## CONVERGENCE OF EMPIRICAL PROCESSES OF

MIXING rv's ON [0, 1]

## By C. S. WITHERS

Applied Mathematics Division, D.S.I.R., Wellington

Conditions are given for the weak convergence of weighted empirical cumulative processes of three types of mixing random variables (rv's) on [0, 1].

1. Introduction. In this section we define three types of mixing conditions and derive a basic lemma concerning them. Section 2 gives a central limit theorem for sums of uniformly bounded strongly-mixing rv's with some examples. Section 3 applies Section 1 and Section 2 to obtain the main results—the convergence of empirical processes to Gaussian processes. This work generalises a theorem of Koul [4] for independent rv's.

Let  $\{X_{iN}, i=1, 2, \dots, n_N\}$  be a sequence of rv's defined on some probability space,  $N=1, 2 \cdots$ . For  $1 \le a \le b \le n_N$  let  $M_{a,b}^N$  be the  $\sigma$ -algebra generated by  $X_{aN}, \dots, X_{bN}$ . Let  $\psi_N, \phi_N, \alpha_N$  be functions on  $\{0, 1, \dots, n_N - 1\}$  such that  $\psi_N(0) = \infty$ ,  $\phi_N(0) = \alpha_N(0) = 1$ . Suppose that for  $1 \le k \le k + i \le n_N$ ,  $A \in M_{1,k}^N$ ,  $B \in M_{k+i,n_N}^N$ 

$$(1) |P(AB) - P(A)P(B)| \le \psi_N(i)P(A)P(B)$$

then we call  $\{X_{iN}\}\ \phi_N$ -mixing. If we replace (1) by

$$(2) |P(AB) - P(A)P(B)| \le \phi_N(i)P(A)$$

or

$$(3) |P(AB) - P(A)P(B)| \le \alpha_N(i)$$

we call  $\{X_{iN}\}$   $\phi_N$ -mixing or  $\alpha_N$ -mixing respectively.

This extension of the usual notions (e.g. see Phillip [5], [6] for  $\psi_N(i) = \psi(i\gamma_N)$ ,  $\phi_N(i) = \phi(i\gamma_N)$ ,  $\alpha_N(i) = \alpha(i)$ ) allows us to obtain C.L.T. results even when  $\psi_N(1) \to \infty$  or  $\sum_i \alpha_N(i) \to \infty$  as  $N \to \infty$ .

LEMMA 1. Suppose  $1 \le k \le k + i \le n_N$ . Let X, Y be real rv's measurable  $M_{1,k}^N$  and  $M_{k+i,n_N}^N$ , respectively. Then each of the following is an upper bound (when finite) for |EXY - EXEY|, for  $\{X_{iN}\} \psi_N$ -mixing,  $(\phi_N$ -mixing,  $\alpha_N$ -mixing respectively).

- (a)  $\psi_N(i)E|X|E|Y|$
- (b)  $2\phi_N^{1/p}(i)E^{1/p}|X|^pE^{1/q}|Y|^q$ , for  $p^{-1}+q^{-1}=1$ , 1
- (c)  $2\phi_N(i)C_2E|X|$ , for  $|Y| \leq C_2$
- (d)  $4\alpha_N(i)C_1C_2$ , for  $|X| \leq C_1$ ,  $|Y| \leq C_2$

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(e) 
$$10\alpha_N(i)^{1-1/p}C_22^{-1/p}E^{1/p}|X|^p$$
, for  $|Y| \le C_2$ ,  $1 \le p < \infty$ 

(f) 
$$k\alpha_N(i)^{1-1/p-1/q}$$
, for  $1 \leq p < \infty$ ,  $1 \leq q < \infty$ ,

where  $k = K(E|X|^p, E|Y|^q)$  and  $K(x, y) = 8((x + y + x^{\frac{1}{2}}y^{\frac{1}{2}})/2)^{1/p+1/q}$ .

