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Phase transitions for chase-escape models on Poisson–Gilbert graphs^{*}

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Abstract

We present results on phase transitions of local and global survival in a two-species model on Poisson–Gilbert graphs. Initially, there is an infection at the origin that propagates on the graph according to a continuous-time nearest-neighbor interacting particle system. The graph consists of susceptible nodes and nodes of a second type, which we call white knights. The infection can spread on susceptible nodes without restriction. If the infection reaches a white knight, this white knight starts to spread on the set of infected nodes according to the same mechanism, with a potentially different rate, giving rise to a competition of chase and escape. We show well-definedness of the model, isolate regimes of global survival and extinction of the infection and present estimates on local survival. The proofs rest on comparisons to the process on trees, percolation arguments and finite-degree approximations of the underlying random graphs.

Keywords: interacting particle systems; random graphs; survival; extinction; percolation; Boolean model.

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1 Setting and main results

In this paper, we pick up a line of research that very recently has attracted some attention, about the survival of some species when chased by another species, see [DJT18] and references therein. To add another interpretation, our motivation for the model stems from applications in device-to-device networks. Imagine a device is infected by

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some malware at time zero, where the device is a vertex in some random geometric graph representing an ad-hoc telecommunication network. In order to stop the malware from spreading into the system like an infection, special devices can be introduced that have the ability to remove the malware from infected neighboring devices. The special devices that carry the patch are sometimes called *white knights*. The white knights are not allowed to simply transfer the patch to any device in their vicinity, but only to malware-carrying devices. This is motivated by the fact that legal regulations do not allow forceful installation of patches without the consent of susceptible devices, unless the device poses a detected threat. However, once the safety hazard is detected, the operator is allowed to take countermeasures. Once the patch is installed, the infected device becomes a white knight itself, creating again chase-escape dynamics where the malware is followed by white knights, see Figure 1 for an illustration. We present more background in Section 2.

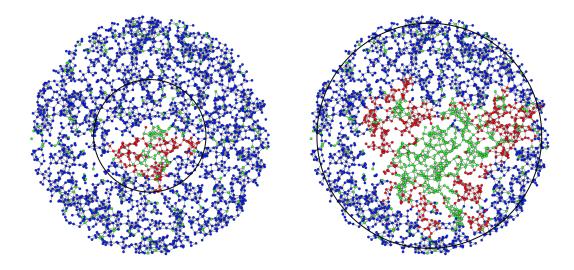


Figure 1: Realization of a random network of nodes that are either infected (red), susceptible (blue) or white knights (green), in a disc at two finite times (left and right). Edges (gray) connect nodes at close proximity and allow for transmission of malware or patching. Circles (black) indicate the maximal distance of the malware to the origin in which the malware was placed initially.

More specifically, we consider a random network of nodes in \mathbb{R}^d given by a homogeneous Poisson point process $X = \{X_i\}_{i \in N}$ with intensity $\mu > 0$, plus an additional node o at the origin. Any two nodes $X_i, X_j \in X \cup \{o\}$ are connected by an edge if and only if $X_j \in B_r(X_i)$, where $B_r(x)$ denotes the ball centered at $x \in \mathbb{R}^d$ with radius r > 0 that we treat as a fixed system parameter. This gives rise to the classical Boolean model or Gilbert graph $g_r(X \cup \{o\})$ from stochastic geometry. Any particle X_i at time $t \ge 0$ can be in one of three states,

$$\xi(t, X_i) = \begin{cases} S, & \text{if } X_i \text{ is susceptible at time } t, \\ I, & \text{if } X_i \text{ is infected at time } t, \\ W, & \text{if } X_i \text{ is a white knight at time } t. \end{cases}$$

For the set of all susceptible, infected and white-knight nodes at time t, we write

respectively

$$S(t) = \{X_i \in X \cup \{o\} : \xi(t, X_i) = S\},\$$

$$I(t) = \{X_i \in X \cup \{o\} : \xi(t, X_i) = I\} \text{ and }\$$

$$W(t) = \{X_i \in X \cup \{o\} : \xi(t, X_i) = W\}.$$

The propagation mechanism is given by a continuous-time Markov jump process with the following transition rates.

- 1. If $\xi(t, X_i) = S$, then X_i becomes infected with rate $\lambda_I \# (I(t) \cap B_r(X_i))$ and
- 2. if $\xi(t, X_i) = I$, then X_i becomes a white knight with rate $\lambda_W \# (W(t) \cap B_r(X_i))$,

where $\lambda_W > 0$ is the *patch rate* and $\lambda_I > 0$ is the *infection rate*. Through a re-scaling on the time axis we can set without loss of generality $\lambda_W = 1$ for the remainder of this manuscript.

We will always assume that the infection starts at the origin, i.e., $I(0) = \{o\}$. Note here that putting an additional node into the network at the origin, amounts to considering the network under the Palm distribution, by Slivnyak–Mecke's theorem. Using this interpretation, the initially infected node is a "typical node" in the system. As for the initial configuration of white knights, we assume them to be an i.i.d. thinning of X with parameter $p \in [0, 1]$. In particular, the initial set of white knights $X_W = W(0)$ then is again a Poisson point process with intensity $\mu_W = p\mu$ and the initial set of susceptible particles $X_S = S(0)$ is a Poisson point process with intensity $\mu_S = (1 - p)\mu$. It is one of the classical result in continuum percolation theory that there exists a unique critical intensity $0 < \mu_r^c < \infty$ such that for $\mu_S > \mu_r^c$, the graph $g_r(X_S \cup \{o\})$ contains a unique infinite component of nodes with probability one and for $\mu_S < \mu_r^c$, the graph $g_r(X_S \cup \{o\})$ contains no infinite component of nodes with probability one, see for example [MR96]. Note that due to scaling relations, without loss of generality, we could set r = 1.

More formally, let us denote by $\mathcal{P} = \mathcal{P}^o \otimes P^{\xi}$ the joint probability distribution of the network model \mathcal{P}^o and the propagation model P^{ξ} with initial configuration ξ . More precisely, by \mathcal{P}^o we denote the joint distribution of the superposition of the independent Poisson point processes X_{S} and X_{W} , with an additional node at the origin. For a given realization $g_r(X_{\mathrm{S}} \cup X_{\mathrm{W}} \cup \{o\})$ of the associated Gilbert graph, P^{ξ} denotes the probability kernel of the Markov propagation model with ξ the initial configuration of states of nodes. ξ is given by

$$\xi(X_i) = \xi(0, X_i) = \begin{cases} I, & \text{for } X_i = o, \\ S, & \text{for } X_i \in X_S, \\ W, & \text{for } X_i \in X_W. \end{cases}$$

Our first result establishes well-definedness of the propagation model for almost-all network realizations. The proof is presented in Section 3.

