Note that the factor in brackets is the same as the last term on the right side of (1) with n replaced by $n - 1$ and $x + 1$ replaced by x. Hence for large x and N,

$$w(n, 1, N, x) \approx \frac{n}{N+n}\left(1 - \frac{x}{N}\right)^{n-1} \cdot \left\{\frac{\left(1 + \dfrac{n}{N-x}\right)^{N-x+n}}{\left(1 + \dfrac{n-1}{N}\right)^{N+n-1}}\right\} \sqrt{1 - \frac{n-1}{N+n-1}}\sqrt{1 + \frac{n-1}{N-x}}.$$

By the same limiting procedure as before,

$$\lim_{x=kN\to\infty} w(n, 1, N, kN) = \frac{n}{N}(1-k)^{n-1}. \tag{5}$$

In any small interval dk, there are Ndk possible values that x can assume; hence the probability that k lies in the interval dk is

$$p(k)\, dk = \frac{n}{N}(1-k)^{n-1}(N\, dk). \tag{6}$$

Therefore, $p(k) = n(1 - k)^{n-1}$. This is exactly the result given by equation (4), but obtained in a somewhat different way.

From the symmetry of the problem, $\lim_{N\to\infty} W(n, 1, N, kN)$ is also the probability that in a large future sample at most a fraction k of the observations will be *less* than the *smallest* observation in the original trial sample of n units. Hence, a life-test of n units may be discontinued as soon as any unit fails and equation (3) will give the probability that in the future at most $100k\%$ of the units will fail in a time shorter than the length of the test. The graphs show W as a function of k for various values of n.

REFERENCE

[1] E. J. Gumbel and H. Von Schelling, "The distribution of the number of exceedances," *Annals of Math. Stat.*, Vol. 21 (1950), pp. 247–262.

CORRECTION TO "THE SAMPLING DISTRIBUTION OF THE RATIO OF TWO RANGES FROM INDEPENDENT SAMPLES"

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In the note mentioned in the title (*Annals of Math. Stat.*, Vol. 21 (1950), pp. 112–116) the distribution given for the above mentioned ratio when the sample values are drawn from a rectangular distribution is correct only when $R \leqq 1$. This is pointed out in an article by P. R. Rider ("The distribution of the