## THE MULTITYPE GALTON-WATSON PROCESS WITH IMMIGRATION<sup>1</sup>

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A multitype branching process with immigration is considered. A necessary and sufficient condition is found for when the process has a limiting stationary distribution. As a corollary to this result we are able to give necessary and sufficient conditions for when the extinction time for an ordinary multitype process has finite expectation.

1. Introduction. The theory of a single type Galton Watson process with immigration has been treated by many authors. Heathcote [4], [5] and Seneta [10] among others have found under certain assumptions, necessary and sufficient conditions for the process to possess a stationary limiting distribution. The most general result of this nature appears in a paper by Foster and Williamson [2]. It is the purpose of this note to extend the results of Heathcote and Seneta to a multitype process. In Section 2, we present a necessary and sufficient condition for the existence of a limiting distribution in the multitype case. As a corollary of this result, we give a necessary and sufficient condition for when the extinction time of a multitype process has finite expectation. This result is a direct generalization of a theorem of Seneta [9].

We now introduce our notation and formally define the process of interest. For any square matrix A we write as usual, A>0 if all elements of A are positive,  $A\ge 0$  if all elements of A are nonnegative. We call X the set of all p tuples  $i=(i_1,\cdots,i_p)'$  where each element  $i_v$  is some nonnegative integer.  $e_\alpha$  denotes the vector with  $i_v=\delta(\alpha,v),\ v=1,\cdots,p$  where  $\delta(\bullet,\bullet)$  is the Kronecker delta. The p-dimensional unit cube of points  $s=(s_1,\cdots,s_p)',\ 0\le s_v\le 1,\ v=1,\cdots,p$ , is denoted by C. For a given  $s\in C$  and  $i\in X$ , we put  $s^i=\prod_{v=1}^p s_v^{i_v}$ . We write Ds for the vector whose components are all equal to  $s,\ 0\le s\le 1$ . For ease of notation we write 0 for D0 and 1 for D1.

For our process we consider a Markov chain  $\{Z_n\}_{n\geq 0}$  defined on A consisting of a p-type Galton Watson process  $\{Y_n = (Y_{n1}, \dots, Y_{np})\}_{n\geq 0}$  (the offspring distribution) augmented by an independent random p-dimensional immigration component at each generation. For the offspring process, we assume given, an offspring vector of p-dimensional probability generating functions (pgf)

$$F(s) = F_1(s) = (F_{11}(s), \dots, F_{1p}(s))$$

$$F_{1j}(s) = \sum_{i \in X} f_j(i)s^i \qquad 1 \le j \le p.$$

with

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The  $f_i(i)$  are interpreted as

$$f_j(i) = P\{Y_1 = i \mid Y_0 = e_j\}.$$
  $i \in X, 1 \le j \le p.$ 

It is well known ([3] page 36) that the pgf vector of the *n*th generation of an ordinary *p*-type branching process is the *n*th functional iterate of F(s),  $F_n(s) = (F_{n1}(s), F_{n2}(s), \dots, F_{np}(s))$ . It is consistent with this to define  $F_0(s) = s$ . For the immigration component, we are given the pgf.

$$B(s) = \sum_{i \in \mathcal{X}} b(i)s^i$$
  $s \in \mathbb{C}$ .

It is assumed that B(0) < 1, i.e. that some immigration occurs with positive probability. Define:

$$m_{ij} = \frac{\partial F_{1i}}{\partial s_i} \left( \mathbf{1} \right) = E\{Y_{1j} \mid Y_0 = e_i\}.$$

Let M be the matrix  $(m_{ij})_{i,j=1}^p$ . It is obvious that  $M \ge 0$ . Throughout this paper, we make the following assumption.

Assumption 1. The elements of M are all finite and M is positive-regular, i.e. there exists an integer  $n_0$  such that  $M^{n_0} > 0$ .

It is proven in [7] that if M is positive-regular then M has an eigenvalue  $\rho$ , which is real, positive, simple and greater in modulus that any other eigenvalue. Let u and v be the right and left eigenvectors of  $\rho$  respectively, normalized so that (u, 1) = 1 and (v, 1) = 1. It is also proven in [7] that u > 0 and v > 0.

