

6. On the Écart between Two “Amounts of Information”

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$$\S 1. \quad d(\lambda_1, \lambda_2; \Lambda) = \sum_{i=0}^{\infty} \Delta P_i \log \left(1 + \frac{\Delta P_i}{P_i} \right)$$

As was shown in the preceding paper the “amount of information”^{2),4)} has been defined by a specified probability space (or distribution), $(R, \mathfrak{X}, \lambda)$, and the partition,¹⁾ Λ , imposed on the space R . And we have conventionally denoted it by $H(\lambda; \Lambda)$. As usual

$$\Lambda : R = \bigcup_{i=0}^{\infty} A_i, \quad A_i \in \mathfrak{X}, \quad A_i \cap A_j = \emptyset \quad (i \neq j).$$

For any two distributions $(R, \mathfrak{X}, \lambda_1)$ and $(R, \mathfrak{X}, \lambda_2)$, providing

(a) $\lambda_1(A_i) = P_i \geq 0, \lambda_2(A_i) = P_i + \Delta P_i \geq 0, \sum_i P_i = \sum_i (P_i + \Delta P_i) = 1$

(b) the series $H(\lambda_1; \Lambda) = \sum_i P_i \log 1/P_i$ and $H(\lambda_2; \Lambda) = \sum_i (P_i + \Delta P_i) \log 1/(P_i + \Delta P_i)$ to converge

(c) $-1 + \alpha \leq \Delta P_i / P_i \leq k; k > 0, 1 > \alpha > 0$ for all i ,

we have directly from the result obtained in the preceding paper

$$0 \leq \Delta H - \sum_{i=0}^{\infty} \Delta P_i \log \frac{1}{P_i + \Delta P_i} \leq \sum_{i=0}^{\infty} \Delta P_i \log \left(1 + \frac{\Delta P_i}{P_i} \right)$$

where $\Delta H = H(\lambda_2; \Lambda) - H(\lambda_1; \Lambda)$.

Denoting $\sum_{i=0}^{\infty} \Delta P_i \log \left(1 + \frac{\Delta P_i}{P_i} \right)$ by $d(\lambda_1, \lambda_2; \Lambda)$, we have easily

$$\begin{aligned} (1.1) \quad (a) \quad & d(\lambda, \lambda; \Lambda) = 0 \\ (b) \quad & d(\lambda_1, \lambda_2; \Lambda) \geq 0 \quad \text{for } \lambda_1 \neq \lambda_2 \\ (c) \quad & d(\lambda_1, \lambda_2; \Lambda) = d(\lambda_2, \lambda_1; \Lambda). \end{aligned}$$

It must be noted that we could not avoid the sign of equality in (b) of (1.1); because even though $\lambda_1 \neq \lambda_2$, we would often have that $\lambda_1(A_i) = \lambda_2(A_i), i = 0, 1, 2, \dots$, for some partitions imposed on R .

To appreciate more fully we consider a distribution $(R, \mathfrak{X}, \lambda_3)$ together with the above $(R, \mathfrak{X}, \lambda_1)$ and $(R, \mathfrak{X}, \lambda_2)$.

Providing again the following

(d) $\lambda_1(A_i) = P_i^{(1)}, \lambda_2(A_i) = P_i^{(2)} = P_i^{(1)} + \Delta P_i^{(1)}, \lambda_3(A_i) = P_i^{(3)} = P_i^{(2)} + \Delta P_i^{(2)}$

(e) $-1 + \alpha \leq \Delta P_i^{(\nu)} / P_i^{(\nu)} \leq k; 1 > \alpha > 0, k > 0, i = 0, 1, 2, \dots, \nu = 1, 2$

we have

$$\begin{aligned} (1.2) \quad & d(\lambda_3, \lambda_1; \Lambda) - \{d(\lambda_1, \lambda_2; \Lambda) + d(\lambda_2, \lambda_3; \Lambda)\} \\ & = \sum_{i=0}^{\infty} (\Delta P_i^{(1)} \log P_i^{(3)} / P_i^{(2)} + \Delta P_i^{(2)} \log P_i^{(2)} / P_i^{(1)}) \end{aligned}$$

and

$$\left. \begin{aligned} & (\Delta P_i^{(1)} \log P_i^{(3)} / P_i^{(2)} + \Delta P_i^{(2)} \log P_i^{(2)} / P_i^{(1)}) \right\} \begin{cases} > 0 \Leftrightarrow \Delta P_i^{(1)} \cdot \Delta P_i^{(2)} > 0 \\ = 0 \Leftrightarrow \Delta P_i^{(1)} \cdot \Delta P_i^{(2)} = 0 \\ < 0 \Leftrightarrow \Delta P_i^{(1)} \cdot \Delta P_i^{(2)} < 0. \end{cases}$$