RESOLVABLE DESIGNS WITH LARGE BLOCKS

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Dedicated to the memory of I. M. Chakravarti

Resolvable designs with two blocks per replicate are studied from an optimality perspective. Because in practice the number of replicates is typically less than the number of treatments, arguments can be based on the dual of the information matrix and consequently given in terms of block concurrences. Equalizing block concurrences for given block sizes is often, but not always, the best strategy. Sufficient conditions are established for various strong optimalities and a detailed study of E-optimality is offered, including a characterization of the E-optimal class. Optimal designs are found to correspond to balanced arrays and an affine-like generalization.

1. Introduction. Block designs arise in comparative experimentation as fundamental devices for improving efficiency when working with heterogeneous experimental units. The blocks are simply a partition of the units into (say) b sets exhibiting homogeneity within sets. Given the blocks, of sizes k_1, k_2, \ldots, k_b , a *block design* is an assignment of v treatments to the $\sum_{j=1}^{b} k_j$ units. Optimality theory for block designs attempts to determine which of the many possible assignments is in some sense best.

In some applications there are restrictions on the collection of possible assignments. A block design is *resolvable* if the blocks can be partitioned into *replicates*, defined as sets of blocks with the property that each treatment is assigned to one unit in each set. The practical impact of, and motivation for, resolvability is to gain orthogonality between treatments and nuisance factors of concern. For instance, resolvability in sequential experimentation, with replicates corresponding to time periods, is used to mitigate time effects. Resolvability can likewise be useful in multi-site experiments and in experiments with multiple individuals handling experimental runs. Notably, the United Kingdom has for some time required the use of resolvable designs in agricultural field trials (see [20]).

The combinatorial study of resolvability in block designs goes back at least as far as the well-known Kirkman's [14] schoolgirl problem. The notion entered the statistical lexicon with Yates' work on square *lattice* designs [29, 30], although the

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term "resolvable design" was introduced by Bose [4]. Yates' lattice designs were extended to rectangular lattices by Harshbarger [11, 12]; see also [2]. Williams [27] and Patterson and Williams [21] introduced a large family of resolvable designs which they termed α -designs. Williams, Patterson and John [28] derived resolvable designs with two replicates from BIBDs (balanced incomplete block designs). Bailey, Monod and Morgan [1] proved strong optimality for the affine resolvable designs introduced by Bose [4]. Resolvable BIBDs have received significant attention in both the combinatorial and the statistical literature; for a summary, see Morgan [17], who surveys the major classes of resolvable designs with many references.

With α -designs, Patterson and Williams [21] provided a flexible method for obtaining reasonably efficient resolvable designs for a wide range of values v, constant block size k and replication r. They also adapted their method to obtain resolvable designs with two different block sizes, k and k-1, a first attempt at addressing the obvious restriction that resolvability with equal block sizes can be achieved only for v a multiple of k. John, Russell, Williams and Whitaker [13], in revisiting that idea, concluded that the α -technique for two block sizes could produce relatively inefficient designs for small v and recommended an interchange algorithm for construction of designs with better efficiency. John et al. [13] also discussed the practical need for resolvable designs with unequal block sizes; for example, about half of 245 experiments examined by Patterson and Hunter [19] had unequal block sizes. See also [20].

To the authors' knowledge, the literature contains no systematic work on determining optimal resolvable block designs when block sizes need not be equal. This paper will undertake such work, for the special case of two blocks per replicate. Only with two blocks in each replicate must a block be large in the sense of containing at least half of the treatments. And with two blocks per replicate, the block sizes must be unequal for any odd v (although this work is not restricted to odd v).

Let $\mathcal{D}(v,r;k_1,k_2)$ denote the class of all resolvable block designs with r replicates of v treatments, each replicate consisting of two blocks, of sizes k_1 and k_2 , where $k_1+k_2=v$. Of special interest is $k_1=k$, $k_2=k-1$ for odd v=2k-1, but here no restrictions are placed on the two block sizes for the general theory. Let k_1 denote the larger block size, $k_1 \geq k_2$, so that with no loss of generality $\frac{v}{2} \leq k_1 \leq v-2$. In the most common applications of resolvable designs, the number of treatments is large relative to the number of replicates; here, $r \leq v-1$ is required, allowing optimality problems to be more easily attacked using the dual of the information matrix. These considerations frame the goal of this paper: to determine the best design $d \in \mathcal{D}(v,r;k_1,k_2)$. An example of a resolvable design in $\mathcal{D}(9,5;5,4)$ is shown in Table 1, with the blocks written as columns. Later, this design will be proven optimal with respect to many useful criteria.

For any resolvable design $d \in \mathcal{D}$, ignoring the replicate grouping leaves an underlying *simple block design* for v treatments in 2r blocks. If the roles of blocks

TABLE 1 A resolvable design in $\mathcal{D}(9,5;5,4)$

1	6	1 2	1 4	1 3	1 2
2	7	4 3	2 5	2 4	3 5
3	8	5 6	3 6	5 6	4 6
4	9	7 8	7 8	8 7	7 9
5		9	9	9	8

and treatments are reversed in this underlying design, another simple block design with 2r treatments in v blocks of size r is produced, called the *dual design*.

Section 2 provides optimality background and establishes equivalences with the dual problem. Section 3 finds conditions for global optimality of designs having equality of block concurrences; for those who wish to jump ahead, the main results there are Theorems 4–7. Section 4 characterizes the E-optimal designs (Theorem 14) and finds the Schur-best of the E-optimal designs (Theorem 15). The main results are applied to special cases with $k_1 - k_2 \le 2$ in Section 5 and designs for these cases are constructed in Section 6. Concluding comments appear in Section 7.

2. Model, information and optimality criteria. Let y_{hjl} denote the yield from the lth experimental unit in block j of replicate h. Thus, the triples (h, j, l) identify the experimental units and the design d corresponds to a map d[h, j, l] from the units to the set of treatments. The standard linear model for the yields incorporates a mean effect μ , replicate effects ρ_h , block effects β_{hj} , treatment effects $\tau_{d[h,j,l]}$ and mean zero, uncorrelated, equivariable random error terms e_{hjl} :

$$y_{hjl} = \mu + \rho_h + \beta_{hj} + \tau_{d[h,j,l]} + e_{hjl},$$

 $h = 1, ..., r; j = 1, 2; l = 1, ..., k_j$. This model may be written in matrix terms as

$$y = \mu 1 + (I_r \otimes 1_v)\rho + L\beta + X_d\tau + e$$
,

where, with the y_{hjl} lexicographically ordered in the $vr \times 1$ yield vector y, the block incidence matrix is $L_{vr \times 2r} = I_r \otimes \binom{1_{k_1}}{0_{k_2}} \binom{1_{k_1}}{1_{k_2}}$ and the design matrix is the $vr \times v$ incidence matrix X_d , for which row (h,j,l) has a 1 in column i if and only if d[h,j,l]=i [i.e., unit (h,j,l) receives treatment i] and all other entries are zero. Replicate effects, block effects, treatment effects and error vectors are $\rho_{r\times 1}$, $\beta_{2r\times 1}$, $\tau_{v\times 1}$ and $e_{vr\times 1}$, respectively. Choice of design is equivalently choice of X_d .

Linear model theory says that the information matrix C_d for estimation of the treatment effects τ is

(1)
$$C_d = X_d'[I - L(L'L)^{-1}L']X_d = rI - N_d D_s^{-1} N_d',$$

where $D_s = L'L = \mathrm{Diag}(k_1, k_2, k_1, k_2, \ldots, k_1, k_2)$ is the diagonal matrix of block sizes. The $v \times 2r$ matrix N_d is the treament/block incidence matrix. The general off-diagonal element $(N_d N'_d)_{i,i'}$ of $N_d N'_d$ is the number of blocks to which both treatments i and i' are assigned, called a *treatment concurrence count*. Notice that the replicate incidence matrix $(I_r \otimes 1_v)$ plays no role in (1); the same form of information matrix is obtained for any simple block design. It follows immediately that the information matrix for the corresponding dual design is

$$(2) C_{\text{dual}} = D_s - \frac{1}{r} N_d' N_d.$$

The off-diagonal elements of N'_dN_d are block concurrence counts.

All treatment contrasts are estimable with design d if and only if C_d has rank v-1. Any such d is said to be *connected*; only connected designs are considered here. Most (but not all) commonly employed optimality criteria, including those to be used here, are functions of the v-1 nonzero eigenvalues of C_d . These will be ordered and labeled $z_{d1} \le z_{d2} \le \cdots \le z_{d,v-1}$.

Much of the optimality work below focuses on minimizing functions of the form

(3)
$$\psi_f(z_d) = \sum_{i=1}^{v-1} f(z_{di}),$$

where f is convex and z_d is the vector of nonzero eigenvalues. If f in (3) is Schur-convex (Schur-convex functions include the convex functions; see [3], Section II.3), then (3) is said to be a *Schur-criterion*. Design d_1 is *Schur-better* than design d_2 (d_2 is *Schur-inferior* to d_1) if $\psi_f(z_{d_1}) \leq \psi_f(z_{d_2})$ for all Schur-convex f with strict inequality for at least one such f. A design optimal with respect to all (i.e., minimizing all) Schur-convex criteria is said to be *Schur-optimal*. If f in (3) satisfies (i) f is continuously differentiable on $(0, \max_{d \in \mathcal{D}} \operatorname{tr}(C_d))$ with f' < 0, f'' > 0 and f''' < 0 and (ii) $\lim_{x \to 0} f(x) = \infty$, then (3) is said to be a *type-1 criterion* (see [7]). A design optimal with respect to all type-1 criteria is said to be *type-1-optimal*. One popular criterion belonging to both families just defined is the A-criterion specified by $f(x) = \frac{1}{x}$. A criterion not of the form (3) (though it can be written as a limit of such criteria) is

$$\psi_E(z_d) = \frac{1}{z_{d1}},$$

called the *E-criterion*. An E-optimal design minimizes (4) or, equivalently, maximizes z_{d1} . For a broader discussion of optimality criteria and their statistical meanings, see [24].