It has been shown in [3] that the value of  $\rho$  determines the asymptotic behavior of the p-type Galton Watson process  $\{Y_n\}$ . As is usual, we will call the  $\{Z_n\}$  process supercritical, critical or subcritical, depending on whether  $\rho > 1 = 1$  or < 1. When  $\rho = 1$ , we will always rule out the case  $F_1(s) = Ms$ ,  $s \in C$ . Define:

(1.1) 
$$P_n(i, s) = \sum_{j \in X} s^j P\{Z_n = j | Z_0 = e_i\} \qquad s \in C.$$

It has been shown in [8] that

$$(1.2) P_n(i, s) = F_{ni}(s) \cdot \prod_{r=0}^{n-1} B(F_r(s)).$$

Our main result gives necessary and sufficient conditions for when  $\lim_{n\to\infty} P_n(i, s)$  is nontrivial.

THEOREM 1. A necessary and sufficient condition for the  $\{Z_n\}$  process to satisfy for each  $j \in X$ 

$$\lim_{n\to\infty} P\{Z_n = j | Z_0 = e_i\} = \prod (j)$$

independent of  $i \in X$  where  $\sum_{j \in X} \prod (j) = 1$  is

(1.3) 
$$\int_0^1 \frac{1 - B(Ds)}{(v, F_1(Ds)) - s} \, ds < \infty.$$

If p = 1, condition (1.3) reduces exactly to the condition given in [2], [9]. Condition (1.3) is the same as for the univariant Galton Watson process with immigration having offspring pgf  $(v, F_1(Ds))$  and immigration pgf B(Ds).

We can obtain two interesting corollaries of Theorem 1.

COROLLARY 1. Assume  $\rho < 1$ . Then the result of Theorem 1 is valid iff

$$\sum_{j \in X - \{\mathbf{0}\}} b(j) \log |j| < \infty \quad (|j| = \sum_{i=1}^{p} j_i).$$

Corollary 1 is a result proven by Quine [8]. Our second Corollary deals with the p-type Galton Watson process  $\{Y_n\}$ . Define:

$$T = \inf\{k : Y_k = \mathbf{0}\}$$

i.e., T is the time to extinction. Then:

COROLLARY 2.

$$E_i(T) = E\{T \mid Y_0 = e_i\} < \infty \qquad \text{for some} \quad i \quad \text{iff} \quad \mathcal{S}_0^1 \frac{1-s}{(v, F_1(Ds))-s} \, ds < \infty \;,$$

in which case  $E_i(T) < \infty$  for all i.

Corollary 2 generalizes a result of Seneta [9] to p-dimensions.

2. Proofs of results. Without loss of generality, we can assume that  $\rho \leq 1$ . If  $\rho > 1$ , then it is a simple matter to show that  $Z_n \to \infty$  w.p.1. To prove Theorem 1, we must show that  $P(s) = \lim_{n \to \infty} P_n(i, s)$  is nonzero. Since we are assuming that  $\rho \leq 1$ ,  $\lim_{n \to \infty} F_{ni}(s) = 1$ , [5]. Therefore, from (1.2)

(2.1) 
$$P(s) = \lim_{n \to \infty} P_n(i, s) = \lim_{n \to \infty} \left[ \prod_{r=0}^{n-1} B(F_r(s)) \right].$$

(2.1) shows that P(s) is indeed independent of the choice of  $Z_0$ . Since  $0 \le B(F_r(s)) \le 1$  for all  $r \ge 0$ , a necessary and sufficient condition for

$$\lim_{n\to\infty} \left[ \prod_{r=0}^{n-1} B(F_r(s)) \right] > 0$$

is  $\sum_{r=0}^{\infty} 1 - B(F_r(s)) < \infty$ . This last sum converges if and only if

$$\sum_{r=0}^{\infty} 1 - B(F_r(\mathbf{0})) < \infty.$$

This is a consequence of the fact that there exists an integer l = l(s) such that  $s \le F_{li}(\mathbf{0})$ ,  $i \le i \le p$  and hence  $F_r(\mathbf{0}) \le F_r(s) \le F_{r+l}(\mathbf{0})$ ,  $r \ge 1$ .

The next series of lemmas will prove that (2.2) holds iff (1.3) is valid. Define:

$$h(s) = \frac{1 - B(Ds)}{(v, F_1(Ds)) - s} \qquad s \in [0, 1).$$

LEMMA 2.1. h(s) is a positive increasing function for  $s \in [0, 1)$ .

