

# INADMISSIBILITY OF THE BEST INVARIANT TEST WHEN THE MOMENT IS INFINITE UNDER ONE OF THE HYPOTHESES

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**1. Introduction.** Let  $(\mathcal{Y}, \mathcal{A}, \lambda_i) (i = 1, 2)$  be probability spaces. For each  $i = 1, 2$  and  $y \in \mathcal{Y}$  let  $F_i(\cdot, y)$  be a distribution function on the real line  $R$  such that  $F_i(\cdot, \cdot)$  is  $\mathcal{B} \times \mathcal{A}$  measurable where  $\mathcal{B}$  is the  $\sigma$ -field of all Borel subsets of the real line  $R$ . Assume the distribution of  $(X, Y) \in R \times \mathcal{Y}$  for  $\theta \in R$  and  $i = 1, 2$  is given by usual extension of

$$P_{i\theta}((X, Y) \in C \times D) = \int_D d\lambda_i(y) \int_C F_i(dx - \theta, y)$$

to measurable subsets of  $R \times \mathcal{Y}$ .

Consider the problem of testing  $H_1: i = 1$  versus  $H_2: i = 2$ . For any level of significance a best invariant test  $\varphi_0$  is of the form

$$(1.1) \quad \begin{aligned} \varphi_0(x, y) &= 1 && \text{if} && \frac{d\lambda_2}{d(\lambda_1 + \lambda_2)}(y) > c \\ &= 0 && \text{if} && \frac{d\lambda_2}{d(\lambda_1 + \lambda_2)}(y) < c. \end{aligned}$$

We restrict attention to the case that the  $F_i(\cdot, y)$  are absolutely continuous with respect to Lebesgue measure for each  $y \in \mathcal{Y}$  and  $i = 1, 2$ . Denote the density of  $F_i(\cdot, y)$  with respect to Lebesgue measure by  $f_i(\cdot, y)$ .

Lehmann and Stein [1] have shown that if  $E_{i0}|X| < \infty$  for  $i = 1, 2$  and if

$$(1.2) \quad \lambda_1\{y: \frac{d\lambda_2}{d(\lambda_1 + \lambda_2)}(y) = c\} = 0$$

then  $\varphi_0$  is admissible. Condition (1.2) guarantees that  $\varphi_0$  is the essentially unique best invariant test at some level. Perng [2; Sections 4 and 5] has given examples showing that, with either the moment condition or (1.2) violated,  $\varphi_0$  may not be admissible. The purpose of this note is to improve Perng's example concerning the moment condition.

Perng has shown that given any  $\delta > 0$  one can construct an example in which  $E_{i0}|X|^\alpha$  is, for  $i = 1, 2$  finite or infinite according as  $\alpha < 1 - \delta$  or  $\alpha \geq 1 - \delta$  and for which  $\varphi_0$  is inadmissible. His example satisfies (1.2). The present example, given in Section 2, also satisfies (1.2) but is such that  $E_{10}|X|^\alpha$  is as in Perng's example while  $E_{20}|X|^\alpha < \infty$  for all  $\alpha > 0$ . This suggests the intuitive idea that knowledge of  $X$  is useful when the distributions of  $X$  under  $H_1$  and  $H_2$  are very different.

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**2. The example.** Let  $\mathcal{Y} = R$  and let  $\lambda_i$  ( $i = 1, 2$ ) have density  $g_i$  with respect to Lebesgue measure where

$$\begin{aligned} g_1(y) = g_2(-y) &= c_1/y^2 && \text{if } y > 1; \\ &= c_2/y^2 && \text{if } y < -1; \\ &= 0 && \text{if } |y| \leq 1 \end{aligned}$$

with  $c_1 + c_2 = 1$ .

Let  $a > 2$  and  $\eta > 0$ . For  $\eta$  sufficiently small,

$$(2.1) \quad [(a - 1 - \eta)/(a - 1 + \eta)]a > 2$$

and

$$(2.2) \quad [(a - 1 - \eta)/(a - 1 + \eta)](a - 1) > 1.$$

Fix  $\epsilon > 0$ . For  $a$  sufficiently close to 2 we have

$$(a - 1)^{1/(1+\epsilon)}(a^{1/(1+\epsilon)} - 1) < 1$$

so that, for  $\eta$  sufficiently small,

$$(2.3) \quad [(a - 1)(a - 1 - \eta)/(a - 1 + \eta)]^{1/(1+\epsilon)} \cdot \{[a(a - 1 + \eta)/(a - 1 - \eta)]^{1/(1+\epsilon)} - 1\} < 1.$$

Fix  $\epsilon, \eta > 0$  and  $a > 2$  satisfying (2.1), (2.2) and (2.3). Let

$$\begin{aligned} f_1(x, y) &= \eta^{-1} && \text{if } y > 1, y^{1+\epsilon} < x < y^{1+\epsilon} + \eta \\ & && \text{or } y < -1, -\eta < x < 0; \\ &= 0 && \text{otherwise,} \end{aligned}$$

and

$$\begin{aligned} f_2(x, y) &= \eta^{-1} && \text{if } y > 1, -\eta < x < 0 \\ & && \text{or } y < -1, 0 < x < \eta; \\ &= 0 && \text{otherwise.} \end{aligned}$$

