

This algorithm is a fast and parsimonious way for representing interaction. For example, if, in their spline bases, f and g have p degrees of freedom, then the minimizing product fg has about p degrees of freedom in it. One adds more multiplicative terms until there is no significant decrease in RSS. Furthermore, the multiplicative terms are easy to interpret.

Unfortunately, numerical results indicate that in the nonindependence case, there are a number of local minima in addition to the global minimum. The algorithm always converges, but it may not converge to the global minimum. This makes the selection of a good starting point important. Our experimental results have been that if we use the starting point given by assuming independence, then the iterates have always converged toward the global minimum.

I am currently working on straightening up the details of this representation of bivariate interaction and hope to go public soon.

REFERENCES

- BREIMAN, L. and FRIEDMAN, J. H. (1985). Estimating optimal transformations for multiple regression and correlation (with discussion). *J. Amer. Statist. Assoc.* **80** 580–619.
- BREIMAN, L. and PETERS, S. (1988). Comparing automatic bivariate smoothers. Technical Report, Dept. Statistics, Univ. California, Berkeley.
- FRIEDMAN, J. H. and STUETZLE, W. (1982). Smoothing of scatterplots. Technical Report, Orion 3, Dept. Statistics, Stanford Univ.
- HASTIE, T. and TIBSHIRANI, R. (1988). Comment on “Monotone regression splines in action” by J. O. Ramsay. *Statist. Sci.* **3** 450–456.
- SMITH, P. (1982). Curve fitting and modelling with splines using statistical variable selection techniques. NASA Contractor Report 66034.

DEPARTMENT OF STATISTICS
UNIVERSITY OF CALIFORNIA
BERKELEY, CALIFORNIA 94720

ZEHUA CHEN, CHONG GU AND GRACE WAHBA¹

University of Wisconsin-Madison

We must begin by thanking the authors for a thought-provoking work. As is well known [Kimeldorf and Wahba (1971) and Wahba (1978)], quadratic penalized likelihood estimates (with nonnegative definite penalty functionals) are Bayes estimates. Let $\mathbf{y} = \mathbf{g} + \boldsymbol{\varepsilon}$ with $\mathbf{g} \sim N(0, \boldsymbol{\Sigma})$ and $\boldsymbol{\varepsilon} \sim N(0, \sigma^2 I)$, then

$$\hat{\mathbf{g}} = \boldsymbol{\Sigma}(\boldsymbol{\Sigma} + \sigma^2 I)^{-1} \mathbf{y} = A\mathbf{y}, \quad \text{say,}$$

which also minimizes $(1/\sigma^2)(\mathbf{y} - \hat{\mathbf{g}})'(\mathbf{y} - \hat{\mathbf{g}}) + \mathbf{g}'\boldsymbol{\Sigma}^+\mathbf{g}$, the resulting smoother matrices are all symmetric nonnegative definite with their eigenvalues in $[0, 1)$. This generalizes to the case where $\boldsymbol{\Sigma}$ is improper, which gives eigenvalues $+1$.

¹Research supported by AFOSR Grant AFOSR 87-0171.