For the current endeavor, it is advantageous to approach the z_{di} , and consequently (3) and (4), through C_{dual} . How this is done is shown in Lemma 1 below. Let $\phi_{dhh'}$ be the block concurrence for (i.e., the number of treatments common to)

the blocks of size k_1 in replicates h and h'. For $h \neq h'$, let $\phi_{dhh'}^* = \phi_{dhh'} - \frac{k_1^2}{k_1 + k_2}$ and define the (symmetric) optimality matrix M_d by

(5)
$$M_{d} = \begin{pmatrix} \frac{k_{1}k_{2}}{v} & \phi_{d12}^{*} & \phi_{d13}^{*} & \cdots & \phi_{d1r}^{*} \\ & \frac{k_{1}k_{2}}{v} & \phi_{d23}^{*} & \cdots & \phi_{d2r}^{*} \\ & & \frac{k_{1}k_{2}}{v} & \cdots & \phi_{d3r}^{*} \\ & & & \ddots & \vdots \\ & & & \frac{k_{1}k_{2}}{v} \end{pmatrix},$$

with eigenvalues $e_{d1} \ge e_{d2} \ge \cdots \ge e_{dr}$.

LEMMA 1. For any $d \in \mathcal{D}(v,r;k_1,k_2)$, the eigenvalues of C_d are 0, $\max\{0,v-r-1\}$ copies of r and $(r-\frac{ve_{d1}}{k_1k_2},r-\frac{ve_{d2}}{k_1k_2},\ldots,r-\frac{ve_{dr}}{k_1k_2})$. Consequently, a ψ_f -optimal design minimizing (3) will equivalently minimize $\sum_{h=1}^r f(r-\frac{ve_{dh}}{k_1k_2})$ and an E-optimal design minimizing (4) will equivalently minimize e_{d1} .

Lemma 1 is proved in Appendix A.1. Provided $r \le v - 1$ (as earlier required in Section 1), C_d has v - r - 1 eigenvalues fixed at r. Working with the dual through M_d not only makes this evident, but allows these structurally fixed eigenvalues to be easily set aside.

With the optimality problem recast in terms of M_d and its eigenvalues, a crucial concept for proofs of Schur-optimality is now defined.

DEFINITION. Let $\{x_i\}_{i=1}^n$ and $\{y_i\}_{i=1}^n$ be nonincreasing sequences of real numbers such that $\sum_{i=1}^n x_i = \sum_{i=1}^n y_i$. If

$$\sum_{i=1}^{l} x_i \le \sum_{i=1}^{l} y_i \quad \text{for all } 1 \le l \le n$$

or, equivalently,

$$\sum_{i=n}^{n-l+1} x_i \ge \sum_{i=n}^{n-l+1} y_i \quad \text{for all } 1 \le l \le n,$$

then $\{y_i\}_{i=1}^n$ is said to majorize $\{x_i\}_{i=1}^n$.

The importance of majorization is evident in the following result (see, e.g., [3], page 40).

THEOREM 2. Let $\{x_i\}_{i=1}^n$ and $\{y_i\}_{i=1}^n$ be nonincreasing sequences of real numbers such that $\sum_{i=1}^n x_i = \sum_{i=1}^n y_i$. Then $\sum_{i=1}^n f(x_i) \leq \sum_{i=1}^n f(y_i)$ for all real-valued convex functions f if and only if $\{y_i\}_{i=1}^n$ majorizes $\{x_i\}_{i=1}^n$.

If the sequences $\{x_i\}_{i=1}^n$ and $\{y_i\}_{i=1}^n$ are written as the elements of vectors x and y, then the fact that y majorizes x is written as $y \succ x$ or $x \prec y$.

Let e_d and e_{d^*} be the vectors of eigenvalues for the optimality matrices for designs d and d^* , respectively. Lemma 3 lists simple majorization facts ([3], page 30) used in subsequent sections. For majorization comparisons, the third of these allows work directly with the e_{dh} rather than the $r - \frac{ve_{dh}}{k_1k_2}$ (see Lemma 1). Thus, if $e_d > e_{d^*}$ and the two vectors are not identical, then d^* is Schur-better than d. Design d^* is Schur-optimal if $e_d > e_{d^*}$ for every $d \in \mathcal{D}$.

LEMMA 3. For real numbers $\{x_i\}_{i=1}^n$ and $\{y_i\}_{i=1}^n$ with $\sum_{i=1}^n x_i = \sum_{i=1}^n y_i$,

- (i) if $x_1 \ge x_2 = x_3 = \dots = x_n$ and $y_1 \ge x_1$, then $\{y_i\}_{i=1}^n$ majorizes $\{x_i\}_{i=1}^n$;
- (ii) if $x_1 = x_2 = \cdots = x_{n-1} \ge x_n$ and $x_n \ge y_n$, then $\{y_i\}_{i=1}^n$ majorizes $\{x_i\}_{i=1}^n$;
- (iii) if $\{y_i\}_{i=1}^n$ majorizes $\{x_i\}_{i=1}^n$, then $\{a-\frac{y_i}{b}\}_{i=1}^n$ majorizes $\{a-\frac{x_i}{b}\}_{i=1}^n$ for any real a,b.
- **3. Equal concurrence designs and global optimality.** Among simple block designs, the BIBDs are Schur-optimal, a result which follows from equality of treatment concurrences inducing complete symmetry of the information matrix. The analogous notion for duals is equality of block concurrences, which this section explores for utility with resolvable designs. A resolvable design $d \in \mathcal{D}(v, r; k_1, k_2)$ having block concurrence counts $\phi_{d12} = \phi_{d13} = \phi_{d23} = \cdots = \phi_{d,r-1,r} = \theta$ for some $k_1 k_2 \le \theta \le k_1$ is called an *equal concurrence design with common concurrence* θ , or ECD(θ). The subsections below will explore, in turn, (i) ECDs with equality of eigenvalues in the optimality matrix, (ii) other ECDs which can be proven Schur-optimal and (iii) Schur-domination of designs with unequal block concurrences by one or more ECDs.
- 3.1. Schur-optimality via equality of eigenvalues. For an ECD(θ), the optimality matrix (5) is

(6)
$$M_d = \left[\frac{p}{v} - \left(\theta - \frac{k_1^2}{v}\right)\right]I + \left(\theta - \frac{k_1^2}{v}\right)J,$$

where I is the $r \times r$ identity matrix, J is the $r \times r$ matrix of ones and $p = k_1 k_2$ is the product of the block sizes. Like a BIBD information matrix, it is completely symmetric, but unlike that matrix, M_d for an ECD(θ) is nonsingular and thus can have two distinct, relevant eigenvalues, rather than just one.

THEOREM 4. Suppose that $\mathcal{D}(v, r; k_1, k_2)$ is a resolvable design setting for which $(k_1 + k_2)|k_1^2$ and define

(7)
$$\theta^* = \frac{k_1^2}{k_1 + k_2} = \frac{k_1^2}{v}.$$

Then $ECD(\theta^*)$'s in \mathcal{D} are Schur-optimal whenever they exist.

TABLE 2
A Schur-optimal ECD(4) in $\mathcal{D}(9, 4; 6, 3)$

1 7	1 5	1 3	3 1
2 8	2 6	2 4	4 2
3 9	3 9	5 9	5 9
4	4	6	6
5	7	7	7
6	8	8	8

PROOF. Given the conditions on \mathcal{D} , the inequalities $k_1 - k_2 \leq \frac{k_1^2}{k_1 + k_2} \leq k_1$ imply that $\theta = \theta^*$ is an admissible value for the common block concurrence of an ECD(θ). For $\theta = \theta^*$, M_d in (6) is $\frac{p}{v}I$. Since the eigenvalues of M_d are all identical, they are majorized by the eigenvalues of every competing design. \square

Corollary 3.4 of [1] established that affine-resolvable designs are Schur-optimal. Theorem 4 generalizes that result when there are two blocks per replicate. Here, the optimality condition is that all pairs of blocks of size k_1 have the same concurrence (7). When $k_1 = k_2$ and $2|k_1$, $ECD(\theta^*)$'s are affine-resolvable designs. It is obvious that (6) has two distinct eigenvalues for any $\theta \neq \theta^*$.

EXAMPLE. Consider the setting $\mathcal{D}(9, 4; 6, 3)$. Then $(k_1 + k_2)|k_1^2$ with $\theta^* = 4$ and if an ECD(4) exists, it is Schur-optimal. In fact, an ECD(4) does exist and is shown in Table 2.

The settings for which k_1^2 is a multiple of $k_1 + k_2$ are relatively sparse (a situation much like that of BIBDs relative to all simple block design settings). For the 1225 pairs $2 \le k_2 < k_1 \le 51$, only 23 meet the divisibility requirement implied by (7). Theorem 4 is thus only a start, albeit an important one.