Elementary integrations yield  $E_{10}|X|^\alpha < \infty$  if, and only if,  $\alpha < 1 - \epsilon/(1 + \epsilon)$  while  $E_{20}|X|^\alpha < \infty$  for all  $\alpha > 0$ . Finally, by (2.2) we can take

$$(2.4) \quad c_1 < c_2 < c_1[(a - 1)(a - 1 - \eta)/(a - 1 + \eta)]^{1/(1+\epsilon)}.$$

Since  $c_2 > c_1$ , the test of the form (1.1) for level  $c_1$  is given by

$$\begin{aligned} \varphi_0(x, y) &= 1 && \text{if } y \geq 0 \\ &= 0 && \text{if } y < 0. \end{aligned}$$

Clearly (1.2) is satisfied so long as  $c$  is chosen so that  $c_1/(c_1 + c_2) < c < c_2/(c_1 + c_2)$ . Hence  $\varphi_0$  is the essentially unique best invariant test at level  $c_1$ .

We now define another test  $\varphi^*$  which will be shown to dominate  $\varphi_0$ . Let

$$\begin{aligned}\varphi^*(x, y) &= 1 && \text{if } y < -1, a - 1 \leq x \leq |y|^{1+\epsilon}; \\ &= 0 && \text{if } y > 1, \max(a - 1, y^{1+\epsilon}) \leq x \leq ay^{1+\epsilon}; \\ &= \varphi_0(x, y) && \text{otherwise.}\end{aligned}$$

Abbreviate  $\varphi_0(X, Y)$  by  $\varphi_0$  and  $\varphi^*(X, Y)$  by  $\varphi^*$  since in the sequel this will cause no confusion. We wish to show that  $E_{1\theta}(\varphi^* - \varphi_0) \leq 0$  and  $E_{2\theta}(\varphi^* - \varphi_0) \geq 0$  with strict inequality in at least one case for some  $\theta$ .

Clearly for  $\theta \leq a - 1$  we have  $E_{2\theta}(\varphi^* - \varphi_0) \geq 0$  and  $E_{1\theta}(\varphi^* - \varphi_0) \leq 0$  with strict inequality in the former for  $a - 1 - \eta < \theta \leq a - 1$  and in latter for  $-\eta < \theta \leq a - 1$ .

Let  $\theta > a - 1$ . In this case the curves  $x = y^{1+\epsilon} + \theta$  and  $x = ay^{1+\epsilon}$  intersect with  $y \geq 1$ . Thus,

$$\begin{aligned}E_{1\theta}(\varphi^* - \varphi_0) &\leq c_2 \int_{(\theta-\eta)^\omega}^\infty y^{-2} dy - c_1 \int_{[(\theta+\eta)/(a-1)]^\omega}^\infty y^{-2} dy \\ &= c_2(\theta - \eta)^{-\omega} - c_1[(a - 1)/(\theta + \eta)]^\omega \\ &\leq (\theta - \eta)^{-\omega} \{c_2 - c_1[(a - 1)(a - 1 - \eta)/(a - 1 + \eta)]^\omega\} \\ &< 0,\end{aligned}$$

where  $\omega = (1 + \epsilon)^{-1}$ . The last inequality results from (2.4). Now from (2.3) and (2.4) we obtain

$$c_2\{[a(a - 1 + \eta)/(a - 1 - \eta)]^\omega - 1\} < c_1.$$

Thus

$$\begin{aligned}E_{2\theta}(\varphi^* - \varphi_0) &\geq c_1 \int_{(\theta+\eta)^\omega}^\infty y^{-2} dy - c_2 \int_{[a/(\theta-\eta)]^\omega}^{\theta^\omega} y^{-2} dy \\ &= c_1[1/(\theta + \eta)]^\omega - c_2\{[a/(\theta - \eta)]^\omega - (1/\theta)^\omega\} \\ &> [1/(\theta + \eta)]^\omega \{c_1 - c_2[a(\theta + \eta)/(\theta - \eta)]^\omega - 1\} \\ &\geq [1/(\theta + \eta)]^\omega \{c_1 - c_2[a(a - 1 + \eta)/(a - 1 - \eta)]^\omega - 1\} \\ &> 0.\end{aligned}$$

This completes the proof.

#### REFERENCES

- [1] LEHMANN, E. L. and STEIN, C. M. (1953). The admissibility of certain invariant statistical tests involving a translation parameter. *Ann. Math. Statist.* **24** 473-479.
- [2] PERNG, S. K. (1967). Inadmissibility of various "good" statistical procedures which are translation invariant. Michigan State University RM-192. Unpublished Ph.D. thesis.