Proposition 1.1 (Well-definedness). P^{ξ} is a well-defined standard continuous-time Markov jump process on $\{I, S, G\}^{X_S \cup X_W \cup \{o\}}$, \mathcal{P}^o -almost surely.

Our main interest lies in the analysis of extinction and survival of the infection. Denote by

$$E = \{ \text{there exists } t \ge 0 \colon \#I(t) = 0 \},\$$

the event of *extinction* of the infection and call E^c the *survival* event. Due to the percolation properties of the underlying network, we further want to distinguish two

types of survival. Let us call

$$L = \{ \text{for all } t \ge 0 \colon \#I(t) > 0 \} \cap \{ \#\Big(\bigcup_{t \ge 0} I(t)\Big) < \infty \},\$$

the event of *local survival*, in which infected nodes are present for all times, but the number of such nodes is finite. On the other hand, let

$$G = \{ \text{for all } t \ge 0 \colon \#I(t) > 0 \} \cap \{ \# \Big(\bigcup_{t \ge 0} I(t) \Big) = \infty \},\$$

denote the event of *global survival*, in which the infection never disappears and additionally reaches infinitely many nodes. All the events E, L and G of course depend on all the model parameters, which we suppress in the notation for convenience.

Let κ_r denote the Lebesgue volume of the ball $B_r(o)$ and assume $\mu_S \kappa_r \ge 1$, then we define the quantity

$$\rho(\mu_{\rm S}\kappa_r) = 2\mu_{\rm S}\kappa_r - 1 - 2\sqrt{(\mu_{\rm S}\kappa_r)^2 - \mu_{\rm S}\kappa_r},$$

and note that $\rho : [1, \infty) \to (0, 1]$, $x \mapsto \rho(x)$ is strictly decreasing. Further note that if $\mu_S \ge \mu_r^c$, then $\mu_S \kappa_r \ge 1$, see for example [Pen91, Equation 6.2]. We are now in the position to state our main result about extinction.

Theorem 1.2 (Global extinction). If $0 \le \mu_{\rm S} < \mu_r^{\rm c}$, then $\mathcal{P}(G) = 0$ for all $\lambda_{\rm I} \ge 0$ and $\mu_{\rm W} \ge 0$. Further, if $\mu_{\rm S} \ge \mu_r^{\rm c}$ and $\lambda_{\rm I} \ge 0$, then there exists $\mu_{\rm W}^{\rm c}(\lambda_{\rm I}, \mu_{\rm S}) < \infty$ such that for all $\mu_{\rm W} > \mu_{\rm W}^{\rm c}(\lambda_{\rm I}, \mu_{\rm S})$ we have that $\mathcal{P}(G) = 0$. Finally, if $\mu_{\rm S} \ge \mu_r^{\rm c}$ and $\lambda_{\rm I} \le \rho(\mu_{\rm S}\kappa_r)$, then $\mu_{\rm W}^{\rm c}(\lambda_{\rm I}, \mu_{\rm S}) = 0$.

Before we present our result about global survival, let us comment on the preceding theorem. In simple terms, Theorem 1.2 says that if the graph of susceptible nodes is insufficiently connected, i.e., $\mu_S < \mu_r^c$, then global survival is impossible for any infection rate and even without any white knights in the system. Next, if the infection is too weak with respect to the intensity of susceptible nodes, i.e., $\lambda_I \leq \rho(\mu_S \kappa_r)$, then global survival is also impossible for any positive intensity of white knights. On the other hand, for any infection rate, sufficiently many white knights in the system lead to global extinction, see the green part in Figure 2 for an illustration. As we will explain later in Section 2, the

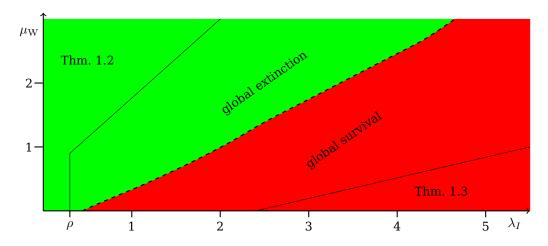


Figure 2: Phase diagram of global survival and extinction in the plane of infection rate vs. white-knight intensity, obtained via simulations for $d = 2, r = 1, \mu_{\rm S} = 3, \lambda_{\rm W} = 1$. Regimes where extinction and survival are covered by the Theorems 1.2 and 1.3 are additionally indicated by solid lines.

specific form of the threshold ρ is due to a comparison to the propagation model on trees. We also present a more detailed explanation and references to preceding research there.

The next result is about global survival, see the red part in Figure 2 for an illustration.

Theorem 1.3 (Global survival). For all $\mu_S > \mu_r^c$ and $\mu_W \ge 0$ there exists $\lambda_I^c(\mu_W, \mu_S) < \infty$ such that for all $\lambda_I > \lambda_I^c(\mu_W, \mu_S)$, we have that $\mathcal{P}(G) > 0$.

In words, Theorem 1.3 states that there is a positive chance for global survival if the underlying graph of susceptible nodes is sufficiently connected and the infection is strong enough to overcome the chasing white knights.

For statements about local survival, let us introduce the notation C_o for the cluster of all nodes in $X = X_S \cup X_W$ for which there exists a path in $g_r(X \cup \{o\})$ connecting them to the origin. Similarly, we denote by C_o^S the cluster of nodes connected to the origin in $g_r(X_S \cup \{o\})$. Further, denote by $\theta(\mu_S) = \mathcal{P}^o(\#C_o^S = \infty)$, the percolation probability of the process of susceptible nodes. Note that the infection can never leave the set C_o^S and hence global survival is impossible if C_o^S is finite and this implies $\mathcal{P}(G) \leq \theta(\mu_S)$.