3.2. Global optimality of other ECDs. Good designs are expected to have eigenvalue structures "close" to that of $ECD(\theta^*)$'s, suggesting this question: Is some equal concurrence design Schur-optimal when $(k_1 + k_2) / k_1^2$? To investigate this question, define the block concurrence parameter $\bar{\theta}$ by

(8)
$$\bar{\theta} = \operatorname{int}\left(\frac{k_1^2}{k_1 + k_2}\right)$$

and write

(9)
$$\gamma = \frac{k_1^2}{k_1 + k_2} - \bar{\theta}.$$

Then $0 \le \gamma < 1$ and a necessary condition for existence of $ECD(\theta^*)$ is $\gamma = 0$. Consequently, γ is called the *block discordancy coefficient*; it measures the departure

of the block sizes from that required for equality of all eigenvalues. The parameter γ will play a pivotal role in the remainder of this paper, as will the concurrence discrepancies defined next.

Define the block concurrence discrepancy matrix $\Delta_d = (\delta_{dhh'})$, where

$$\delta_{dhh'} = \begin{cases} \phi_{dhh'} - \bar{\theta}, & \text{if } h \neq h', \\ 0, & \text{if } h = h'. \end{cases}$$

For $h \neq h'$, the off-diagonal elements $\delta_{dhh'}$ will be referred to as block concurrence discrepancies. Rewritten in terms of block discrepancies and the discordancy coefficient, the general optimality matrix (5) becomes

(10)
$$M_d = \frac{p}{v}I - \gamma(J - I) + \Delta_d.$$

For ECD(θ^*), $\gamma = 0$, $\Delta_d = 0$ and $M_d = \frac{p}{v}I$. If $\gamma \neq 0$, (10) shows that the form of optimal design may depend both on the magnitude of γ and the values of the discrepancies. Matrix (10) is completely symmetric for, and only for, ECDs, in which case Δ_d is an integer multiple of J - I.

The ECDs which are combinatorially closest to $ECD(\theta^*)$ are those with either $\theta = \bar{\theta}$ or $\theta = \bar{\theta} + 1$. Not surprisingly, these are strong competitors in the optimality race. $ECD(\bar{\theta})$ s have $\phi_{dhh'} = \bar{\theta}$ for each $1 \le h \ne h' \le r$, so $\Delta_d = 0$ and the eigenvalues of M_d are

(11)
$$\xi_1(\gamma) = \frac{p}{\nu} + \gamma \quad \text{and} \quad \xi_2(\gamma) = \frac{p}{\nu} - (r-1)\gamma,$$

with frequencies r-1 and 1, respectively. The two eigenvalues satisfy $\xi_1(\gamma) \ge \xi_2(\gamma)$. If $\phi_{dhh'} = \bar{\theta} + 1$ for every $h \ne h'$, then $\text{ECD}(\bar{\theta} + 1)$'s have $\Delta_d = (J - I)$ and the eigenvalues of M_d are

(12)
$$\xi_1(\gamma - 1) = \frac{p}{v} - (1 - \gamma) \quad \text{and}$$
$$\xi_2(\gamma - 1) = \frac{p}{v} + (r - 1)(1 - \gamma),$$

with frequencies r-1 and 1, respectively, and with $\xi_2(\gamma-1) \ge \xi_1(\gamma-1)$.

THEOREM 5. For $0 \le \gamma \le \frac{1}{2}$, $ECD(\bar{\theta})$'s are type-1-optimal and E-optimal.

PROOF. The eigenvalues of the information matrix for any design $d \in \mathcal{D}(v,r;k_1,k_2)$ are $0 < z_{d1} \le z_{d2} \le \cdots \le z_{dr}$ and v-r-1 copies of r, and $\sum_{i=1}^r z_{di} = r(r-1)$ is constant for all designs in \mathcal{D} . For an $\mathrm{ECD}(\bar{\theta})$, call it \bar{d} , the $z_{\bar{d}i}$, following from (11), have the form $z_{\bar{d}1} = z_{\bar{d}2} = \cdots = z_{\bar{d},v-2} \le z_{\bar{d},v-1}$. Theorem 2.3 of [7] thus gives the result if it can also be shown that \bar{d} minimizes $\sum_{i=1}^{v-1} z_{di}^2$ over \mathcal{D} .

For $d \in \mathcal{D}(v, r; k_1, k_2)$ with optimality matrix $(M_d)_{hh'} = (\delta_{dhh'} - \gamma)$ having trace tr $M_d = \frac{pr}{v}$,

$$\operatorname{tr} C_d^2 = \sum_{i=1}^{v-1} z_{di}^2 = (v - r - 1)r^2 + \sum_{h=1}^r \left(r - \frac{ve_{dh}}{p}\right)^2$$

$$= (v - 1)r^2 - \frac{2vr}{p} \operatorname{tr} M_d + \frac{v^2}{p^2} \operatorname{tr} M_d^2$$

$$= (v - 3)r^2 + r + \frac{2v^2}{p^2} \sum_{h \le h'} (\delta_{dhh'} - \gamma)^2,$$

so that $\operatorname{tr} C_d^2$ is minimized by designs that minimize $\sum \sum_{h < h'} (\delta_{dhh'} - \gamma)^2$. Since $\delta_{dhh'}$ is integral, the unique minimum of $\operatorname{tr} C_d^2$ on $0 \le \gamma < \frac{1}{2}$ is at $\delta_{dhh'} \equiv 0$, achieved only by $\operatorname{ECD}(\bar{\theta})$. For $\gamma = \frac{1}{2}$, any values $\delta_{dhh'} \in \{0, 1\}$ minimize $\operatorname{tr} C_d^2$.

Now define the F-criterion as the value of the largest eigenvalue of C_d that is not constrained by the setting to equal r, that is,

$$\psi_F(C_d) = z_{dr} = r - \frac{ve_{dr}}{n}.$$

Minimizing $\psi_F(C_d)$ over \mathcal{D} is equivalent to maximizing e_{dr} . This criterion can be important in establishing Schur-optimality, as shown next.

THEOREM 6. An $ECD(\bar{\theta})$ is Schur-better than a competitor with a different set of eigenvalues if and only if it is F-equivalent or better than that competitor. Consequently, $ECD(\bar{\theta})$'s are Schur-optimal if and only if they are F-optimal.

PROOF. Follows from (11) and Lemma 3(i). \Box

A result of a similar flavor holds for $ECD(\bar{\theta} + 1)$'s using the E-criterion. As pointed out by Kunert [15], page 385, designs with eigenvalues z_{di} in the form of Lemma 3(ii) are Schur-best whenever they are E-optimal. For the current problem, this is stated as follows.

THEOREM 7. An $ECD(\bar{\theta}+1)$ is Schur-better than a competitor with a different set of eigenvalues if and only if it is E-equivalent or better than that competitor. Consequently, $ECD(\bar{\theta}+1)$'s are Schur-optimal if and only if they are E-optimal.

In establishing necessary and sufficient conditions for Schur-optimality for their respective ECDs in terms of a single eigenvalue, Theorems 6 and 7 provide simple tests for comparing these designs to any other design. An immediate consequence is that an ECD(θ) with $\theta \notin \{\bar{\theta}, \bar{\theta}+1\}$ is Schur-inferior to at least one of these two competitors. Thus, among ECDs, only these two competitors remain. They are compared to one another in the following corollary.

COROLLARY 8. $ECD(\bar{\theta})$'s are Schur-better than $ECD(\bar{\theta}+1)$'s if and only if $\gamma \leq \frac{1}{r}$, and $ECD(\bar{\theta}+1)$'s are Schur-better than $ECD(\bar{\theta})$'s if and only if $\gamma \geq \frac{r-1}{r}$.

Simply use (11) and (12) in applying Theorems 6 and 7. \Box Proof.

Though design nonexistence can play a role, Corollary 8 says that one does not expect to find a Schur-optimal design for $\frac{1}{r} \le \gamma \le \frac{r-1}{r}$ (Theorem 5 makes for an interesting juxtaposition). Schur-domination can nonetheless be used to eliminate many competitors, as shown next.

3.3. Schur-inferiority of designs lacking equality of block concurrences. ven the results of the preceding subsection, the remaining question from a global optimality perspective is if (and when) designs outside the ECD class can be preferable. This question can be effectively pursued by application of Theorems 6 and 7, once workable bounds for the largest and smallest eigenvalues e_{d1} and e_{dr} of the optimality matrix are in place.

LEMMA 9. Let $d \in \mathcal{D}(v, r; k_1, k_2)$ have concurrence discrepancy matrix $\Delta_d = (\delta_{dhh'})$ and optimality matrix M_d . Then:

(i)
$$\delta_{d12} \leq 0 \Rightarrow e_{d1} \geq \frac{p}{v} + \gamma - \delta_{d12}$$
 and $e_{dr} \leq \frac{p}{v} - \gamma + \delta_{d12}$

(i)
$$\delta_{d12} \leq 0 \Rightarrow e_{d1} \geq \frac{p}{v} + \gamma - \delta_{d12}$$
 and $e_{dr} \leq \frac{p}{v} - \gamma + \delta_{d12}$;
(ii) $\delta_{d12} > 0 \Rightarrow e_{d1} \geq \frac{p}{v} - \gamma + \delta_{d12}$ and $e_{dr} \leq \frac{p}{v} + \gamma - \delta_{d12}$.

