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# Corrigendum to "Classification with asymmetric label noise: Consistency and maximal denoising"\*

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**Abstract:** We point out a flaw in Lemma 15 of [1]. We also indicate how the main results of that section are still valid using a modified argument.

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The proof of Lemma 15 in [1] is invalid because it uses that the distribution of Y given  $\tilde{Y}$  is independent of X. Unfortunately, this is not true even if  $\tilde{Y}$  given Y is independent of X, as can easily be verified using Bayes rule.

Fortunately, the other results [1] that depend on Lemma 15 can still be proved, using an analogue to Lemma 15 for the label flipping model [2, 3]. We state this alternate lemma and offer a concise proof, using the notation of [1].

**Lemma 1.** Let  $(X, Y, \tilde{Y})$  be jointly distributed. Assume  $\tilde{Y}$  given Y is independent of X, denote  $\rho_i = \Pr(\tilde{Y} \neq i | Y = i)$  and assume  $\rho_0 + \rho_1 < 1$ . Then for any f,

$$(1 - \rho_0 - \rho_1)(R_P(f) - R_P^*) = R_{\tilde{P},\alpha}(f) - R_{\tilde{P},\alpha}^*$$

where

$$\alpha = \frac{1}{2} + \frac{1}{2}(\rho_0 - \rho_1).$$

*Proof.* We have

$$\begin{split} \tilde{\eta}(x) &= \Pr(\tilde{Y} = 1 | X = x) \\ &= \Pr(\tilde{Y} = 1 | Y = 1, X = x) \Pr(Y = 1 | X = x) \\ &+ \Pr(\tilde{Y} = 1 | Y = 0, X = x) \Pr(Y = 0 | X = x) \\ &= (1 - \rho_1) \eta(x) + \rho_0 (1 - \eta(x)) \\ &= (1 - \rho_0 - \rho_1) \eta(x) + \rho_0. \end{split}$$

\*Main article 10.1214/16-EJS1193.

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It follows that

$$\eta(x) - \frac{1}{2} = \frac{\tilde{\eta}(x) - \alpha}{1 - \rho_0 - \rho_1}$$

where

$$\alpha = \frac{1}{2} + \frac{1}{2}(\rho_0 - \rho_1),$$

and therefore

$$(1 - \rho_0 - \rho_1)(R_P(f) - R_P^*) = (1 - \rho_0 - \rho_1)\mathbb{E}_X[|\eta(X) - \frac{1}{2}|\mathbf{1}_{\{u(f(X))\neq u(\eta(X) - \frac{1}{2})\}}]$$
  
=  $\mathbb{E}_X[|\tilde{\eta}(X) - \alpha|\mathbf{1}_{\{u(f(X))\neq u(\tilde{\eta}(X) - \alpha)\}}]$   
=  $R_{\tilde{P},\alpha}(f) - R_{\tilde{P},\alpha}^*$ 

where we have used Eqns. (20) and (21) from [1] (Note that the left-hand side of Eqn. (21) should be  $R_{P,\alpha}(f) - R_{P,\alpha}^*$ ).

**Remark 1.**  $\pi_0 + \pi_1 < 1$  is equivalent to  $\rho_0 + \rho_1 < 1$  provide  $q := \Pr(Y = 1)$  and  $\tilde{q} := \Pr(\tilde{Y} = 1)$  satisfy  $0 < q, \tilde{q} < 1$ . To see this, suppose  $\rho_0 + \rho_1 < 1$ . Bayes' rule gives

$$\pi_0 = \frac{\rho_1 q}{\rho_1 q + (1 - \rho_0)(1 - q)}$$

and

$$\pi_1 = \frac{\rho_0(1-q)}{\rho_0(1-q) + (1-\rho_1)q}$$

and algebra leads to

$$1 - \pi_0 - \pi_1 = (1 - \rho_0 - \rho_1) \frac{q(1 - q)}{[\rho_1 q + (1 - \rho_0)(1 - q)][\rho_0 (1 - q) + (1 - \rho_1)q]}.$$

The ratio on the right is positive provided 0 < q < 1, and thus  $\pi_0 + \pi_1 < 1$ . The reverse implication uses identical reasoning, and establishes that if  $\pi_0 + \pi_1 < 1$ , then

$$\rho_0 = \frac{\pi_1 \tilde{q}}{\pi_1 \tilde{q} + (1 - \pi_0)(1 - \tilde{q})}$$

and

$$\rho_1 = \frac{\pi_0(1-\tilde{q})}{\pi_0(1-\tilde{q}) + (1-\pi_1)\hat{q}}$$

and  $\rho_0 + \rho_1 < 1$ . In light of this remark, it should be stipulated that  $0 < q, \tilde{q} < 1$  throughout Section 7.

**Remark 2.** The remainder of the arguments in Section 7 now carry forward with this new  $\alpha$ . The only difference is that a different estimator for  $\alpha = \frac{1}{2} + \frac{1}{2}(\rho_0 - \rho_1)$  is needed. Such an estimator is obtained from the formulas for  $\rho_0$  and  $\rho_1$  in terms of  $\pi_0, \pi_1$ , and  $\tilde{q}$  in the previous remark. Since  $\tilde{q}$  can be estimated easily from the noisy training data, and  $\pi_0$  and  $\pi_1$  can be estimated in the same way as described in the paper, we still get a consistent estimator for  $\alpha$ , and this estimator will still satisfy Proposition 17 (the rate of convergence of  $\hat{\alpha}$  to  $\alpha$ ) under (C'). Thus, the proofs of the main consistency results still hold.

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### Corrigendum

**Remark 3.** The alternate lemma described above actually allows us to drop the condition (A') in the consistency results, and return to the weaker condition (A).

**Remark 4.** In light of Remark 1 above, the assumption in the co-training theorem of Section 8 can be replaced by the equivalent condition that the sum of the false positive and false negative rates of h is less than 1. In light of Remark 3, the last sentence of Section 8 should be modified so that assuming (A') is no longer required.

## References

- G. Blanchard, M. Flaska, G. Handy, S. Pozzi, and C. Scott. Classification with asymmetric label noise: Consistency and maximal denoising. *Elec*tronic Journal of Statistics, 10:2780–2824, 2016. MR3549019
- [2] N. Natarajan, I. S. Dhillon, P. Ravikumar, and A. Tewari. Learning with noisy labels. In Advances in Neural Information Processing Systems 26, 2013.
- [3] N. Natarajan, I. S. Dhillon, P. Ravikumar, and A. Tewari. Cost-sensitive learning with noisy labels. *Journal of Machine Learning Research*, 18 (155):1-33, 2018. URL http://jmlr.org/papers/v18/15-226.html.