PROOF. For (a) to (d) see Lemmas 1, 2 of Phillip [5] and page 171 of Billingsley [1]. (f) is proved by an easy generalisation of Lemma 1.3 of Ibragimov [3], who gives  $K(x, y) = 4 + 2(x + y + x^{\frac{1}{2}}y^{\frac{1}{2}})$  for p = q. An alternative value,  $K(x, y) = 10x^{1/p}y^{1/q}$ , is given by Lemma 1 of Deo [2], and thus implies (e). (These authors all consider the stationary case).

## 2. A C.L.T. for uniformly bounded rv's.

THEOREM 1. Let  $\{X_{iN}\}$  be real rv's satisfying (3). Suppose  $EX_{iN}=0$ ,  $|X_{iN}|\leq C<\infty$ ,  $1\leq i\leq n_N$ , and  $t_N^2=E(\sum_{j=1}^n X_{jN})^2$ ,  $N\geq 1$ , where  $n=n_N\to\infty$  as  $N\to\infty$ . Let  $k=k_N$ ,  $p=p_N$ ,  $q=q_N$  be sequences of positive integers such that  $k(p+q)\leq n$ , and

(4) 
$$t_N^{-2}q^2k \sum_{1}^{k-1} \alpha_N(jp) \to 0$$

$$t_N^{-2}kq \sum_{0}^{q-1} \alpha_N(j) \to 0$$

(6) 
$$t_N^{-2}q(p+q) \sum_{1}^{k-1} \alpha_N(jp) \to 0$$

(7) 
$$t_N^{-2}(n-k(p+q)) \sum_{j=1}^{n-k(p+q)} \alpha_N(j) \to 0$$

$$k\alpha_N(q) \to 0$$
,  $t_N^{-4}k\{p \sum_{j=0}^{p-1} j^2\alpha_N(j) + p^2(\sum_{j=0}^{p-1} \alpha_N(j))^2\} \to 0$ , as  $N \to \infty$ . Then

(8) 
$$t_N^{-1} \sum_{j=1}^{n} X_{jN} \to_{\mathscr{L}} N(0, 1).$$

PROOF. This follows from the proof of Theorem 1.6 of Ibragimov, [2]: define  $S_N''$  analogously to  $S_N''$  on line 4, page 359. Equations (4) and (5) deal with the cases  $i \neq j$ , i = j respectively for the first term on the R.H.S. of line 4. Equations (6) and (7) ensure that the second and third terms on the R.H.S. of line 4 tend to zero. Hence  $ES_N''^2 \rightarrow 0$ .

Finally, in place of  $E|\sum_{i=1}^{p} x_i|^4 \leq MC^4p^3/1np$  we have used

$$E|\sum_{1+j}^{p+j} X_{iN}|^4 \leq 4! \ 4C^4 \{3p \ \sum_{0}^{p-1} (i+1)^2 \alpha_N(i) + 4p^2 (\sum_{0}^{p-1} \alpha_N(i))^2\}$$

(cf. Lemma 4, page 172 of [1].)

COROLLARY 1. Let  $\{X_{iN}\}$  be real uniformly bounded rv's with mean zero satisfying (3). Suppose  $E(\sum_{i=1}^{n} X_{jN})^{2}/n \to T < \infty$  and

(9) 
$$n = n_N \to \infty$$
 as  $N \to \infty$ .

Then any of the following sets of conditions are sufficient to ensure that  $n^{-\frac{1}{2}} \sum_{j=1}^{n} X_{jN} \to \mathcal{N}(0, T)$ .

- (a)  $\max_{1 \le i \le n} \alpha_N(i) = o(n^{-\frac{1}{2}})$
- (b)  $\max_{1 \le i \le n} i \log i \alpha_N(i) = o(e^{\frac{2}{3}}/\log^{\frac{1}{3}}e)$ , where  $e = \log n_N$ .
- (c)  $\sum_{i=1}^{n-1} \alpha_N(i) \le k < \infty$  and  $n_N^{1-a} \alpha_N([n_N^b]) \to 0$ , where 0 < 2b < a < 1-b,  $n = n_N$ .