PROOF. The leading 2×2 minor of M_d , which is $M_{d11} = (\frac{p}{v} + \gamma - \delta_{d12})I$ $(\gamma - \delta_{d12})J$, has eigenvalues $\frac{p}{v} + \gamma - \delta_{d12}$ and $\frac{p}{v} - \gamma + \delta_{d12}$. A Sturmian separation theorem ([22], page 64) provides the bounds. \square

Define design d to be an almost equal concurrence design, or AECD, if $\phi_{dhh'} \in$ $\{\bar{\theta}, \bar{\theta}+1\}$ for all $h \neq h'$ and if each value is attained for some h, h'. If any $\phi_{dhh'}$ is not in $\{\bar{\theta}, \bar{\theta}+1\}$, then d is an unequal concurrence design, or UCD. In terms of discrepancies, AECDs have all $\delta_{dhh'} \in \{0, 1\}$, while UCDs have some $\delta_{dhh'} \leq$ -1 or ≥ 2 . Depending on γ , AECDs can be optimal in at least some senses (see Section 4), necessarily ruling out global optimality of ECDs for some γ . The next few results will show that ECDs often dominate UCDs.

COROLLARY 10. Suppose $d \in \mathcal{D}(v, b; k_1, k_2)$ is a UCD with $\delta_{dhh'} \leq -\alpha$ for some $1 \le h \ne h' \le r$ and some integer $\alpha \ge 1$. Then:

- (i) $ECD(\bar{\theta})$'s are Schur-better than d if $\gamma \leq \frac{\alpha}{r-2}$;
- (ii) $ECD(\bar{\theta} + 1)$'s are Schur better than d if $\gamma \ge \frac{r \alpha 1}{r}$.

PROOF. For the UCD d as described in the corollary, take $\delta_{d12} \leq -\alpha$. Then from Lemma 9, $e_{d1} \ge \frac{p}{v} + \gamma - \alpha$ and $\frac{p}{v} - \gamma + \alpha \ge e_{dr}$. By Theorem 6, an ECD($\bar{\theta}$) is Schur-better than d if $\xi_2(\gamma) \ge \frac{p}{v} - \gamma + \alpha \ge e_{dr} \iff \gamma \le \frac{\alpha}{r-2}$. By Theorem 7, an $ECD(\bar{\theta}+1)$ is Schur-better than d if $e_{d1} \ge \frac{p}{v} + \gamma - \alpha \ge \xi_2(\gamma - 1) \iff \gamma \ge 1$ $\frac{r-\alpha-1}{r}$.

COROLLARY 11. When $r \le 4$, all UCDs with some $\delta_{dhh'} \le -1$ are Schurinferior to an ECD, and when r = 5 or 6, UCDs with some $\delta_{dhh'} \leq -2$ are Schurinferior to an ECD.

The next two corollaries are similarly shown.

COROLLARY 12. Suppose $d \in \mathcal{D}(v, b; k_1, k_2)$ is a UCD with $\delta_{dhh'} \geq \alpha$ for some $1 \le h \ne h' \le r$ and some integer $\alpha \ge 2$. Then:

- (i) $ECD(\bar{\theta})$'s are Schur-better than d if $\gamma \leq \frac{\alpha}{r}$; (ii) $ECD(\bar{\theta}+1)$'s are Schur better than d if $\gamma \geq \frac{r-\alpha-1}{r-2}$.

COROLLARY 13. When $r \leq 4$, all UCDs with some $\delta_{dhh'} \geq 2$ are Schurinferior to an ECD, and when r = 5 or 6, UCDs with some $\delta_{dhh'} \geq 3$ are Schurinferior to an ECD.

Corollaries 11 and 13 say that optimal (with respect to any convex criterion) designs in settings $\mathcal{D}(v, r; k_1, k_2)$ with $r \le 4$ must be an $ECD(\bar{\theta})$, an $ECD(\bar{\theta} + 1)$ or an AECD. Optimal designs in settings with r = 5 or 6 must have block concurrence discrepancies $\delta_{dhh'} \in \{-1, 0, 1, 2\}$ for all $1 \le h \ne h' \le r$. It is unlikely that such global statements can be much improved. The importance of these results lies in the prevalence of small r in the application of resolvable designs.

4. E-optimality. Sufficient conditions, which are necessary given existence, will be developed for E-optimality of designs in $\mathcal{D}(v, r; k_1, k_2)$. The main results, Theorems 14 and 15, will be stated after first defining a subclass where E-optimal designs will be found.

Corollary 16 below will remove the UCDs from E-contention, so that only $ECD(\bar{\theta})$'s, $ECD(\bar{\theta}+1)$'s and AECDs need be considered. These designs have all $\delta_{dhh'} \in \{0, 1\}$ and so have Δ_d which is the adjacency matrix of a simple, undirected graph on r vertices. Any of these designs for which $\Delta_d - J$ is (with suitable ordering of replicates) of the form

(13)
$$\begin{pmatrix} -J_{t_1} & & & & \\ & -J_{t_2} & & & \\ & & \ddots & & \\ & & & -J_{t_n} \end{pmatrix}$$

for positive integers n and $t_1 \le t_2 \le \cdots \le t_n$ $(\sum_{i=1}^n t_i = r)$ is said to be groupaffine. For group-affine designs, concurrences $\phi_{dhh'}$ are constant $(=\theta)$ within groups of sizes t_1, \ldots, t_n and are constant $(=\bar{\theta}+1)$ between groups. A groupaffine design is said to be *uniform* if, for given n, the range of group sizes t_i is at most one. For any group-affine design with the number of groups $n \le n_{\gamma} = \inf(\frac{1}{1-\gamma})$, let $t_{(d)}$ denote the vector of its group sizes t_i arranged in increasing order and with n_{γ} elements, padding with zeros as necessary. For example, if r=7, $\gamma=\frac{7}{9}$ and d has t_i values 1, 3 and 3, then $n_{\gamma}=4$ and $t_{(d)}=(0,1,3,3)$. Now the main results can be stated. The phrase "E-Schur-optimal" means "Schur-optimal within the class of all E-optimal designs."

THEOREM 14. Any group-affine design is E-optimal if $n \le n_{\gamma}$ and $\gamma \le \frac{r-1}{r}$. If any such design exists, then all such designs form the class \mathcal{D}_E of all E-optimal designs. $ECD(\bar{\theta}+1)$'s are Schur-optimal for $\gamma \ge \frac{r-1}{r}$.

THEOREM 15. For $d_1, d_2 \in \mathcal{D}_E$, if $t_{(d_1)} \prec t_{(d_2)}$, then d_1 is Schur-better than d_2 . Design $d^* \in \mathcal{D}_E$ is E-Schur-optimal if it is uniform with $n = n_{\gamma}$.

Theorem 14 characterizes the class of E-optimal block concurrence structures for $\mathcal{D}(v,r;k_1,k_2)$. Although not much discussed in the literature, the class of E-optimal designs in a given simple block design setting often contains a variety of designs with different information matrices (see, e.g., [18] and [26]). Theorem 14 reveals the same situation for resolvable design settings. This allows other design criteria to be brought to bear: one should choose the best of the E-optimal designs. Theorem 15 does this. In a very strong sense, E-Schur-optimal designs are the best of the E-optimal designs.

The proofs of Theorems 14 and 15 depend on a series of technical results that are developed in the remainder of this section. So as not to overly disrupt the flow of the main line of proof, the longer of the "sub-proofs" are delayed until Appendix A. The first task is to rule out UCDs.

COROLLARY 16. For all $r \ge 2$ and $0 \le \gamma < 1$, $ECD(\bar{\theta})$'s are E-better than UCDs.

PROOF. The maximum eigenvalue of the optimality matrix for an $ECD(\bar{\theta})$ is $\xi_1(\gamma) = \frac{p}{v} + \gamma$. For the UCD d, suppose that $\delta_{d12} \leq -\alpha$ for some integer $\alpha \geq 1$. Then by Lemma 9, $e_{d1} \geq \frac{p}{v} + \gamma - \delta_{d12} > \xi_1(\gamma)$ and $ECD(\bar{\theta})$'s are E-better than d. If $\delta_{d12} \geq \alpha$ for some integer $\alpha \geq 2$, then another application of the lemma gives $e_{d1} \geq \frac{p}{v} - \gamma + \delta_{d12} > \xi_1(\gamma)$ and, again, $ECD(\bar{\theta})$'s are E-better than d. \square

Completing the proof of Theorem 14 is a matter of minimizing e_{d1} over all d for which every $\delta_{dhh'} \in \{0, 1\}$. Before providing the details, here is a sketch of what will be done. First, a lower bound for e_{d1} is found in terms of the largest eigenvalue of a specific projection $P\Delta_d P$ of the discrepancy matrix. This bound

implies that $ECD(\bar{\theta})$'s are E-superior to all designs for which $P\Delta_d P$ has a positive eigenvalue. The next step is thus to find a necessary and sufficient condition on Δ_d so that $P\Delta_d P$ is nonpositive definite; this condition turns out to be exactly the definition of group-affine design. Finally, the general form of M_d for groupaffine designs is examined in detail to determine which of these competitors are E-optimal, producing the conditions of Theorem 14.

LEMMA 17. For $d \in \mathcal{D}(v, r; k_1, k_2)$ and with $P = (I - \frac{1}{r}J)$, let u_{d1} and u_{dr} be the maximum and minimum eigenvalues of $P \Delta_d P$, respectively. Then:

- (i) if $u_{d1} > 0$ then $e_{d1} \ge \frac{p}{v} + \gamma + u_{d1}$;
- (ii) $e_{dr} \leq \frac{p}{v} + \gamma + u_{dr}$.