Lemma 1.4 (Local survival). For all parameters,

$$\exp(-\mu\kappa_r) \le \mathcal{P}(L) \le (1 - \theta(\mu_{\rm S})) \exp(-\mu_{\rm W}\kappa_r).$$

In particular, if the process of susceptible nodes is subcritical, we have $\mathcal{P}(G) = 0$, $\mathcal{P}(L) \leq \exp(-\mu_W \kappa_r)$ and thus $\exp(-\mu \kappa_r) \leq \mathcal{P}(E^c) \leq \exp(-\mu_W \kappa_r)$. If the process of susceptible nodes is supercritical but the other parameters guarantee global extinction as described in Theorem 1.2, then for the survival probability $\exp(-\mu \kappa_r) \leq \mathcal{P}(E^c) \leq (1 - \theta(\mu_S)) \exp(-\mu_W \kappa_r)$. Note that $\mu \mapsto \theta(\mu)$ tends to one exponentially fast, see [PP96], and hence local survival is exponentially unlikely for dense networks both in μ_S and μ_W .

In the next section we explain the strategy of the proofs, and comment on related results in the literature. The proofs are presented in Section 3.

2 Strategy of proofs

The study of epidemic models defined in terms of interacting particle systems with some additional randomness coming from an environment has, by now, a long history, see for example the early works [Lig92, And92]. The consideration of such processes on random graphs has attracted attention more recently, see for example [Dur10a, Dur10b] and references therein. Here, a particularly interesting class is given by random graphs with a prescribed degree distribution, see for instance [CD09, MVY13, MMVY16, MV16], dense Erdös–Rényi graphs [dALRR15], and Galton–Watson trees [Kor16]. The literature on interacting particle systems on random geometries, i.e., where the random graphs are embedded in space, and in particular do not obey any degree bounds, is much sparser and younger. Notable here is the work [MS16], which establishes existence of a subcritical phase of the contact process on Poisson–Gilbert graphs and Poisson–Delaunay tessellations.

As mentioned in the beginning, propagation models analogous to the one presented in this manuscript have been studied in recent years on various fixed networks such as trees [Kor05], complete graphs [Kor15], and lattices [DJT18] also via simulations [TKL18], and mainly motivated by applications in probabilistic biology and rumor spreading. Typically, there the infection is represented by red particles and white knights by blue particles. Apart from our generalization towards random geometries, another difference in the analysis presented here is that, in the initial configuration, the white knights form a Poisson point process of infinitely many nodes, whereas in the preceding works on fixed geometries there is only a finite number of white knights present in the system. This leads to the new phenomenon that we can have almost-sure extinction, even for $\lambda_I > 1$, as the infection will constantly discover white knights. Nevertheless, our proofs are partially based on the results for fixed networks, in particular the tree considered by Kordzakhia [Kor05], since there it is possible to derive explicit bounds for the infection rate by balancing numbers of paths compared with propagation along one path, which is then in fact one dimensional. More precisely, let us consider our propagation model on a fixed connected graph H that includes a root $\{o\}$ and a generic point $\{o'\}$ which is only connected to the root by a single edge. Let the starting configuration be given by

$$\xi'(x) = \begin{cases} W, & \text{for } x = o', \\ I, & \text{for } x = o, \\ S, & H \setminus \{o, o'\}. \end{cases}$$

Let $\lambda_{\rm I}^{\rm c}(H)$ denote the critical rate for (global) extinction of the infection based on the propagation model as explained above, where the underlying Gilbert graph and the initial condition ξ are replaced by H and ξ' . Then, the following result is proved in [DJT18, Theorem 1, Corollary 2].

Lemma 2.1 (Extinction on fixed networks). Let $\Gamma_n(H)$ denote the set of self-avoiding paths of length n in H, starting from the root. If there exists $k \in \{2, 3, ...\}$ such that for all $n \in \mathbb{N}$ we have that $\#\Gamma_n(H) \leq k^n$, then

$$\lambda_{\mathrm{I}}^{\mathrm{c}}(H) \ge \lambda_{\mathrm{I}}^{\mathrm{c}}(\mathbb{T}_k) = 2k - 1 - 2\sqrt{k^2 - k},$$

where \mathbb{T}_k is the rooted *k*-ary tree.

The next result establishes existence of a constant, the *connective constant*, bounding the number of self-avoiding paths for almost-all realizations of the Poisson–Gilbert graph. **Lemma 2.2** (Connective constant). For all $\mu_{\rm S}$, r > 0 and all $\gamma > \mu_{\rm S}\kappa_r$ we have

$$\limsup_{n\uparrow\infty} n^{-1}\log \#\Gamma_n\big(g_r(X_{\mathcal{S}}\cup\{o\})\big) \le \log\gamma,\tag{2.1}$$

for \mathcal{P}^{o} -almost all X_{S} .

We present the proof of Lemma 2.2 in Section 3. Now the last statement of Theorem 1.2 is an immediate consequence of the following proposition, which leverages Lemma 2.2 and the proof idea of Lemma 2.1 to the setting of infinitely-many white knights.

Proposition 2.3. If $\mu_{\rm S} \ge \mu_r^{\rm c}$, $\lambda_{\rm I} \le \rho(\mu_{\rm S}\kappa_r)$ and $\mu_{\rm W} > 0$, then $\mathcal{P}(G \cap \{\#C_o = \infty\}) = 0$.

Proof of Theorem 1.2, first and last statement. First of all, if $\mu_{\rm S} < \mu_r^{\rm c}$, then, as mentioned above, $\mathcal{P}(G) \leq \theta(\mu_{\rm S}) = 0$. For the last statement, let $\mu_{\rm S} \geq \mu_r^{\rm c}$, $\lambda_{\rm I} \leq \rho(\mu_{\rm S}\kappa_r)$ and $\mu_{\rm W} > 0$, then using Proposition 2.3, we have that

$$\mathcal{P}(G) = \mathcal{P}(G \cap \{\#C_o = \infty\}) + \mathcal{P}(G \cap \{\#C_o < \infty\}) = 0,$$

since global survival is impossible on finite clusters.

The technique used for the proof of Proposition 2.3, which is based on the works on trees, fail in the regimes of large λ_{I} , since the discovery of a new white knight on a one-dimensional path not only stops the infection on this path, but also creates new white knights. However, with the help of percolation arguments we prove in Section 3 the following proposition, which immediately implies the second statement of Theorem 1.2.

Proposition 2.4. For all $\lambda_{\rm I} \geq 0$ and $\mu_{\rm S} \geq 0$, there exists $\mu_{\rm W}^{\rm c}(\lambda_{\rm I}, \mu_{\rm S}) < \infty$ such that for all $\mu_{\rm W} > \mu_{\rm W}^{\rm c}(\lambda_{\rm I}, \mu_{\rm S})$, we have that $\mathcal{P}(G \cap \{\#C_o = \infty\}) = 0$.

