

**TABLE OF THE ASYMPTOTIC DISTRIBUTION
OF THE SECOND EXTREME**

BY E. J. GUMBEL AND J. ARTHUR GREENWOOD

New York City and Manhattan Life Insurance Company

The asymptotic distributions of the extreme values taken from an initial distribution of the exponential type are now widely used, for example in flood control [6] and in problems connected with the breaking strength of material [1]. Therefore, the corresponding distribution of the penultimate (and of the second) value may also be of practical interest.

Let $F(x)$ be the initial probability; let $f(x) = F'(x)$ be the initial density (distribution). Let n be a large sample size; let the rank m ($m \ll n$) be counted from the top. Finally, let the parameters u_m and α_m be defined as the solutions of

$$(1) \quad F(u_m) = 1 - m/n; \quad \alpha_m = nf(u_m)/m.$$

Then the asymptotic distribution $\varphi_m(x_m)$ of the m th largest value x_m is [2]

$$\varphi_m(x_m) = \frac{m^m}{\Gamma(m)} \alpha_m \exp[-my_m - me^{-y_m}],$$

where

$$y_m = \alpha_m(x_m - u_m),$$

provided that the initial distribution is of the exponential type. The asymptotic distribution ${}_m\varphi(m, x)$ of the m th smallest value is

$${}_m\varphi(m, x) = \varphi_m(-x_m).$$

The probability function $\Phi_m(x_m)$ is obtained from

$$\begin{aligned} \Phi_m(x_m) &= \frac{m^m}{\Gamma(m)} \int_{-\infty}^{y_m} \exp[-my - me^{-y}] dy \\ &= \frac{1}{\Gamma(m)} \int_{m\alpha_m^{-1}e^{-y_m}}^{\infty} z^{m-1} e^{-z} dz, \end{aligned}$$

whence

$$(2) \quad \Phi_m(x_m) = 1 - I(t_m, m - 1),$$

where

$$t_m = \sqrt{m} e^{-y_m}$$

and I is the incomplete Gamma function ratio of Karl Pearson [5]. In the special case $m = 2$, the probability function of the penultimate value is

$$(3) \quad \Phi_2(x_2) = 1 - I(\sqrt{2} e^{-y_2}, 1).$$

The modal penultimate value is, of course, u_2 , and the intervals corresponding