

is minimized by the choice  $c = c^*$ . A design which minimizes  $d(x_0)$  can then be obtained easily from  $c^*$  in a manner described in [8]; for our present considerations, we need only mention that  $d(x_0) = [m(c^*)]^{-2}$ , which can be used to tell us whether or not  $x_0 \in B$ .

Finally, we remark that the Chebyshev approximation problem just described in terms of the  $g_i$ 's can be rewritten as a "modified Chebyshev problem" in terms of the original  $f_i$ 's, namely, to minimize

$$\max_x \left| \left[ 1 + \sum_2^k c_i f_i(x_0) \right] f_1(x) / f_1(x_0) - \sum_2^k c_i f_i(x) \right|.$$

For computational purposes, it is often convenient to solve this problem by first solving the restricted Chebyshev problem of minimizing

$$\max_x \left| f_1(x) / f_1(x_0) - r^{-1} \sum_2^k c_i f_i(x) \right|$$

subject to  $\sum_2^k c_i f_i(x_0) = r - 1$ , then multiplying the resulting minimum by  $r$  and minimizing with respect to  $r$ .

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## A CONTOUR-INTEGRAL DERIVATION OF THE NON-CENTRAL CHI-SQUARE DISTRIBUTION

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The brief discussion which follows presents a contour-integral derivation of the non-central chi-square distribution. Although this distribution is well

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known, the particular mode of derivation in this paper may have interest pedagogically and may serve as an example of the utility of the contour integral approach.

Consider the expression

$$(1) \quad \rho^2 = x_1^2 + x_2^2 + x_3^2 + \dots + x_k^2,$$

where the  $x_i$  are independent and each normally distributed,  $x_i : n(a_i, \sigma)$ .

The Fourier integral  $\phi_j(t)$  of the random variable  $x_j^2$  is

$$(2) \quad \phi_j(t) = \frac{1}{(1 - 2it\sigma^2)^{\frac{1}{2}}} \exp\left(\frac{ita_j^2}{1 - 2it\sigma^2}\right),$$

and the Fourier integral  $\phi(t)$  of  $\rho^2$  is, therefore,

$$(3) \quad \phi(t) = \frac{1}{(1 - 2it\sigma^2)^{k/2}} \exp\left(\frac{it}{1 - 2it\sigma^2}\right) \sum_{j=1}^k a_j^2.$$

The probability density function of  $\rho^2$ ,  $f(\rho^2)$ , is then given by the Fourier transform of (3),

$$(4) \quad f(\rho^2) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp(-it\rho^2) \frac{1}{(1 - 2it\sigma^2)^{k/2}} \exp\left(\frac{itr^2}{1 - 2it\sigma^2}\right) dt,$$

where  $r^2 = \sum_{j=1}^k a_j^2$ . In (4) let  $z = (\rho/r)(1 - 2it\sigma^2)$ , which yields

$$(5) \quad f(\rho^2) = \frac{1}{4\pi i \sigma^2} \left(\frac{\rho}{r}\right)^{\frac{1}{2}(k-2)} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] \cdot \int_{\rho/r-i\infty}^{\rho/r+i\infty} \exp\left[\frac{\rho r}{2\sigma^2} \left(\frac{z^2 + 1}{z}\right)\right] \frac{dz}{z^{k/2}}.$$

At this point, it is convenient and interesting (but not necessary) to write (5) in the equivalent forms,

$$(6) \quad f(\rho^2) = \frac{1}{4\pi i \sigma^2} \left(\frac{\rho}{r}\right)^{\frac{1}{2}(k-2)} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] \cdot \sum_{n=0}^{\infty} \left(\frac{\rho r}{2\sigma^2}\right)^n \frac{1}{n!} \int_{\rho/r-i\infty}^{\rho/r+i\infty} \exp\left(\frac{\rho r}{2\sigma^2} z\right) \frac{dz}{z^{n+k/2}} \quad (k \text{ odd}),$$

$$(7) \quad f(\rho^2) = \frac{1}{4\pi i \sigma^2} \left(\frac{\rho}{r}\right)^{\frac{1}{2}(k-2)} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] \cdot \sum_{n=-\infty}^{\infty} I_n\left(\frac{\rho r}{\sigma^2}\right) \int_{\rho/r-i\infty}^{\rho/r+i\infty} z^{n-k/2} dz \quad (k \text{ even}),$$

where  $I_n(\rho r/\sigma^2)$  is the modified Bessel function of the first kind, of order  $n$ .