PROOF. Since (v, 1) = 1,  $(v, F_1(Ds))$  is a legitimate 1-dimensional pgf. Furthermore,

$$\frac{d}{ds}(v, F_1(Ds))\Big|_{s=1-} = (vM, \mathbf{1})$$
$$= \rho(v, \mathbf{1}) = \rho.$$

Since  $\rho$  is assumed to be less than or equal to one, we can appeal to the argument given in [2] to prove the result.  $\Box$ 

Define:

$$g(s) = (v, F_1(s) - s) \qquad s \in C.$$

LEMMA 2.2. g(s) is decreasing in s,  $(s \in C)$  i.e. if  $s_1 \leq s_2$  then  $g(s_1) \geq g(s_2)$ .

PROOF. Joffe and Spitzer [6] showed that if Assumption 1 holds then there exists a matrix E(s) such that

$$(2.3) 1 - F_1(s) = (M - E(s))(1 - s) s \in C.$$

The matrix E(s) has the properties that

- (a)  $0 < E(s) \le M \text{ if } s > 0$ , and
- (b) E(s) is "monotone", i.e.  $t \ge s \Rightarrow E(t) \le E(s)$ .

Therefore:

$$g(s) = (v, F_1(s) - s) = (v, 1 - s) - (v, 1 - F_1(s))$$
  
=  $(v, 1 - s) - (v, (M - E(s))(1 - s))$   
=  $(1 - \rho)(v, 1 - s) + (v, E(s)(1 - s))$ .

However, both of the final terms on the right are decreasing in s.  $\Box$  Define:

$$\alpha_n = \min_{1 \le i \le p} F_{ni}(\mathbf{0})$$
 and  $\beta_n = \max_{1 \le i \le p} F_{ni}(\mathbf{0})$ .

LEMMA 2.3. There exist constants  $N_0$ ,  $K_1$  and  $K_2$  such that for all  $n > N_0$ 

(2.4) 
$$\int_{\alpha_n}^{\alpha_{n+1}} h(s) ds \leq K_1(1 - B(F_{n+1}(\mathbf{0})))$$

and

(2.5) 
$$\int_{\beta_n}^{\beta_n+1} h(s) ds \ge K_2(1 - B(F_n(\mathbf{0}))) .$$

PROOF. It follows from the properties of  $F_n(\mathbf{0})$  that the  $\{\alpha_n\}$  and  $\{\beta_n\}$  are increasing sequences of real numbers. Also, since  $\rho \leq 1$ ,  $\lim_{n\to\infty} F_n(\mathbf{0}) = 1$ , [6]. Therefore,

$$\lim_{n\to\infty} \alpha_n \uparrow 1$$
 and  $\lim_{n\to\infty} \beta_n \uparrow 1$ .

We will first establish (2.4).

By Lemma 2.1

$$(2.6) \qquad \qquad \int_{\alpha_n}^{\alpha_n+1} h(s) \, ds \leq h(\alpha_{n+1})(\alpha_{n+1} - \alpha_n) \, .$$

By Lemma 2.2

(2.7) 
$$(v, F_1(D\alpha_{n+1}) - D\alpha_{n+1}) \ge (v, F_1(F_{n+1}(\mathbf{0})) - F_{n+1}(\mathbf{0}))$$
$$= (v, F_{n+2}(\mathbf{0}) - F_{n+1}(\mathbf{0})) .$$

Also:

$$(2.8) \frac{1 - B(D\alpha_{n+1})}{1 - \alpha_{n+1}} \le \frac{1 - B(D\beta_{n+1})}{1 - \beta_{n+1}} \le \frac{1 - B(F_{n+1}(0))}{1 - \beta_{n+1}}.$$

From (2.6), (2.7) and (2.8), we obtain

$$\int_{\alpha_n}^{\alpha_{n+1}} h(s) ds \leq \left[1 - B(F_{n+1}(\mathbf{0}))\right] \left[\frac{1 - \alpha_{n+1}}{1 - \beta_{n+1}}\right] \left[\frac{\alpha_{n+1} - \alpha_n}{(v, F_{n+2}(\mathbf{0}) - F_{n+1}(\mathbf{0}))}\right].$$

To prove (2.4) we need to show that for some  $N_0$ .

$$\sup_{n\geq N_0} \left[ \frac{1-\alpha_n}{1-\beta_n} \right] < \infty$$

and

(2.10) 
$$\sup_{n \ge N_0} \left[ \frac{\alpha_{n+1} - \alpha_n}{(v, F_{n+2}(\mathbf{0}) - F_{n+1}(\mathbf{0}))} \right] < \infty.$$

