# A NOTE ON STATIONARY GAUSSIAN SEQUENCES

#### By Chandrakant M. Deo

# University of Ottawa

Let  $\{\xi_n\}$  be a stationary Gaussian sequence with  $E(\xi_0) = 0$ ;  $E(\xi_0 \xi_n) = r_n$ . If  $n^{\alpha}r_n \to 0$  for some  $\alpha > 1$ , then Strassen's functional law of iterated logarithm applies to  $\{\xi_n\}$ .

1. Main result. Let  $\{\xi_n: -\infty < n < \infty\}$  be a stationary Gaussian sequence of random variables on a probability space  $(\Omega, \mathcal{F}, P)$ . Let  $E\xi_0 = 0$  and  $S_n = \sum_{i=1}^n \xi_i$ . For n > 3,  $\omega \in \Omega$ , let  $g_n(\cdot, \omega)$  be the function on [0, 1] defined by

(1) 
$$g_n(j/n, \omega) = (2n \log \log n)^{-\frac{1}{2}} S_j, \qquad j = 0, 1, \dots, n$$

and  $g_n(\:\!\!\bullet\:\!\!, \omega)$  is linear on the subintervals [(j-1)/n, j/n],  $j=1,2,\cdots,n$ . For a nonnegative number  $\sigma$  let  $K_\sigma$  denote the set of absolutely continuous functions on [0,1] which vanish at zero and whose derivatives have  $L_2$ -norms less than or equal to  $\sigma$ . We say that  $\{\xi_n\}$  satisfies Strassen's law of iterated logarithm if there exists a  $\sigma \geq 0$ , and a set  $\Omega_0$  with  $P(\Omega_0) = 1$  such that for each  $\omega \in \Omega_0$  the sequence  $\{g_n(\:\!\!\!\bullet\:\!\!\!, \omega)\}$  is precompact in the metric space C[0,1] and has  $K_\sigma$  as the set of its limit points.

For  $n \ge 1$ , let  $r_n = (E\xi_0^2)^{-1}E(\xi_0\xi_n)$ . The object of this note is to give a simple sufficient condition in terms of the correlation sequence  $\{r_n\}$ , for  $\{\xi_n\}$  to obey Strassen's law.

The case of strong-mixing, stationary Gaussian sequences was considered in Deo (1973). For a positive integer n, let  $\alpha_n$  denote the sup |P(AB) - P(A)P(B)| where the supremum is taken over all sets A, B such that A is in the  $\sigma$ -field generated by  $\{\xi_j : j \leq 0\}$  and B in the  $\sigma$ -field generated by  $\{\xi_j : j \geq n\}$ . It is shown in [2] that if

(2) 
$$\sum \alpha_n^p < \infty \qquad \text{for some } 0 < p < 2,$$

then Strassen's law applies to  $\{\xi_n\}$ . Furthermore, the condition (2) is satisfied whenever  $\{\xi_n\}$  has a strictly positive spectral density which satisfies a Hölder condition of order greater than  $\frac{1}{2}$ .

In this note it is shown that another sufficient condition for Strassen's law to apply to  $\{\xi_n\}$  is

$$\lim_{n\to\infty} n^{\alpha} r_n = 0 \qquad \text{for some } \alpha > 1.$$

Also we give examples to show that conditions (2) and (3) overlap but neither of them implies the other. It is also shown that (3) cannot be weakened to  $\alpha > 0$  (at least not with the norming used in the definition of  $g_n$ 's and with

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Strassen's set  $K_{\sigma}$  as the limit set). The main proposition here is derived as a consequence of a very general recent theorem of I. Berkes (1973).

Under (3), the series  $\sum_{n=1}^{\infty} r_n$  converges absolutely. Let

(4) 
$$\sigma^2 = E(\xi_0^2) \{ 1 + 2 \sum_{n=1}^{\infty} r_n \}.$$

PROPOSITION. Let (3) hold. Then the sequence  $\{\xi_n\}$  obeys Strassen's law with  $\sigma$  defined in (4).

