

SOME THEOREMS IN AN EXTENDED RENEWAL THEORY, IV

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1. Consider a sequence of non-negative independent random variables $X_1, X_2, \dots$ and let $N(t)$ be the number of sums $X_1, X_1 + X_2, \dots$ which are less than t . In other words $N(t)$ is the random variable such that

$$(1.1) \quad \sum_{k=1}^{N(t)} X_k < t \leq \sum_{k=1}^{N(t)+1} X_k.$$

The existence of

$$(1.2) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n E(X_k) = m,$$

with some conditions on $X_1, X_2, \dots$, ensures

$$(1.3) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T [H(t+h) - H(t)] dt = \frac{h}{m},$$

where

$$(1.4) \quad H(t) = E\{N(t)\} = \sum_{n=1}^{\infty} P(X_1 + \dots + X_n < t).$$

This problem is a renewal theorem in a wide sense and has been treated by Kawata [6] and extended by the author [3], [4]. Moreover, we have as an extension of (1.3) above that

$$(1.5) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T dt \int_0^t \phi(t-u) dH(u) = \frac{1}{m} \int_0^{\infty} \phi(u) du$$

where $\phi(u)$ is a Baire function integrable over $(0, \infty)$. This fact has been demonstrated by the author [5] in a somewhat extended form. By the similar way as in Theorem 1 of [5], we can prove the following

THEOREM 1. *Let $X_k, k=1, 2, \dots$, be non-negative, independent random variables having finite mean values. If there exist positive constants L and K such that*

$$E(X_k) \geq L, \quad \text{Var}(X_k) \leq K \quad \text{for } k=1, 2, \dots$$

and

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n E(X_k) = m, \quad m > 0$$

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exists, then

$$\int_0^T dt \int_0^t \phi(t-u) dN(u)$$

is a random variable, i. e. a measurable function defined on the probability space and it follows that

$$(1.6) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T dt \int_0^t \phi(t-u) dN(u) = \frac{1}{m} \int_0^\infty \phi(u) du$$

with probability 1, where $\phi(u)$ is a Baire function integrable over $(0, \infty)$.

Proof. $\phi(x)$ being a continuous function,

$$\int_0^T dt \int_0^t \phi(t-u) dN(u)$$

is measurable, because this double integral may be considered as the limit of the Riemann sum $\sum \phi(t-u) \Delta N(u) \Delta t$. Using the transfinite induction, we can see the measurability of the integral in which $\phi(u)$ is any Baire function integrable over $(0, \infty)$. Applying Theorem 1 in [2] to the sequence X_k , $k=1, 2, \dots$, we get

$$\lim_{T \rightarrow \infty} \frac{N(T)}{T} = \frac{1}{m} \quad (\text{a. s.}),$$

from which (1.6) is obtained by the similar way as in Theorem 1 of [5].

In this Theorem 1, if $\phi(u)$ is the characteristic function of the semi-closed interval $(0, h]$, (1.6) becomes

$$(1.7) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T [N(t) - N(t-h)] dt = \frac{h}{m} \quad (\text{a. s.}),$$

where $N(t-h) = 0$ for $t-h < 0$. So (1.6) may be considered to show an asymptotic property of $N(t+h) - N(t)$, i. e. the renewal number in $(t, t+h]$ as $t \rightarrow \infty$. In the next section, we assume the law of iterated logarithm on X_k , $k=1, 2, \dots$, to make Theorem 1 more complete.

2. In this section we make the following assumptions:

(2.1) X_k , $k=1, 2, \dots$, are non-negative independent random variables,

(2.2) $0 < m < \infty$ and $\frac{1}{n} \sum_{k=1}^n E(X_k) = m + o\left(\sqrt{\frac{\log \log n}{n}}\right)$ ($n \rightarrow \infty$),

(2.3) $B_n \equiv \sum_{k=1}^n \text{Var}(X_k) \rightarrow \infty$ ($n \rightarrow \infty$),

(2.4) $X_k - E(X_k)$, $k=1, 2, \dots$, obey the law of iterated logarithm, and

(2.5) $\phi(x)$ is a Baire function integrable over $(0, \infty)$ and

$$\int_x^\infty |\phi(u)| du = O\left(\frac{\log \log x}{x}\right) \quad (x \rightarrow \infty).$$

In the first place, we shall prepare the following Lemma which is found as Theorem 4 in [2].