ECP 25 (2020), paper 25.

Proof of Theorem 1.2, second statement. In the setting of Proposition 2.4, we have $\mathcal{P}(G) = \mathcal{P}(G \cap \{\#C_o = \infty\}) + \mathcal{P}(G \cap \{\#C_o < \infty\}) = 0.$

Before we comment on the proof of Theorem 1.3, let us make some remarks on Figure 2. First, the linear dependence of $\lambda_{\rm I} \mapsto \mu_{\rm W}^{\rm c}(\mu_{\rm S}, \lambda_{\rm I})$ can be heuristically understood as follows. Consider the rescaled process of white knights with intensity $\alpha \mu_{\rm W}$ and patch rate $1/\alpha$ for $\alpha > 0$. Then, the rescaled chase-escape model, in the limit as α tends to infinity, converges to a *contact process* on $g_r(X_{\rm S} \cup \{o\})$ with recovery rate given by $\mu_{\rm W}\kappa_r$ and infection rate $\lambda_{\rm I}$. Then, the linear dependence in $\lambda_{\rm I}$ emerges through a re-scaling on the time axis. As a second remark, we note that our simulations, leading to the phase-transition line in Figure 2, indicate that global survival is even possible in a situation where the infection is strictly slower than the patching, i.e., $\lambda_{\rm I} < 1$. This is proven for trees [Kor05], but only conjectured on the lattice [DJT18], and also not covered by our results. The simulations also suggest monotonicity and uniqueness of the phase-separating line.

The main challenge for the proof of Theorem 1.3 lies in the fact that almost-all network realizations have an unbounded degree. Although large degrees should support survival of the infection, lack of monotonicity prevents us from using this idea directly. See also our comments on monotonicity below. However, we can estimate the network by graphs of bounded degree and use discrete percolation arguments for the approximations. The proof of Theorem 1.3 is presented in Section 3.

Finally the quantitative statements about local survival as presented in Lemma 1.4, are a simple consequence of void-space probabilities, let us give the proof here as well.

Proof of Lemma 1.4. First, note that local survival is only possible if $\#C_o^S < \infty$ since otherwise, with probability one, there exists a white knight in the set $B_r(C_o^S) = \bigcup_{X_i \in C_o^S} B_r(X_i)$, which is eventually reached by the infection. Consequently, the infection can only survive if it escapes towards infinity on C_o^S . In particular, local survival is only possible if no white knights are in $B_r(C_o^S)$, i.e.,

$$\mathcal{P}(L) = \mathcal{P}(L \cap \{\#C_o^S < \infty\}) = \mathbb{E}\left[\exp(-\mu_W |B_r(C_o^S)|)\mathbb{1}\{\#C_o^S < \infty\}\right]$$

$$\leq (1 - \theta(\mu_S))\exp(-\mu_W \kappa_r).$$

On the other hand, the infection survives if the origin is isolated. This gives the lower bound. $\hfill \Box$

Let us finish this section by commenting on monotonicity properties of the phase diagram as sketched in Figure 2. Both, simulations and common sense suggest existence of a unique phase-separating curve and several monotonicities depending on the parameters. For example that additional infected nodes or an increase in the infection rate should increase the probability for the infection to survive. However, to prove existence, uniqueness and monotonicities is challenging, mainly because of the existence of configurations that exhibit counterintuitive effects, standing in the way of coupling arguments. Let us give one example here. Note that white knights can only act towards perviously infected nodes, and therefore an increase in infected nodes also benefits the spread of white knights. To illustrate this, consider the nearest-neighbor graph on $\mathbb N$ as presented in Figure 3 with a white knight at node 1, an infection at node 3 and all other nodes being susceptible. Imagine node 2 and its associated edges were absent, then the infection would spread towards infinity unstopped for any positive infection rate. Now, if we add an infected node at position 2, then more infections are in the system. Still, if $\lambda_{\rm I} < 1$, the infection will now go extinct. Other examples can be constructed to also showcase configurations where an increase of $\lambda_{\rm I}$ leads to a decrease for the probability of survival of the infection.



Figure 3: Illustration of a configuration of a white knight (green, position 1), an infected node (red, position 3) and susceptible nodes (blue) for which adding an infected node at position 2 would stop the infection that would otherwise propagate to infinity.

Let us finally mention that in the survival regime it is reasonable to believe that there exists an infection speed α depending on all the parameters of the system, such that as $t \to \infty$ we have $B_{\alpha t(1-\epsilon)}(o) \subset I(t) \subset B_{\alpha t(1+\epsilon)}(o)$ with probability one, conditioned on the event that the origin is connected to infinity.

In the following section we present all remaining proofs.

3 Proofs

In the sequel, we denote by \mathcal{E}^o and E^{ξ} the expectations associated to \mathcal{P}^o and P^{ξ} , respectively. We abbreviate for balls, $B_n(o)$ by B_n and for boxes, $Q_n(o)$ by Q_n . For any $A \subset \mathbb{R}^d$ we write $A^c = \mathbb{R}^d \setminus A$.

Proof of Proposition 1.1. Note that for positive intensities, \mathcal{P}^{o} -almost surely, all the Poisson–Gilbert graphs $g_r(X_S \cup \{o\})$, $g_r(X_W \cup \{o\})$ and $g_r(X \cup \{o\})$ have an unbounded degree, potentially leading to blow-ups in finite time for the propagation model. In particular, well-definedness of the propagation model on the Gilbert graphs can not be guaranteed using the standard conditions based on generators. For example there exists no finite bound on the transition rate uniformly in the nodes, as required in [Lig85, Proposition 3.2 Chapter 3].

However, in our case, we can establish well-definedness due to the fact that in our initial condition there is only one infection present. More precisely, consider the process $(\xi_n(t,\cdot))_{t\geq 0}$ on $g_r((X \cup \{o\}) \cap B_n)$ with $n \in r\mathbb{N}$, defined via the same rates as presented above. Then, for \mathcal{P}^o -almost-all network realizations, the process $(\xi_n(t,\cdot))_{t\geq 0}$ is well-defined and standard as a finite state space Markov process. Next, let

$$\tau_n = \inf\{t > 0 \colon I_{n+1}(t) \not\subset B_n\}$$

denote the time at which the infection, based on the process $\xi_{n+1}(t, \cdot)$, hits the boundary of B_n . Then, for all $t < \tau_n$ our original process $\xi(t, \cdot)$ coincides with $\xi_n(t, \cdot)$ and is thus well-defined.

