

# ON THE DISTRIBUTIONS OF THE RATIOS OF THE ROOTS OF A COVARIANCE MATRIX AND WILKS' CRITERION FOR TESTS OF THREE HYPOTHESES<sup>1</sup>

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**1. Introduction.** Let  $\mathbf{X}(p \times n)$  be a matrix variate with columns independently distributed as  $N(\mathbf{0}, \Sigma)$ . Then the distribution of the latent roots,  $0 < w_1 \leq \dots \leq w_p < \infty$ , of  $\mathbf{X}\mathbf{X}'$  is first considered in this paper for deriving the distributions of the ratios of individual roots  $w_i/w_j$  ( $i < j = 2, \dots, p$ ). In particular, the distributions of such ratios are derived for  $p = 2, 3$  and 4. The use of these ratios in testing the hypothesis  $\delta\Sigma_1 = \Sigma_2$ ,  $\delta > 0$  unknown, has been pointed out, where  $\Sigma_1$  and  $\Sigma_2$  are the covariance matrices of two  $p$ -variate normal populations.

Further, the non-central distributions of Wilks' criterion,  $\Lambda = W^{(p)} = \prod_{i=1}^p (1 - c_i)$ , are obtained in the following cases: (1) test of  $\Sigma_1 = \delta\Sigma_2$ ,  $\delta > 0$  known, (2) MANOVA and (3) Canonical correlation, where  $c_i$ 's stand for latent roots of a matrix arising in each of the situations. The density functions are given in terms of Meijer's  $G$ -function [12] and for  $p = 2$ , the density and distribution functions are explicitly evaluated. For Case (2), Pillai and Al-Ani [15] have derived the density for  $p = 2, 3$  and 4 using some results on Mellin transforms [2, 3, 4], and Jouris [9] has shown by induction that the  $G$ -function can be expressed in an alternate form than given in the paper; this latter form includes as special cases the results of Pillai and Al-Ani [15].

**2. Distribution of ratios of the roots of a covariance matrix.** The distribution of the latent roots,  $0 < w_1 \leq w_2 \leq \dots \leq w_p < \infty$  of  $\mathbf{X}\mathbf{X}'$  depends only upon the latent roots of  $\Sigma$  and can be given in the form (James [6])

$$(2.1) \quad K(p, n) |\Sigma|^{-\frac{1}{2}n} |\mathbf{W}|^m \{ \exp(-\frac{1}{2} \text{tr } \mathbf{W}) \} \\ \cdot \prod_{i>j} (w_i - w_j) {}_0F_0(\frac{1}{2}(\mathbf{I}_p - \Sigma^{-1}), \mathbf{W}), \quad 0 < w_1 \leq w_2 \leq \dots \leq w_p < \infty,$$

where

$$m = \frac{1}{2}(n - p - 1), \quad K(p, n) = \Pi^{\frac{1}{2}p^2} / \{ 2^{\frac{1}{2}pn} \Gamma_p(\frac{1}{2}n) \Gamma_p(\frac{1}{2}p) \}, \\ \mathbf{W} = \text{diag}(w_1, \dots, w_p),$$

$$(2.2) \quad {}_pF_q(a_1, \dots, a_p; b_1, \dots, b_q; \mathbf{S}, \mathbf{T}) \\ = \sum_{k=0}^{\infty} \sum_{\kappa} [(a_1)_{\kappa} \dots (a_p)_{\kappa}] / [(b_1)_{\kappa} \dots (b_q)_{\kappa}] \cdot C_{\kappa}(\mathbf{S}) C_{\kappa}(\mathbf{T}) / [C_{\kappa}(\mathbf{I}_p) k!]$$

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