LEMMA. *Under the conditions (2.1)–(2.4), we have*

$$(2.6) \quad \lim_{T \rightarrow \infty} \frac{\left| N(T) - \frac{T}{m} \right|}{\sqrt{2T \log \log T}} \leq \frac{\sqrt{B}}{\sqrt{m^3}} \quad (\text{a. s.}),$$

where

$$B = \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(X_k).$$

THEOREM 2. *Under the conditions (2.1)–(2.5), we have*

$$(2.7) \quad \lim_{T \rightarrow \infty} \frac{\left| \int_0^T dt \int_0^t \phi(t-u) dN(u) - \frac{T}{m} \int_0^\infty \phi(u) du \right|}{\sqrt{2T \log \log T}} \leq \frac{3\sqrt{B}}{\sqrt{m^3}} \|\phi\| + \frac{\|\phi\| + C/2}{m}$$

with probability 1, where

$$B = \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(X_k), \quad \|\phi\| = \int_0^\infty |\phi(u)| du$$

and

$$C = \overline{\lim}_{x \rightarrow \infty} \frac{x}{\log \log x} \int_x^\infty |\phi(u)| du.$$

Proof. We have

$$(2.8) \quad \begin{aligned} & \int_0^T dt \int_0^t \phi(t-u) dN(u) - \frac{T}{m} \int_0^\infty \phi(u) du \\ &= \left[N(T) - \frac{T}{m} \right] \int_0^\infty \phi(v) dv - \int_0^{T - \sqrt{2T \log \log T}} dN(u) \int_{T-u}^\infty \phi(v) dv \\ & \quad - \int_{T - \sqrt{2T \log \log T}}^T dN(u) \int_{T-u}^\infty \phi(v) dv \\ &= P + Q + R, \quad \text{say.} \end{aligned}$$

Then, we have

$$(2.9) \quad \lim_{T \rightarrow \infty} \frac{|P|}{\sqrt{2T \log \log T}} \leq \frac{\sqrt{B}}{\sqrt{m^3}} \|\phi\|$$

by (2.6), because $|P| \leq |N(T) - T/m| \cdot \|\psi\|$.

In the next place, we have for an arbitrary small $\varepsilon > 0$ with probability 1 that

$$\begin{aligned} \frac{|Q|}{\sqrt{2T \log \log T}} &\leq \frac{N(T)}{T} \sqrt{\frac{T}{2 \log \log T}} \int_{\sqrt{2T \log \log T}}^{\infty} |\psi(v)| dv \\ &< (C + \varepsilon) \cdot \frac{N(T)}{T} \cdot \frac{\log \log \sqrt{2T \log \log T}}{2 \log \log T} < \frac{C + 2\varepsilon}{2} \cdot \frac{N(T)}{T} \end{aligned}$$

for sufficiently large T . So we have

$$(2.10) \quad \lim_{T \rightarrow \infty} \frac{|Q|}{\sqrt{2T \log \log T}} \leq \frac{C}{2m} \quad (\text{a. s.}).$$

By (2.6), we have for an arbitrary small $\varepsilon > 0$ with probability 1 that

$$N(T) < \frac{T}{m} + \left(\frac{\sqrt{B}}{\sqrt{m^3}} + \varepsilon \right) \sqrt{2T \log \log T}$$

and

$$\frac{1}{m} (T - \sqrt{2T \log \log T}) - \left(\frac{\sqrt{B}}{\sqrt{m^3}} + \varepsilon \right) \sqrt{2T \log \log T} < N(T - \sqrt{2T \log \log T})$$

for sufficiently large T , so that we get with probability 1 that

$$\begin{aligned} |R| &\leq \|\psi\| [N(T) - N(T - \sqrt{2T \log \log T})] \\ &\leq \|\psi\| \left[\frac{1}{m} + 2 \left(\frac{\sqrt{B}}{\sqrt{m^3}} + \varepsilon \right) \right] \sqrt{2T \log \log T} \end{aligned}$$

for sufficiently large T , which implies

$$(2.11) \quad \lim_{T \rightarrow \infty} \frac{|R|}{\sqrt{2T \log \log T}} \leq \left(\frac{1}{m} + \frac{2\sqrt{B}}{\sqrt{m^3}} \right) \|\psi\|.$$

Summing up (2.8), (2.9), (2.10) and (2.11), we get (2.7) with probability 1.