- (d)  $\max_{1 \leq i < n} i^2 \alpha_N(i) = o(n^2)$ .
- (e)  $\sum_{1}^{n-1} i^2 \alpha_N(i) = o(n_N^{\frac{1}{2}})$ , and  $\max_{1}^{n-1} \alpha_N(i) \leq K < \infty$ .
- (f)  $\sum_{1}^{j} i^{2} \alpha_{N}(i) \leq K j^{r}$ ,  $1 \leq j < n_{N}$ , where  $r < \frac{3}{2}$ .
- (g)  $\sum_{i=1}^{n-1} i^2 \alpha_N(i) \leq Kn^r$  and  $\sum_{i=1}^{n-1} \alpha_N(i) < Kn^d$ ,

where either  $0 \le d \le \frac{1}{12}$ ,  $r < \frac{4}{3} - 2d$  or  $\frac{1}{12} \le d \le \frac{3}{10}$ ,  $r < \frac{3}{2} - 4d$ .

PROOF. If T = 0 conditions (a)—(g) are redundant. If T > 0 this is just a matter of checking the conditions of Theorem 1 with  $k = \lfloor n/(p+q) \rfloor$ , and  $p = \lfloor n^a \rfloor$ ,  $q = \lfloor n^b \rfloor$  where

- (a)  $a = \frac{1}{2}$ ,  $b = \frac{1}{4}$
- (d)  $a = \frac{5}{9}, b = \frac{1}{3}$
- (e)  $a = \frac{1}{2}$ ,  $b = \frac{1}{4}$
- (f)  $a = \frac{2}{3}$ ,  $6b = (4 + \varepsilon)/(1 + \varepsilon)$  where  $\varepsilon = 3/r 2$
- (g)  $a = 1 \varepsilon 2d$ ,  $b = \min(\frac{2}{3}, \frac{3}{4} d)$  where  $\varepsilon > 0$  is small.

For (b), use  $p = [n^{\frac{1}{2}} \cdot e^{\frac{1}{4}} \cdot \log^{\frac{1}{4}} e]$ ,  $q = [n^{\frac{1}{4}} \cdot e^{-\frac{1}{2}} \cdot \log^{-\frac{1}{2}} e]$ , and in (e)—(g) apply inequalities such as  $\sum_{1}^{n-1} \alpha_{N}(i) \leq K$ , K an integer  $\Rightarrow \sum_{1}^{n} i^{2} \alpha_{N}(i) \leq K \sum_{p=K+1}^{n} i^{2} = O(p)$ .

For (f) one uses  $\sum_{1}^{k} \alpha_{N}(jp) \leq \sup \sum_{1}^{k} \beta_{j} = \sum_{1}^{k} Kj^{-2}p^{r-2}(j^{r} - (j-1)^{r})$ , where the sup is taken over  $\{\sum_{1}^{j} i^{2}\beta_{i} \leq Kj^{r}p^{r-2}, \beta_{j} \geq 0, j = 1, \dots, k\}$ .

NOTE. Of course there is no loss in assuming K = 1 in (e). By (b), M in Theorem 1.6 of [2] can be improved to  $o(e^2/\ln e)^{\frac{1}{2}}$  where  $e = \ln n$ .

As an example we show that Theorem 3.2 of Serfling [8] holds with the assumption of strict stationarity removed provided that  $A^2$  given by (3.8) of [8] is well defined. This follows from

COROLLARY 2. Suppose for  $j=1,2,\{X_{iN}^{(j)},i=1,\cdots,n^{(j)}\}$  are  $\alpha_N$ -mixing sequences of real rv's with  $\alpha_N=\alpha_N^{(j)},\,n^{(j)}=n_N^{(j)}$  and  $n^{(1)}/n^{(2)}\to C<\infty$  as  $N\to\infty$ . Suppose that the two sequences are independent, that for each  $i,\,X_{iN}^{(j)},\,$  has continuous cdf  $F_N^{(j)},\,$  that

$$\sum_{1}^{n^{(j)}-1}\alpha_{N}^{(j)} \leq K < \infty$$
 ,

and  $\alpha_N^{(j)}$  satisfies one of the sets of conditions (a)—(e) of Corollary 1, j=1,2. Let  $U_N$  be the two-sample Wilcoxon statistic

