

INADMISSIBILITY OF THE BEST INVARIANT TEST IN THREE OR MORE DIMENSIONS

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Consider the problem where one observes (X, Y) with X an arbitrary random quantity and Y a p -dimensional random vector, and where there are two hypotheses, H_i , under which $(X, Y - \eta)$ is distributed according to P_i ($i = 0, 1$) for some unknown point $\eta \in R^p$. Lehmann and Stein [1] show that if $p = 1$, then under certain reasonable conditions, the best invariant test of H_0 versus H_1 is admissible. Here we present an example showing that the analogous result does not hold if $p \geq 3$. The example was suggested by certain problems in the recovery of interblock information in the randomized designs considered in [2]. Although the present example is somewhat artificial and not directly applicable to the above problems, it is not unlikely that similar methods might also work there.

Let η be an unknown point in R^p and let $\varepsilon_i = (-1)^i$ for $i = 0, 1$. Suppose $X = (W, V)$ and Y are distributed as follows under H_i : W is normally distributed with mean ε_i and variance 1, V is independent of W and uniformly distributed over the surface of the unit sphere in R^p , and $Y = \eta + \varepsilon_i V$. The problem is invariant under translation: $(W, V, Y) \rightarrow (W, V, Y + c)$ for $c \in R^p$. An invariant test is one depending only on (W, V) and the best invariant test of given size accepts H_0 if and only if $W > K$. We shall take $K = 0$ for ease of computation. We want to show that for sufficiently small $a > 0$ and sufficiently large b , we have for all η

$$(1) \quad P_{0,\eta} \left\{ W + \frac{aV'Y}{b + \|Y\|^2} > 0 \right\} > P_0\{W > 0\},$$

$$(2) \quad P_{1,\eta} \left\{ W + \frac{aV'Y}{b + \|Y\|^2} > 0 \right\} < P_1\{W > 0\},$$

where V and Y are column vectors, V' denotes the transpose of V and $\|Y\|^2 = Y'Y$. Because of the symmetry of the problem under interchange of the two hypotheses, we need only prove (1).

We use the identity

$$(3) \quad \frac{1}{A+B} = \frac{1}{A} - \frac{B}{A^2} + \frac{B^2}{A^2(A+B)}$$

with

$$(4) \quad A = b + \|\eta\|^2, \quad B = 2\eta'V + 1,$$

Received January 19, 1970; revised September 28, 1970.

¹ This work facilitated by a grant from the National Science Foundation (GS-2044X).

² This work facilitated by a grant from the National Science Foundation (GS-8985).

and note that (under H_0)

$$(5) \quad \|Y\|^2 = \|\eta + V\|^2 = \|\eta\|^2 + 2\eta'V + 1$$

(since $\|V\|^2 = 1$). Thus,

$$(6) \quad \frac{V'Y}{b + \|Y\|^2} = \frac{1 + \eta'V}{b + \|\eta\|^2} \left\{ 1 - \frac{2\eta'V + 1}{b + \|\eta\|^2} + \frac{(2\eta'V + 1)^2}{(b + \|\eta\|^2)(b + \|\eta + V\|^2)} \right\} \\ = \frac{1}{b + \|\eta\|^2} \left\{ 1 + \eta'V - \frac{2(\eta'V)^2}{b + \|\eta\|^2} + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right\}$$

where the remainder is uniform in both η and V ; this follows from the fact that $\|V\| = 1$ and the results

$$(7) \quad \left| \frac{\eta'V}{b + \|\eta\|^2} \right| \leq \frac{\|\eta\|}{b + \|\eta\|^2} = O\left(\frac{1}{b^{\frac{1}{2}}}\right),$$

$$(8) \quad \frac{(1 + 2V'\eta)^2}{b + \|\eta + V\|^2} = \frac{(2V'V + 2V'\eta - 1)^2}{b + \|\eta + V\|^2} \leq \frac{8[V'(V + \eta)]^2 + 2}{b + \|\eta + V\|^2} \\ \leq \frac{8\|V + \eta\|^2 + 2}{b + \|\eta + V\|^2} = O(1).$$

Therefore

$$(9) \quad P_{0,\eta} \left\{ W + \frac{aV'Y}{b + \|Y\|^2} > 0 \right\} = E\Phi \left(1 + \frac{aV'Y}{b + \|Y\|^2} \right) \\ = E\Phi \left(1 + \frac{a}{b + \|\eta\|^2} \left[1 + \eta'V - \frac{2(\eta'V)^2}{b + \|\eta\|^2} + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right] \right) \\ = \Phi(1) + \frac{a\Phi'(1)}{b + \|\eta\|^2} E \left[1 + \eta'V - \frac{2(\eta'V)^2}{b + \|\eta\|^2} + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right] \\ + \frac{a^2}{2(b + \|\eta\|^2)^2} E\Phi''(U) \left[1 + \eta'V - \frac{2(\eta'V)^2}{b + \|\eta\|^2} + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right]^2$$

where U is a value between 1 and $(1 + aV'Y/(b + \|Y\|^2))$. But

$$(10) \quad E(\eta'V) = 0 \quad \text{and} \quad E(\eta'V)^2 = \frac{1}{p} \|\eta\|^2,$$

and since $\Phi''(x)$ is uniformly bounded, the final expectation in (9) is bounded by

$$(11) \quad KE \left[3 + \|\eta\| + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right]^2 \leq (1 + \|\eta\|^2)O(1)$$

(again uniformly in η and V). Therefore,

$$(12) \quad P_{0,\eta} \left\{ W + \frac{aV'Y}{b + \|Y\|^2} > 0 \right\} = \Phi(1) + \frac{a\Phi'(1)}{b + \|\eta\|^2} \left[1 - \frac{(2/p)\|\eta\|^2}{b + \|\eta\|^2} + O\left(\frac{1}{b^{\frac{1}{2}}}\right) \right. \\ \left. + \frac{a(1 + \|\eta\|^2)}{b + \|\eta\|^2} O(1) \right] > \Phi(1) = P_{0,\eta}\{W > 0\}$$

for all $\eta \in R^p$ if $p \geq 3$ provided $a > 0$ is sufficiently small and b sufficiently large.

Some comments about the inessentiality of special features of this example should be made.

(1) We have not really used the fact that $W \sim N(\varepsilon_i, 1)$, but only that (i) the distribution of W under H_0 is the same as the distribution of $-W$ under H_1 , (ii) the condition $W > 0$ is equivalent to rejection of the likelihood ratio test (based on W alone) for some critical level, and (iii) W has a density under H_0 which is positive at zero and has a uniformly bounded first derivative.

(2) We have strongly used the condition that the distribution of V is spherically symmetric (to obtain (10)), but the condition $\|V\|^2 = 1$ could be circumvented with somewhat greater care (both in the definition of the improved test and in the calculations).

(3) The assumption of a singular distribution for Y , that is $Y = \eta + \varepsilon_i V$, is also inessential since η can be replaced by $\eta + Z$ (where Z has an arbitrary distribution independent of (V, W) and the same under H_0 and H_1). Using conditional distributions (given Z), the above proof shows that (12) holds for the conditional probability given Z and hence, for the unconditional probability.

REFERENCES

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