

REMARK CONCERNING TWO-STATE SEMI-MARKOV PROCESSES

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Let $\{X_1, Y_1, X_2, Y_2, \dots\}$ be a sequence of independent non-negative random variables where the X 's have a common distribution function F and the Y 's, a common distribution function G . Define

$$\begin{aligned}
 S_0 &= 0, \\
 S_k &= \sum_{i=1}^k (X_i + Y_i), & k = 1, 2, \dots, \\
 S'_k &= S_k + X_{k+1}, & k = 0, 1, 2, \dots,
 \end{aligned}$$

and, for each $t, 0 < t < \infty$,

$$\begin{aligned}
 Z(t) &= 1 && \text{if } S_k < t \leq S'_k \text{ for some } k = 0, 1, \dots, \\
 &= 0 && \text{otherwise.}
 \end{aligned}$$

$Z(t)$ is a two-state Semi-Markov Process (see Lévy [2], Pyke [3] and [5], Smith [6]). Such a process arises in work sampling, and also in counter models as treated, e.g., by Pyke [4].

Let $P(t) = P(Z(t) = 1)$. From a result of Smith ([7], Theorem 1) it follows that, if $H = F * G$ (the distribution function of $X + Y$) is non-lattice, then

$$\lim_{t \rightarrow \infty} P(t) = \begin{cases} \frac{EX}{EX + EY} & \text{if } EX < \infty, \quad EY < \infty \\ 0 & EX < \infty, \quad EY = \infty. \end{cases}$$

Our remark is directed toward the behavior of $P(t)$ for large t , allowing that both EX and EY can be infinite. It consists in the following observation:

$$(1) \quad \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T P(t) dt = \lim_{s \rightarrow 0} \frac{1 - f(s)}{1 - f(s)g(s)}, \quad s > 0,$$

if either of the limits exist, where $f(s)$ and $g(s)$ are the Laplace-Stieltjes transforms of F and G . The truth of this remark is established by starting from a convolution representation of $P(t)$ (see [4]), taking Laplace transforms, and bringing to bear the Abelian and Tauberian theorems on p. 182 and p. 192 of [9].

If $f(s)$ and $g(s)$ are known, (1) is directly applicable. The following examples demonstrate the applicability of (1) for other situations.

(a) Suppose $g(s) = f^k(s)$ for some $k = 1, 2, \dots$. Then the limit in (1) exists and is equal to $1/(k + 1)$.

Received October 3, 1960; revised November 3, 1960.

¹ Work sponsored by the Office of Naval Research under Contract number Nonr 266 (55) (Nr 042-099).



(b) Suppose, as $t \rightarrow \infty$,

$$\int_0^t (1 - F(x)) dx \sim \frac{A}{\Gamma(2 - \alpha)} t^{1-\alpha}, \quad 0 < \alpha \leq 1, \quad A > 0,$$

$$\int_0^t (1 - G(x)) dx \sim \frac{B}{\Gamma(2 - \beta)} t^{1-\beta}, \quad 0 < \beta \leq 1, \quad B > 0.$$

It can be shown, using the Abelian theorem on p. 182 of [9], that the limit in (1) is $A/(A + B)$ (if $\alpha = \beta$), 1 (if $\alpha > \beta$), and 0 (if $\alpha < \beta$), a result also obtainable from [8].

The limit (1) could be studied from the point of view of Darling and Kac [1]. Possibly, their results would yield conditions on F and G for (1) to hold.

The behavior of $P(t)$ itself, for large t , does not seem to be ascertainable by the method given here.

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AN EXAMPLE OF AN ANCILLARY STATISTIC AND THE COMBINATION OF TWO SAMPLES BY BAYES' THEOREM

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1. Origin of the example. In [1], an example was given in which a fiducial distribution served as a distribution *a priori* to be combined with a different set of data (not capable of yielding probability statements), by Bayes' Theorem. In [2], it was shown that this procedure of combining samples, when each sample yielded a fiducial distribution, could lead to a contradiction. In [3], an attempt