$$U_N = (n^{(1)}n^{(2)})^{-1} \sum_{i=1}^{n^{(1)}} \sum_{j=1}^{n^{(2)}} s(X_{jN}^{(2)} - X_{iN}^{(1)})$$

where s(u) = -1, 0, 1 according as u < 0, = 0, > 0. Let  $\gamma_N = 2 \int_{N}^{\infty} F_N^{(1)} dF_N^{(2)} - 1$ . Then

$$n^{(1)}(U_N-\gamma_N)\to_{\mathscr{Q}}N(0,4V)$$

where

$$\begin{split} V &= \lim_{N \to \infty} n^{(1)-1} \operatorname{Var} \, \textstyle \sum_{1}^{n^{(1)}} x_i + C \lim_{N \to \infty} n^{(2)-1} \operatorname{Var} \, \textstyle \sum_{1}^{n^{(2)}} y_i \,, \\ x_i &= F_N^{(2)} (X_{iN}^{(1)}) \,, \\ y_i &= F_N^{(1)} (X_{iN}^{(2)}) \end{split}$$

provided V exists. Further, if  $\{x_i\}$ ,  $\{y_i\}$  do not depend on N and for all i, j

Cov 
$$(x_i, x_{i+j}) = \text{Cov}(x_1, x_{i+j})$$
  
Cov  $(y_i, y_{i+j}) = \text{Cov}(y_1, y_{1+j})$ 

then V exists and equals

$$\operatorname{Var} x_1 + 2 \sum_{1}^{\infty} \operatorname{Cov} (x_1, x_{1+i}) + C(\operatorname{Var} y_1 + 2 \sum_{1}^{\infty} \operatorname{Cov} (y_1, y_{1+i}))$$
.

PROOF. This is immediate on examining that of Theorem 3.2; the requirement  $C \neq 0$  is unnecessary.

3. Empirical processes. Theorems 2, 3, 4, 5, respectively, concern independent,  $\alpha_N$ -dependent,  $\phi_N$ -dependent, and  $\phi_N$ -dependent samples. For definition of C,  $(D, \mathcal{D})$  see [1].

Let  $(C_{iN}, \dots, C_{n_NN})$  be constants and let

$$\sigma_N^2 = n_N^{-1} \sum_{i=1}^{n_N} C_{iN}^2$$
.

Suppose  $\{X_{iN}\}\$  have cdfs  $\{F_{iN}\}\$  on [0, 1], and  $n_N \to \infty$  as  $N \to \infty$ . Let

$$g_{Nq}(t) = n_N^{-1} \sum_{i=1}^{n_N} \left| \frac{C_{iN}}{\sigma_N} \right|^q F_{iN}(t) , \qquad 0 \le t \le 1 ,$$

and  $L_N(t) = \sigma_N^{-1} n_N^{-\frac{1}{2}} \sum_{i=1}^{n_N} C_{iN}(I(X_{iN} \le t) - F_{iN}(t))$ . Let  $r, \{t_i\}$  be numbers depending on  $\delta > 0$  such that

(10) 
$$0 = t_0 < t_1 \cdots < t_r = 1,$$
$$t_i - t_{i-1} \ge \delta, i = 2, \cdots, r - 1.$$

THEOREM 2. (Koul, Theorem 2.2 of [2]). Suppose for  $N \ge 1 \{X_{iN}\}$  are independent for  $N \ge 1$ ,

(11) 
$$\max_{i=1}^{n_N} C_{iN}^2 / (n_N \sigma_N^2) \to 0 ,$$

(12) 
$$EL_{N}(s)L_{N}(t) \to K(s, t) , \qquad 0 \leq s, t \leq 1, \ as \ N \to \infty ,$$

and

(13) 
$$\lim \sup_{N\to\infty} \sup_{0\leq t\leq 1-\delta} \left(g_{N2}(t+\delta)-g_{N2}(t)\right)\to 0 \qquad \text{as } \delta\to 0.$$

Then

$$(14) L_N \to_{\mathscr{D}} L \quad \text{in} \quad (D, \mathscr{D})$$

where L is a zero-mean Gaussian process such that  $P(L \in C[0, 1]) = 1$  and EL(s)L(t) = K(s, t).