What remains to be shown is that \mathcal{P} -almost surely $\lim_{n\uparrow\infty} \tau_n = \infty$. To show this, we will use a large-deviation argument to establish a bound on the minimal asymptotic speed of the infection process and apply the Borel–Cantelli lemma. More precisely, first note that for any $\alpha > 0$ and $n \in r\mathbb{N}$, we have that

$$P^{\xi}\left(\tau_{n} < n/\alpha\right) \le \#\Gamma_{n/r}P^{\xi}\left(S_{n/r} < n/\alpha\right),\tag{3.1}$$

where $\#\Gamma_m$ is the number of self-avoiding paths in $g_r(X_S \cup \{o\})$ of length $m \in \mathbb{N}$ starting in $\{o\}$ and S_m is the sum of m independent exponentially-distributed random variables with parameter λ_I . For the estimate (3.1) we used that, in order for the infection to reach B_n^c in less than n/α time, at least n/r infection events happened along at least one of the self-avoiding paths of length n/r in less than n/α time. Now, using Lemma 2.2 and large-deviation bounds for independent exponentially-distributed random variables, we obtain for $\alpha > r\lambda_I$ and \mathcal{P}^o -almost all network realizations the bound

$$\limsup_{m\uparrow\infty} \frac{1}{m} \log\left(\#\Gamma_m P^{\xi}\left(S_m < \frac{mr}{\alpha}\right)\right) \le \log\gamma - \frac{\lambda_{\mathrm{I}}r}{\alpha} + 1 + \log\frac{\lambda_{\mathrm{I}}r}{\alpha} = -C_{\gamma,\lambda_{\mathrm{I}},r}(\alpha).$$
(3.2)

In particular, for $\alpha > \alpha_c$ with $\alpha_c = \inf\{\alpha > r\lambda_I: -C_{\gamma,\lambda_I,r}(\alpha) > 0\}$, there exists $n_o \in r\mathbb{N}$ such that

$$\sum_{n \in r \mathbb{N}} P^{\xi} \left(\tau_n < n/\alpha \right) < n_o + \sum_{n \in r \mathbb{N} : n > n_o} \exp(-nC_{\gamma, \lambda_{\mathrm{I}}, r}(\alpha)/2) < \infty.$$

Finally, by an application of the Borel–Cantelli lemma, \mathcal{P}^{o} -almost surely, we have that $P^{\xi} \left(\limsup_{n\uparrow\infty} \{\tau_n < n/\alpha\} \right) = 0$ and thus almost surely under \mathcal{P} , for all but finitely many n, we have $\tau_n \ge n/\alpha$, which finishes the proof.

Let us note that the above proof of existence of the process also derives bounds on the minimal speed of propagation of the infection in terms of solutions to fixed-point equations determining the critical speed α_c .

Proof of Lemma 2.2. First, we compute the expectation of $\#\Gamma_n$ by multiple applications of Slivnyak–Mecke's theorem. More precisely, let X_{i_0} be an alternative notation for o, then for any $n \in \mathbb{N}$ we have that

$$\begin{aligned} \mathcal{E}^{o}[\#\Gamma_{n}] &= \mathcal{E}^{o}\Big[\sum_{i_{1},\dots,i_{n}\in N}^{\neq}\prod_{k=1}^{n}\mathbb{1}\{|X_{i_{k}}-X_{i_{k-1}}| < r\}\Big] \\ &= \mu_{\mathrm{S}}\int \mathcal{E}^{o}\Big[\sum_{i_{1},\dots,i_{n-1}\in N}^{\neq}\prod_{k=1}^{n-1}\mathbb{1}\{|X_{i_{k}}-X_{i_{k-1}}| < r\}\mathbb{1}\{|x-X_{i_{n-1}}| < r\}\Big]\delta x \\ &= \mu_{\mathrm{S}}\kappa_{r}\,\mathcal{E}^{o}\Big[\sum_{i_{1},\dots,i_{n-1}\in N}^{\neq}\prod_{k=1}^{n-1}\mathbb{1}\{|X_{i_{k}}-X_{i_{k-1}}| < r\}\Big] = (\mu_{\mathrm{S}}\kappa_{r})^{n}, \end{aligned}$$

where the \neq sign indicates that we sum over mutually distinct indices. Next, using the Markov inequality, we have

$$\mathcal{P}^{o}(\#\Gamma_{n} \geq \gamma^{n}) \leq \gamma^{-n} \mathcal{E}^{o}[\#\Gamma_{n}] = (\mu_{\mathrm{S}} \kappa_{r} / \gamma)^{n}$$

and thus, for $\gamma > \mu_{S}\kappa_{r}$ we obtain $\sum_{n \in \mathbb{N}} \mathcal{P}^{o}(\#\Gamma_{n} \ge \gamma^{n}) < \infty$. Hence, the Borel–Cantelli lemma yields that $\#\Gamma_{n} > \gamma^{n}$ only for finitely-many n. This completes the proof. \Box

Proof of Proposition 2.3. It suffices to consider the critical and supercritical regimes where $\mu_S \ge \mu_r^c$. Let us abbreviate $A = \bigcup_{t\ge 0} I(t)$. First, note that it suffices to prove that for almost-all network realizations we have that $E^{\xi}[\#A] < \infty$ for $\lambda_I \le \rho(\mu_S \kappa_r)$ and any $\mu_W > 0$, as this implies $P^{\xi}(G) = 0$. For this, we want to bound #A from above by numbers of infected nodes along one-dimensional paths. We have to be careful here since by the lack of monotonicity, as explained at the end of Section 2, we cannot simply remove white knights in order to produce a situation as in Lemma 2.1. However, note that for every $x \in A$ there must exist a self-avoiding path $\varphi_n(x) \subset g_r(X_S \cup \{o\})$ of some finite length n = n(x), started at the origin, such that x received its infection along $\varphi_n(x)$. Let $I_n \subset A$ denote the set of nodes in A that were infected at some time along a path of length n. Then, for \mathcal{P}^o -almost all X with $X_W \cap B_r(o) \neq \emptyset$, i.e., realizations of the network where the origin is directly adjacent to a white knight, we have

$$E^{\xi}[\#A] \le \sum_{n \ge 0} E^{\xi}[\#I_n] \le \sum_{n \ge 0} \#\Gamma_n P'(o \text{ infects } n \text{ on } \mathbb{N} \cup \{o\} \cup \{o'\}), \tag{3.3}$$

where P' denotes the distribution of the process described in Lemma 2.1 for $H = \mathbb{N} \cup \{o\} \cup \{o'\}$. Here, we used that more neighboring white knights along a fixed path lead to even smaller probability of survival.

