

DENSITY ESTIMATION BY ORTHOGONAL SERIES¹

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There has been much discussion of density estimation by kernel methods (e.g., Whittle (1958), Parzen (1962), Watson and Leadbetter (1963)). Thus, given a random sample x_1, \dots, x_n from the density $f(x)$, the estimator

$$(1) \quad \hat{f}_n = 1/n \sum_{k=1}^n \delta_n(x - x_k)$$

has minimum integrated mean square error (M.I.S.E.) if

$$(2) \quad \phi_{\delta_n} = |\phi_f|^2 / \{n^{-1}[1 + (n-1)|\phi_f|^2]\}$$

where $\phi_{\delta_n} = \int e^{ixt} \delta_n(x) dt$, $\phi_f = \int e^{ixt} f(x) dx$. Whittle introduced Bayesian arguments to find the optimum kernel; he assumed a covariance function for the values of $f(x)$ at different x values.

It is obvious that orthogonal series estimates could be used if it is assumed that $f(x) = \sum_0^\infty \alpha_m \varphi_m(x)$, where $(\varphi_m(x))$ is an orthonormal basis. Several papers (Cencov (1962), van Ryzin (1966), Schwartz (1967), and Kronmal and Tarter (1968)) have considered this possibility. The estimator that springs to mind is

$$(3) \quad f_n^*(x) = \sum_0^\infty \lambda_m(n) \alpha_m \varphi_m(x)$$

where

$$(4) \quad \alpha_m = n^{-1} \sum_{k=1}^n \varphi_m(x_k).$$

and the sequence $\{\lambda_m(n)\}$ is chosen to improve the properties of $f_n^*(x)$ e.g., to make $f_n^*(x)$ a M.I.S.E. estimator in its class. The papers of Cencov, van Ryzin and Schwartz use a special sequence $\{\lambda_m(n)\}$; they set $\lambda_m(n) = 1$, $m = 1, \dots, M(n)$, $\lambda_m(n) = 0$, $m > M(n)$, and concern themselves, in part, with the determination of $M(n)$. Now

$$\begin{aligned} J = E \int (f(x) - f_n^*(x))^2 dx &= \sum_0^\infty E(\alpha_m - \lambda_m(n) \alpha_m)^2 \\ &= \sum_0^\infty \{\alpha_m^2 (1 - \lambda_m(n))^2 + n^{-1} \lambda_m^2(n) \text{var}(\varphi_m(x))\}. \end{aligned}$$

Hence

$$(5) \quad \lambda_m(n) = \alpha_m^2 / [\alpha_m^2 + n^{-1} \text{var}(\varphi_m(x))],$$

i.e.,

$$(5') \quad \lambda_m(n) = \{\alpha_m^2 / E(\varphi_m^2)\} / \{n^{-1}[1 + (n-1) \alpha_m^2 / E(\varphi_m^2)]\}.$$

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