3. In this section, we shall give some results on the process $X(t)$ defined as follows:

$$(3.1) \quad \begin{aligned} X(t) &= t - \sum_{k=1}^{N(t)} X_k & \text{if } N(t) \geq 1, \\ &= X_0 + t & \text{if } N(t) = 0 \end{aligned}$$

where X_0 is a non-negative random variable. Doob [1] has shown that

$$(3.2) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T X(t) dt = \frac{E(X_k^2)}{2E(X_k)} \quad (\text{a. s.}),$$

when X_k , $k = 1, 2, \dots$, are non-negative, mutually independent, identically distributed random variables.

Even in the case where X_k , $k = 1, 2, \dots$, do not necessarily have the same distribution, we have

$$(3.3) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T X(t) dt = \frac{b}{2m} \quad (\text{a. s.})$$

by assuming the existence of the limits

$$m = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n E(X_k) \quad \text{and} \quad b = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n E(X_k^2)$$

with some additional conditions on $X_1, X_2, \dots$. This fact has been noticed by the author [2]. In the following, we shall try to make (3.3) above more complete. In addition to the assumptions (2.1)–(2.4), we assume the following conditions:

$$(3.4) \quad X_k^2 - E(X_k^2), \quad k = 1, 2, \dots, \quad \text{obey the law of iterated logarithm}$$

and

$$(3.5) \quad \frac{1}{n} \sum_{k=1}^n E(X_k^2) = b + o\left(\sqrt{\frac{\log \log n}{n}}\right) \quad (n \rightarrow \infty).$$

THEOREM 3. *Under the conditions (2.1)–(2.4), (3.4) and (3.5), we have*

$$(3.6) \quad \lim_{T \rightarrow \infty} \frac{\left| \int_0^T X(t) dt - \frac{b}{2m} T \right|}{\sqrt{2T \log \log T}} \leq \frac{b\sqrt{B}}{2\sqrt{m^3}} + \frac{\sqrt{V}}{2\sqrt{m}}$$

with probability 1, where

$$B = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(X_k) \quad \text{and} \quad V = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(X_k^2).$$

Proof. Since

$$\begin{aligned} \int_0^T X(t) dt &= \int_0^{X_1} (X_0 + t) dt + \int_{X_1}^{X_1+X_2} (t - X_1) dt + \dots + \int_{S_{N(T)}}^T (t - S_{N(T)}) dt \\ &= X_0 X_1 + \frac{1}{2} \sum_{k=1}^{N(T)} X_k^2 + \frac{1}{2} (T - S_{N(T)})^2, \end{aligned}$$

we have

$$(3.7) \quad X_0 X_1 + \frac{1}{2} \sum_{k=1}^{N(T)} X_k^2 < \int_0^T X(t) dt \leq X_0 X_1 + \frac{1}{2} \sum_{k=1}^{N(T)+1} X_k^2.$$

By (3.4), we have for an arbitrary small $\varepsilon > 0$ with probability 1 that

$$\begin{aligned}
(3.8) \quad & \sum_{k=1}^{N(T)} E(X_k^2) - (1 + \varepsilon) \sqrt{2V_{N(T)} \log \log V_{N(T)}} \\
& < \sum_{k=1}^{N(T)} X_k^2 \\
& < \sum_{k=1}^{N(T)} E(X_k^2) + (1 + \varepsilon) \sqrt{2V_{N(T)} \log \log V_{N(T)}}
\end{aligned}$$

for sufficiently large T , where $V_n = \sum_{k=1}^n \text{Var}(X_k^2)$. On the other hand, (3.5) gives