The proof of Lemma 17 appears in Appendix A.2. Lemma 17 combined with Theorems 6 and 7 immediately gives the following corollary:

COROLLARY 18. For $d \in \mathcal{D}(v, r; k_1, k_2)$ and with $P = (I - \frac{1}{r}J)$, let u_{d1} and u_{dr} be the maximum and minimum eigenvalues of $P\Delta_d P$.

- (i) If $\gamma < -\frac{u_{dr}}{r}$, then $ECD(\bar{\theta})$'s are Schur-better than d.
- (ii) If $u_{d1} > 0$ and $\gamma > (\frac{r u_{d1} 1}{r})$, then $ECD(\bar{\theta} + 1)$'s are Schur-better than d.
- (iii) If $u_{d1} > 0$, then $ECD(\bar{\theta})$'s are E-better, but not necessarily Schur-better, than d.

Parts (i) and (ii) of Corollary 18 provide alternative routes (cf. Corollaries 10 and 12) for establishing Schur-domination of ECDs. Part (iii) is the key part of the proof of Theorem 14. If d is not to be eliminated by $ECD(\bar{\theta})$, then $P\Delta_d P$ cannot have a positive eigenvalue, that is, must be nonpositive definite. Recall that the current E-competitors are all d for which Δ_d is the adjacency matrix of a simple, undirected graph.

LEMMA 19. Let A be the adjacency matrix for a simple undirected graph of r vertices. PAP is nonpositive definite if and only if A-J may be written (possibly after vertex permutation) in the form (13) for some positive integers n and $t_1 \le t_2 \le \cdots \le t_n$ with $\sum_{i=1}^n t_i = r$.

Combining Lemma 19 with Corollaries 16 and 18(iii), it has now been shown that existence of an ECD($\bar{\theta}$) implies that an E-optimal design is a group-affine design. Group-affine designs include the ECD($\bar{\theta}$)'s (put n=1, $\Delta_d=0$) and the ECD($\bar{\theta}+1$) (put n=r, $\Delta_d=J-I$) at the extremes for n. The proof of Lemma 19 appears in Appendix A.3.

It remains to determine which of the group-affine designs are actually E-best, requiring knowledge of e_{d1} for this class. For any group-affine design, write M_d =

 $(\frac{p}{n} + \gamma)I + \Delta_d - \gamma J$, where, for given $t_1 \leq \cdots \leq t_n$,

$$\Delta_{d} - \gamma J = (-1) \begin{pmatrix} \gamma J_{t_{1},t_{1}} & (\gamma - 1)J_{t_{1},t_{2}} & \cdots & (\gamma - 1)J_{t_{1},t_{n}} \\ (\gamma - 1)J_{t_{2},t_{1}} & \gamma J_{t_{2},t_{2}} & \cdots & (\gamma - 1)J_{t_{2},t_{n}} \\ \vdots & \vdots & \ddots & \vdots \\ (\gamma - 1)J_{t_{n},t_{1}} & (\gamma - 1)J_{t_{n},t_{2}} & \cdots & \gamma J_{t_{n},t_{n}} \end{pmatrix}$$

$$\equiv (-1)H_{d}.$$

An E-optimal design will maximize the minimum eigenvalue of H_d . Clearly, H_d has $\sum_{i=1}^n (t_i-1) = r-n$ eigenvalues of zero (corresponding to eigenvectors which are orthogonal contrasts within groups of sizes t_1, \ldots, t_n). So all of these designs have H_d with at least one eigenvalue of zero, except $\text{ECD}(\bar{\theta}+1)$, for which the eigenvalues of H_d are 1 (frequency r-1) and $1+r(\gamma-1)$. $\text{ECD}(\bar{\theta}+1)$ is therefore E-optimal if $1+r(\gamma-1)\geq 0$, that is, if $\gamma\geq \frac{r-1}{r}$ (in which case it is Schur-optimal; see Theorem 7). Otherwise, all designs for which H_d is nonnegative definite are E-optimal. Needed now are the eigenvalues of H_d other than zero. Let D_t be the diagonal matrix with diagonal elements (t_1,\ldots,t_n) . Lemma 20 is proved in Appendix A.4.

LEMMA 20. The eigenvalues of H_d specified by (14) are 0 (with frequency r-n) and the eigenvalues of $D_t^{1/2}ED_t^{1/2}$, where $E=I_n-(1-\gamma)J_n$.

Consequently, H_d is nonnegative definite if and only if $D_t^{1/2}ED_t^{1/2}$ is nonnegative definite. This is so if and only if E is nonnegative definite, that is, if and only if $1-\gamma \leq \frac{1}{n}$. Thus, for any $\gamma \leq \frac{r-1}{r}$, the E-optimal designs are all group-affine designs for which the number of groups n is no larger than $\frac{1}{1-\gamma}$, that is, for which $n \leq n_\gamma$. This completes the proof of Theorem 14.

To prove Theorem 15, recall the definition of $t_{(d)}$ given just prior to Theorem 14 and note that now only $d \in \mathcal{D}_E$ is being considered. It is shown in Appendix A.4 that the eigenvalues of H_d , aside from $r - n_{\gamma}$ zeros, are the eigenvalues of $E^{1/2}D_tE^{1/2}$. Thus, the problem is to show that $t_{(d_1)} \prec t_{(d_2)}$ implies the eigenvalues of $E^{1/2}D_{t_{(d_1)}}E^{1/2}$ are majorized by those of $E^{1/2}D_{t_{(d_2)}}E^{1/2}$.

Now, $t_{(d_1)} \prec t_{(d_2)}$ is equivalent to $t_{(d_1)} = St_{(d_2)}$, where S is a doubly stochastic matrix; $S = \sum_{i=1}^m a_i Q_i$ for permutation matrices Q_i and positive numbers a_i summing to 1 ([3], pages 33 and 37). Thus,

$$\begin{split} E^{1/2}D_{t_{(d_1)}}E^{1/2} &= E^{1/2}D_{St_{(d_2)}}E^{1/2} \\ &= E^{1/2} \Bigg[\sum_i a_i \, Q_i \, D_{t_{(d_2)}} \, Q_i \, \Bigg] E^{1/2} \\ &= \sum_i a_i \, Q_i \big[E^{1/2} D_{t_{(d_2)}} E^{1/2} \big] Q_i \, , \end{split}$$

the last equality following because $E^{1/2} = (I - \frac{1}{n}J) + \sqrt{\frac{1}{n} - (1 - \gamma)}J$ commutes with any permutation matrix. This shows that $E^{1/2}D_{t_{(d_1)}}E^{1/2}$ is a symmetrization of $E^{1/2}D_{t_{(d_2)}}E^{1/2}$ and thus the eigenvalues of the first matrix are majorized by those of the second (this follows from [3], page 69). Since $t_{(d^*)} \prec t_{(d)}$ for every $d \in \mathcal{D}_E$, d^* is E-Schur-optimal and Theorem 15 is proved.

5. Special cases: $(k_1 - k_2) \le 2$. In this section, the three important special cases of k_1 and k_2 being equal or nearly so, $k_2 \in \{k_1, k_1 - 1, k_1 - 2\}$, are investigated in light of the results in Sections 3 and 4. Put $k_2 = k_1 - m$ so that $(k_1 - k_2) \le 2$ says m = 0, 1, or 2, and for any m,

(15)
$$\frac{k_1^2}{k_1 + k_2} = \frac{k_1^2}{2k_1 - m} = \frac{k_1}{2} + \frac{m}{4} + \frac{m^2}{4(2k_1 - m)}.$$

Recall that $\bar{\theta}$ is the integer part of (15). The values for $\gamma = \frac{k_1^2}{k_1 + k_2} - \bar{\theta}$ in the corollaries below are easily found using (15).

COROLLARY 21. For $k_1 = k_2$,

- (i) if $2|k_1$, then $\gamma = 0$, $(k_1 + k_2)|k_1^2$ and $ECD(\theta^*)$'s are Schur-optimal;
- (ii) if $2 / k_1$, then $\gamma = \frac{1}{2}$ and $ECD(\bar{\theta})$'s are E-Schur and type-1-optimal.

When $k_1 = k_2$ and $2|k_1$, Schur-optimality also follows from [1], Corollary 3.4. For $2/k_1$, the result follows from Theorems 5 and 15.

COROLLARY 22. For $k_1 - k_2 = 1$,

- (i) if 2|k₁, then γ = v+1/4v and ECD(θ)'s are E-Schur and type-1-optimal.
 (ii) if 2|k₁, then γ = 3v+1/4v and uniform group-affine designs with four groups are E-Schur-optimal if r > 5. For r < 4, $ECD(\bar{\theta} + 1)$'s are Schur-optimal.

As an example, the design in Table 1 is E-Schur-optimal for 9 treatments in 5 replicates with block sizes 4 and 5. The design consisting of the first four replicates is Schur-optimal.

COROLLARY 23. For $k_1 - k_2 = 2$,

- (i) if $2|k_1$, then $\gamma = \frac{v+2}{2v}$, and uniform group-affine designs are E-Schuroptimal if the number of groups is 3 for v = 6 and 2 for $v \ge 10$. For r = 2, $ECD(\theta + 1)$'s are Schur-optimal;
 - (ii) if $2 / k_1$, then $\gamma = \frac{1}{v}$, and $ECD(\bar{\theta})$'s are Schur-optimal.

