

We can characterize (2) as saying: Let $\hat{\pi}$ be the linear estimate of minimum variance of π , which is a linear combination of π_1, \dots, π_k , and let $\hat{\nu}(\pi)$ be the estimate of the variance of $\hat{\pi}$ based on s^2 . Then the confidence interval statements

$$\hat{\pi} - [kF_{\alpha} \hat{\nu}(\pi)]^{1/2} \leq \pi \leq \hat{\pi} + [kF_{\alpha} \hat{\nu}(\pi)]^{1/2},$$

simultaneously for all π , are correct with probability $1 - \alpha$. This result is contained in [1] and [2].

The quantity $D^2 = \sum \pi_i^2$ is a conventional measure of the "distance" of the null hypothesis that $\pi_1 = \dots = \pi_k = 0$ from the true state of nature. The power of the analysis of variance test depends only on D^2/σ^2 . Hence it would be useful for the experimenter to obtain some information about D .

Making use of the triangle inequality, it follows from (1) that

$$1 - \alpha \leq \Pr\{(\sum \hat{\pi}_i^2)^{1/2} - (kF_{\alpha} s^2)^{1/2} \leq D \leq (\sum \hat{\pi}_i^2)^{1/2} + (kF_{\alpha} s^2)^{1/2}\}.$$

The quantity $\sum \hat{\pi}_i^2 = Q_1^2$ is what the experimenter calculates as the "sum of squares due to hypothesis." Hence, instead of just making the statements about the functions π , we can make the simultaneous estimates

$$\begin{aligned} \hat{\pi} - [kF_{\alpha} \hat{\nu}(\pi)]^{1/2} \leq \pi \leq \hat{\pi} + [kF_{\alpha} \hat{\nu}(\pi)]^{1/2}, & \quad \text{for all } \pi = \sum a_i \pi_i, \\ Q_1 - (kF_{\alpha} s^2)^{1/2} \leq D \leq Q_1 + (kF_{\alpha} s^2)^{1/2}, & \end{aligned}$$

with the probability of being correct equal to $1 - \alpha$.

REFERENCES

[1] H. SCHEFFÉ, "A method for judging all contrasts in the analysis of variance," *Biometrika*, Vol. 40 (1953), pp. 87-104.
 [2] S. N. ROY AND R. C. BOSE, "Simultaneous confidence interval estimation," *Ann. Math. Stat.*, Vol. 24 (1953), pp. 513-536.

AN INEQUALITY ON POISSON PROBABILITIES

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This note proves an inequality concerning the exponential series or Poisson distribution, however one prefers to view the matter. Specifically, it will be shown that if $[\lambda]$ is the greatest integer not exceeding λ ,

$$1) \quad \sum_{j=0}^{[\lambda]} \frac{\lambda^j}{j!} > \begin{cases} e^{\lambda-1} & \text{for all } \lambda \geq 0; \\ \frac{1}{2}e^{\lambda} & \text{for all integral } \lambda > 0. \end{cases}$$

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