

## ON THE MEDIAN OF THE DISTRIBUTION OF EXCEEDANCES

BY K. SARKADI

*Mathematical Institute of the Hungarian Academy of Sciences,  
Budapest, Hungary*

The distribution of exceedances may be defined by the following formula (see, e.g. [2]), where  $x$  corresponds to the number of exceedances

$$(1) \quad w(n_1, m, n_2, x) = \frac{\binom{n_1 + n_2 - m - x}{n_1 - m} \binom{x + m - 1}{m - 1}}{\binom{n_1 + n_2}{n_1}},$$

$$x = 0, 1, \dots, n_2; n_1, n_2, m \leq n$$

given natural numbers.

There are known—among others—the following two fundamentally equivalent models or representations of this distribution [2], [3], [4]:

*A. Exceedances.* We have two random samples of sizes  $n_1$  and  $n_2$ , respectively, from the same continuous distribution. The number of exceedances is defined as the number of elements of the second sample which surpass at least  $n_1 - m + 1$  elements of the first, for a fixed natural number  $m \leq n_1$ . The distribution of the number of exceedances is given by formula (1).

*B. Pascal model without replacement.* An urn contains  $n_1$  black and  $n_2$  red balls. We draw balls from the urn until we have drawn  $m$  black balls. The distribution of the number of the red balls drawn is given by (1).

In [1], Gumbel proved that, for  $n_1 = n_2 = n$ , the median of the number of exceedances is  $m - 1$ , more precisely that

$$(2) \quad W(n, m, n, m - 1) = \frac{1}{2},$$

where  $W$  is the cumulated form of  $w$ ,

$$W(n_1, m, n_2, x) = \sum_{y=0}^x w(n_1, m, n_2, y)$$

In this paper a simple proof of this result is given.

We shall, in fact, prove the following more general result:

$$(3) \quad W(n_1, m_1, n_2, m_2 - 1) + W(n_2, m_2, n_1, m_1 - 1) = 1.$$

In terms of model *A*, the  $m_1$ th element of the first sample exceeds the  $m_2$ th element of the second sample if and only if the number of exceedances ( $y$ ) takes one of the values  $0, 1, \dots, m_2 - 1$ , these possibilities being mutually exclusive.

Received May 20, 1959; revised August 10, 1959.