PROOF. Assume first  $\sigma > 0$ ; with this assumption there is no loss of generality in taking  $\sigma = 1$ . Let  $\delta_n = \sup_{j \ge n} |r_j|$ . Then (3) is equivalent to

(5) 
$$\lim_{n\to\infty} n^{\alpha} \delta_n = 0 \qquad \text{for some } \alpha > 1.$$

Clearly we can assume  $1 < \alpha < 2$ . We first show that there exists a constant C (depending only on the sequence  $\{\xi_n\}$ ) such that for all positive integers m, n, p we have

(6) 
$$|E\{(mn)^{-\frac{1}{2}} \sum_{i=1}^{m} \xi_i \sum_{j=m+p}^{m+p+n-1} \xi_j\}| \leq C[\min(m,n)]^{1-\alpha}$$

and

(7) 
$$|E\{n^{-\frac{1}{2}} \sum_{i=1}^{n} \xi_i\}^2 - 1| \leq C n^{1-\alpha}.$$

To prove (6) assume for the sake of definiteness  $m \le n$ . The left side of (6) is  $\le$ 

$$\begin{split} (mn)^{-\frac{1}{2}}E(\xi_0^{\,2}) \, \sum_{i=1}^m \, \sum_{j=m_1+p}^{m_1+p+n-1} |r_{j-i}| & \leq E(\xi_0^{\,2})m^{-1} \{ \sum_{j=1}^m j\delta_j + m \, \sum_{j=m+1}^\infty \delta_j \} \\ & \leq E(\xi_0^{\,2})m^{-1} \{ C_1 \, \sum_{j=1}^m j^{1-\alpha} + C_2 \, m m^{1-\alpha} \} \\ & \leq E(\xi_0^{\,2})m^{-1} \{ C_3 m^{2-\alpha} + C_2 m^{2-\alpha} \} \\ & \leq C_4 m^{1-\alpha} \, . \end{split}$$

Here,  $C_1$ ,  $C_2$ ,  $C_3$ ,  $C_4$ , are positive constants which do not depend on m, n, p. Similarly, the left side of (7) is equal to

$$\begin{split} |E(\xi_0^{\,2}) \,+\, 2E(\xi_0^{\,2}) \, & \textstyle \sum_{j=1}^{n-1} (1-j/n) r_j - 1| \\ & = |E(\xi_0^{\,2}) \,+\, 2E(\xi_0^{\,2}) \, \textstyle \sum_{n=1}^{j-1} (1-j/n) r_j - \sigma^2| \\ & \leq 2E(\xi_0^{\,2}) \, \Big\{ \frac{1}{n} \, \textstyle \sum_{j=1}^{n-1} j \delta_j + \, \textstyle \sum_{j=n}^{\infty} \delta_j \Big\} \, \, \leq C_5 \, n^{1-\alpha} \,. \end{split}$$

Thus (6) and (7) hold with  $C = \max(C_4, C_5)$ .

Now to prove the proposition let us apply Theorem 2 of Berkes (1973); and toward this end verify conditions A and  $B_3$  in that paper. Verification of condition A is straightforward in view of the easily-checked fact that for all choices of positive integer n and numbers  $a_1, a_2, \dots, a_n$  the random variable  $(\sum a_i^2)^{-\frac{1}{2}} \sum_{i=1}^n a_i \xi_i$  has normal distribution with mean zero and variance at most equal to  $\{E(\xi_0^2) + DE(\xi_0^2) \sum_{k=1}^\infty k^{-\alpha}\}$  where  $D = \sup_k k^{\alpha} \delta_k$ .

It remains to verify the condition  $B_3$  of [1]. In the notation there the left side of (2.6) is equal to

(8) 
$$\left| \exp\left( -\frac{1}{2} \sum_{i=1}^{r} \lambda_i^2 \rho(i, i) - \frac{1}{2} \sum_{i \neq j} \lambda_i \lambda_j \rho(i, j) \right) - \exp\left( -\frac{1}{2} \sum_{i=1}^{r} \lambda_i^2 \right) \right|$$

where we have written, for  $1 \le i \le r$  and  $1 \le j \le r$ ;

$$\rho(i,j) = E \{ n[(t_i{'} - t_i)(t_j{'} - t_j)]^{-\frac{1}{2}} \sum_{\substack{k = [nt_i]+1}}^{[nt_i{'}]} \xi_k \sum_{\substack{l = [nt_j]+1}}^{[nt_j{'}]} \xi_l \} \cdot$$

By (6) and (7),

$$|\rho(i,j)| \leq C(nt)^{1-\alpha} \qquad 1 \leq i \neq j \leq r,$$

and

$$(10) |\rho(i,i)-1| \leq C(nt)^{1-\alpha} 1 \leq i \leq r$$

where  $t = \min_{1 \le i \le r} (t_i' - t_i)$ .