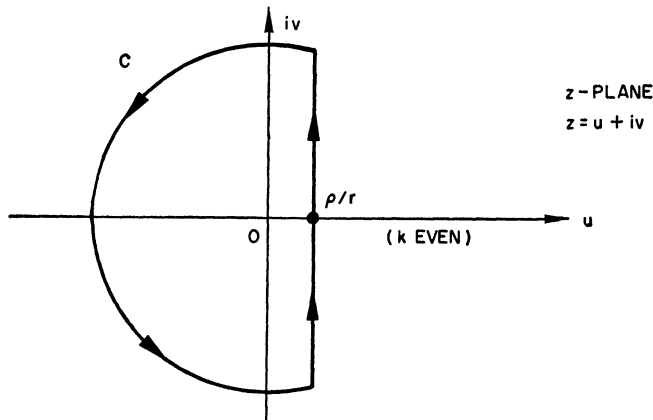


FIG. 1

Since the integrand in (7) has no branch points, the integration can be performed by considering the vertical path  $(\rho/r) - i\infty$  to  $(\rho/r) + i\infty$ , to be part of a complex contour  $C$  in the  $z$ -plane as shown in Figure 1. Equation (7) is then easily evaluated as

$$\begin{aligned}
 f(\rho^2) &= \frac{1}{2\sigma^2} \left(\frac{\rho}{r}\right)^{\frac{1}{2}k-1} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] I_{(k/2-1)}\left(\frac{\rho r}{\sigma^2}\right) \\
 (8) \qquad &= \frac{1}{2\sigma^2} \left(\frac{\rho}{r}\right)^{m-1} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] \sum_{j=0}^{\infty} \frac{\left(\frac{\rho r}{2\sigma^2}\right)^{2j+m-1}}{\Gamma(j+1)\Gamma(j+m)}.
 \end{aligned}$$

Since the integrand in expression (6) has a branch point at the origin, a somewhat different contour is required as shown in Figure 2. For reasons which will become evident later, the integral

$$(9) \qquad \int_{\rho/r-i\infty}^{\rho/r+i\infty} \exp\left(\frac{\rho r}{2\sigma^2} z\right) \frac{dz}{z^{\lambda+1}}$$

will be considered, where  $n + (k/2)$  has been parameterized in the form of  $\lambda + 1$ . Thus,

$$\begin{aligned}
 (10) \qquad \lim_{\substack{R \rightarrow \infty \\ d \rightarrow 0}} \int_c \exp\left(\frac{\rho r}{2\sigma^2} z\right) \frac{dz}{z^{\lambda+1}} &= \lim_{R \rightarrow \infty} \left\{ \int_{c_1} + \int_{c_2} + \int_{c_3} + \int_{c_7} + \int_{c_8} \right\} \\
 &\quad + \lim_{d \rightarrow 0} \int_{c_5} + \lim_{\substack{R \rightarrow \infty \\ d \rightarrow 0}} \left\{ \int_{c_4} + \int_{c_6} \right\} = 0.
 \end{aligned}$$