Note 1. Koul gives (incorrectly)  $t_i - t_{i-1} \leq \delta$  in (10).

Koul required  $F_{iN}$  continuous (which is not necessary from his proof),  $n_N = N$  and replaced (13) by the stronger condition

$$\lim\sup\nolimits_{N\to\infty} \max\nolimits_{1\leq i\leq N_{\boldsymbol{n}}} \max\nolimits_{1\leq \boldsymbol{j}\leq r} \left(F_{iN}(t_{\boldsymbol{j}}) - F_{iN}(t_{\boldsymbol{j}-1})\right) \to 0 \qquad \text{as} \quad \delta\to 0 \ ,$$

which is equivalent to

(15) 
$$\lim \sup_{N\to\infty} \max_{1\leq i\leq n} \sup_{t} (F_{iN}(t+\delta) - F_{iN}(t)) \to 0 \quad \text{as } \delta \to 0.$$

(For example if  $C_{iN} \equiv 1$ , and

$$F_{iN}(t) = 1$$
  $t = 1$   
=  $(1 - k_{iN})t$ ,  $0 \le t < 1$ ,

where  $n_N^{-1} \sum_{i=1}^{n_N} k_{iN} \to 0$ ,  $0 \le k_{iN} \le 1$ , but  $\max_{1 \le i \le N_n} k_{iN} \to 0$  then (13) holds but not this stronger condition.)

Note 2. If (16) holds (13) is equivalent to (13) with  $g_{Nq}$  replacing  $g_{N2}$  for any q > 0. The latter holds if  $g_{Nq}(t) \to g(t)$ ,  $0 \le t \le 1$ , where g is continuous.

THEOREM 3. Suppose for  $N \ge 1$   $\{X_{iN}\}$  are  $\alpha_N$ -mixing, (12), (13),

(16) 
$$\max_{1 \le i \le n} |C_{iN}| \le k_0 \sigma_N \qquad \text{where } k_0 < \infty,$$

for some b in  $(\frac{1}{2}, 1)$  and some d < (1 + b)/2

(17) 
$$\sum_{0}^{n-1} (j+1)^{2} \alpha_{N}(j)^{1-b} \leq k_{3} n^{d}, \qquad N \geq 1 \quad \text{where} \quad k_{3} < \infty$$

(18) 
$$\sum_{0}^{n-1} \alpha_N(j)^{1-b} < k_2, \qquad N \ge 1 \quad \text{where} \quad k_2 < \infty,$$

and

(19) 
$$g_{N1}$$
 is a continuous, strictly increasing function,  $N \ge 1$ . Then (14) holds.

Note 3. (19) can be removed if (13) is strengthened to  $\sup_t |H_N(t) - g(t)| \to 0$  as  $N \to \infty$  where  $H_N = g_{N1}$  or  $n^{-1} \sum_{i=1}^n F_{iN}$ , and  $g \in C$ .

Since our draft report [10], Deo [2] has published a special case of Theorem 3, viz  $\alpha_N = \alpha$ ,  $C_{iN} = 1$ ,  $F_{iN} = F$ , d = 0.

PROOF. For convenience we suppress N. Using d=0 in Corollary 1(g), we conclude that the finite-dimensional distributions of  $L_N$  converge to those of L. (One can also prove this under stronger conditions by adapting Theorem 19.4 of [1], based on Rosen [7]). By Theorems 15.1, 15.5 of [1] it suffices to prove (19.51) of [1] for  $L_N$ . By Note 2 with q=1 this is so if (19.51) holds for  $\bar{L}_N=L_N(g_1^{-1})$ . For  $0 \le s < t \le 1$  set

$$\begin{split} & \Delta_i = F_i(g_1^{-1}(t)) - F_i(g_1^{-1}(s)) \\ & Z_i = I(s < g_1(X_i) \le t) - \Delta_i, \ 1 \le i \le n \ . \end{split}$$

