

**ON THE ASYMPTOTIC POWER OF THE ONE-SAMPLE  
KOLMOGOROV-SMIRNOV TESTS<sup>1</sup>**

BY DANA QUADE

*University of North Carolina*

**1. Introduction.** Let  $X_1, X_2, \dots, X_n$  be a random sample of  $n$  observations from some unknown distribution function  $F$ , and let

$$\begin{aligned} F_n(x) &= 0, & x < X_{(1)}, \\ &= i/n, & X_{(i)} \leq x < X_{(i+1)}, \\ &= 1, & x \geq X_{(n)}, \end{aligned}$$

where  $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$  are the ordered observations. If  $H$  is some completely specified continuous distribution function, we may reject the hypothesis that  $F = H$  for large values of  $K_n = \sup_{-\infty < x < \infty} n^{\frac{1}{2}} |F_n(x) - H(x)|$  or  $K_n^+ = \sup_{-\infty < x < \infty} n^{\frac{1}{2}} [F_n(x) - H(x)]$ ; specifically, if  $K_n \geq d_n(\alpha)$  where  $P\{K_n \geq d_n \mid F = H\} = \alpha$  or if  $K_n^+ \geq d_n^+(\alpha)$  where  $P\{K_n^+ \geq d_n^+ \mid F = H\} = \alpha$ . These probabilities do not depend on the true underlying distribution function, so long as it is continuous. The test based on  $K_n$  was first proposed in 1933 by Kolmogorov [9], and the related  $K_n^+$  test was later suggested by Smirnov [12]; they are called respectively the two-sided and one-sided one-sample Kolmogorov-Smirnov tests of goodness of fit. For a fuller expository treatment we refer to the paper by Darling [5] which also includes an extensive bibliography.

The power of the  $K_n$  test when  $F$  is equal to some alternative continuous distribution function  $G$  is

$$P\{K_n \geq d_n(\alpha) \mid F = G\} = P\{\sup_{-\infty < x < \infty} n^{\frac{1}{2}} |F_n(x) - H(x)| \geq d_n(\alpha) \mid F = G\},$$

and if  $\{G_n\}$  is some sequence of alternative distributions, we may define the asymptotic power against  $\{G_n\}$  to be

$$\lim_{n \rightarrow \infty} P\{\sup_{-\infty < x < \infty} n^{\frac{1}{2}} |F_n(x) - H(x)| \geq d_n(\alpha) \mid F = G_n\}$$

if this limit exists. Following Doob [7], we introduce the stochastic process  $Z_n(t) = n^{\frac{1}{2}}(F_n[F^{-1}(t)] - t)$ ,  $0 \leq t \leq 1$ ; then the asymptotic power may be rewritten in the form

$$\lim_{n \rightarrow \infty} P\{\sup_{0 < t < 1} |Z_n(t) - n^{\frac{1}{2}}(H[G_n^{-1}(t)] - t)| \geq d_n(\alpha) \mid F = G_n\}.$$

We may omit the condition  $F = G_n$  in this expression, since all probability state-

---

Received 4 February 1964; revised 16 September 1964.

<sup>1</sup> Supported in part by the United States Air Force through the Air Force Office of Scientific Research of the Air Research and Development Command, under Contract No. AF 49 (638)-261 and in part by the National Institutes of Health, Division of General Medical Sciences, under Public Health Service Research Grant GM-10397-02.