Since  $\lim_{n\to\infty} F_n(0) = 1$ , there exists an integer  $n_1$  such that  $F_{n_1}(0) > 0$ . Set  $N_0 = n_1 + n_0$  and let  $n > N_0$ . If we apply (2.3) to the pgf  $F_{n_0}(s)$  we obtain the existence of a matrix  $E_{n_0}(s)$  such that

$$(2.11) 1 - F_{n_0}(s) = (M_{n_0} - E_{n_0}(s))(1 - s) s \in C$$

where

$$M_{n_0} = \left(\frac{\partial F_{n_0 i}}{\partial s_i}(1)\right)_{1 \le i, j \le p}$$

Put  $s = F_{n-n_0}(\mathbf{0})$ . It follows from the properties of  $E_{n_0}(s)$  that

$$(2.12) M_{n_0} - E_{n_0}(F_{n_1}(\mathbf{0})) \le M_{n_0} - E_{n_0}(F_{n-n_0}(\mathbf{0})) \le M_{n_0} n \ge N_0.$$

Due to Assumptions 1, our choice of  $n_0$  and  $n_1$ , and (2.12) we can assert that there exists positive constants  $L_1$  and  $L_2$  such that

$$L_1 \leq M_{n_0} - E_{n_0}(F_{n-n_0}(\mathbf{0})) \leq L_2$$
  $n \geq N_0$ .

Using (2.11) it is not difficult to show that

$$\frac{1-\alpha_n}{1-\beta_n} \leq \frac{L_2}{L_1} < \infty \qquad \qquad n \geq N_0.$$

This proves (2.9). To prove (2.10) we argue as follows. Assume  $\alpha_n = F_{ni}(\mathbf{0})$  where i depends on n. Then

$$\gamma_n = \frac{\alpha_{n+1} - \alpha_n}{(v, F_{n+2}(\mathbf{0}) - F_{n+1}(\mathbf{0}))} \le \frac{F_{n+1i}(\mathbf{0}) - F_{ni}(\mathbf{0})}{v_i [F_{n+2i}(\mathbf{0}) - F_{n+1i}(\mathbf{0})]} \qquad n \ge N_0.$$

The Mean Value Theorem for p-dimensions ([1], page 117) gives for each  $1 \le i \le p$ 

$$F_{n+1i}(\mathbf{0}) - F_{ni}(\mathbf{0}) = \sum_{j=1}^{p} \frac{\partial F_{n_0 i}}{\partial S_{ij}} (S_n) (F_{n-n_0+1j}(\mathbf{0}) - F_{n-n_0 j}(\mathbf{0}))$$

and

$$F_{n+2i}(\mathbf{0}) - F_{n+1i}(\mathbf{0}) = \sum_{j=1}^{n} \frac{\partial F_{n_0+1i}}{\partial s_j} (T_n) (F_{n-n_0+1j}(\mathbf{0}) - F_{n-n_0j}(\mathbf{0})) \qquad n \ge N_0$$

where

$$F_{n-n_0}(\mathbf{0}) \le S_n, T_n \le F_{n-n_0+1}(\mathbf{0}).$$

It follows by monotonicity that

$$\frac{\partial F_{n_0i}}{\partial s_j}(S_n) \leq \max_{1 \leq i, j \leq p} \frac{\partial F_{n_0i}}{\partial s_j}(\mathbf{1}) = L_4$$

and

$$\frac{\partial F_{n_0+1i}}{\partial s_i}(T_n) \ge \min_{1 \le i, j \le p} \frac{\partial F_{n_0+1i}}{\partial s_i}(F_{n_1}(\mathbf{0})) = L_3 \qquad n \ge N_0.$$

Due to Assumption 1 and our choices of  $n_0$  and  $n_1$ ,  $0 < L_3 < L_4 < \infty$ . Thus

$$\gamma_n \leq \frac{L_4}{L_3} \frac{1}{\min_{1 \leq i \leq n} v_i} < \infty$$
  $n \geq N_0$ .