The Schur-optimality in part (ii) of Corollary 23 follows from Theorem 6 and Lemma 9.

6. ECDs, balanced arrays and design construction. Balanced arrays were introduced by Chakravarti [5, 6] as a useful device for fractional factorial designs and they have since been investigated by a plethora of authors, including, of late, Kuriki [16], Fuji-Hara and Miyamoto [9], Sinha, Dhar and Kageyama [25] and Ghosh and Teschmacher [10]. Here, only strength-two arrays on two symbols will be needed. A balanced array of strength 2, BA(N, m, 2), on the symbols 0 and 1, is an $N \times m$ array with the property that for any selection of two columns, the N pairs formed by the rows are (0,0), (0,1), (1,0) and (1,1), with frequencies μ_0 , μ_1 , μ_1 and μ_2 , respectively. Two-symbol orthogonal arrays of strength two are the special case $\mu_0 = \mu_1 = \mu_2$. When the number of 0's in each column is specified (as is the case below), the μ_i 's are all determined by $\theta = \mu_0$ and the array will be denoted BA(N, m, 2; θ).

Bailey, Monod and Morgan [1] demonstrate the combinatorial equivalence between affine resolvable designs and orthogonal arrays. Their method can be used to express any resolvable design as a combinatorial array, as follows. Given a resolvable design for v treatments in r replicates, each consisting of s blocks, label the blocks within a replicate $0, \ldots, s-1$. Construct a $v \times r$ array by identifying rows of the array with treatments of the resolvable design, and columns with replicates: symbol $j \in \{0, 1, \ldots, s-1\}$ is placed in row i, column k if and only if treatment k is in block k of replicate k. Theorem 24 is evident.

THEOREM 24. Each $ECD(\theta)$ in $\mathcal{D}(v, r; k_1, k_2)$ is equivalent to a $BA(v, r, 2; \theta)$.

A group-affine design with n groups of sizes t_1, \ldots, t_n is a juxtaposition (BA₁, BA₂, ..., BA_n) of balanced arrays BA_i = BA($v, t_i, 2; \theta$) so that any two columns from different BA_i's form a BA($v, 2, 2; \theta + 1$). Call such a juxtaposition a *grouped balanced array*, denoted by GBA($v, (t_1, \ldots, t_n), 2; \theta$). Constructions for the designs in Corollaries 21–23 of Section 5 are now listed. These are stated in terms of *Hadamard matrices*, that is, orthogonal matrices for which every element is 1 or -1. The *Hadamard conjecture* says that a Hadamard matrix exists for every order a multiple of four. Existence is known for all such orders up to 664 and for infinitely many other values (see [8]). A Hadamard matrix is said to be *standardized* if the first row and column are all ones; this can always be achieved.

THEOREM 25. If $k_1 = k_2$ is even, then a Schur-optimal $ECD(\theta^*)$ corresponds to an $OA(v, r, 2; \frac{v}{2})$. Existence of a Hadamard matrix of order v implies existence of the OA for every $r \le v - 1$.

THEOREM 26. If $k_1 = k_2$ is odd, then a type-1-optimal $ECD(\bar{\theta})$ corresponds to a $BA(v, r, 2; \frac{v-2}{4})$. Existence of a Hadamard matrix of order v+2 implies existence of the BA for every $r \leq \frac{v}{2}$.

PROOF. The value of $\bar{\theta}$ is $\operatorname{int}(\frac{(v/2)^2}{v}) = \frac{v-2}{4}$. Given the standardized Hadamard matrix, permute columns (except the first) so that the second row has 1 in the first $\frac{v+2}{2}$ columns. Delete the first two rows and the first $\frac{v+2}{2}$ columns, then replace -1 by 0 throughout. Clearly, the result is $\operatorname{BA}(v,\frac{v}{2},2;\frac{v-2}{4})$. \square

THEOREM 27. If $k_1 = k_2 + 1$ is even, then a type-1-optimal $ECD(\bar{\theta})$ corresponds to a $BA(v, r, 2; \frac{v+1}{4})$. Existence of a Hadamard matrix of order v+1 implies existence of the BA for every $r \leq v$.

PROOF. The value of $\bar{\theta}$ is $\operatorname{int}(\frac{((v+1)/2)^2}{v}) = \frac{v+1}{4}$. Given the standardized Hadamard matrix, delete the first row and column, then replace -1 by 0 throughout. The result is $\operatorname{BA}(v,v,2;\frac{v+1}{4})$. \square

THEOREM 28. If $k_1=k_2+1$ is odd, then a type-1-optimal $ECD(\bar{\theta}+1)$ with $r\leq 4$ corresponds to a $BA(v,r,2;\frac{v+3}{4})$, which always exists. For $r\geq 5$ and $v\geq 9$, an E-Schur-optimal group-affine design corresponds to a $GBA(v,(t_1,t_2,t_3,t_4),2;\frac{v-1}{4})$ with $t_4-t_1\leq 1$. Existence of a Hadamard matrix of order v+3 with a $4\times r$ submatrix of the form

$$\begin{pmatrix} 1_{t_1} & 1_{t_2} & 1_{t_3} & -1_{t_4} \\ 1_{t_1} & 1_{t_2} & -1_{t_3} & 1_{t_4} \\ 1_{t_1} & -1_{t_2} & 1_{t_3} & 1_{t_4} \\ -1_{t_1} & 1_{t_2} & 1_{t_3} & 1_{t_4} \end{pmatrix}$$

implies existence of the GBA.

PROOF. The value of $\bar{\theta}$ is $\operatorname{int}(\frac{((v+1)/2)^2}{v}) = \frac{v-1}{4}$. For $r \leq 4$, take $\operatorname{OA}(v-1,r,2)$ on $\{0,1\}$ and add one row of 0's to get $\operatorname{BA}(v,r,2;\frac{v+3}{4})$. Given the assumed Hadamard matrix, delete v+3-r columns not containing the submatrix, then delete the submatrix and add a row of 1's. Replacing -1 by 0 throughout gives the GBA. \Box

The maximum r admitted for a given v by the Hadamard construction in Theorem 28 depends on the particular Hadamard matrix chosen. All nonisomorphic Hadamard matrices are known up through order 28 and, for these, a complete search has produced (v, r) = (9, 5), (13, 5), (17, 9), (21, 9), (25, 13); the first four of these appear in Appendix B. A search of a few known Hadamard matrices (compiled by N. J. A. Sloane at www.research.att.com/~njas/hadamard/) of orders up to 48 has further produced (v, r) = (29, 13), (33, 16), (37, 18), (41, 24), (45, 19). Clearly there is room for more work to be done here.

THEOREM 29. If $k_1 = k_2 + 2$ is even, for $v \ge 10$ an E-Schur-optimal group-affine design corresponds to a GBA(v, (t_1, t_2) , 2; $\frac{v+2}{4}$) with $t_2 - t_1 \le 1$. Existence

of a Hadamard matrix of order v + 2 implies existence of the GBA for every $t_1 + t_2 = r \le \frac{v}{2} + 1$.

PROOF. The value of $\bar{\theta}$ is $\operatorname{int}(\frac{((v+2)/2)^2}{v}) = \frac{v+2}{4}$. Given the standardized Hadamard matrix of order v+2=4h, permute columns (except the first) so that the first three rows are

$$\begin{pmatrix} 1_h & 1_h & 1_h & 1_h \\ 1_h & -1_h & 1_h & -1_h \\ 1_h & -1_h & -1_h & 1_h \end{pmatrix}.$$

Now delete the first 2h columns, delete the first three rows, then add one row of -1's and finally replace -1 by 0 throughout. The result is $GBA(v, (\frac{v+2}{4}, \frac{v+2}{4}), 2; \frac{v+2}{4})$ with two juxtaposed $BA(v, \frac{v+2}{4}, 2; \frac{v+2}{4})$ being the first h and last h columns. For smaller r, delete $int(h-\frac{r}{2})$ columns from the first component BA and $int(h-\frac{r+1}{2})$ from the second. \square

THEOREM 30. If $k_1 = k_2 + 2$ is odd, a Schur-optimal $ECD(\bar{\theta})$ corresponds to a $BA(v, r, 2; \frac{v+4}{4})$. Existence of a Hadamard matrix of order v implies existence of the BA for every $r \le v - 1$.

PROOF. The value of $\bar{\theta}$ is $\operatorname{int}(\frac{((v+2)/2)^2}{v}) = \frac{v+4}{4}$. Given the standardized Hadamard matrix, delete the first row and column, add a row of -1's and then replace -1 by 0 throughout. The result is $BA(v, v-1, 2; \frac{v+4}{4})$. \square

7. Comments. By exploiting properties of the dual, optimality theory for resolvable designs with two blocks per replicate has been developed. Section 3 establishes conditions for Schur-optimality of equal concurrence designs depending on the block discordancy coefficient γ . For γ where Schur-optimality is not established, the class of competitors has been significantly narrowed via Schur-ordering for small r. Section 4 characterizes the class of all E-optimal designs whenever any group-affine design exists and further determines the Schur-best of the E-optimal designs. Sections 5 and 6 apply these results for the important cases $k_1 - k_2 \le 2$, including explicit design constructions employing an equivalence with balanced arrays. It is evident from the constructions that many other designs, eliminated by the Schur-domination argument in Theorem 15, do exist.