Thus, the expression in (8) is dominated by

(11) 
$$\left| \exp\left(-\frac{1}{2}\sum_{i=1}^{r}\lambda_{i}^{2}(\rho(i,i)-1)-\frac{1}{2}\sum_{i\neq j}\lambda_{i}\lambda_{j}\rho(i,j)\right)-1\right|.$$

For any real u,  $|e^{-u} - 1| \le |u|e^{|u|}$ . Using this and (9) and (10) we see that the expression (11) is at most  $Le^L$  where L stands for

$$\frac{1}{2}C(nt)^{1-\alpha}\sum_{i=1}^{r}\lambda_{i}^{2}+\frac{1}{2}C(nt)^{1-\alpha}\sum_{i\neq j}|\lambda_{i}||\lambda_{j}|.$$

Write  $||\lambda||^2 = \sum_{i=1}^{r} \lambda_i^2$ . Then L itself is dominated by

$$C(nt)^{1-\alpha}||\lambda||^2(1+r^2).$$

Thus we have shown that the left side of (2.6) in [1] is at most

(12) 
$$C(nt)^{1-\alpha}(1+r^2)||\lambda||^2 \exp\left\{C(nt)^{1-\alpha}(1+r^2)||\lambda||^2\right\}.$$

For  $||\lambda||^2 \le C^{-1}(1+r^2)^{-1}(nt)^{(\alpha-1)/2}$ , and nt > 1, (12) is at most  $(nt)^{(\alpha-1)/2}$ . Thus the left side of (2.6) in [1] is at most max  $(2, (nt)^{(1-\alpha)/2})$  whenever  $||\lambda||^2 \le C^{-1}(1+r^2)^{-1}(nt)^{(\alpha-1)/2}$ . This verifies the condition  $B_3$  of [1] and the proof of the proposition is complete when  $\sigma > 0$ . The degenerate case  $\sigma = 0$  for which we need to show that  $g_n$  converges uniformly to zero for almost all  $\omega$  can be handled along the lines of Lemma 9 in [2]. We omit the details.

# 2. Examples.

EXAMPLE 1. Let  $\{\xi_n\}$  have spectral density

$$h(\lambda) = M + 2 \sum_{n=1}^{\infty} 2^{-3n/4} \cos 2^n \lambda , \qquad -\pi < \lambda < \pi ;$$

where M is chosen to make h strictly positive. It is known that this Weirstrass function satisfies Hölder condition of order  $\frac{3}{4}$ ; see e.g. page 47 of Zygmund (1968). In this case therefore the condition (2) is satisfied but not (3).

EXAMPLE 2. If the spectral density of  $\{\xi_n\}$  vanishes on a subinterval of  $(-\pi, \pi)$  and has two bounded derivatives then the condition (3) is satisfied. However, such a process cannot be strong-mixing because it is not even purely non-deterministic.

Example 3. Let  $\frac{1}{2} < \alpha < 1$ , and  $\{\eta_n : -\infty < n < \infty\}$  be an i.i.d. sequence of standard normal variables. Define

$$\xi_j = \sum_{k=j+1}^{\infty} \eta_k / (k-j)^{\alpha}, \qquad -\infty - j < \infty.$$

Then  $E(\xi_0 \xi_n) = O(n^{-\alpha})$ . However, the variance of  $S_n = \sum_{j=1}^n \xi_j$  is easily seen to be of the order of magnitude  $n^{3-2\alpha}$  and thus Strassen's law clearly cannot apply with the norming used in the definition of  $g_n$  functions. In this case, the proper norming would be to divide by  $[2 \operatorname{Var}(S_n) \log \log n]^{\frac{1}{2}}$  rather than by  $[2n \log \log n]^{\frac{1}{2}}$ . However, even with this different norming, it is unlikely that Strassen's set  $K_\sigma$  will appear as the limit set. This is because the finite-dimensional distributions of the sequence  $\{(\operatorname{Var} S_n)^{-\frac{1}{2}} S_{[nt]} : 0 \le t \le 1\}$  converge, not to those of the Brownian motion, but to those of the Gaussian process  $\{\zeta(t) : 0 \le t \le 1\}$  with the covariance function

$$E(\zeta(s)\zeta(t)) = \text{const.} [s^{3-2\alpha} + t^{3-2\alpha} - |t - s|^{3-2\alpha}], \quad 0 \le s, t \le 1.$$

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DEPARTMENT OF MATHEMATICS UNIVERSITY OF OTTAWA OTTAWA, ONT. KIN6N5