An application of Lemma 2.1, as presented in [DJT18, Theorem 1 and Corollary 2], yields that

$$P'(o \text{ infects } n \text{ on } \mathbb{N} \cup \{o\} \cup \{o'\}) \leq C(\lambda_{\mathrm{I}})\lambda_{\mathrm{I}}'^n n^{-3/2},$$

for $\lambda'_{\rm I} = 4\lambda_{\rm I}/(1+\lambda_{\rm I})^2$, which is derived via the reflection principle in case of $\lambda_{\rm I} < 1$. Note that by assumption, $\gamma > 1$ and thus for $\lambda_{\rm I} \le 2\gamma - 1 - 2\sqrt{\gamma^2 - \gamma} < 1$ in particular $\lambda'_{\rm I} < 1$. Hence, for such $\lambda_{\rm I}$, the right-hand side of (3.3) is finite and thus $E^{\xi}[\#A] < \infty$ for \mathcal{P}^o -almost all X with $X_{\rm W} \cap B_r(o) \neq \emptyset$.

To finish the proof let us consider the network realizations where $X_{W} \cap B_{r}(o) = \emptyset$, i.e., where the origin is not adjacent to a white knight. Denote the set of all connected finite subsets of nodes that contain at most one node adjacent to a white knight by

$$\mathcal{J} = \{ J \subset X_{\rm S} \cup \{o\} : o \in J, \#J < \infty, J \text{ connected}, \#\{X_i \in J : \operatorname{dist}(X_i, X_W) < r\} = 1 \},\$$

where $\operatorname{dist}(x, B) = \inf\{|x - y| : y \in B\}$ denotes the distance of a point $x \in \mathbb{R}^d$ to a set $B \subset \mathbb{R}^d$. For any $J \in \mathcal{J}$ we define ξ^J to be the configuration where the infected nodes are precisely given by J, i.e., we define for all $X_i \in X$

$$\xi^{J}(X_{i}) = \begin{cases} I, & \text{if } X_{i} \in J \\ S, & \text{if } X_{i} \in X_{S} \setminus J \\ W, & \text{if } X_{i} \in X_{W}. \end{cases}$$

Then, for \mathcal{P}^{o} -almost-all X with $\#C_{o} = \infty$ we define $\tau = \inf\{t \ge 0 : \xi(t, \cdot) = \xi^{J} \text{ for some } J \in \mathcal{J}\}$, the stopping time at which the process explores the white knights for the first time. With these definitions, note that

$$\sum_{J\in \mathcal{J}} P^{\xi}\left(\xi(\tau)=\xi^{J}\right)=1,$$

since every realization of the infection propagation on an infinite cluster eventually reaches a white knight in finite time. Furthermore, the node that explored the white knight is unique and will be denoted by o_J . Therefore, for \mathcal{P}^o -almost-all X with $\#C_o = \infty$ we have

$$P^{\xi}(G) = \sum_{J \in \mathcal{J}} P^{\xi} \left(G \mid \xi(\tau) = \xi^J \right) P^{\xi}(\xi(\tau) = \xi^J).$$

Now, due to the strong Markov property, we have

$$P^{\xi}(G \mid \xi(\tau) = \xi^{J}) = P^{J}(G \mid \xi(\tau) = \xi^{J}) \le P^{J}(G)/P^{J}(\xi(\tau) = \xi^{J}) = 0,$$

where P^J denotes the infection process not started in o, but in o_J , constructed on the same probability space as P^{ξ} . The last equality holds, as $P^J(\xi(\tau) = \xi^J) > 0$ and $P^J(G) = 0$, since with these definitions, there is a white knight next to the origin and the first part of the proof can be applied again. Now, since \mathcal{J} is countable, we can conclude that indeed for \mathcal{P}^o -almost all X we have $P^{\xi}(G) = 0$.

Proof of Proposition 2.4. The proof is based on a percolation argument. We call a node $X_i \in X_S$ an *open node* if

- 1. X_i is not isolated in $g_r(X_S \cup \{o\})$, and
- 2. once X_i is infected, then it transmits its infection towards at least one of its neighbors in X_S (regardless if neighbors are already infected or patched) before X_i is directly patched by a neighboring white knight in X_W .

We call X_i a *closed node* otherwise. For example X_i is an open node if it is not isolated in $g_r(X_S \cup \{o\})$ but isolated in $g_r(X_W \cup \{X_i\})$. Note that a node is labeled open or closed based on its neighborhood at initial time. If we would do the labelling at some later time, there would be potentially more white knights or less susceptible nodes in the neighborhood. Hence, it would be less likely for the node to be open. This construction is designed particularly to create a time-monotone behavior (with respect to the probability for a node to be open), in an otherwise notoriously non-monotone model. An infinite self-avoiding path of open nodes in $g_r(X_S \cup \{o\})$ does not guarantee global survival of the infection, since we do not require that the infection propagates to infinity on this path. However, absence of an infinite self-avoiding path of open nodes implies absence of global survival in the realization. Indeed, if there is no path to infinity of nodes that are able to infect at least one neighboring node from the initially susceptible nodes, then in particular there is no path to infinity of nodes that are able to infect neighboring nodes that are still susceptible at the time at which the infection arrives. Let us emphasize that this comes from the fact that less susceptible nodes or more white knights in the neighborhood make it even harder to transmit the infection towards at least one neighboring susceptible node before being patched by a neighboring white knight.