This proves (2.10). The proof of (2.5) is done in a similar way and the details will be omitted.  $\square$ 

We are now in a position to prove Theorem 1. Since the  $\{\alpha_n\}$  and the  $\{\beta_n\}$  are increasing,

$$\int_0^1 h(s) \, ds < \infty \iff \sum_{n=1}^\infty \int_{\alpha_n}^{\alpha_{n+1}} h(s) \, ds < \infty$$
$$\iff \sum_{n=1}^\infty \int_{\beta_n}^{\beta_n+1} h(s) \, ds < \infty$$

However, by Lemma 2.3

$$\sum_{n=1}^{\infty} \int_{\beta_n}^{\beta_{n+1}} h(s) ds < \infty \Rightarrow \sum_{n=1}^{\infty} \{1 - B(F_n(\mathbf{0}))\} < \infty$$

and

$$\sum_{n=1}^{\infty} \left\{ 1 - B(F_n(\mathbf{0})) \right\} < \infty \Rightarrow \sum_{n=1}^{\infty} \int_{\alpha_n}^{\alpha_{n+1}} h(s) \, ds < \infty.$$

All that remains to be proven is that  $\sum_{j \in X} \prod_{j \in X} \prod_{j \in X} p(j) = 1$  or equivalently  $\lim_{s \to 1^-} P(s) = 1$ . To do this, one can proceed exactly as in [8]. The details are omitted.

We now turn to the proofs of Corollary 1 and Corollary 2. Consider first Corollary 1. It is easy to show that

$$\lim_{s \to 1} \frac{\frac{1 - B(Ds)}{(v, F_1(Ds)) - s}}{\frac{1 - B(Ds)}{1 - s}} = (1 - p)^{-1}.$$

Therefore, if  $\rho < 1$ ,

$$\int_0^1 h(s) ds < \infty \qquad \text{iff} \quad \int_0^1 \frac{1 - B(Ds)}{1 - s} ds < \infty .$$

But

$$\int_{0}^{1} \frac{1 - B(Ds)}{1 - s} ds = \sum_{j \in X} h(j) \int_{0}^{1} \frac{1 - s^{|j|}}{1 - s} ds 
= \sum_{j \in X} h(j) (\sum_{i=1}^{|j|} 1/i).$$

Therefore:

$$\int_0^1 \frac{1 - B(Ds)}{1 - s} \, ds < \infty \qquad \text{iff} \quad \sum_{j \in X - \{\mathbf{0}\}} b(j) \log |j| < \infty.$$

This proves Corollary 1. To prove Corollary 2, we observe that

$$E_i(T) = \sum_{n=1}^{\infty} P\{T \ge n \mid Y_0 = e_i\}$$
  
=  $\sum_{n=1}^{\infty} [1 - F_{ni}(\mathbf{0})].$ 

Choose  $B(s) = s_i$ . Then,

$$\sum_{n=1}^{\infty} 1 - B(F_n(0)) = \sum_{n=1}^{\infty} 1 - F_{ni}(0).$$

So by Theorem 1,

$$\sum_{n=1}^{\infty} 1 - F_{ni}(\mathbf{0}) < \infty$$
 iff  $\int_0^1 \frac{1-s}{(v, F_1(Ds)) - s} ds$ .

## REFERENCES

- [1] APOSTLE, T. (1957). Mathematical Analysis. Addison-Wesley, London.
- [2] FOSTER, J. and J. A. Williamson (1971). Limit theorems for the Galton process with time dependent immigration. Z. Wahrscheinlichkeitstheorie und Verw. Gebiete 20 227-235.
- [3] HARRIS, T. E. (1963). The Theory of Branching Processes. Springer-Verlag, Berlin.
- [4] HEATHCOTE, C. R. (1965). A branching process allowing immigration. J. Roy. Statist. Soc. Ser. B 27 138-143.
- [5] HEATHCOTE, C. R. (1966). Corrections and comments on the paper, "A branching process allowing immigration." J. Roy. Statist. Soc. Ser. B 28 213-217.
- [6] JOFFE, A. and F. SPITZER (1967). On multitype branching process with  $\rho \le 1$ . J. Math. Anal. Appl. 19 409-430.
- [7] KARLIN, S. (1966). A First Course in Stochastic Processes. Academic Press, New York.
- [8] QUINE, M. P. (1970). The multitype Galton-Watson process with immigration. J. Appl. Probability 7 411-422.
- [9] Seneta, E. (1967). The Galton Watson process with mean one. J. Appl. Probability 4 489-495.
- [10] SENETA, E. (1968). The stationary distribution of a branching process allowing immigration: a remark on the critical case. J. Roy. Statist. Soc. Ser. B 30 176-179.
- [11] YANG, Y. S. (1972). On branching processes allowing immigration. J. Appl. Probability 9 24-31.

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