For large $(r \ge v)$ replication, the problem can be quite different and optimality work would proceed based on treatment, rather than block, concurrences. Here is a simple construction for that case: given a BIBD for v treatments in b blocks of size k, there is a Schur-optimal resolvable design in $\mathcal{D}(v,b;k,v-k)$ formed by the blocks of the BIBD and their complements. While it is possible that these designs are equivalent to balanced arrays (when the starting BIBD is symmetric), it is also true that the resulting $\phi_{dhh'}$ can be widely dispersed, for they are determined by the block concurrence counts for the underlying BIBD.

The optimality route is clear for $r \le v-1$ and $\gamma \le \frac{1}{2}$, but for larger γ there are still unanswered questions. For instance, A- and E-optimality need not coincide and the problem of determining an A-best design remains open. For $r \le 4$, the authors have solved the A-optimality problem in its entirety, including construction, and for larger r, have done this for the special cases of $k_1 - k_2 \le 2$. These and related results will be reported elsewhere.

APPENDIX A: PROOFS

A.1. Proof of Lemma 1. Let $D_s^{1/2}$ be the diagonal matrix of square roots of block sizes and write $B_d = N_d D_s^{-1/2}$. Multiplying C_d by $\frac{1}{r}$ and right and left multiplying C_{dual} by $D_s^{-1/2}$, equations (1) and (2) become

(16)
$$\frac{1}{r}C_d = I - \frac{1}{r}N_dD_s^{-1}N_d' = I - \frac{1}{r}B_dB_d' = C_d^*$$

and

(17)
$$D_s^{-1/2}C_{\text{dual}}D_s^{-1/2} = I - \frac{1}{r}D_s^{-1/2}N_d'N_dD_s^{-1/2} = I - \frac{1}{r}B_d'B_d = C_{\text{dual}}^*$$

Since (1) and (16) differ only by a constant and since the nonzero eigenvalues of $B_d B'_d$ and $B'_d B_d$ are identical, the eigenvalues of C_d can be found from those of C^*_{dual} in (17), which depends on block concurrences and block sizes.

Continuing, write $C_{\text{dual}}^* = I - \frac{1}{r} A_d$ for $A_d = B_d' B_d = D_s^{-1/2} N_d' N_d D_s^{-1/2}$. Regardless of the design d, $A_d D_s^{1/2} 1 = r D_s^{1/2} 1$; r is an eigenvalue of A_d corresponding to the zero eigenvalue common to C_{dual} and C_d . One term of the spectral decomposition of A_d is then

(18)
$$\frac{r(D_s^{1/2}1)(D_s^{1/2}1)'}{(D_s^{1/2}1)'(D_s^{1/2}1)} = \frac{1}{(k_1 + k_2)} \left[J \otimes \begin{pmatrix} k_1 & \sqrt{k_1 k_2} \\ \sqrt{k_1 k_2} & k_2 \end{pmatrix} \right],$$

where J is an $r \times r$ matrix of ones. Subtract (18) from A_d and denote the result A_d^* . Then a bit of manipulation, employing the fact that all four block concurrence counts for blocks in replicate h with blocks in replicate h' are determined by $\phi_{dhh'}$, gives

(19)
$$A_d^* = \frac{1}{k_1 k_2} M_d \otimes \begin{pmatrix} k_2 & -\sqrt{k_1 k_2} \\ -\sqrt{k_1 k_2} & k_1 \end{pmatrix}.$$

Since the eigenvalues of the 2×2 matrix in (19) are 0 and $k_1 + k_2$, the b = 2r eigenvalues of A_d^* are r copies of 0 and $\frac{v}{k_1 k_2}$ times the r eigenvalues of M_d . The eigenvalues of C_d as stated in the lemma are now immediate.

A.2. Proof of Lemma 17. The lower bound on e_{d1} follows from

$$e_{d1} = \max_{x'x=1} x' M_d x$$

$$= \max_{x'x=1} x' \left[\left(\frac{p}{v} + \gamma \right) I - \gamma J + \Delta_d \right] x$$

$$\geq \max_{x'x=1} x' \left[\left(\frac{p}{v} + \gamma \right) I - \gamma J + \Delta_d \right] x$$

$$= \frac{p}{v} + \gamma + \max_{x'x=1} x' \Delta_d x$$

$$= \frac{p}{v} + \gamma + \max_{x'x=1} x' P \Delta_d P x$$

$$= \frac{p}{v} + \gamma + \max_{x'x=1} x' P \Delta_d P x$$

$$= \frac{p}{v} + \gamma + \max_{x'x=1} x' P \Delta_d P x$$

$$= \frac{p}{v} + \gamma + u_{d1}.$$

The penultimate equality holds since $u_{d1} > 0$ and 1 is an eigenvector of $P \Delta_d P$ with eigenvalue 0. Likewise,

$$e_{dr} = \min_{x'x=1} x' M_d x$$

$$= \min_{x'x=1} x' \left[\left(\frac{p}{v} + \gamma \right) I - \gamma J + \Delta_d \right] x$$

$$\leq \frac{p}{v} + \gamma + \min_{\substack{x'x=1 \\ x'1=0}} x' P' \Delta_d P x$$

$$= \frac{p}{v} + \gamma + u_{dr}.$$

The last equality holds provided $u_{dr} < 0$, for reasons similar to those above. If $u_{dr} > 0$, the bound still holds, since $e_{dr} \le \frac{\operatorname{tr}(M_d)}{r} = \frac{p}{v} \le \frac{p}{v} + \gamma + u_{dr}$.

A.3. Proof of Lemma 19. Suppose that A - J is of the suggested form. Write $t_{(i)}$ for $t_1 + t_2 + \cdots + t_i$ and $t_{(0)} = 0$. Then PAP = P(A - J)P, implying that

$$\max \operatorname{eig}(PAP) = \max_{x'x=1} x' P(A - J) P x$$

$$= \max_{y=Px, x'x=1} y' (A - J) y$$

$$= \max_{y=Px, x'x=1} - \sum_{i=1}^{t} \left(\sum_{j=t_{(i-1)}+1}^{t_{(i)}} y_j \right)^2,$$

which is clearly nonpositive.

Now, subscripting by the dimension, suppose that $P_rA_rP_r$ is nonpositive definite (npd). Exhaustive enumeration shows that $A_r - J_r$ must have the form (13) for r = 3, 4, 5. Assuming nonpositivity implies that the form must hold for a given r, it will be shown that the same implication holds for r + 1. Denoting the upper left-hand side $r \times r$ submatrix of A_{r+1} by A_r (which is itself an adjacency matrix), it is claimed that $P_{r_1}A_{r+1}P_{r+1}$ is npd $\Rightarrow P_rA_rP_r$ is npd. If not, then there exists x such that $x'P_rA_rP_rx > 0$. Write $y' = (x'P_r, 0)$. Then $y'1 = 0 \Rightarrow y'P_{r+1}A_{r+1}P_{r+1}y = y'A_{r+1}y = x'A_rx > 0$, a contradiction.

Since $P_r A_r P_r$ is npd, the induction hypothesis says that

$$A_{r+1} - J_{r+1} = \begin{pmatrix} -J_{t_1} & & & \\ & -J_{t_2} & & \\ & & \ddots & \\ & & & -J_{t_n} \\ \hline & a' & & -1 \end{pmatrix}$$

for some vector $a_{r\times 1}$ of 0's and -1's. Indeed, by the induction hypothesis, every $s\times s$ principal minor of $A_{r+1}-J_{r+1}$ must have the form (13), so a may be partitioned as $a'=(a'_1,a'_2,\ldots,a'_n)$, where a_i is either -1_{t_i} or 0_{t_i} . If n=1, the proof is done. If n=2 and $a_2=-1_{t_2}$, then

$$A_{r+1} - J_{r+1} = \begin{pmatrix} -J_{t_1} & 0 & a_1 \\ 0 & -J_{t_2} - 1_{t_2} \\ a'_1 & -1'_{t_2} & -1 \end{pmatrix}.$$

If now $a_1 = -1_{t_1}$, then $A_{r+1} - J_{r+1}$ contains a principal minor,

$$M_3 = (-1) \begin{pmatrix} 1 & 0 & 1 \\ 0 & 1 & 1 \\ 1 & 1 & 1 \end{pmatrix},$$

which is not of the form (13), contradicting the fact that the result holds for r = 3. Thus, at most one of a_1, a_2 is nonzero. It follows immediately that for n > 2, the same statement holds for any pair $a_i, a_{i'}$. Permuting rows and columns as needed, it can be assumed that $a_i = 0_{t_i}$ for i < n and consequently that $A_{r+1} - J_{r+1}$ has the form (13).