In order to show absence of an infinite self-avoiding path of open nodes, let us discretize space into boxes $Q_{3r}(3rz)$ of side-length 3r, centered at 3rz with $z \in \mathbb{Z}^d$ and use results on lattice percolation. The site $z \in \mathbb{Z}^d$ is called an *open site*, if one of the following events happens,

$$\begin{split} A_m(z) &= \{ \# \big(X_{\rm S} \cap Q_{3r}(3rz) \big) \ge m \}, \text{ or } \\ B_n(z) &= \{ \text{there exists } X_i \in X_{\rm S} \cap Q_{3r}(3rz) \colon \# \big(X_{\rm S} \cap B_r(X_i) \big) \ge n+1 \}, \text{ or } \\ C_{k,\mu_{\rm W}}(z) &= \{ \text{there exists } x \in Q_{3r}(3rz) \colon \# (B_r(x) \cap X_{\rm W}) < k \} \\ D(z) &= \{ Q_{3r}(3rz) \text{ contains an open node} \}. \end{split}$$

Otherwise $z \in \mathbb{Z}^d$ is called a *closed site*. Now, suppressing the dependence on z if z = o, we have that

$$\mathcal{P}(o \text{ is an open site}) = 1 - \mathcal{P}(D^{c} \mid A_{m}^{c} \cap B_{n}^{c} \cap C_{k,\mu_{w}}^{c}) \mathcal{P}(A_{m}^{c} \cap B_{n}^{c} \cap C_{k,\mu_{w}}^{c}).$$

If a node X_i has $n \ge 1$ neighbors in X_S and $k \ge 0$ neighbors in X_W , the probability for X_i to be a closed node is given by $k/(k + n\lambda_I)$. Hence, we can bound

$$\mathcal{P}(D^{c} \mid A_{m}^{c} \cap B_{n}^{c} \cap C_{k,\mu_{W}}^{c}) \\ \geq \mathcal{E}^{o} \Big[\prod_{X_{i} \in X_{S} \cap Q_{3r}(o)} \frac{\#(X_{W} \cap B_{r}(X_{i}))}{\#(X_{W} \cap B_{r}(X_{i})) + (\#(X_{S} \cap B_{r}(X_{i})) - 1)\lambda_{I}} \mid A_{m}^{c} \cap B_{n}^{c} \cap C_{k,\mu_{W}}^{c} \Big] \\ \geq \Big(\frac{k}{k + n\lambda_{I}} \Big)^{m},$$

and

$$\mathcal{P}(A_m^{\mathsf{c}} \cap B_n^{\mathsf{c}} \cap C_{k,\mu_{\mathsf{W}}}^{\mathsf{c}}) \ge 1 - \left(\mathcal{P}(A_m) + \mathcal{P}(B_n) + \mathcal{P}(C_{k,\mu_{\mathsf{W}}})\right)$$

Now, for any $\lambda_{\rm I} \ge 0$ and $\varepsilon > 0$, there exist $m_o, n_o \in \mathbb{N}$ such that for all $m > m_o$ and $n > n_o$ we have

$$\mathcal{P}(A_m) < \varepsilon$$
 and $\mathcal{P}(B_n) < \varepsilon$

and $k_o(m,n) \in \mathbb{N}$ such that for all $k > k_o(m,n)$ we have

$$\left(\frac{k}{k+n\lambda_{\rm I}}\right)^m > 1-\varepsilon.$$

ECP 25 (2020), paper 25.

http://www.imstat.org/ecp/

Finally, we can then pick $\mu_{\rm W}^{\rm c}(k)$ sufficiently large, such that for all $\mu_{\rm W} > \mu_{\rm W}^{\rm c}(k)$ also

$$\mathcal{P}(C_{k,\mu_{\mathrm{W}}}) < \varepsilon$$

Together, the probability for an open site can be made arbitrarily small, since

$$\mathcal{P}(o \text{ is an open site}) \leq 1 - (1 - \varepsilon)(1 - 3\varepsilon) = 4\varepsilon - 3\varepsilon^2.$$

The random field of good and bad sites constitutes a two-dependent site-percolation model on \mathbb{Z}^d that can be dominated by an independent site-percolation model, using the domination-by-product-measures result [LSS97, Theorem 0.0]. Then, for sufficiently large μ_W we have absence of percolation of good sites. As the side-length of the boxes is larger than 2r, this also excludes the possibility of continuum percolation of open nodes in the Gilbert graph, which then implies absence of global survival of the infection process also on the event $\{\#C_o = \infty\}$.

Proof of Theorem 1.3. Similar to the proof of Proposition 2.4 we will use a percolation argument for a suitably chosen discretization of \mathbb{R}^d to show existence of an infinite cluster of nodes that transmit their infection faster than any cure attempt of neighboring white knights. More precisely, we call $X_i \in X_S$ an open node if X_i transmits the infection to all its neighbors in $g_r(X_S)$ before an attempt to cure X_i has been made by any neighbor in $g_r(X_S \cup X_W)$. We call X_i a closed node otherwise. With this definition, using also the strong Markov property, for the global survival of the infection, it suffices to prove existence of an infinite cluster of open nodes, which is connected to the origin, with positive probability.

We aim to achieve this by choosing the infection rate $\lambda_{\rm I}$ large, however, due to the unbounded degree of the graphs, there is no globally sufficiently large infection rate such that for \mathcal{P}^{o} -almost all graphs the survival rate is above a fixed $\epsilon > 0$. We therefore introduce a discretization of \mathbb{R}^{d} into boxes and distinguish those boxes in which the degrees are bounded. For this, let $g_{r,\lambda_{\rm I}}^{m,n}(X_{\rm S} \cup X_{\rm W})$ denote the Gilbert graph with connectivity threshold r > 0 and vertex set

$$\{X_i \in X_{\mathcal{S}} \colon \#(X_{\mathcal{W}} \cap B_r(X_i)) \le m, \#(X_{\mathcal{S}} \cap B_r(X_i)) \le n+1 \text{ and } X_i \text{ is an open node}\}.$$
(3.4)

This is a Gilbert graph based on a dependently thinned Poisson point process, where vertices are removed if they have too many neighbors in $X_{\rm S}$ and $X_{\rm W}$ or transmit their infection too slowly. Let $\theta_{m,n,\lambda_{\rm I}}(\mu_{\rm S},\mu_{\rm W})$ denote the associated percolation probability. Our parameters are such that $\theta(\mu_{\rm S}) > 0$. We claim that

$$\lim_{m,n,\lambda_{\rm I}\uparrow\infty}\theta_{m,n,\lambda_{\rm I}}(\mu_{\rm S},\mu_{\rm W}) = \theta(\mu_{\rm S}).$$
(3.5)