By Lemma 1(e) with  $i_1 \le i_2 = i_1 + i \le i_3 = i_2 + j \le i_4 = i_3 + k$ ,

(20) 
$$|EZ_{i_1} \cdots Z_{i_4}|/10\Delta_{i_1}^b$$

$$\leq \min \left\{ \alpha(i)^{1-b}, \alpha(k)^{1-b}, \alpha(j)^{1-b} + 10\alpha(i)^{1-b}\alpha(k)^{1-b}\Delta_{i_3}^b \right\}.$$

Therefore

$$E|\bar{L}_N(t) - \bar{L}_N(s)|^4 \le 10.4! (hk_0^3 \cdot 3k_3n^{d-1} + 10k_0^2k_2^2h^2)$$

where  $h = k_0^{1-b}(t-s)^b$ . Let  $g = \min(2b, b+2-2d)$ , and

$$R_{c} = 10.4! (3k_0^3k_3(2\varepsilon^{-1})^{2-2d} + 10k_0^2k_2^2)k_0^{1-b}$$
.

If

$$\varepsilon n^{-\frac{1}{2}} < 2(t-s)$$

then

(22) 
$$E|\bar{L}_N(t) - \bar{L}_N(s)|^4 \leq R_{\varepsilon}(t-s)^g.$$

Therefore (21)  $\Rightarrow$  (22) with  $\bar{L}_N$  replaced by  $\bar{L}_{Ni} = L_{Ni}(g_1^{-1}), i = 1, 2$  where  $L_{N1}(t) = \sigma_N^{-1} n^{-\frac{1}{2}} \sum_{C_i \geq 0} C_i(I(X_i \leq t) - F_i(t)), \text{ and } L_{N2}(t) = L_N(t) - L_{N1}(t).$  Hence by Theorem 12.2 of [1], for  $m = 1, 2, \cdots$ 

$$P(M_{m_i} \geq \varepsilon) \leq K_{\varepsilon} \delta^g$$

where  $K_{\epsilon} = R_{\epsilon} \cdot K_{4,g}^{1} \cdot \epsilon^{-4}$  and

$$M_{mj} = \max_{i=1}^{m} |\bar{L}_{Nj}(s+i\delta/m) - \bar{L}_{Nj}(s)|, \qquad j=1,2,$$

whenever

$$\varepsilon n^{-\frac{1}{2}} \le 2\delta/m \ .$$

For  $s \le t \le s + p$ ,

$$|L_N(t) - L_N(s)| \le |L_{N_1}(s+p) - L_{N_2}(s)| + |L_{N_2}(s+p) - L_{N_2}(s)| + n^{\frac{1}{2}}(g_1(s+p) - g_1(s)).$$

Hence if  $V_N = \sup_{s \le t \le s + \delta} |\bar{L}_N(t) - \bar{L}_N(s)|$  then  $V_N \le 3M_{m1} + 3M_{m2} + n^{\frac{1}{2}}\delta/m$  (c.f. (22.17), (22.18) of [1].) Hence if (23) holds and  $n^{\frac{1}{2}}\delta < \varepsilon m$  then

$$(25) P(V_N \ge 7\varepsilon) \le 2K_{\varepsilon} \delta^g.$$

Choosing m satisfying  $r^{-1} < m \varepsilon n^{-\frac{1}{2}} \le 2r^{-1}$  where  $r = [\delta]^{-1}$ , and using Corollary 8.3 of [1] with  $t_i = i/r$ , (19.51) for  $\bar{L}_N$  now follows.

THEOREM 4. If for some d < 1 the conditions of Theorem 3 hold with " $\alpha_N$ -mixing" replaced by " $\phi_N$ -mixing" and  $\alpha_N(j)^{1-b}$  is replaced by  $\phi_N(j)$  in (17) and (18), then (14) holds.