$$(3.9) \quad \sum_{k=1}^{N(T)} E(X_k^2) = bN(T) + o(\sqrt{N(T) \log \log N(T)}) \quad (T \rightarrow \infty) \quad (\text{a. s.})$$

which implies with (3.7) and (3.8) with probability 1 that

$$\begin{aligned}
(3.10) \quad & \frac{b}{2} \left[N(T) - \frac{T}{m} \right] - \frac{1}{2} (1 + \varepsilon) \sqrt{2V_{N(T)} \log \log V_{N(T)}} + o(\sqrt{N(T) \log \log N(T)}) \\
& < \int_0^T X(t) dt - \frac{b}{2m} T \\
& < \frac{b}{2} \left[N(T) - \frac{T}{m} \right] + X_0 X_1 + \frac{b}{2} + \frac{1}{2} (1 + \varepsilon) \sqrt{2V_{N(T)+1} \log \log V_{N(T)+1}} \\
& \quad + o(\sqrt{N(T) \log \log N(T)})
\end{aligned}$$

for sufficiently large T . Noticing that we have

$$\lim_{T \rightarrow \infty} \frac{N(T) \log \log N(T)}{T \log \log T} = \frac{1}{m} \quad (\text{a. s.})$$

and

$$\varlimsup_{T \rightarrow \infty} \frac{V_{N(T)} \log \log V_{N(T)}}{T \log \log T} \leq \frac{V}{m} \quad (\text{a. s.})$$

by the similar method as in the proof of Theorem 4 of [2], we can show (3.6) from (3.10) and (2.6).

REMARK. Assuming the weaker condition instead (3.5) that

$$(3.5') \quad \frac{1}{n} \sum_{k=1}^n E(X_k^2) = b + O\left(\sqrt{\frac{\log \log n}{n}}\right) \quad (n \rightarrow \infty),$$

we can maintain Theorem 3 by only adding

$$C_1 = \frac{1}{2} \varlimsup_{n \rightarrow \infty} \sqrt{\frac{n}{\log \log n}} \left| \frac{1}{n} \sum_{k=1}^n E(X_k^2) - b \right|$$

to the right side of (3.6).

4. In this section, we make the following assumptions:

(4.1) $X_k, Y_k, \dots, Z_k, \quad k=1, 2, \dots,$ are non-negative independent random variables,

(4.2) $0 < m_x, m_y, \dots, m_z < \infty$ and

$$\begin{aligned} \frac{1}{n} \sum_{k=1}^n E(X_k) &= m_x + o\left(\sqrt{\frac{\log \log n}{n}}\right), \\ \frac{1}{n} \sum_{k=1}^n E(Y_k) &= m_y + o\left(\sqrt{\frac{\log \log n}{n}}\right), \\ &\dots\dots\dots, \\ \frac{1}{n} \sum_{k=1}^n E(Z_k) &= m_z + o\left(\sqrt{\frac{\log \log n}{n}}\right) \quad (n \rightarrow \infty), \end{aligned}$$

(4.3) $B_{x,n} \equiv \sum_{k=1}^n \text{Var}(X_k) \rightarrow \infty,$
 $B_{y,n} \equiv \sum_{k=1}^n \text{Var}(Y_k) \rightarrow \infty,$
 $\dots\dots\dots,$
 $B_{z,n} \equiv \sum_{k=1}^n \text{Var}(Z_k) \rightarrow \infty \quad (n \rightarrow \infty),$

(4.4) $X_k - E(X_k), \quad k=1, 2, \dots; \quad Y_k - E(Y_k), \quad k=1, 2, \dots; \quad \dots;$
 $Z_k - E(Z_k), \quad k=1, 2, \dots,$ obey the law of iterated logarithm respectively,

(4.5) $\phi(x, y, \dots, z)$ is a non-decreasing function defined on the domain:
 $x > 0, y > 0, \dots, z > 0$ and satisfies Lipschitz condition,
i.e. that there exist positive constants $A, B, \dots, C$ such that

$$0 \leq \phi(x+h, y+k, \dots, z+l) - \phi(x, y, \dots, z) \leq Ah + Bk + \dots + Cl$$

for all positive $x, y, \dots, z, h, k, \dots, l$

and

(4.6) $\phi(x, y, \dots, z)$ is positive homogeneous and there exists a positive constant γ such that

$$\phi(x, y, \dots, z) \geq \gamma \cdot \min(x, y, \dots, z).$$

As examples of the function ϕ satisfying the conditions (4.5) and (4.6), we can give $\max(x, y, \dots, z)$, $\min(x, y, \dots, z)$ and $\sqrt[p]{x^p + y^p + \dots + z^p}$ with $p \geq 1$.

