-1}) \geq m] > 1 - \psi(y) - \delta \quad \text{if } Z_0^\infty = \{0\}.$$

By an obvious coupling (discard all offspring at a distance greater than ρ from the parent), given the same initial value the point process $Z_n^{\infty,\rho}$ converges weakly to Z_n^∞ as $\rho \rightarrow \infty$, in the sense of, for example, Aldous and Steele [1]. Therefore, it is possible to take ρ to be so big that (29) and (30) are still true with the process (Z_n^∞) replaced by $(Z_n^{\infty,\rho})$. Similarly, by Lemma 4 we can then find d_0 such that for $d \geq d_0$ they still hold with $(Z_n^{\infty,\rho})$ replaced by the image under L of $(Z_n^{d,\rho})$; that is, there exist ρ and d_0 such that, for $d \geq d_0$,

$$(31) \quad P[Z_k^{d,\rho}(L^{-1}(A_{1,1})) > m \text{ and } Z_k^{d,\rho}(L^{-1}(A_{1,-1})) > m] > 1 - \delta, \\ \text{if } |Z_0^{d,\rho}(L^{-1}(A_{0,0}))| \geq m,$$

and

$$(32) \quad P[Z_{k_1}^{d,\rho}(L^{-1}(A_{1,1})) > m \text{ and } Z_{k_1}^{d,\rho}(L^{-1}(A_{1,-1})) > m] > 1 - \psi(y) - \delta, \\ \text{if } Z_0^{d,\rho} = \{0\}.$$

The value of ρ remains fixed for the remainder of the proof.

Choose k_2 so large that for a Galton–Watson process (Z_n) with generating function G_y , the total number of progeny in the first k generations (from m ancestors) or in the first k_1 generations (from one ancestor) is unlikely to

exceed k_2 , that is, so that

$$(33) \quad \max \left(P \left[\sum_{n=0}^k Z_n > k_2 \mid Z_0 = m \right], P \left[\sum_{n=0}^{k_1} Z_n > k_2 \mid Z_0 = 1 \right] \right) < \delta.$$

We now describe an algorithm which may be viewed as a version of the modified algorithm for generating F'_∞ in Section 5, taking B_1 to be B^ρ and B_2 to be the unit ball B . The aim is to show that the algorithm continues forever, so that $|F'_\infty| = \infty$, with probability at least $1 - \psi(y) - 2\varepsilon$; this will give us (28) by the corollary (18) of Lemma 1. The algorithm consists of a sequence of steps, indexed by the sites (i, j) of the lattice \mathcal{L} . Each step consists of a copy of the BRW $(Z_n^{d,\rho})$, run for a finite number of generations. The steps are to be performed in the order $(0, 0), (1, -1), (1, 1), (2, -2), (2, 0), (2, 2), (3, -3), (3, -1), \dots$. Step $(0, 0)$ is special since here the BRW runs for k_1 generations and starts from a single point at 0. In each subsequent step (i, j) , the BRW runs for k generations and starts from a set of m points in $L^{-1}(A_{i,j})$. Initially, we set each site (i, j) of \mathcal{L} to be “vacant,” but we shall change its status to “occupied” if step (i, j) is “successful.”

Step $(0, 0)$ is to run a BRW $Z_n^{d,\rho}, n = 0, 1, \dots, k_1$, with $Z_0^{d,\rho} = \{0\}$. Order the particles of the BRW as follows. Particles of an earlier generation are to precede those of a later one. Within generation $n, n > 0$, particles with distinct parents inherit the ordering placed on their parents, while siblings are ordered by modulus.

Now modify this branching process as follows. Consider successively each point X of the BRW after generation 0 (in the ordering given, starting with the point of smallest modulus in $Z_1^{d,\rho}$), and remove X (along with its descendants) if it lies in any translate of B centered at an (unremoved) particle that precedes X in the ordering, and is not a sibling or parent of X . If at the end of this procedure more than k_2 particles remain, remove all but the first k_2 of the remaining particles in the ordering. Let the set F consist of all remaining particles in generations $0, 1, 2, \dots, k_1$. This corresponds to the set F_i in the algorithm of Section 5, at the end of step $(0, 0)$, while G_i at this point consists of all remaining particles except those in generation k_1 .

Step $(0, 0)$ is deemed to be “successful” if (i) $Z_{k_1}^{d,\rho}(L^{-1}(A_{1,1})) \geq m$ and $Z_{k_1}^{d,\rho}(L^{-1}(A_{1,-1})) \geq m$; (ii) no particle has cause to be removed; and (iii) no particle of generation k_1 lies within a distance less than $3/4$ of its parent or siblings. If step $(0, 0)$ is successful, change the status of site $(0, 0)$ to “occupied.” If it is unsuccessful, the algorithm is terminated.

Assuming step $(0, 0)$ to have been successful, proceed with step $(1, -1)$ as follows. Let the set $S_{1,-1}$ consist of the m particles in $L^{-1}(A_{1,-1})$ of the k_1 th generation of step $(0, 0)$ with smallest modulus. Such a set exists by condition (i) for the success of step $(0, 0)$. Run another BRW, this time of k generations, again denoted $(Z_n^{d,\rho})$, this time with $Z_0^{d,\rho} = S_{1,-1}$. Again, order the particles of the new BRW by the same rule as in step $(0, 0)$. Consider successively each particle X after generation 0 of the new BRW, in the ordering given, and

remove X (along with its descendants) if it lies in a translate of B centered at any unremoved particle of the BRW preceding X in the ordering, other than its parent and siblings, or at any point of the set F created in step $(0, 0)$, other than the parent of X . Then remove all but the first k_2 of the remaining particles of this new BRW. Add all the remaining particles of this new BRW to the set F . The new set F corresponds to the set F_i in the algorithm of Section 5, at the end of step $(1, -1)$.