PROOF. Instead of (20) one uses

$$|EZ_{i_1} \cdot \cdot \cdot|/(4\Delta_{i_1}) \leq \min \left\{ \phi(i), \, \phi(k), \, \phi(j) \, + \, 4\phi(i)\phi(k)\Delta_{i_3} \right\}.$$

COROLLARY 3. The condition  $\sum_{0}^{\infty} i^{2}\phi(i)^{\frac{1}{2}} < \infty$  of Theorems 22.1, 22.2 of [1] can be weakened to

$$\sum_{i=0}^{n} i^2 \phi(i) = O(n^d)$$
 for some  $d < 1$ .

(This improves Sen [8] who showed that  $\sum i\phi(i)^{\frac{1}{2}} < \infty$  was sufficient.)

THEOREM 5. Suppose for  $N \ge 1$   $\{X_{iN}\}$  are  $\psi_N$ -mixing. Suppose for  $N \ge 1$  (12), (13), (16) and

$$\sum_{i=1}^{n-1} i \psi_N(i) \leq k_2 < \infty \qquad \text{for} \quad N \geq 1.$$

Then (14) holds.

PROOF. Here we avoid assuming (19) by proving for  $s \le t \le u$ 

$$E|L_N(t) - L_N(s)|^2|L_N(u) - L_N(t)|^2 \le k(g_{N1}(u) - g_{N1}(s))^2$$
,  $N \ge 1$ 

where  $k < \infty$ . This is done by breaking the L.H.S. into 29 separate sums and applying Lemma 1(a). The proof now proceeds as for Theorem 2.

Note 4. With obvious changes (e.g. replacing  $\sum_{i=1}^{n-1} \text{ by } \int_{0}^{n} di$ ) the results in this paper apply to processes  $\{X_{i,N}, 1 \le i \le n_N\}$  where i, N, n vary continuously.

EXAMPLE 1. Let  $F_0$  be a cdf on [0, 1], and  $\psi \ge 0$ , a function on [0, 1] such that  $0 < x < 1 \Rightarrow \psi(x) < \infty$ . Suppose the conditions of Theorem 2 or 3 or 4 or 5 hold and

$$\int_0^1 (x-x^2)^b \psi(x) \, dx < \infty , \quad \text{(with } b=1 \text{ for Theorems 2, 4, 5)} ,$$
 
$$\limsup_N \max_i \sup_t \frac{dF_{iN}}{dF_0}(t) < \infty ,$$

and

$$\limsup_{N} \int \delta_{N}^{2} \psi(F_{0}) dF_{0} < \infty$$

where  $\delta_N = n^{\frac{1}{2}} (n^{-1} \sigma_N^{-1} \sum_{1}^{n} C_{iN} F_{iN} - F_0) \to \delta$  uniformly as  $N \to \infty$ . Let  $F_N(x) = n^{-1} \sigma_N^{-1} \sum_{1}^{n} C_{iN} 1(X_{iN} \le x)$ . Then  $A_N = n \int (F_N - F_0)^2 \psi(F_0) dF_0 \to_{\mathscr{L}} A = \int (L + \delta)^2 \psi(F_0) dF_0$  where L is given by (14).

PROOF. The condition on  $\{dF_{iN}/dF_0\}$  implies for some  $C < \infty$   $EL_N^2 \le C(F_0 - F_0^2)^b$  in [0, 1], N large (with b = 1 if using Theorems 2, 4, 5.) Hence, for  $\varepsilon > 0$  there exists  $u \in (0, \frac{1}{2})$  and  $N_0$  such that

$$(\int_0^u + \int_{1-u}^1) E(L_N + \delta_N)^2 \psi(F_0) dF_0 < \varepsilon^2, \qquad N_0 \le N \le \infty$$

where  $L_{\infty} = L$ ,  $\delta_{\infty} = \delta$ . Therefore by Theorem 5.5 of [1],

$$B_n = \int_u^{1-u} (L_N + \delta_N)^2 \psi(F_0) dF_0 \rightarrow_{\mathscr{L}} B = \int_u^{1-u} (L + \delta)^2 \psi(F_0) dF_0$$
.