The author [3], [5] has discussed some properties of the renewal type on W_n , where

$$(4.7) \quad W_n = \phi\left(\sum_{k=1}^n X_k, \sum_{k=1}^n Y_k, \dots, \sum_{k=1}^n Z_k\right),$$

that is, asymptotic properties of $M(t)$ defined as follows:

$$(4.8) \quad W_{M(t)} < t \leq W_{M(t)+1}.$$

It can be shown that $M(t)$ is uniquely defined with probability 1 for each $t > 0$. We shall derive more complete properties on $M(t)$ than those in [3], [5].

THEOREM 4. *Under the conditions (4.1)–(4.6), we have*

$$(4.9) \quad \lim_{T \rightarrow \infty} \frac{\left| M(T) - \frac{T}{\phi(m_x, m_y, \dots, m_z)} \right|}{\sqrt{2T \log \log T}} \leq \frac{A\sqrt{B_x} + B\sqrt{B_y} + \dots + C\sqrt{B_z}}{\sqrt{\phi(m_x, m_y, \dots, m_z)^3}}$$

with probability 1, where

$$B_x = \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(X_k), \quad B_y = \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(Y_k), \quad \dots, \quad B_z = \overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \text{Var}(Z_k).$$

Proof. By (4.2)–(4.6), we have for an arbitrary small $\varepsilon > 0$ with probability 1 that

$$\begin{aligned} & n\phi(m_x, m_y, \dots, m_z) + o(\sqrt{n \log \log n}) \\ & - (1 + \varepsilon)(A\sqrt{2B_{x,n} \log \log B_{x,n}} + B\sqrt{2B_{y,n} \log \log B_{y,n}} + \dots \\ & \quad + C\sqrt{2B_{z,n} \log \log B_{z,n}}) \\ & < W_n \\ & < n\phi(m_x, m_y, \dots, m_z) + o(\sqrt{n \log \log n}) \\ & \quad + (1 + \varepsilon)(A\sqrt{2B_{x,n} \log \log B_{x,n}} + B\sqrt{2B_{y,n} \log \log B_{y,n}} + \dots \\ & \quad + C\sqrt{2B_{z,n} \log \log B_{z,n}}) \end{aligned}$$

for sufficiently large n . On the other hand, we have

$$\lim_{T \rightarrow \infty} \frac{M(T)}{T} = \frac{1}{\phi(m_x, m_y, \dots, m_z)} \quad (\text{a. s.}).$$

(See Theorem 3 in [3].) So, using the similar method as in Theorem 4 of [2], we have (4.9) which was to be proved.

COROLLARY. *Under the conditions (4.1)–(4.6) and (2.5), we have*

$$\begin{aligned} & \lim_{T \rightarrow \infty} \frac{\left| \int_0^T dt \int_0^t \phi(t-u) dM(u) - \frac{T}{\phi(m_x, m_y, \dots, m_z)} \int_0^\infty \phi(u) du \right|}{\sqrt{2T \log \log T}} \\ (4.10) \quad & \leq \frac{3(A\sqrt{B_x} + B\sqrt{B_y} + \dots + C\sqrt{B_z})}{\sqrt{\phi(m_x, m_y, \dots, m_z)^3}} \cdot \|\psi\| \\ & \quad + \frac{1}{m} \left(\|\psi\| + \frac{1}{2} \overline{\lim}_{x \rightarrow \infty} \frac{x}{\log \log x} \int_x^\infty |\phi(u)| du \right) \quad (\text{a. s.}). \end{aligned}$$

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