Step $(1, -1)$ is deemed to be a success if, for this BRW, (i) $Z_k^{d,\rho}(L^{-1}(A_{2,0})) \geq m$ and $Z_k^{d,\rho}(L^{-1}(A_{2,-2})) \geq m$; (ii) no particle of this BRW has cause to be removed; and (iii) no particle of generation k lies within a distance $3/4$ of its parent or siblings. If this step is a success, change the status of $(1, -1)$ to “occupied.”

Now perform step $(1, 1)$, and continue in the same manner. At each stage, pick the next site (i, j) in the ordering of \mathcal{L} which has not already been considered and for which there is an oriented path of occupied sites in \mathcal{L} from $(0, 0)$ to $(i-1, j-1)$ [which we call event $E_-(i, j)$], or to $(i-1, j+1)$ [which we call event $E_+(i, j)$]. If no such site (i, j) exists, the algorithm terminates. For any (i, j) having been thus picked, if event $E_+(i, j)$ has occurred let $S_{i,j}$ consist of the m particles in $L^{-1}(A_{i,j})$ with smallest modulus out of those in generation k (or k_1 in the case $i = 1$) of the BRW of step $(i-1, j+1)$. If event $E_-(i, j)$ has occurred but $E_+(i, j)$ has not, let $S_{i,j}$ be defined similarly using the BRW of step $(i-1, j-1)$ instead of $(i-1, j+1)$.

Then in step (i, j) , run a new BRW, again denoted $(Z_n^{d,\rho})$, for k generations as before, now with $Z_0^{d,\rho} = S_{i,j}$. Again order the particles of the new BRW as in the case of step $(0, 0)$, and consider successively each particle X subsequent to generation 0, removing X (along with its descendants) if it lies in the translate of B centered at any unremoved particle of the BRW preceding X in the ordering, other than its parent or siblings, or at any particle of the set F existing at the start of step (i, j) , other than the parent of X . Then remove all but the first k_2 of the remaining particles of the BRW, and add the remaining particles to the set F .

Step (i, j) is “successful” if (i) it produces at least m sites of the k th generation in $L^{-1}(A_{i+1,j+1})$ and at least m sites of the k th generation in $L^{-1}(A_{i+1,j-1})$ and if it satisfies conditions (ii) and (iii) described for step $(1, -1)$ above. If step (i, j) is “successful,” the status of site (i, j) is changed to “occupied.”

The resulting sequence of BRW’s is equivalent to a version of the algorithm for generating F'_∞ described in Section 5. If the algorithm never terminates, then F'_∞ is infinite.

For sufficiently large d , the probability that step $(0, 0)$ is successful exceeds $1 - \psi(y) - 3\delta$, by (32), (33) and Lemma 3. We now show that if d is sufficiently large, then, for each (i, j) other than $(0, 0)$ in \mathcal{L} , given that (i, j) is picked at all, the probability that step (i, j) satisfies conditions (i)–(iii) to be “successful” exceeds $1 - 5\delta$.

Condition (ii) fails if (a) the total progeny of the k -generation BRW of step (i, j) exceeds k_2 , or (b) one of the particles in the BRW of step (i, j) lies in

the translate of B centered at one of its predecessors (other than its parent or siblings) in that BRW, or (c) one of the particles in the BRW of step (i, j) lies in a translate of B centered at a particle already in F as a result an earlier step of the algorithm. By (33) and Lemma 3,

$$(34) \quad P[(\text{ii}) \text{ fails by (a) or (b) in step } (i, j)] < 2\delta.$$

If condition (ii) fails by (c), one of the first k_2 particles in the BRW of step (i, j) lies in a translate of B centered at a particle created at an earlier step of the algorithm, say step (i', j') . Since the steps of the projected BRW $(L(Z_n^{d, \rho}))$ are of length at most ρ , it is impossible for this to happen unless $|x - y| \leq (2k + k_1)\rho$, for some $x \in A_{i, j}$ and $y \in A_{i', j'}$. The number of such (i', j') is bounded above by some constant K , independent of i, j and d , and, for each such (i', j') , at most k_2 sites are added to F . Using the fact that the previous successful step $(i - 1, j + 1)$ or $(i - 1, j - 1)$ satisfies conditions (ii) and (iii), we have

$$(35) \quad P[(\text{ii}) \text{ fails by (c) only, in step } (i, j)] \leq k_2^2 K \sup_{|x| \geq 3/4} \{P[|\mathbf{X} - x| \leq 1]\},$$

where \mathbf{X} is uniform on B^ρ . By Lemma 3, for sufficiently large d , this upper bound is at most δ . Also, by (31), for large enough d we have

$$(36) \quad P[(\text{ii}) \text{ holds but (i) fails in step } (i, j)] \leq \delta,$$

and by Lemma 3, for large d ,

$$(37) \quad P[(\text{iii}) \text{ fails in step } (i, j)] \leq \delta.$$

It follows from (34)–(37) that for sufficiently large d the probability that step (i, j) is successful exceeds $1 - 5\delta$. A comparison with oriented percolation with parameter $1 - 5\delta$ shows that the branching random walk algorithm continues forever, with probability at least $(1 - \psi(y) - 3\delta)(1 - \varepsilon)$, the first factor being the lower bound on the probability of success in step $(0, 0)$. Then (28) follows by Lemma 1 and the choice of δ . \square

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