A.4. Eigenvalue equations for H_d and proof of Lemma 20. Any vector of the form $(c'_1, c'_2, \ldots, c'_n)$, where $c_i \in \Re^{t_i}$ is either a contrast vector or the zero vector, is an eigenvector of H_d with eigenvalue 0. Consequently, all other eigenvectors are of the form $e' = (x_1 1'_{t_1}, x_2 1'_{t_2}, \ldots, x_n 1'_{t_n})$ for some scalars x_1, \ldots, x_n . The first equation in the system $H_d e = \lambda e$ is $\gamma t_1 x_1 - (1 - \gamma) t_2 x_2 - \cdots - (1 - \gamma) t_n x_n = \lambda x_1$; the other equations may be written similarly in order to see that $H_d e = \lambda e$ are equivalent to $[D_t - (1 - \gamma) 1_n t'] x = \lambda x$, where $t = (t_1, \ldots, t_n)'$ and

 $x = (x_1, ..., x_n)'$. Thus, the remaining eigenvalues of H_d are the right eigenvalues of $D_t - (1 - \gamma) 1_n t'$. Now, D_t is positive definite and E is nonnegative definite for $\gamma \ge \frac{n-1}{n}$, so both have symmetric square root matrices and D_t is invertible. Thus,

$$|D_{t} - (1 - \gamma)1_{n}t' - \lambda I| = 0$$

$$\iff |D_{t} - (1 - \gamma)1_{n}t' - \lambda I||D_{t}^{-1}| = 0$$

$$\iff |E - \lambda D_{t}^{-1}| = 0$$

$$\iff |D_{t}^{1/2}ED_{t}^{1/2} - \lambda I| = 0 \quad \text{(proving Lemma 20)}$$

$$\iff |E^{1/2}D_{t}E^{1/2} - \lambda I| = 0 \quad \text{(needed in Theorem 15)},$$

the last step following because GG' and G'G have the same eigenvalues for any square G. Now, in the proof of Theorem 15 some elements of t are allowed to be zero (without loss of generality, $t_1 = \cdots = t_z = 0$ for integer $z \ge 1$), in which case D_t is not invertible. In this case there are z additional zero eigenvalues plus a reduced system of n-z equations in t_{z+1}, \ldots, t_n . It is easy to see that $|D_t^{1/2}ED_t^{1/2}-\lambda I|=|0_{z,z}-\lambda I_z||\tilde{D}_t^{1/2}\tilde{E}\tilde{D}_t^{1/2}|$, where \tilde{E} and \tilde{D}_t are the lower-right submatrices of E and D_t of order n-z. Thus, the n eigenvalues sought are still those of $D_t^{1/2}ED_t^{1/2}$ and thus of $E^{1/2}D_tE^{1/2}$.

APPENDIX B: UNIFORM GBAs WITH FOUR GROUPS

GBA(9, (1, 1, 1, 2), 2; 2)	GBA(13, (1, 1, 1, 2), 2; 3)
11111	11111
$0\ 0\ 0\ 0\ 0$	11111
01001	01100
01010	00010
00110	10010
10010	00001
00101	1 0 0 0 1
11100	00110
10001	10100
	00101
	01010
	01001
	1 1 0 0 0

GBA(17, (2, 2, 2, 3), 2; 4)	GBA(21, (2, 2, 2, 3), 2; 5)
111111111	111111111
000000000	11111111
$1\ 0\ 0\ 1\ 0\ 0\ 1\ 0\ 1$	$1\ 1\ 0\ 0\ 0\ 0\ 1\ 1\ 0$
001101100	$1\ 0\ 1\ 0\ 1\ 0\ 1\ 0\ 0$
$1\ 0\ 0\ 0\ 1\ 1\ 0\ 0\ 1$	$0\ 1\ 1\ 0\ 1\ 0\ 0\ 1$
101110010	000101100
010010110	001010010
010111100	$1\ 0\ 0\ 0\ 0\ 0\ 1\ 1$
$1\ 0\ 0\ 0\ 0\ 1\ 1\ 1\ 0$	010101010
$1\ 1\ 1\ 0\ 0\ 1\ 0\ 0\ 1$	$0\ 1\ 0\ 0\ 0\ 0\ 1\ 0\ 1$
101010100	$1\ 0\ 0\ 1\ 0\ 1\ 0\ 0\ 1$
000101011	$0\ 1\ 1\ 0\ 0\ 1\ 0\ 0\ 1$
$0\ 1\ 1\ 0\ 0\ 1\ 0\ 1\ 0$	$1\ 0\ 1\ 0\ 0\ 1\ 0\ 0\ 1$
110100010	010011010
$0\ 1\ 0\ 1\ 1\ 0\ 0\ 0\ 1$	$0\ 1\ 1\ 1\ 0\ 0\ 1\ 0\ 0$
001010011	001001110
$0\ 1\ 1\ 0\ 0\ 0\ 1\ 0\ 1$	$1\ 0\ 1\ 1\ 0\ 0\ 0\ 1\ 0$
	$1\ 0\ 0\ 0\ 1\ 1\ 1\ 0\ 0$
	000110101
	$1\ 1\ 0\ 1\ 1\ 0\ 0\ 0\ 0$
	000110011

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REFERENCES

- BAILEY, R. A., MONOD, H. and MORGAN, J. P. (1995). Construction and optimality of affineresolvable designs. *Biometrika* 82 187–200. MR1332849
- [2] BAILEY, R. A. and SPEED, T. P. (1986). Rectangular lattice designs: Efficiency factors and analysis. Ann. Statist. 14 874–895. MR0856795
- [3] BHATIA, R. (1997). Matrix Analysis. Springer, New York. MR1477662
- [4] BOSE, R. C. (1942). A note on the resolvability of balanced incomplete block designs. Sankhyā 6 105–110. MR0008064
- [5] CHAKRAVARTI, I. M. (1956). Fractional replication in asymmetrical factorial designs and partially balanced arrays. Sankhyā 17 143–164. MR0085677
- [6] CHAKRAVARTI, I. M. (1961). On some methods of construction of partially balanced arrays. Ann. Math. Statist. 32 1181–1185. MR0130770

- [7] CHENG, C. S. (1978). Optimality of certain asymmetrical experimental designs. *Ann. Statist.* **6** 1239–1261. MR0523760
- [8] CRAIGEN, R. and KHARAGHANI, H. (2007). Hadamard matrices and Hadamard designs. In *Handbook of Combinatorial Designs*, 2nd ed. (C. J. Colbourn and J. H. Dinitz, eds.) 273–280. Chapman and Hall/CRC, Boca Raton, FL.
- [9] FUJI-HARA, R. and MIYAMOTO, N. (2000). Balanced arrays from quadratic functions. J. Statist. Plann. Inference 84 285–293. MR1747509
- [10] GHOSH, S. and TESCHMACHER, L. (2002). Comparisons of search designs using search probabilities. J. Statist. Plann. Inference 104 439–458. MR1906265
- [11] HARSHBARGER, B. (1946). Preliminary report on the rectangular lattices. *Biometrics* 2 115–119. MR0026784
- [12] HARSHBARGER, B. (1949). Triple rectangular lattices. Biometrics 5 1–13. MR0030486
- [13] JOHN, J. A., RUSSELL, K. G., WILLIAMS, E. R. and WHITAKER, D. (1999). Resolvable designs with unequal block sizes. *Aust. N. Z. J. Stat.* **41** 111–116. MR1705928
- [14] KIRKMAN, T. P. (1851). Query VI on "Fifteen young ladies" Lady's and Gentleman's Diary No. 148 48.
- [15] KUNERT, J. (1985). Optimal repeated measurements designs for correlated observations and analysis by weighted least squares. *Biometrika* **72** 375–389. MR0801776
- [16] KURIKI, S. (1993). On existence and construction of balanced arrays. Discrete Math. 116 137–155. MR1222122
- [17] MORGAN, J. P. (1996). Nested designs. In *Design and Analysis of Experiments* (S. Ghosh and C. R. Rao, eds.) 939–976. North-Holland, Amsterdam. MR1492587
- [18] MORGAN, J. P. (2007). Optimal incomplete block designs. J. Amer. Statist. Assoc. 102 655–663.
- [19] PATTERSON, H. D. and HUNTER, E. A. (1983). The efficiency of incomplete block designs in National List and Recommended List cereal variety trials. *J. Agricultural Sci.* **101** 427–433.
- [20] PATTERSON, H. D. and SILVEY, V. (1980). Statutory and recommended list trials of crop varieties in the United Kingdom. J. Roy. Statist. Soc. Ser. A 143 219–252.
- [21] PATTERSON, H. D. and WILLIAMS, E. R. (1976). A new class of resolvable incomplete block designs. *Biometrika* 63 83–92. MR0403109
- [22] RAO, C. R. (1973). Linear Statistical Inference and Its Applications, 2nd ed. Wiley, New York. MR0346957
- [23] RECK, B. H. (2002). Nearly balanced and resolvable block designs. Ph.D. dissertation, Old Dominion Univ.
- [24] SHAH, K. R. and SINHA, B. K. (1989). *Theory of Optimal Designs. Lecture Notes in Statist.* **54**. Springer, New York. MR1016151
- [25] SINHA, K., DHAR, V. and KAGEYAMA, S. (2002). Balanced arrays of strength two from block designs. J. Combin. Des. 10 303–312. MR1918802
- [26] UDDIN, N. and MORGAN, J. P. (1992). Optimal block designs with maximum blocksize and minimum replication constraints. *Comm. Statist. Theory Methods* 21 179–195. MR1159160
- [27] WILLIAMS, E. R. (1975). A new class of resolvable designs. Ph.D. dissertation, Univ. Edinburgh.
- [28] WILLIAMS, E. R., PATTERSON, H. D. and JOHN, J. A. (1976). Resolvable designs with two replications. J. Roy. Statist. Soc. Ser. B 38 296–301. MR0654162
- [29] YATES, F. (1936). A new method or arranging variety trials involving a large number of varieties. J. Agricultural Sci. 26 424–455.

[30] YATES, F. (1940). Lattice squares. J. Agricultural Sci. 30 672-687.

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