If (3.5) holds, then this implies that for sufficiently large m, n and $\lambda_{\rm I}$ also $g_{r,\lambda_{\rm I}}^{m,n}(X_{\rm S} \cup X_{\rm W})$ is supercritical and the infection survives globally. In order to prove (3.5), first note that for all $m, n, \lambda_{\rm I}$ we have that $\theta_{m,n,\lambda_{\rm I}}(\mu_{\rm S}, \mu_{\rm W}) \leq \theta(\mu_{\rm S})$ since $g_{r,\lambda_{\rm I}}^{m,n}(X_{\rm S} \cup X_{\rm W})$ is based on a thinning of the vertices in $g_r(X_{\rm S})$. To show the reverse direction, consider boxes $Q_s(sz)$ for $z \in \mathbb{Z}^d$ with s > r and define for any set $A \subset \mathbb{R}^d$ the diameter of A by $\operatorname{diam}(A) = \sup\{|x-y|: x, y \in A\}$. We say that a site $z \in \mathbb{Z}^d$ is a good site if all the following events happen

$$\begin{split} A_m(z) &= \{ \text{for all } X_i \in X_{\mathrm{S}} \cap Q_{3s}(sz) \colon \# \big(X_{\mathrm{W}} \cap B_r(X_i) \big) \leq m \}, \\ B_n(z) &= \{ \text{for all } X_i \in X_{\mathrm{S}} \cap Q_{3s}(sz) \colon \# \big(X_{\mathrm{S}} \cap B_r(X_i) \big) \leq n+1 \}, \\ C_{\lambda_{\mathrm{I}}}(z) &= \{ \text{all } X_i \in X_{\mathrm{S}} \cap Q_{3s}(sz) \text{ are good nodes} \}, \\ D_s(z) &= \{ g_r(X_{\mathrm{S}}) \text{ contains a unique cluster } Z \text{ in } Q_s(sz) \text{ with } \operatorname{diam}(Z) \geq s/2 \}, \text{ and} \\ E_s(z) &= \{ g_r(X_{\mathrm{S}}) \text{ contains a unique cluster } Z \text{ in } Q_{3s}(sz) \text{ with } \operatorname{diam}(Z) \geq s/2 \}. \end{split}$$

Otherwise z is called a *bad site*. In particular, if o is connected in $g_r(X_S)$ to ∂Q_s and o is contained in an infinite cluster of good sites $z \in \mathbb{Z}^d$, then the infection survives globally. Indeed, since o is connected in $g_r(X_S)$ to ∂Q_s , there is a cluster Z_o with $\operatorname{diam}(Z_o) \ge s/2$ in $Q_s(o)$. Further, let z be a neighbor of o in the infinite component of good sites. Then, there exists a unique cluster Z_z in $Q_s(z)$ again with $\operatorname{diam}(Z_z) \ge s/2$. Since Z_o and Z_z are also unique in $Q_{3s}(z)$ respectively $Q_{3s}(o)$, Z_o and Z_z must be connected in $Q_{3s}(o) \cup Q_{3s}(z)$. This can be iterated along the path of good sites to infinity. By the goodness of that path, also there is no thinning, and hence the infection can globally survive. Now we can estimate,

$$0 < \theta(\mu_{\rm S}) \leq \mathcal{P}(o \text{ is part of a finite cluster of good sites}) + \mathcal{P}(G),$$

and it suffices to show that the percolation probability for the process of good sites can be pushed arbitrarily close to one as the parameters s, m, n and $\lambda_{\rm I}$ tend to infinity. Note that by the definition of goodness of nodes, goodness of sites, and since s > r, the process of good sites is a 3-dependent percolation process. Using the domination-by-product measure result [LSS97, Theorem 0.0], it suffices to bound the 3-dependent percolation process from below by a supercritical Bernoulli percolation process with parameter arbitrarily close to one, in the usual sense of evaluations of increasing events. In other words, it suffices to show that

$$\limsup_{s\uparrow\infty}\limsup_{m\uparrow\infty}\limsup_{n\uparrow\infty}\limsup_{\lambda_{I}\uparrow\infty}\mathcal{P}(o \text{ is a bad site})=0.$$

For this, we can bound the probability for a bad site by

$$\begin{split} \mathcal{P}(o \text{ is a bad site}) &= 1 - \mathcal{P}(A_m \cap B_n \cap C_{\lambda_{\mathrm{I}}} \cap D_s \cap E_s) \\ &\leq \mathcal{P}(A_m \cap B_n \cap C_{\lambda_{\mathrm{I}}}^{\mathrm{c}}) + \mathcal{P}(A_m^{\mathrm{c}}) + \mathcal{P}(B_n^{\mathrm{c}}) + \mathcal{P}(D_s^{\mathrm{c}}) + \mathcal{P}(E_s^{\mathrm{c}}), \end{split}$$

where we suppressed the dependence on z = o. Now, using the large-deviation estimates in [PP96, Theorem 2], we can choose s sufficiently large such that

$$\mathcal{P}(D_s^c) < \varepsilon$$
 and $\mathcal{P}(E_s^c) < \varepsilon$.

Next, for given s, we can choose n and m sufficiently large such that also

$$\mathcal{P}(A_m^c) < \varepsilon$$
 and $\mathcal{P}(B_n^c) < \varepsilon$,

by convergence in bounded domains. Finally, for given s, m and n, note that, under the events A_m and B_n , the probability of a node $X_i \in X_S \cap Q_{3s}(o)$ to be an open node is bounded from below by

$$\left(\frac{\lambda_{\mathrm{I}}}{\lambda_{\mathrm{I}} + \#(B_r(X_i) \cap X_{\mathrm{S}}) - 1 + \#(B_r(X_i) \cap X_{\mathrm{W}})}\right)^{\#(B_r(X_i) \cap X_{\mathrm{S}}) - 1} \ge \left(\frac{\lambda_{\mathrm{I}}}{\lambda_{\mathrm{I}} + n + m}\right)^n.$$

Moreover, by the neighbor constraint imposed by the event B_n , there is also a maximal number of nodes that can be contained in $Q_{3s}(o)$, i.e., there exists $k = k(s, n) \in \mathbb{N}$ such that $\#(X_S \cap Q_{3s}(o)) < k$. This implies that

$$\mathcal{P}(A_m \cap B_n \cap C_{\lambda_{\mathrm{I}}}^{\mathrm{c}}) \le 1 - \left(\frac{\lambda_{\mathrm{I}}}{\lambda_{\mathrm{I}} + n + m}\right)^{nk}$$

where we used that the indicators that nodes are open is a family of independent random variable indexed by the nodes in $X_{\rm S}$. In particular, for given s, m and n, we can now choose $\lambda_{\rm I}$ sufficiently large such that also

$$\mathcal{P}(A_m \cap B_n \cap C^{\mathbf{c}}_{\lambda_{\mathbf{I}}}) < \varepsilon,$$

which concludes the proof.

ECP 25 (2020), paper 25.

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