

# THE TREATMENT OF TIES IN THE WILCOXON TEST<sup>1</sup>

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**1. Introduction.** Let  $(X_1, \dots, X_n)$  be a sample of  $n$  independent observations from a distribution  $F$ , and  $(Y_1, \dots, Y_m)$  be a sample of independent observations from  $G$ . Then, if all  $m + n$  observations are different, the Wilcoxon test will reject the hypothesis  $F = G$ , when the sum  $S_{nm}$  of the ranks  $R_i$  of the  $X_i$  is too small or too large.

For the case with a positive probability of ties two procedures have been proposed. One is to order the tied observations randomly, the other is to replace  $S_{nm}$  by  $S'_{nm} = \sum_{i=1}^n R'_i$ . Here  $R'_i = \text{midrank}(X_i) = \frac{1}{2}[N_1(i) + N_2(i) + 1]$ .  $N_1(i)$  is the number of observations smaller than  $X_i$  and  $N_2(i)$  is the number of observations (including  $X_i$ ) not larger than  $X_i$ .

If there are only finitely many values  $\xi_k$  at which ties may occur and if  $p_k = P\{X_1 = \xi_k\}$ , then as shown by Putter [3] under certain regularity conditions the asymptotic relative efficiency of the "randomized" with respect to the mid-rank test is  $1 - \sum_{k=1}^n p_k^3$ . Using a slight modification of Putter's argument this note will show that this conclusion is still true if  $p_k = P\{X_1 = \xi_k\} > 0$  and  $q_k = P\{Y_1 = \xi_k\} > 0$  for infinitely many values  $\xi_k$ . The result is illustrated by applying it to certain parametric families of distributions, for which the efficiency of the midrank test has been investigated by Chanda [1]. Putter's notation will be used throughout the paper.

**2. The basic theorem.** Following Putter, let for

$$k = 1, 2, \dots, p_k = P\{X_1 = \xi_k\} > 0, \quad q_k = P\{Y_1 = \xi_k\} > 0;$$

$U_k =$  number of  $X$ 's equal to  $\xi_k$ ,  $V_k =$  number of  $Y$ 's equal to  $\xi_k$ ;  $U = (U_1, U_2, \dots)$ ,  $V = (V_1, V_2, \dots)$ ,  $W = U + V$ ;  $S_{nm}^0 =$  any statistic whose distribution is that of  $S_{nm}$  under  $F = G$ ;  $\mu_{nm} = ES_{nm}^0 = n(n + m + 1)/2$ ,  $\sigma_{nm}^2 = \text{Var } S_{nm}^0 = nm(n + m + 1)/12$ ;  $T_{nm}^0 = (S_{nm}^0 - \mu_{nm})/\sigma_{nm}$ .

Then the following theorem connects the asymptotic distributions of  $S_{nm}$  and of  $S'_{nm}$ .

**THEOREM 1.** *If  $m/n$  converges to a positive number  $c$  as  $m, n \rightarrow \infty$ , then we have for any pair  $(F, G)$  of distributions with common discontinuities  $\xi_k$ ,  $k = 1, 2, \dots$*

$$(2.1) \quad \sigma_{U_k V_k}^2 / \sigma_{nm}^2 = a_k^2 \rightarrow_P b_k^2 = (1 + c)^{-1} p_k q_k(\theta) [p_k + c q_k(\theta)]$$

$$(2.2) \quad (S_{nm} - ES_{nm}) / \sigma_{nm} = T_{nm} \rightarrow_{\mathcal{L}} N(0, b^2)$$

$$(2.3) \quad (S'_{nm} - ES_{nm}) / \sigma_{nm} = T'_{nm} \rightarrow_{\mathcal{L}} N(0, \bar{b}^2),$$

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