Hence,

$$\begin{aligned} |P(A_N \leq x + \varepsilon) - P(A \leq x + \varepsilon)| \\ &\leq |P(A_N \leq x + \varepsilon) - P(B_N \leq x)| + |P(B_N \leq x) - P(B \leq x)| \\ &+ |P(B \leq x) - P(A \leq x + \varepsilon)| \\ &\leq \varepsilon + P(x < A_N \leq x + \varepsilon) + \varepsilon + P(x < A \leq x + \varepsilon) . \end{aligned}$$

Hence  $A_n \to_{\mathscr{L}} A$ .

EXAMPLE 2. Let  $a(\cdot)$  be continuous and nondecreasing on [0, 1] such that a(s) > 0 for s > 0. Consider the cdf

$$F(s, u) = a(s)/a(u), \quad 0 \le s < u \le 1$$
  
= 1,  $\quad 0 < u \le s \le 1.$ 

When testing  $H_{0N}$ :  $\{F_{iN}(s) \equiv F(s, i/n_N)\}$  for  $\{X_{iN}\}$  independent, an asymptotically  $\alpha$ -level test is to reject  $H_0 \Leftrightarrow$ 

$$\int_0^1 L_{0N}^2(s) \cdot \frac{(\int_s^1 a^{-1})}{(k^{-1}A(s) + ka(s))^4} da(s) > t_{\alpha},$$

where  $L_{0N}$  denotes  $L_N$  with expectations under  $H_0$ , and  $P(\int_0^1 (W^0)^2 > t_\alpha) = \alpha$ .  $W^0$ , W are the Brownian-Bridge and Wiener process,  $\{C_{iN}\}$  satisfy (16),  $A(s) = \int_s^1 a^{-1} - a(s) \int_s^1 a^{-2}$ , and k > 0 is an arbitrary constant.

PROOF.  $n^{-1} \sum_{1}^{n} F(s, i/n) \to s + a(s) \int_{s}^{1} a^{-1} \to (13)$  and (12) holds with K(s, t) = a(s)A(t),  $s \le t$  so that by Theorem 2 under  $H_0$ ,  $L_{0N} \to_{\mathscr{L}} A \cdot W(a/A)$ . Finally on uses Theorem 5.1 of [1] and expresses W in terms of  $W^0$ .

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## REFERENCES

- [1] BILLINGSLEY, PARTRICK (1968). Convergence of Probability Measures. Wiley, New York.
- [2] DEO, CHANDRAKANT M. (1973). A note on empirical processes of strong-mixing sequences. Ann. Probability 1 870-875.
- [3] IBRAGIMOV, I. A. (1962). Some limit theorems for stationary processes. *Theor. Probability Appl.* 7 349–382.
- [4] KOUL, HIRA LAL (1970). Some convergence theorems for ranks and weighted empirical cumulatives. *Ann. Math. Statist.* 41 1768-1773.
- [5] PHILLIP, WALTER (1969). The central limit problem for mixing sequences of random variables. Z. Wahrscheinlichkeitstheorie und Verw. Gebiete 12 155-171.
- [6] PHILLIP, WALTER (1969). The remainder in the central limit theorem for mixing stochastic processes. Ann. Math. Statist. 40 601-609.
- [7] ROSEN, B. (1967). On the central limit theorem for sums of dependent rv's. Z. Wahrscheinlichkeitstheorie und Verw. Gebiete 7 48-82.
- [8] SEN, PRANAB (1971). A note on weak convergence of empirical processes for sequences of φ-mixing random variables. Ann. Math. Statist. 42 2131-2133.
- [9] SERFLING, R. J. (1968). The Wilcoxon two-sample statistic on strongly mixing processes. Ann. Math. Statist. 39 1202-1209.
- [10] WITHERS, C. S. (1973). Convergence of empirical processes of mixing random variables on [0, 1] I. Tech. Report 15, Applied Mathematics Division, D.S.I.R., Wellington.

APPLIED MATHEMATICS DIVISION
DEPT. OF SCIENTIFIC AND INDUSTRIAL RESEARCH
P.O. Box 8030
WELLINGTON, NEW ZEALAND