

DISCUSSION

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As stated by the author in Section 3, his emphasis in the study of

$$(1) \quad Dy = \Gamma(\theta) = \sum_{\alpha=1}^s \theta_{\alpha} A_{\alpha}$$

is different from mine [Anderson (1969, 1970, 1973)]. A brief exposition of this other point of view might put the present paper into a larger perspective. I shall use the notation of the present paper as well as the assumption of $Ey = 0$ and corresponding modifications in statements.

Anderson considered N observations from $N(0, \Gamma(\theta))$; in the present paper $N = 1$. We suppose A_1, \dots, A_s to be symmetric matrices and that for some vector $\theta = (\theta_1, \dots, \theta_s)'$ the matrix $\sum_{\alpha=1}^s \theta_{\alpha} A_{\alpha}$ is positive definite. Then the derivative equations for the maximum likelihood estimate of θ are

$$(2) \quad \text{tr} \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} A_b = y' \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} A_b \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} y, \quad b = 1, \dots, s.$$

If there are several vectors $\hat{\theta}$ satisfying (2), the vector minimizing $|\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha}|$ is taken. Rewriting (2) as

$$(3) \quad \sum_{c=1}^s \text{tr} \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} A_b \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} A_c \hat{\theta}_c = y' \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} A_b \left(\sum_{\alpha=1}^s \hat{\theta}_{\alpha} A_{\alpha} \right)^{-1} y, \quad b = 1, \dots, s,$$

suggests an iterative procedure for solving (2). Let the vector $\hat{\theta}^{(0)}$ be an initial estimate, and define $\Gamma(\hat{\theta}^{(i)}) = \sum_{\alpha=1}^s \hat{\theta}_{\alpha}^{(i)} A_{\alpha}$, $i = 0, 1, \dots$. The i th stage of the iteration is solving the linear equations

$$(4) \quad \sum_{c=1}^s \text{tr} \Gamma^{-1}(\hat{\theta}^{(i-1)}) A_b \Gamma^{-1}(\hat{\theta}^{(i-1)}) A_c \hat{\theta}_c^{(i)} = y' \Gamma^{-1}(\hat{\theta}^{(i-1)}) A_b \Gamma^{-1}(\hat{\theta}^{(i-1)}) y, \quad b = 1, \dots, s.$$

This is the method of scoring.

If the matrices $\{A_{\alpha}\}$ commute [the author's (c)], there exists an orthogonal matrix B such that $A_{\alpha} = B \Lambda_{\alpha} B'$ where Λ_{α} is diagonal with diagonal elements $\lambda_{t\alpha}$, $t = 1, \dots, n$. Then

$$(5) \quad B' \Gamma(\theta) B = \sum_{\alpha=1}^s \theta_{\alpha} \Lambda_{\alpha}$$

¹Work sponsored by the Army Research Office Contract No. DAAG 29-85-0239 at Stanford University and written while the author was a National Research Council Resident Research Associate at the Naval Postgraduate School at Monterey, California.

