

# ON A MODIFICATION OF CERTAIN RANK TESTS

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**1. Introduction.** This note is concerned with a modification of certain well-known rank tests based on the theory of Chernoff-Savage [2]. Let  $X_1, X_2, \dots, X_m$  and  $Y_1, Y_2, \dots, Y_n$  be two independent samples from populations with continuous distribution functions  $F$  and  $G$ . Our interest is to test the hypothesis  $H_0: F \equiv G$ . For the problem for location, we assume  $G(x) = F(x + \theta)$  and want to test  $H_0: \theta = 0$  against the alternative  $H_1: \theta > 0$ . On the other hand, the problem for scale is to test  $H_0: \theta = 1$  against  $H_2: \theta > 1$  under the restrictions  $G(x) = F(x/\theta)$  and  $F(0) = \frac{1}{2}$ . There are well-known test statistics, Wilcoxon's  $U$  [4] for the location problem and Mood's  $M$  [3] and Ansari-Bradley's  $S$  [1] for the scale problem. We shall modify these statistics to raise the asymptotic efficiency of the corresponding tests.

**2. Test statistics.** We adopt the following statistics for  $k > 0$ ,

$$(1) \quad mU_k = \sum_{i=1}^N (i/N)^k Z_i, \quad N = m + n,$$

$$(2) \quad mS_k = \sum_{i=1}^p (i/N)^k Z_i + \sum_{i=p+1}^N \{(N+1-i)/N\}^k Z_i, \quad p = [\frac{1}{2}(N+1)],$$

$$(3) \quad mM_k = \sum_{i=1}^N |i/N - (N+1)/2N|^k Z_i,$$

where  $Z_i$  is 1(0) if the  $i$ th smallest among the combined sample is  $X(Y)$ . We may directly apply the theorem of Chernoff-Savage for  $U_k$  and  $S_k$  and show their asymptotic normality with the mean  $\mu_k(\theta)$  and the variance  $\sigma_k^2(\theta)$ ,

$$(4) \quad \begin{aligned} \mu_k(\theta) &= \int_{-\infty}^{\infty} H(x)^k dF(x), && \text{for } U_k \\ &= \int_{-\infty}^0 H(x)^k dF(x) + \int_0^{\infty} \{1 - H(x)\}^k dF(x), && \text{for } S_k \end{aligned}$$

and under the hypothesis  $H_0$

$$(5) \quad \begin{aligned} N\sigma_k^2 &= [(1 - \lambda_N)/\lambda_N]k^2/(2k+1)(k+1)^2, && \text{for } U_k, \\ &= [(1 - \lambda_N)/\lambda_N]k^2/4^k(2k+1)(k+1)^2, && \text{for } S_k, \end{aligned}$$

where  $H(x)$  is the combined population distribution function and  $\lambda_N = m/N$ . Though the assumption (4) of their theorem (see [2]) is not satisfied for  $M_k$ , we may similarly prove its asymptotic normality from the fact that the following

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