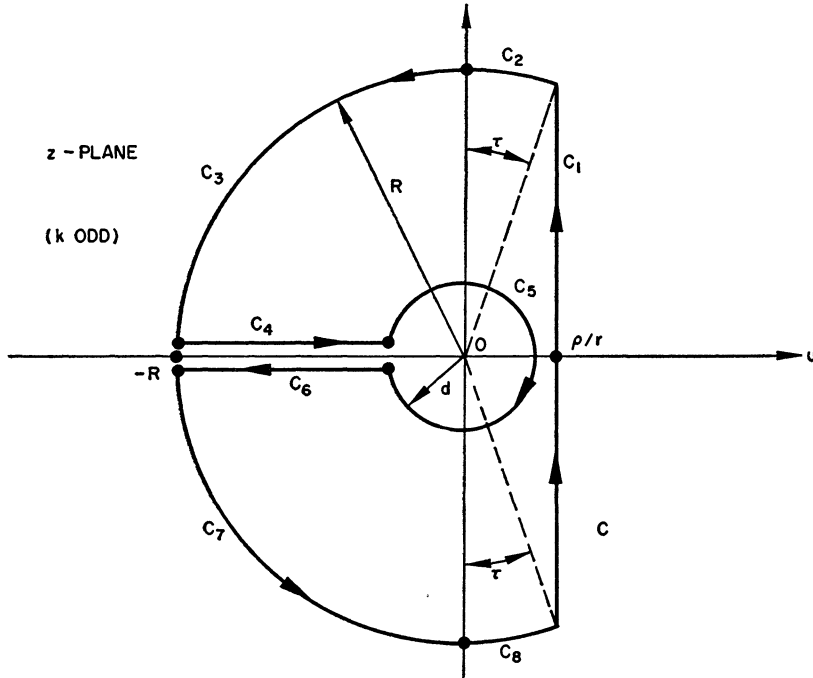


FIG. 2

It can be shown that

$$(11) \quad \lim_{R \rightarrow \infty} \int_{c_2} = \lim_{R \rightarrow \infty} \int_{c_8} = 0 \quad \text{for } \lambda > -1,$$

$$(12) \quad \lim_{R \rightarrow \infty} \int_{c_3} = \lim_{R \rightarrow \infty} \int_{c_7} = 0 \quad \text{for } \lambda > -1,$$

and

$$(13) \quad \lim_{d \rightarrow 0} \int_{c_5} = 0 \quad \text{for } \lambda < 0.$$

From (10), (11), (12), and (13), the following can be written:

$$(14) \quad \int_{(\rho/r)-i\infty}^{(\rho/r)+i\infty} \exp\left(\frac{\rho r}{2\sigma^2} z\right) \frac{dz}{z^{\lambda+1}} = -2i \sin(\lambda\pi) \int_0^\infty \exp\left(-\frac{\rho r}{2\sigma^2} s\right) \frac{ds}{s^{\lambda+1}} \\ = \frac{2\pi i}{\Gamma(\lambda+1)} \left(\frac{\rho r}{2\sigma^2}\right)^\lambda.$$

The equality in (14) is valid for  $-1 < \lambda < 0$  and, in addition, both sides of (14) are holomorphic functions of  $\lambda$  throughout the region  $-\infty < \lambda < \infty$ . Therefore,

by analytic continuation [see MacRobert [2], page 122], the equality in (14) holds for  $-\infty < \lambda < \infty$  and namely for the  $n + k/2$  under consideration in expression (6).

Expression (6) can now be written as

$$(15) \quad f(\rho^2) = \left(\frac{1}{2\sigma^2}\right)^{k/2} (\rho^2)^{\frac{1}{2}(k-2)} \exp\left[-\frac{1}{2\sigma^2}(\rho^2 + r^2)\right] \sum_{n=0}^{\infty} \frac{(\rho^2 r^2 / 2^2 \sigma^4)^n}{\Gamma(n+1)\Gamma(n + \frac{1}{2}k)} \quad (k \text{ odd}).$$

But (8), with  $k$  even, can be written in exactly the same form as (15). Thus, (15) is the density function for  $\rho^2$ , with  $k$  even or odd.

Letting  $\gamma = r^2/2\sigma^2$  and  $\chi'^2 = \rho^2/\sigma^2$ , (15) can also be written as

$$(16) \quad f\left(\frac{\rho^2}{\sigma^2}\right) d\rho^2 = f(\chi'^2)\sigma^2 d\chi'^2 = \frac{1}{2} \left(\frac{\chi'^2}{2}\right)^{\frac{1}{2}(k-2)} \exp(-\gamma) \exp\left(-\frac{\chi'^2}{2}\right) \sum_{n=0}^{\infty} \frac{(\frac{1}{2}\chi'^2\gamma)^n d\chi'^2}{n!\Gamma(n + \frac{1}{2}k)}.$$

Equation (16) is the non-central chi-square distribution, and the Fourier integral derivation of Equation (16) is different from that usually found in the literature—see Mann [3], pages 65–68 and Anderson [1], pages 112 and 113.

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A CHARACTERIZATION OF THE INVERSE GAUSSIAN DISTRIBUTION

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**1. Introduction and summary.** M. C. K. Tweedie [2] defined the inverse Gaussian distributions via the density functions

$$(1) \quad \begin{aligned} f(x; m, \lambda) &= [\lambda/(2\pi x^3)]^{\frac{1}{2}} \exp[-\lambda(x - m)^2/(2m^2x)] && \text{for } x > 0 \\ &= 0 && \text{for } x \leq 0. \end{aligned}$$

The parameters  $\lambda$  and  $m$  are positive. The corresponding densities reflected about the origin, and with  $\lambda$  and  $m$  negative, may also be considered as in the Inverse Gaussian family. The characteristic function of the Inverse Gaussian

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